



**Answer the following questions:**

**Q1: [3+2+3]**

An oil drilling company drills at a large number of locations in search of oil. The probability of success at any location is 0.25 and the locations may be regarded as independent.

- What is the probability that the driller will experience 1 success if 10 locations are drilled?
- The driller feels that he will go bankrupt if he drills 10 times before experiencing his first success. What is the probability that he will go bankrupt?
- What is the probability that he will get the first success on the 10<sup>th</sup> trial?

**Q2: [5+4]**

a) The joint probability density function of the two random variables  $X$  and  $Y$

is  $f(x,y)=8xy$ ,  $0 \leq x \leq y \leq 1$ . Find  $f_{X|Y}(x|\frac{1}{2})$  and  $\rho(X,Y)$

b) Let  $X$  and  $Y$  two random variables have the joint normal (bivariate normal) distribution. What value of  $\alpha$  that minimizes the variance of  $Z=\alpha X+(1-\alpha)Y$ ? Simplify your result when  $X$  and  $Y$  are independent.

**Q3: [5+3]**

a) Let  $X = \begin{cases} 0 & \text{if } N = 0 \\ \xi_1 + \xi_2 + \dots + \xi_N & \text{if } N > 0 \end{cases}$  be a random sum and assume that  $E(\xi_k) = \mu$ ,  $E(N) = \nu$  and  $\text{Var}(\xi_k) = \sigma^2$ ,  $\text{Var}(N) = \tau^2$

Prove that  $E(X) = \mu\nu$  and  $\text{Var}(X) = \nu\sigma^2 + \mu^2\tau^2$

b) The number of accidents occurring in a factory in a week is a Poisson random variable with mean 2. The number of individuals injured in different accidents is independently distributed, each with mean 3 and variance 4. Determine the mean and variance of the number of individuals injured in a week.

---

## The Model Answer

### Q1: [3+2+3]

a) This implies that  $n=10$ ,  $p=0.25$  and  $X=1$

$$\begin{aligned}\therefore \text{pr}(x=1) &= \binom{10}{1} p^1 q^9 \\ &= 10 \times 0.25 \times 0.75^9 \\ &= 0.1877\end{aligned}$$

b) The probability that he will go bankrupt is given by

$$\begin{aligned}\text{pr}(x=0) &= \binom{10}{0} p^0 q^{10} \\ &= 0.25^0 \times 0.75^{10} \\ &= 0.0563\end{aligned}$$

c) What is the probability that he will get the first success on the 10<sup>th</sup> trial?

$$\begin{aligned}\text{pr}(x=10) &= p(1-p)^9 \\ &= 0.25(0.75)^9 \\ &= 0.0188\end{aligned}$$

### Q2: [5+4]

$$\text{a) } f_{x|y}(x,y) = \frac{f_{x,y}(x,y)}{f_y(y)}$$

$$f_{x,y}(x,y) = 8xy, \quad 0 \leq x \leq y \leq 1$$

$$\begin{aligned}\therefore f_Y(y) &= \int_{-\infty}^{\infty} f(x,y) dx \\ &= \int_0^y 8xy dx \\ &= 8y \left[ \frac{x^2}{2} \right]_0^y\end{aligned}$$

$$\therefore f_Y(y) = 4y^3, \quad 0 \leq y \leq 1$$

$$\begin{aligned}\therefore f_{x|y}(x|y) &= \frac{8xy}{4y^3} \\ &= \frac{2x}{y^2}\end{aligned}$$

$$\therefore f_{x|y}(x|\frac{1}{2}) = 8x, \quad 0 \leq x \leq 1$$

$$\begin{aligned}\therefore f_X(x) &= \int_{-\infty}^{\infty} f(x,y)dy \\ &= \int_x^1 8xydy, \quad x \leq y \leq 1 \\ &= 8x \left[ \frac{y^2}{2} \right]_x^1\end{aligned}$$

$$\therefore f_X(x) = 4x(1-x^2), \quad 0 \leq x \leq 1$$

$$\begin{aligned}E(X) &= \int_{-\infty}^{\infty} xf(x)dx \\ &= \int_0^1 x \cdot 4x(1-x^2)dx \\ &= 4 \int_0^1 (x^2 - x^4)dx\end{aligned}$$

$$\therefore E(X) = 4 \left[ \frac{x^3}{3} - \frac{x^5}{5} \right]_0^1 = \frac{8}{15}$$

$$\begin{aligned}E(X^2) &= \int_0^1 x^2 \cdot 4x(1-x^2)dx \\ &= 4 \int_0^1 (x^3 - x^5)dx\end{aligned}$$

$$\therefore E(X^2) = \frac{1}{3}$$

$$\text{Var}(X) = E(X^2) - E^2(X)$$

$$\therefore \text{Var}(X) = \frac{1}{3} - \left(\frac{8}{15}\right)^2 = \frac{11}{225}$$

Similarly,

$$E(Y) = \int_0^1 y(4y^3)dy$$

$$\therefore E(Y) = \frac{4}{5}$$

$$E(Y^2) = \int_0^1 y^2 (4y^3) dy$$

$$\therefore E(Y^2) = \frac{2}{3}$$

$$E(XY) = \int_0^1 \int_0^y xy(8xy) dx dy$$

$$= 8 \int_0^1 y^2 \left[ \frac{x^3}{3} \right]_0^y dy$$

$$\therefore E(XY) = \frac{8}{3} \int_0^1 y^5 dy = \frac{4}{9}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$$

$$\therefore \text{Var}(X) = \frac{4}{9} - \left(\frac{8}{15}\right)\left(\frac{4}{5}\right) = \frac{4}{225}$$

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

$$= \frac{4/225}{\sqrt{11/225} \sqrt{2/75}} = \frac{4}{\sqrt{66}}$$

$$\therefore \rho(X, Y) \approx 0.49$$

b)  $Z = \alpha X + (1 - \alpha)Y$

$$\text{Var}(Z) = \alpha^2 \sigma_X^2 + 2\alpha(1 - \alpha)\rho\sigma_X\sigma_Y + (1 - \alpha)^2 \sigma_Y^2$$

$$\therefore \text{Var}(Z) = \alpha^2 \sigma_X^2 + (2\alpha - 2\alpha^2)\rho\sigma_X\sigma_Y + (1 - 2\alpha + \alpha^2)\sigma_Y^2$$

To get  $\alpha^*$  that minimizes  $\text{Var}(Z)$  let  $\frac{\partial V}{\partial \alpha} = 0$

$\Rightarrow$

$$2\alpha\sigma_X^2 + (2 - 4\alpha)\rho\sigma_X\sigma_Y + (-2 + 2\alpha)\sigma_Y^2 = 0$$

$$\therefore \alpha = \alpha^* = \frac{\sigma_Y^2 - \rho\sigma_X\sigma_Y}{\sigma_X^2 - 2\rho\sigma_X\sigma_Y + \sigma_Y^2}, \quad -1 < \rho < 1$$

For independent random variables X and Y,  $\rho = 0$

Consequently,  $\alpha^* = \frac{\sigma_Y^2}{\sigma_X^2 + \sigma_Y^2}$

### Q3: [5+3]

i) To prove that  $E(X) = \mu v$

$$\text{a) } \therefore E(X) = \sum_{n=0}^{\infty} E[X|N=n]P_N(n) \quad \text{Def. of Total Expectation}$$

$$\therefore E(X) = \sum_{n=1}^{\infty} E[\xi_1 + \xi_2 + \dots + \xi_n | N=n]P_N(n) \quad \text{Def. of Random Sum}$$

$$\therefore E(X) = \sum_{n=1}^{\infty} E[\xi_1 + \xi_2 + \dots + \xi_n | N=n]P_N(n) \quad \text{Prop. of Conditional Expectation}$$

$$\therefore E(X) = \sum_{n=1}^{\infty} E[\xi_1 + \xi_2 + \dots + \xi_n]P_N(n) \quad \text{where } N \text{ is independent of } \xi_1, \xi_2, \dots$$

$$\therefore E(\xi_k) = \mu, \quad k=1, 2, \dots, n$$

$$\therefore E(X) = \sum_{n=1}^{\infty} n\mu P_N(n)$$

$$\therefore E(X) = \mu \sum_{n=1}^{\infty} nP_N(n)$$

$$\therefore E(X) = \mu E(N) = \mu v$$

i) To prove that  $\text{Var}(X) = v\sigma^2 + \mu^2\tau^2$

$$\text{Var}(X) = E[(X - \mu v)^2]$$

$$= E[X - N\mu + N\mu - v\mu]^2$$

$$\text{Var}(X) = E[(X - N\mu)^2] + E[\mu^2(N - v)^2] + 2E[\mu(X - N\mu)(N - v)] \quad (1)$$

$$\begin{aligned} \therefore E[(X - N\mu)^2] &= \sum_{n=0}^{\infty} E[(X - N\mu)^2 | N=n]P_N(n) \\ &= \sum_{n=1}^{\infty} E[(\xi_1 + \xi_2 + \dots + \xi_n - n\mu)^2 | N=n]P_N(n) \end{aligned}$$

$$\therefore E[(X - N\mu)^2] = \sum_{n=1}^{\infty} E[(\xi_1 + \xi_2 + \dots + \xi_n - n\mu)^2]P_N(n)$$

$$\therefore \text{Var}(\xi_k) = E(\xi_k - \mu)^2 = \sigma^2, \quad k=1, 2, \dots, n$$

$$\begin{aligned} \therefore E[(X - N\mu)^2] &= \sum_{n=1}^{\infty} n\sigma^2 P_N(n) \\ &= \sigma^2 \sum_{n=1}^{\infty} nP_N(n) \\ \therefore E[(X - N\mu)^2] &= \nu\sigma^2, \text{ where } \sum_{n=1}^{\infty} nP_N(n) = \nu \quad (2) \end{aligned}$$

$$\begin{aligned} E[\mu^2(N - \nu)^2] &= \mu^2 E[(N - \nu)^2] \\ \therefore E[\mu^2(N - \nu)^2] &= \mu^2 \text{Var}(N) = \mu^2 \tau^2 \quad (3) \end{aligned}$$

Also,

$$\begin{aligned} E[\mu(X - N\mu)(N - \nu)] &= \mu \sum_{n=1}^{\infty} E[(X - n\mu)(n - \nu) | N = n] P_N(n) \\ &= \mu \sum_{n=1}^{\infty} (n - \nu) E[(X - n\mu) | N = n] P_N(n) \\ &= 0 \quad (4) \end{aligned}$$

where  $E[(X - n\mu) | N = n] = E(X - n\mu)$  independent prop.

$$\begin{aligned} &= E(\xi_1 + \xi_2 + \dots + \xi_n - n\mu) \\ &= n\mu - n\mu = 0 \end{aligned}$$

Substitute (2), (3) and (4) in (1), we get

$$\text{Var}(X) = \nu\sigma^2 + \mu^2\tau^2$$

b)

$N \sim \text{Poisson}(2)$

$N$  is the # of accidents in a week

$\xi_k$  is the # of individuals injured for  $k$ th accident

$$E(\xi_k) = 3, \quad \text{var}(\xi_k) = 4$$

$$E(N) = 2, \quad \text{var}(N) = 2$$

$$\therefore E(X) = \mu\nu = 3(2) = 6$$

$$\text{var}(X) = \nu\sigma^2 + \mu^2\tau^2$$

$$\therefore \text{var}(X) = 2(4) + 9(2) = 26$$