## Chapter 3

## Fundamental Sampling Distributions

## Department of Statistics and Operations Research



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## Plan

(1) Random sampling and statistics
(2) Sampling Distribution of Means and the Central Limit Theorem
(3) Sampling Distribution of the Difference between Two Means

4 Sampling Distribution of the Variance
(5) The Student's Distribution
(6) The Fisher Distribution
(7) The Fisher with Two Sample Variances
(8) Sampling Distribution of Proportions and the Central Limit
(9) Sampling Distribution of the Difference between Two Proportions

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## Definitions

(1) A population consists of the totality of the observations with which we are concerned.
(2) A sample is a subset of a population.
(3) Any function of the random variables constituting a random sample is called a statistic.

- Sample mean: $\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}$.
- Sample median: $\widetilde{X}=\left\{\begin{array}{l}x_{\frac{n+1}{2}}, \text { if } n \text { is odd, } \\ \frac{1}{2}\left(x_{\frac{n}{2}}+x_{\frac{n}{2}+1}\right), \text { if } n \text { is even. }\end{array}\right.$
- Sample variance: $S^{2}=\frac{1}{n-1} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}$.

The computed value of $S^{2}$ for a given sample is denoted by $s^{2}$.

## Theorem

If $S^{2}$ is the variance of a random sample of size $n$, we may write

$$
S^{2}=\frac{1}{n-1}\left[\sum_{i=1}^{n} X_{i}^{2}-n \bar{X}^{2}\right]
$$

- Sample standard deviation: $S=\sqrt{S^{2}}$


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## Definition

The probability distribution of a statistic is called a sampling distribution.

## Theorem

If $X_{1}, X_{2}, \ldots, X_{n}$ are independent random variables having normal distributions with means $\mu_{1}, \mu_{2}, \ldots, \mu_{n}$ and variances $\sigma_{1}^{2}, \sigma_{2}^{2}, \ldots, \sigma_{n}^{2}$, respectively, then the random variable $Y=a_{1} X_{1}+a_{2} X_{2}+\cdots+a_{n} X_{n}$ has a normal distribution with mean

$$
\mu_{Y}=a_{1} \mu_{1}+a_{2} \mu_{2}+\cdots+a_{n} \mu_{n}
$$

and variance

$$
\sigma_{Y}^{2}=a_{1}^{2} \sigma_{1}^{2}+a_{2}^{2} \sigma_{2}^{2}+\cdots+a_{n}^{2} \sigma_{n}^{2}
$$

Suppose that a random sample of $n$ observations is taken from a normal population with mean $\mu$ and variance $\sigma^{2}$. Each observation $X_{i}, i=1,2, \ldots, n$, of the random sample will then have the same normal distribution. Hence, we conclude that

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i}
$$

has a normal distribution with mean

$$
\mu_{\bar{X}}=\frac{1}{n}\{\mu+\mu+\ldots+\mu\}=\frac{1}{n} \sum_{i=1}^{n} \mu=\mu
$$

and variance

$$
\sigma_{\bar{X}}^{2}=\frac{1}{n^{2}}\left\{\sigma^{2}+\sigma^{2}+\ldots+\sigma^{2}\right\}=\frac{1}{n^{2}} \sum_{i=1}^{n} \sigma^{2}=\frac{\sigma^{2}}{n}
$$

## Corollary

If $X_{1}, X_{2}, \ldots, X_{n}$ are independent random variables having normal distributions with means $\mu$ and variances $\sigma^{2}$, then the sample mean $\bar{X}$ is normally distributed with mean equal to $\mu$ and standard deviation equal to $\sigma / \sqrt{n}$. Consequently the random variable

$$
Z=\frac{(\bar{X}-\mu)}{\sigma / \sqrt{n}}
$$

is a standard normal distribution.

## Theorem (Central limit theorem)

If $\bar{X}$ is the mean of a random sample of size $n$ taken from a population with mean $\mu$ and finite variance $\sigma^{2}$, then the limiting form of the distribution of

$$
Z=\frac{(\bar{X}-\mu)}{\sigma / \sqrt{n}}
$$

as $n \rightarrow \infty$, is the standard normal distribution $N(0,1)$.
The normal approximation for $\bar{X}$ will generally be good if $n \geq 30$.

## Example

An electrical firm manufactures light bulbs that have a length of life that is approximately normally distributed, with mean equal to 800 hours and a standard deviation of 40 hours. Find the probability that a random sample of 16 bulbs will have an average life of less than 775 hours.

## Solution

Here $\mu=800, \sigma=40$ and $n=16$. The random variable $\bar{X}$ is normally distributed with mean $\mu_{\bar{X}}=\mu=800$ and standard deviation $\sigma_{\bar{X}}=\sigma_{X} / \sqrt{n}=10$.
Then $(\bar{X}-800) / 10$ is a standard normal distribution $N(0,1)$. Hence,

$$
P(\bar{X}<775)=P\left(\frac{\bar{X}-800}{10}<\frac{775-800}{10}\right)=P(Z<-2.5)=0.0062
$$

## Example

Traveling between two campuses of a university in a city via shuttle bus takes, on average, 28 minutes with a standard deviation of 5 minutes. In a given week, a bus transported passengers 40 times. What is the probability that the average transport time was more than 30 minutes?

## Solution

In this case, $\mu=28$ and $\sigma=3$. We need to calculate the probability $\operatorname{Pr}(\bar{X}>30)$ with $n=40$. Hence,

$$
\begin{aligned}
P(\bar{X}>30) & =P\left(\frac{\bar{X}-28}{5 / \sqrt{40}} \geq \frac{30-28}{5 / \sqrt{40}}\right) \\
& =P(Z \geq 2.53)=1-P(Z \leq 2.53) \\
& =1-0.9925=0.0075
\end{aligned}
$$

There is only a slight chance that the average time of one bus trip will exceed 30 minutes.

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## Theorem

If independent samples of size $n_{1}$ and $n_{2}$ are drawn at random from two populations, discrete or continuous, with means $\mu_{1}$ and $\mu_{2}$ and variances $\sigma_{1}^{2}$ and $\sigma_{2}^{2}$, respectively, then the sampling distribution of the differences of means, $\bar{X}_{1}-\bar{X}_{2}$, is approximately normally distributed with mean and variance given by

$$
\mu_{\bar{X}_{1}-\bar{X}_{2}}=\mu_{1}-\mu_{2} \text { and } \sigma_{\bar{X}_{1}-\bar{X}_{2}}^{2}=\frac{\sigma_{1}^{2}}{n_{1}}+\frac{\sigma_{2}^{2}}{n_{2}}
$$

Hence,

$$
Z=\frac{\left(\bar{X}_{1}-\bar{X}_{2}\right)-\left(\mu_{1}-\mu_{2}\right)}{\sqrt{\sigma_{1}^{2} / n_{1}+\sigma_{2}^{2} / n_{2}}}
$$

is approximately a standard normal variable.
If both $n_{1}$ and $n_{2}$ are greater than or equal to 30, the normal approximation for the distribution of $\bar{X}_{1}-\bar{X}_{2}$ is good.

## Example

Eighteen specimens are painted using type A, and the drying time, in hours, is recorded for each. The same is done with type $B$. The population standard deviations are both known to be 1.0. Assuming that the mean drying time is equal for the two types of paint, find $P\left(\bar{X}_{A}-\bar{X}_{B}>1.0\right)$, where $\bar{X}_{A}$ and $\bar{X}_{B}$ are average drying times for samples of size $n_{A}=n_{B}=18$.

## Solution

From the sampling distribution of $\bar{X}_{A}-\bar{X}_{B}$, we know that the distribution is approximately normal with mean
$\mu_{\bar{X}_{A}-\bar{X}_{B}}=\mu_{A}-\mu_{B}=0$ and variance $\sigma_{\bar{X}_{A}-\bar{X}_{B}}^{2}=\frac{\sigma_{A}^{2}}{n_{A}}+\frac{\sigma_{B}^{2}}{n_{B}}=1 / 9$.
Corresponding to the value $\bar{X}_{A}-\bar{X}_{B}=1.0$, we have

$$
z=\frac{1-\left(\mu_{A}-\mu_{B}\right)}{\sqrt{1 / 9}}=\frac{1-0}{\sqrt{1 / 9}}=3
$$

so

$$
\operatorname{Pr}(Z>3.0)=1-P(Z<3.0)=1-0.9987=0.0013
$$

## Example

The television picture tubes of manufacturer $A$ have a mean lifetime of 6.5 years and a standard deviation of 0.9 year, while those of manufacturer $B$ have a mean lifetime of 6.0 years and a standard deviation of 0.8 year. What is the probability that a random sample of 36 tubes from manufacturer $A$ will have a mean lifetime that is at least 1 year more than the mean lifetime of a sample of 49 tubes from manufacturer $B$ ?

## Solution

We are given the following information:

$$
\begin{array}{cc}
\text { Population } 1 & \text { Population 2 } \\
\mu_{1}=6.5 & \mu_{2}=6.0 \\
\sigma_{1}=0.9 & \sigma_{2}=0.8 \\
n_{1}=36 & n_{2}=49
\end{array}
$$

If we use, the sampling distribution of $\bar{X}_{1}-\bar{X}_{2}$ will be approximately normal and will have a mean and standard deviation

$$
\mu_{\bar{X}_{1}-\bar{X}_{2}}=6.5-6.0 \text { and } \sigma_{\bar{X}_{1}-\bar{X}_{2}}=\sqrt{\frac{0.81}{36}+\frac{0.64}{49}}=0.189
$$

Hence,

$$
\begin{aligned}
\operatorname{Pr}\left(\bar{X}_{1}-\bar{X}_{2}\right. & \geq 1.0)=P(Z>2.65)=1-P(Z<2.65) \\
& =1-0.9960=0.0040
\end{aligned}
$$

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## Theorem

(1) If $X_{1}, X_{2}, \ldots, X_{n}$ an independent random sample that have the same standard normal distribution then $X=\sum_{i=1}^{n} X_{i}^{2}$ is chi-squared distribution, with degrees of freedom $\nu=n$.
(2) The mean and variance of the chi-squared distribution $\chi^{2}$ with $\nu$ degrees of freedom are $\mu=\nu$ and $\sigma^{2}=2 \nu$.

| c | C |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.995 | 0.99 | 10.98 | 0.975 | 0.95 | 0.90 | 0.80 | 10.75 | 0.70 | 10. 50 |
| 1 | 0.04393 | $0.0{ }^{3} 157$ | $0.0{ }^{3} 628$ | $0.0{ }^{3} 982$ | 0.00393 | 0.0158 | 0.0642 | 0.102 | 0.148 | 0.455 |
| 2 | 0.0100 | 0.0201 | 0.0404 | 0.0506 | 0.10 .3 | 0.211 | 0.446 | 0.575 | 0.713 | 1.386 |
| 3 | 0.0717 | 0.115 | 0.185 | 0.216 | 0.35 .2 | 0.584 | 1.005 | 1.213 | 1.424 | 2.366 |
| 1 | 0.207 | 0.297 | 0.429 | 0.484 | 0.711 | 1.064 | 1. 6.49 | 1.923 | 2.195 | 3.357 |
| 5 | 0.412 | 0.55 .4 | 0.752 | 0.831 | 1.145 | 1.610 | 2.343 | 2.675 | 3.000 | 4.35 .1 |
| 6 | 0.676 | 0.872 | 1. 134 | 1.237 | 1. 635 | 2.204 | 3.070 | 3.455 | 3.828 | 5.348 |
| 7 | 0.989 | 1.239 | 1.564 | 1.690 | 2.167 | 2.833 | 3.822 | 4.255 | 4.671 | 6.346 |
| 8 | 1.344 | 1.647 | 2.032 | 2.180 | 2.733 | 3.490 | 4.594 | 5.071 | 5.527 | 7.344 |
| 9 | 1.735 | 2.088 | 2.532 | 2.700 | 3.325 | 4.168 | 5.380 | 5.899 | 6.393 | 8.343 |
| 10 | 2.156 | 2.558 | 3.059 | 3.247 | 3.940 | 4.865 | 6.179 | 6.737 | 7.267 | 9.342 |
| 11 | 2.603 | 3.053 | 3.609 | 3.816 | 4.575 | 5.578 | 16.989 | 7.584 | 8.148 | 10.341 |
| 12 | 3.074 | 3.571 | 4.178 | 4.404 | 5.226 | 6.304 | 7.807 | 8.438 | 9.034 | 11.340 |
| 13 | 3.565 | 4.107 | 4.765 | 5.009 | 5.892 | 7.041 | 8.634 | 9.299 | 9.926 | 12.340 |
| 14 | 4.075 | 4.660 | 5.368 | 5.629 | 6.571 | 7.790 | 9.467 | 10. 165 | 10.821 | 13.339 |
| 15 | 4.601 | 5.229 | 5.985 | 6.262 | 7.261 | 8.547 | 10.307 | 11.037 | 11.721 | 14.339 |
| 16 | 5.142 | 5.812 | 6.614 | 16.908 | 7.962 | 9.312 | 11.152 | 11.912 | 12.624 | 15.338 |
| 17 | 5.6997 | 6.408 | 7.255 | 7.564 | 8.672 | 10.085 | 12.002 | 12.792 | 13.531 | 16.3388 |
| 18 | 6.265 | 7.015 | 7.906 | 8.231 | 9.390 | 10.865 | 12.857 | 13.675 | 14.440 | 17.338 |
| 19 | 6.844 | 7.633 | 8.567 | 8.907 | 10.117 | 11.651 | 13.716 | 14.562 | 15.352 | 18.338 |
| 20 | 7.434 | 8.260 | 9.237 | 9.591 | 10.85.1 | 12.443 | 14.578 | 15.452 | 16.266 | 19.337 |
| 21 | 8.034 | 8.897 | 9.915 | 10.283 | 11.591 | 13.240 | 15.445 | 16.344 | 17.182 | 20.337 |
| 22 | 8.643 | 9.542 | 10.600 | 10.982 | 12.338 | 14.041 | 16.314 | 17.240 | 18.101 | 21.337 |
| 23 | 9.260 | 10.196 | 11.293 | 11.689 | 13.091 | 14.848 | 17.187 | 18.137 | 19.021 | 22.337 |
| 24 | 9.886 | 10.856 | 11.992 | 12.401 | 13.848 | 15.659 | 18.062 | 19.037 | 19.943 | 23.337 |
| 25 | 10.520 | 11.524 | 12.697 | 13.120 | 14.611 | 16.473 | 18.940 | 19.939 | 20.867 | 24.337 |
| 26 | 11.160 | 12.198 | 13.409 | 13.844 | 15.379 | 17.292 | 19.820 | 20.843 | 21.792 | 25.336 |
| 27 | 11.808 | 12.878 | 14.125 | 14.573 | 16.151 | 18.114 | 20.703 | 21.749 | 22.719 | 26.336 |
| 28 | 12.461 | 13.56 .5 | 14.847 | 15.308 | 16.928 | 18.939 | 21.588 | 22.657 | 23.647 | 27.336 |
| 29 | 13.121 | 14.256 | 15.574 | 16.047 | 17.708 | 19.768 | 22.475 | 23.567 | 24.577 | 28.336 |
| 30 | 13.787 | 14.953 | 16.306 | 16.791 | 18.493 | 20.599 | 23.364 | 24.478 | 25.508 | 29.336 |
| 40 | 20.707 | 22.164 | 23.838 | 24.433 | 26.509 | 29.051 | 32.345 | 33.66 | 34.872 | 39.335 |
| 50 | 27.991 | 29.707 | 31.664 | 32.357 | 34.764 | 37.689 | 41.449 | 42.942 | 44.313 | 49.335 |
| 6.0 | 35.534 | 37.485 | 39.699 | 40.482 | 13.188 | 46.459 | 50.641 | 52.294 | 53.809 | 59.335 |

Figure: Table A. 5 Critical Values of the Chi-Squared Distribution

Table A. 5 (continued) Critical Values of the Chi-Squared Distribution

| $v$ |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.30 | 0.25 | 0.20 | 0.10 | 0.05 | 0.025 | 0.02 | 0.01 | 0.005 | 0.001 |
| 1 | 1.074 | 1.323 | 1.642 | 2.706 | 3.841 | 5.024 | 5.412 | 6.635 | 7.879 | 10.827 |
| 2 | 2.408 | 2.773 | 3.219 | 4.605 | 5.991 | 7.378 | 7.824 | 9.210 | 10.597 | 13.815 |
| 3 | 3.665 | 4.108 | 4.642 | 6.251 | 7.815 | 9.348 | 9.837 | 11.345 | 12.838 | 16.266 |
| 4 | 4.878 | 5.385 | 5.989 | 7.779 | 9.488 | 11.143 | 11.668 | 13.277 | 14.860 | 18.466 |
| 5 | 6.064 | 6.626 | 7.289 | 9.236 | 11.070 | 12.832 | 13.388 | 15.086 | 16.750 | 20.515 |
| 6 | 7.231 | 7.841 | 8.558 | 10.645 | 12.592 | 14.449 | 15.033 | 16.812 | 18.548 | 22.457 |
| 7 | 8.383 | 9.037 | 9.803 | 12.017 | 14.067 | 16.013 | 16.622 | 18.475 | 20.278 | 24.321 |
| 8 | 9.524 | 10.219 | 11.030 | 13.362 | 15.507 | 17.5 .35 | 18.168 | 20.090 | 21.95 .5 | 26.124 |
| 9 | 10.656 | 11.389 | 12.242 | 14.684 | 16.919 | 19.023 | 19.679 | 21.666 | 23.589 | 27.877 |
| 10 | 11.781 | 12.549 | 13.442 | 15.987 | 18.307 | 20.483 | 21.161 | 23.209 | 25.188 | 29.588 |
| 11 | 12.899 | 13.701 | 14.631 | 17.275 | 19.675 | 21.920 | 22.618 | 24.725 | 26.757 | 31.264 |
| 12 | 14.011 | 14.845 | 15.812 | 18.549 | 21.026 | 23.337 | 24.054 | 26.217 | 28.300 | 32.909 |
| 13 | 15.119 | 15.984 | 16.985 | 19.812 | 22.362 | 24.736 | 25.471 | 27.688 | 29.819 | 34.527 |
| 14 | 16.222 | 17.117 | 18.151 | 21.064 | 23.685 | 26.119 | 26.873 | 29.141 | 31.319 | 36.124 |
| 15 | 17.322 | 18.245 | 19.311 | 22.307 | 24.996 | 27.488 | 28.259 | 30.578 | 32.801 | 37.698 |
| 16 | 18.418 | 19.369 | 20.465 | 23.542 | 26.296 | 28.845 | 29.633 | 32.000 | 34.267 | 39.252 |
| 17 | 19.511 | 20.489 | 21.615 | 24.769 | 27.587 | 30.191 | 30.995 | 33.409 | 35.718 | 40.791 |
| 18 | 20.601 | 21.605 | 22.760 | 25.989 | 28.869 | 31.526 | 32.346 | 34.805 | 37.156 | 42.312 |
| 19 | 21.689 | 22.718 | 23.900 | 27.204 | 30.144 | 32.852 | 33.687 | $36.191$ | 38.582 | $43.819$ |
| 20 | 22.775 | 23.828 | 25.038 | 28.412 | 31.410 | 34.170 | 35.020 | 37.566 | 39.997 | 45.314 |
| 21 | 23.858 | 24.935 | 26.171 | 29.615 | 32.671 | 35.479 | 36.343 | 38.932 | 41.401 | 46.796 |
| 22 | 24.939 | 26.039 | 27.301 | 30.813 | 33.924 | 36.781 | 37.659 | 40.289 | 42.796 | 48.268 |
| 23 | 26.018 | 27.141 | 28.429 | 32.007 | 35.172 | 38.076 | 38.968 | 41.638 | 44.181 | 49.728 |
| 24 | 27.096 | 28.241 | 29.553 | 33.196 | 36.415 | 39.364 | 40.270 | 42.980 | 45.558 | 51.179 |
| 25 | 28.172 | 29.339 | 30.675 | 34.382 | 37.652 | 40.646 | 41.566 | 44.314 | 46.928 | 52.619 |
| 26 | 29.246 | 30.435 | 31.795 | 35.563 | 38.885 | 41.923 | 42.856 | 45.642 | 48.290 | 54.0 .51 |
| 27 | 30.319 | 31.528 | 32.912 | 36.741 | 40.113 | 43.195 | 44.140 | 46.963 | 49.645 | 55.475 |
| 28 | 31.391 | 32.620 | 34.027 | 37.916 | 41.337 | 44.461 | 45.419 | 48.278 | 50.994 | 56.892 |
| 29 | 32.461 | 33.711 | 35.139 | 39.087 | 42.557 | 45.722 | 46.693 | 49.588 | 52.335 | 58.301 |
| 30 | 33.530 | 34.800 | 36.250 | 40.256 | 43.773 | 46.979 | 47.962 | 50.892 | 53.672 | 59.702 |
| 40 | 44.165 | 45.616 | 47.269 | 51.805 | 55.758 | 59.342 | 60.436 | 63.691 | 66.766 | 73.403 |
| 50 | 54.723 | 56.334 | 58.164 | 63.167 | 67.505 | 71.420 | 72.613 | 76.154 | 79.490 | 86.660 |
| 60 | 65.226 | 66.981 | 68.972 | 74.397 | 79.082 | 83.298 | 84.58 | 88.379 | 91.952 | 99.608 |

## Example

For a chi-squared distribution, find
(a) $\chi_{0.025}^{2}$ when $\nu=15$;
(b) $\chi_{0.01}^{2}$ when $\nu=7$;
(c) $\chi_{0.05}^{2}$ when $\nu=24$.

Solution
(a) 27.488 .
(b) 18.475 .
(c) 36.415 .

## Example

For a chi-squared distribution $X$, find $\chi_{\alpha}^{2}$ such that
(a) $P\left(X>\chi_{\alpha}^{2}\right)=0.99$ when $\nu=4$;
(b) $P\left(X>\chi_{\alpha}^{2}\right)=0.025$ when $\nu=19$;
(c) $P\left(37.652<X<\chi_{\alpha}^{2}\right)=0.045$ when $\nu=25$.

Solution
(a) $\chi_{\alpha}^{2}=\chi_{0.99}^{2}=0.297$.
(b) $\chi_{\alpha}^{2}=\chi_{0.025}^{2}=32.852$.
(c) $\chi_{0.05}^{2}=37.652$. Therefore, $\alpha=0.05-0.045=0.005$. Hence, $\chi_{\alpha}^{2}=\chi_{0.005}^{2}=46.928$.

## Theorem 23

If $S^{2}$ is the variance of a random sample of size $n$ taken from a normal population having the variance $\sigma^{2}$, then the statistic

$$
\chi^{2}=\frac{(n-1) S^{2}}{\sigma^{2}}=\sum_{i=1}^{n} \frac{\left(X_{i}-\bar{X}\right)^{2}}{\sigma^{2}}
$$

has a chi-squared distribution with $\nu=n-1$ degrees of freedom.

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## Theorem

Let $Z$ be a standard normal random variable and $V$ a chi-squared random variable with $\nu$ degrees of freedom. If $Z$ and $\nu$ are independent, then the distribution of the random variable $T$, where

$$
T=\frac{Z}{\sqrt{V / \nu}}
$$

This is known as the $t$-distribution with $\nu$ degrees of freedom.

## Corollary

Let $X_{1}, X_{2}, \ldots, X_{n}$ be independent random variables that are all normal with mean $\mu$ and standard deviation $\sigma$. Let

$$
\bar{X}=\frac{1}{n} \sum_{i=1}^{n} X_{i} \quad \text { and } \quad S^{2}=\frac{1}{n-1} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}
$$

Then the random variable $T=\frac{\bar{X}-\mu}{S / \sqrt{n}}$ has a $t$-distribution with $\nu=n-1$ degrees of freedom.

| 0 | 0. 40 | 0.330 | 0.20 | 0.15 | 0.10 | 0.0 .5 | 10.1025 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 0.325 | 0.727 | 1.376 | 1.96 .3 | 3.078 | 6.314 | 12.706 |
| 2 | 0.289 | 0.617 | 1.061 | 1.386 | 1.886 | 2.920 | 4.303 |
| 3 | 0.277 | 0.584 | 0.978 | 1.250 | 1. 6.38 | 2.3553 | 3.182 |
| 1 | 0.271 | 0.569 | 0.941 | 1.190 | 1. 5.3 .3 | 2.132 | 2.776 |
| 5 | 0.267 | 0.559 | 0.920 | 1.156 | 1.476 | 2.015 | 2.571 |
| 6 | 0.265 | 0.553 | 0.906 | 1.1.34 | 1.440 | 1.943 | 2.447 |
| 7 | 0. 216.3 | 0.549 | 0. 8.896 | 1.119 | 1. 415 | 1. 895 | 2.3 .65 |
| 8 | 0.262 | 0.546 | 0.889 | 1. 108 | 1.397 | 1.860 | 2.306 |
| 9 | 0.261 | 0.543 | 0.883 | 1.100 | 1.383 | 1.833 | 2.262 |
| 10 | 0. 2650 | 0.542 | 0.879 | 1. 0993 | 1.372 | 1.812 | 2.228 |
| 11 | 0.250 | 0.540 | 0.876 | 1.088 | 1.3363 | 1.796 | 2.201 |
| 12 | 0.259 | 0.539 | 0.873 | 1.083 | 1.356 | 1.782 | 2.179 |
| 13 | 0.259 | 0.53 .8 | 0.870 | 1.079 | 1.350 | 1.771 | 2.160 |
| 14 | 0.258 | 0.537 | 0.8688 | 1.076 | 1.3.45 | 1.761 | 2.145 |
| 15 | 0.258 | 0.536 | 0.866 | 1.074 | 1.3 .41 | 1.753 | 2.131 |
| 16 | 0.258 | 0.535 | 0.865 | 1.071 | 1.337 | 1. 746 | 2.120 |
| 17 | 0.257 | 0.53 .4 | 0.863 | 1.069 | $1.333: 3$ | 1. 740 | 2.110 |
| 18 | 0.257 | 0.53 .1 | 0.86 .2 | 1.06:7 | 1.330 | 1.734 | 2.101 |
| 19 | 0.257 | $0.53: 3$ | 0.861 | 1.066 | 1.328 | 1.729 | 2.0993 |
| 210 | 0.257 | 0.5.3.3 | 0.860 | 1. 064 | 1.325 | 1.725 | 2.086 |
| 21 | 0.257 | 0.5.32 | 0.859 | 1.06.3 | 1.323 | 1.721 | 2.080 |
| 22 | 0.256 | 0.532 | 0.858 | 1.06.1 | 1.321 | 1.717 | 2.074 |
| 23 | 0.256 | 0.532 | 0.858 | 1.060 | 1.319 | 1.714 | 2.069 |
| 24 | 0.256 | 0.5.31 | 0.857 | 1.059 | 1.318 | 1.711 | 2.064 4 |
| 25 | 0. 25.5 | 0.531 | 0.856 | 1.058 | 1.316 | 1.708 | 2.060 |
| 26 | 0.256 | 0.531 | 0.856 | 1.058 | 1.315 | 1.706 | 2.056 |
| 27 | 0.256 | 0.531 | 0.855 | 1.057 | 1.314 | 1.703 | 2.052 |
| 28 | 0. 0.556 | 0.5.30 | 0.855 | 1.0.56 | 1.31 .3 | 1.701 | 2.048 |
| 29 | 0.256 | 0.530 | 0.854 | 1.0 .55 | 1.311 | 1.699 | 2.045 |
| 30 | 0.256 | 0.530 | 0.854 | 1.0.5 | 1.310 | 1.697 | 2.042 |
| 40 | 0.255 | 0.529 | 0.851 | 1.050 | 1.30:3 | 1.684 | 2.021 |
| 60 | 0. 2.5 .4 | 0.527 | 0.848 | 1.045 | 1. 2.296 | 1.671 | 2.000 |
| 120 | 0.254 | 0.526 | 0.845 | 1.041 | 1.289 | 1.658 | 1.980 |
| $\infty$ | 0.253 | 0.524 | 0.842 | 1.036 | 1.282 | 1.645 | 1.960 |

Figure: Table A. 4 Critical Values of the t-Distribution

| $v$ | a |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0.02 | 0.015 | 0.01 | 0.0075 | 0.005 | 0.0025 | 0.0005 |
| 1 | 15.894 | 21.205 | 31.821 | 42.433 | 63.656 | 127.321 | 636.578 |
| 2 | 4.849 | 5.643 | 6.965 | 8.073 | 9.925 | 14.089 | 31.600 |
| 3 | 3.482 | 3.896 | 4.541 | 5.047 | 5.841 | 7.453 | 12.924 |
| 4 | 2.999 | 3.298 | 3.747 | 4.088 | 4.604 | 5.598 | 8.610 |
| 5 | 2.757 | 3.003 | 3.365 | 3.6344 | 4.032 | 4.773 | 6.869 |
| 6 | 2.612 | 2.829 | 3.143 | 3.372 | 3.707 | 4.317 | 5.959 |
| 7 | 2.517 | 2.715 | 2.998 | 3.203 | 3.499 | 4.029 | 5.408 |
| 8 | 2.449 | 2.634 | 2.896 | 3.085 | 3.355 | 3.833 | 5.041 |
| 9 | 2.398 | 2.574 | 2.821 | 2.998 | 3.250 | 3.690 | 4.781 |
| 10 | 2.359 | 2.527 | 2.764 | 2.932 | 3.169 | 3.581 | 4.587 |
| 11 | 2.328 | 2.491 | 2.718 | 2.879 | 3.106 | 3.497 | 4.437 |
| 12 | 2.303 | 2.461 | 2.681 | 2.836 | 3.055 | 3.428 | 4.318 |
| 13 | 2.282 | 2.436 | 2.650 | 2.801 | 3.012 | 3.372 | 4.221 |
| 14 | 2.264 | 2.415 | 2.624 | 2.771 | 2.977 | 3.326 | 4.140 |
| 15 | 2.249 | 2.397 | 2.602 | 2.746 | 2.947 | 3.286 | 4.073 |
| 16 | 2.235 | 2.382 | 2.583 | 2.724 | 2.921 | 3.252 | 4.015 |
| 17 | 2.224 | 2.368 | 2.567 | 2.706 | 2.898 | 3.222 | 3.965 |
| 18 | 2.214 | 2.356 | 2.552 | 2.689 | 2.878 | 3.197 | 3.922 |
| 19 | 2.205 | 2.346 | 2.539 | 2.674 | 2.861 | 3.174 | 3.883 |
| 20 | 2.197 | 2.336 | 2.528 | 2.661 | 2.845 | 3.153 | 3.850 |
| 21 | 2.189 | 2.328 | 2.518 | 2.649 | 2.831 | 3.135 | 3.819 |
| 22 | 2.183 | 2.320 | 2.508 | 2.639 | 2.819 | 3.119 | 3.792 |
| 23 | 2.177 | 2.313 | 2.500 | 2.629 | 2.807 | 3.104 | 3.768 |
| 24 | 2.172 | 2.307 | 2.492 | 2.620 | 2.797 | 3.091 | 3.745 |
| 25 | 2.167 | 2.301 | 2.485 | 2.612 | 2.787 | 3.078 | 3.725 |
| 26 | 2.162 | 2.296 | 2.479 | 2.605 | 2.779 | 3.067 | 3.707 |
| 27 | 2.158 | 2.291 | 2.473 | 2.598 | 2.771 | 3.057 | 3.689 |
| 28 | 2.154 | 2.286 | 2.467 | 2.592 | 2.763 | 3.047 | 3.674 |
| 29 | 2.150 | 2.282 | 2.462 | 2.586 | 2.756 | 3.038 | 3.660 |
| 30 | 2.147 | 2.278 | 2.457 | 2.581 | 2.750 | 3.030 | 3.646 |
| 40 | 2.123 | 2.250 | 2.423 | 2.542 | 2.704 | 2.971 | 3.551 |
| 60 | 2.099 | 2.223 | 2.390 | 2.504 | 2.660 | 2.915 | 3. 460 |
| 120 | 2.076 | 2.196 | 2.358 | 2.468 | 2.617 | 2.860 | 3.373 |
| $\infty$ | 2.054 | 2.170 | 2.326 | 2.432 | 2.576 | 2.807 | 3.290 |

The $t$-value with $\nu=14$ degrees of freedom that leaves an area of 0.025 to the left, and therefore an area of 0.975 to the right, is

$$
t_{0.975}=-t_{0.025}=-2.145
$$

## Example

Find $\operatorname{Pr}\left(-t_{0.025}<T<t_{0.05}\right)$.

## Solution

Since $t_{0.05}$ leaves an area of 0.05 to the right, and $-t_{0.025}$ leaves an area of 0.025 to the left, we find a total area of $1-0.05-0.025=0.925$ between $-t_{0.025}$ and $t_{0.05}$.
Hence

$$
\operatorname{Pr}\left(-t_{0.025}<T<t_{0.05}\right)=0.925
$$

## Example

Find $k$ such that $\operatorname{Pr}(k<T<-1.761)=0.045$ for a random sample of size 15 selected from a normal distribution with $T=\frac{\bar{X}-\mu}{S / \sqrt{n}}$.

## Solution

From Table A. 4 we note that 1.761 corresponds to $t_{0.05}$ when $\nu=14$. Therefore, $-t_{0.05}=-1.761$. Since $k$ in the original probability statement is to the left of $-t_{0.05}=-1.761$, let $k=-t_{\alpha}$. Then, by using figure, we have

$$
0.045=0.05-\alpha, \text { or } \alpha=0.005
$$

Hence, from Table A. 4 with $\nu=14$, $k=-t_{0.005}=-2.977$ and $\operatorname{Pr}(-2.977<T<-1.761)=0.045$.

## Plan

(1) Random sampling and statistics

2 Sampling Distribution of Means and the Central Limit Theorem
(3) Sampling Distribution of the Difference between Two Means
a Sampling Distribution of the Variance
(5) The Student's Distribution
(6) The Fisher Distribution
(7) The Fisher with Two Sample Variances
(8) Sampling Distribution of Proportions and the Central Limit
(0) Sampling Distribution of the Difference between Two Proportions

The statistic $F$ is defined to be the ratio of two independent chi-squared random variables, each divided by its number of degrees of freedom.

## Theorem 31

The random variable

$$
F=\frac{U / \nu_{1}}{V / \nu_{2}}
$$

where $U$ and $V$ are independent random variables having chi-squared distributions with $\nu_{1}$ and $\nu_{2}$ degrees of freedom, respectively, is the $F$-distribution with $\nu_{1}$ and $\nu_{2}$ degrees of freedom (d.f.).

| Table | A-6 C | cal | of the $F$-Distribution |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $v_{2}$ | $f_{0.05}\left(w_{1}, v_{2}\right)$ |  |  |  |  |  |  |  |  |
|  | $v_{1}$ |  |  |  |  |  |  |  |  |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| 1 | 161.45 | 199.50 | 215.71 | 224.58 | 230.16 | 233.99 | 236.77 | 238.88 | 240.54 |
| 2 | 18.51 | 19.00 | 19.16 | 19.25 | 19.30 | 19.33 | 19.35 | 19.37 | 19.38 |
| 3 | 10.13 | 9.55 | 9.28 | 9.12 | 9.01 | 8.94 | 8.89 | 8.85 | 8.81 |
| 4 | 7.71 | 6.94 | 6.59 | 6.39 | 6.26 | 6.16 | 6.09 | 6.04 | 6.00 |
| 5 | 6.61 | 5.79 | 5.41 | 5.19 | 5.05 | 4.95 | 4.88 | 4.82 | 4.77 |
| 6 | 5.99 | 5.14 | 4.76 | 4.53 | 4.39 | 4.28 | 4.21 | 4.15 | 4.10 |
| 7 | 5.59 | 4.74 | 4.35 | 4.12 | 3.97 | 3.87 | 3.79 | 3.73 | 3.68 |
| 8 | 5.32 | 4.46 | 4.07 | 3.84 | 3.69 | 3.58 | 3.50 | 3.44 | 3.39 |
| 9 | 5.12 | 4.26 | 3.86 | 3.63 | 3.48 | 3.37 | 3.29 | 3.23 | 3.18 |
| 10 | 4.96 | 4.10 | 3.71 | 3.48 | 3.33 | 3.22 | 3.14 | 3.07 | 3.02 |
| 11 | 4.84 | 3.98 | 3.59 | 3.36 | 3.20 | 3.09 | 3.01 | 2.95 | 2.90 |
| 12 | 4.75 | 3.89 | 3.49 | 3.26 | 3.11 | 3.00 | 2.91 | 2.85 | 2.80 |
| 13 | 4.67 | 3.81 | 3.41 | 3.18 | 3.03 | 2.92 | 2.83 | 2.77 | 2.71 |
| 14 | 4.60 | 3.74 | 3.34 | 3.11 | 2.96 | 2.85 | 2.76 | 2.70 | 2.65 |
| 15 | 4.54 | 3.68 | 3.29 | 3.06 | 2.90 | 2.79 | 2.71 | 2.64 | 2.59 |
| 16 | 4.49 | 3.63 | 3.24 | 3.01 | 2.85 | 2.74 | 2.666 | 2.59 | 2.54 |
| 17 | 4.45 | 3.59 | 3.20 | 2.96 | 2.81 | 2.70 | 2.61 | 2.55 | 2.49 |
| 18 | 4.41 | 3.55 | 3.16 | 2.93 | 2.77 | 2.66 | 2.5 .8 | 2.51 | 2.46 |
| 19 | 4.38 | 3.52 | 3.13 | 2.90 | 2.74 | 2.63 | 2.54 | 2.48 | 2.42 |
| 20 | 4.35 | 3.49 | 3.10 | 2.87 | 2.71 | 2.60 | 2.51 | 2.45 | 2.39 |
| 21 | 4.32 | 3.47 | 3.07 | 2.84 | 2.68 | 2.57 | 2.49 | 2.42 | 2.37 |
| 22 | 4.30 | 3.44 | 3.05 | 2.82 | 2.66 | 2.55 | 2.46 | 2.40 | 2.34 |
| 23 | 4.28 | 3.42 | 3.03 | 2.80 | 2.64 | 2.53 | 2.44 | 2.37 | 2.32 |
| 24 | 4.26 | 3.40 | 3.01 | 2.78 | 2.62 | 2.51 | 2.42 | 2.36 | 2.30 |
| 25 | 4.24 | 3.39 | 2.99 | 2.76 | 2.60 | 2.49 | 2.40 | 2.34 | 2.28 |
| 26 | 4.23 | 3.37 | 2.98 | 2.74 | 2.59 | 2.47 | 2.39 | 2.32 | 2.27 |
| 27 | 4.21 | 3.35 | 2.96 | 2.73 | 2.57 | 2.46 | 2.37 | 2.31 | 2.25 |
| 28 | 4.20 | 3.34 | 2.95 | 2.71 | 2.56 | 2.45 | 2.36 | 2.29 | 2.24 |
| 29 | 4.18 | 3.33 | 2.93 | 2.70 | 2.55 | 2.43 | 2.35 | 2.28 | 2.22 |
| 30 | 4.17 | 3.32 | 2.92 | 2.69 | 2.53 | 2.42 | 2.33 | 2.27 | 2.21 |
| 40 | 4.08 | 3.23 | 2.84 | 2.61 | 2.45 | 2.34 | 2.25 | 2.18 | 2.12 |
| 60 | 4.00 | 3.15 | 2.76 | 2.53 | 2.37 | 2.25 | 2.17 | 2.10 | 2.04 |
| 120 | 3.92 | 3.07 | 2.68 | 2.45 | 2.29 | 2.18 | 2.09 | 2.02 | 1.96 |
| $\infty$ | 3.84 | 3.00 | 2.60 | 2.37 | 2.21 | 2.10 | 2.01 | 1.94 | 1.88 |

Figure : Table A. 6 Critical Values of the E-Distribution
$f_{\text {O. } 05}\left(v_{1}, v_{2}\right)$

| $v_{2}$ | $v_{1}$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 10 | 12 | 15 | 20 | 24 | 30 | 40 | 60 | 120 | $\infty$ |
| 1 | 241.88 | 243.91 | 245.95 | 248.01 | 249.05 | 250.10 | 251.14 | 252.20 | 253.25 | 254.31 |
| 2 | 19.40 | 19.41 | 19.43 | 19.45 | 19.45 | 19.46 | 19.47 | 19.48 | 19.49 | 19.50 |
| 3 | 8.79 | 8.74 | 8.70 | 8.66 | 8.64 | 8.62 | 8.59 | 8.57 | 8.55 | 8.53 |
| 4 | 5.96 | 5.91 | 5.86 | 5.80 | 5.77 | 5.75 | 5.72 | 5.69 | 5.66 | 5.63 |
| 5 | 4.74 | 4.68 | 4.62 | 4.56 | 4.53 | 4.50 | 4.46 | 4.43 | 4.40 | 4.36 |
| 6 | 4.06 | 4.00 | 3.94 | 3.87 | 3.84 | 3.81 | 3.77 | 3.74 | 3.70 | 3.67 |
| 7 | 3.64 | 3.57 | 3.51 | 3.44 | 3.41 | 3.38 | 3.34 | 3.30 | 3.27 | 3.23 |
| 8 | 3.35 | 3.28 | 3.22 | 3.15 | 3.12 | 3.08 | 3.04 | 3.01 | 2.97 | 2.93 |
| 9 | 3.14 | 3.07 | 3.01 | 2.94 | 2.90 | 2.86 | 2.83 | 2.79 | 2.75 | 2.71 |
| 10 | 2.98 | 2.91 | 2.85 | 2.77 | 2.74 | 2.70 | 2.66 | 2.62 | 2.58 | 2.54 |
| 11 | 2.85 | 2.79 | 2.72 | 2.65 | 2.61 | 2.57 | 2.53 | 2.49 | 2.45 | 2.40 |
| 12 | 2.75 | 2.69 | 2.62 | 2.54 | 2.51 | 2.47 | 2.43 | 2.38 | 2.34 | 2.30 |
| 13 | 2.67 | 2.60 | 2.53 | 2.46 | 2.42 | 2.38 | 2.34 | 2.30 | 2.25 | 2.21 |
| 14 | 2.60 | 2.53 | 2.46 | 2.39 | 2.35 | 2.31 | 2.27 | 2.22 | 2.18 | 2.13 |
| 15 | 2.54 | 2.48 | 2.40 | 2.33 | 2.29 | 2.25 | 2.20 | 2.16 | 2.11 | 2.07 |
| 16 | 2.49 | 2.42 | 2.35 | 2.28 | 2.24 | 2.19 | 2.15 | 2.11 | 2.06 | 2.01 |
| 17 | 2.45 | 2.38 | 2.31 | 2.23 | 2.19 | 2.15 | 2.10 | 2.06 | 2.01 | 1.96 |
| 18 | 2.41 | 2.34 | 2.27 | 2.19 | 2.15 | 2.11 | 2.06 | 2.02 | 1.97 | 1.92 |
| 19 | 2.38 | 2.31 | 2.23 | 2.16 | 2.11 | 2.07 | 2.03 | 1.98 | 1.93 | 1.88 |
| 20 | 2.35 | 2.28 | 2.20 | 2.12 | 2.08 | 2.04 | 1.99 | 1.95 | 1.90 | 1.84 |
| 21 | 2.32 | 2.25 | 2.18 | 2.10 | 2.05 | 2.01 | 1.96 | 1.92 | 1.87 | 1.81 |
| 22 | 2.30 | 2.23 | 2.15 | 2.07 | 2.03 | 1.98 | 1.94 | 1.89 | 1.84 | 1.78 |
| 23 | 2.27 | 2.20 | 2.13 | 2.05 | 2.01 | 1.96 | 1.91 | 1.86 | 1.81 | 1.76 |
| 24 | 2.25 | 2.18 | 2.11 | 2.03 | 1.98 | 1.94 | 1.89 | 1.84 | 1.79 | 1.73 |
| 25 | 2.24 | 2.16 | 2.09 | 2.01 | 1.96 | 1.92 | 1.87 | 1.82 | 1.77 | 1.71 |
| 26 | 2.22 | 2.15 | 2.07 | 1.99 | 1.95 | 1.90 | 1.85 | 1.80 | 1.75 | 1.69 |
| 27 | 2.20 | 2.13 | 2.06 | 1.97 | 1.93 | 1.88 | 1.84 | 1.79 | 1.73 | 1.67 |
| 28 | 2.19 | 2.12 | 2.04 | 1.96 | 1.91 | 1.87 | 1.82 | 1.77 | 1.71 | 1.65 |
| 29 | 2.18 | 2.10 | 2.03 | 1.94 | 1.90 | 1.85 | 1.81 | 1.75 | 1.70 | 1.64 |
| 30 | 2.16 | 2.09 | 2.01 | 1.93 | 1.89 | 1.84 | 1.79 | 1.74 | 1.68 | 1.62 |
| 40 | 2.08 | 2.00 | 1.92 | 1.84 | 1.79 | 1.74 | 1.69 | 1.64 | 1.58 | 1.51 |
| 60 | 1.99 | 1.92 | 1.84 | 1.75 | 1.70 | 1.65 | 1.59 | 1.53 | 1.47 | 1.39 |
| 120 | 1.91 | 1.83 | 1.75 | 1.66 | 1.61 | 1.55 | 1.50 | 1.43 | 1.35 | 1.25 |
| $\infty$ | 1.83 | 1.75 | 1.67 | 1.57 | 1.52 | 1.46 | 1.39 | 1.32 | 1.22 | 1.00 |

Table A. 6 (continued) Critical Values of the $F$-Distribution

| $v_{2}$ | $f_{0.01}\left(v_{1}, v_{2}\right)$ |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $v_{1}$ |  |  |  |  |  |  |  |  |
|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| 1 | 4052.18 | 4999.50 | 5403.35 | 5624.58 | 5763.65 | 5858.99 | 5928.36 | 5981.07 | 6022.47 |
| 2 | 98.50 | 99.00 | 99.17 | 99.25 | 99.30 | 99.33 | 99.36 | 99.37 | 99.39 |
| 3 | 34.12 | 30.82 | 29.46 | 28.71 | 28.24 | 27.91 | 27.67 | 27.49 | 27.35 |
| 4 | 21.20 | 18.00 | 16.69 | 15.98 | 15.52 | 15.21 | 14.98 | 14.80 | 14.66 |
| 5 | 16.26 | 13.27 | 12.06 | 11.39 | 10.97 | 10.67 | 10.46 | 10.29 | 10.16 |
| 6 | 13.75 | 10.92 | 9.78 | 9.15 | 8.75 | 8.47 | 8.26 | 8.10 | 7.98 |
| 7 | 12.25 | 9.55 | 8.45 | 7.85 | 7.46 | 7.19 | 6.99 | 6.84 | 6.72 |
| 8 | 11.26 | 8.65 | 7.59 | 7.01 | 6.63 | 6.37 | 6.18 | 6.03 | 5.91 |
| 9 | 10.56 | 8.02 | 6.99 | 6.42 | 6.06 | 5.80 | 5.61 | 5.47 | 5.35 |
| 10 | 10.04 | 7.56 | 6.55 | 5.99 | 5.64 | 5.39 | 5.20 | 5.06 | 4.94 |
| 11 | 9.65 | 7.21 | 6.22 | 5.67 | 5.32 | 5.07 | 4.89 | 4.74 | 4.63 |
| 12 | 9.33 | 6.93 | 5.95 | 5.41 | 5.06 | 4.82 | 4.64 | 4.50 | 4.39 |
| 13 | 9.07 | 6.70 | 5.74 | 5.21 | 4.86 | 4.62 | 4.44 | 4.30 | 4.19 |
| 14 | 8.86 | 6.51 | 5.56 | 5.04 | 4.69 | 4.46 | 4.28 | 4.14 | 4.03 |
| 15 | 8.68 | 6.36 | 5.42 | 4.89 | 4.56 | 4.32 | 4.14 | 4.00 | 3.89 |
| 16 | 8.53 | 6.23 | 5.29 | 4.77 | 4.44 | 4.20 | 4.03 | 3.89 | 3.78 |
| 17 | 8.40 | 6.11 | 5.18 | 4.67 | 4.34 | 4.10 | 3.93 | 3.79 | 3.68 |
| 18 | 8.29 | 6.01 | 5.09 | 4.58 | 4.25 | 4.01 | 3.84 | 3.71 | 3.60 |
| 19 | 8.18 | 5.93 | 5.01 | 4.50 | 4.17 | 3.94 | 3.77 | 3.63 | 3.52 |
| 20 | 8.10 | 5.85 | 4.94 | 4.43 | 4.10 | 3.87 | 3.70 | 3.56 | 3.46 |
| 21 | 8.02 | 5.78 | 4.87 | 4.37 | 4.04 | 3.81 | 3.64 | 3.51 | 3.40 |
| 22 | 7.95 | 5.72 | 4.82 | 4.31 | 3.99 | 3.76 | 3.59 | 3.45 | 3.35 |
| 23 | 7.88 | 5.66 | 4.76 | 4.26 | 3.94 | 3.71 | 3.54 | 3.41 | 3.30 |
| 24 | 7.82 | 5.61 | 4.72 | 4.22 | 3.90 | 3.67 | 3.50 | 3.36 | 3.26 |
| 25 | 7.77 | 5.57 | 4.68 | 4.18 | 3.85 | 3.63 | 3.46 | 3.32 | 3.22 |
| 26 | 7.72 | 5.53 | 4.64 | 4.14 | 3.82 | 3.59 | 3.42 | 3.29 | 3.18 |
| 27 | 7.68 | 5.49 | 4.60 | 4.11 | 3.78 | 3.56 | 3.39 | 3.26 | 3.15 |
| 28 | 7.64 | 5.45 | 4.57 | 4.07 | 3.75 | 3.53 | 3.36 | 3.23 | 3.12 |
| 29 | 7.60 | 5.42 | 4.54 | 4.04 | 3.73 | 3.50 | 3.33 | 3.20 | 3.09 |
| 30 | 7.56 | 5.39 | 4.51 | 4.02 | 3.70 | 3.47 | 3.30 | 3.17 | 3.07 |
| 40 | 7.31 | 5.18 | 4.31 | 3.83 | 3.51 | 3.29 | 3.12 | 2.99 | 2.89 |
| 60 | 7.08 | 4.98 | 4.13 | 3.65 | 3.34 | 3.12 | 2.95 | 2.82 | 2.72 |
| 120 | 6.85 | 4.79 | 3.95 | 3.48 | 3.17 | 2.96 | 2.79 | 2.66 | 2.56 |
| $\infty$ | 6.63 | 4.61 | 3.78 | 3.32 | 3.02 | 2.80 | 2.64 | 2.51 | 2.41 |


| $v_{2}$ | $f_{\mathrm{O.O1}}\left(v_{1}, v_{2}\right)$ |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $v_{1}$ |  |  |  |  |  |  |  |  |  |
|  | 10 | 12 | 15 | 20 | 24 | 30 | 40 | 60 | 120 | $\infty$ |
| 1 | 6055.85 | 6106.32 | 6157.28 | 6208.73 | 6234.63 | 6260.65 | 6286.78 | 6313.03 | 6339.39 | 6365.86 |
| 2 | 99.40 | 99.42 | 99.43 | 99.45 | 99.46 | 99.47 | 99.47 | 99.48 | $99.49$ | 99.50 |
| 3 | 27.23 | 27.05 | 26.87 | 26.69 | 26.60 | 26.50 | 26.41 | 26.32 | 26.22 | 26.13 |
| 4 | 14.55 | 14.37 | 14.20 | 14.02 | 13.93 | 13.84 | 13.75 | 13.65 | 13.56 | 13.46 |
| 5 | 10.05 | 9.89 | 9.72 | 9.55 | 9.47 | 9.38 | 9.29 | 9.20 | 9.11 | 9.02 |
| 6 | 7.87 | 7.72 | 7.56 | 7.40 | 7.31 | 7.23 | 7.14 | 7.06 | 6.97 | 6.88 |
| 7 | 6.62 | 6.47 | 6.31 | 6.16 | 6.07 | 5.99 | 5.91 | 5.82 | 5.74 | 5.65 |
| 8 | 5.81 | 5.67 | 5.52 | 5.36 | 5.28 | 5.20 | 5.12 | 5.03 | 4.95 | 4.86 |
| 9 | 5.26 | 5.11 | 4.96 | 4.81 | 4.73 | 4.65 | 4.57 | 4.48 | 4.40 | 4.31 |
| 10 | 4.85 | 4.71 | 4.56 | 4.41 | 4.33 | 4.25 | 4.17 | 4.08 | 4.00 | 3.91 |
| 11 | 4.54 | 4.40 | 4.25 | 4.10 | 4.02 | 3.94 | 3.86 | 3.78 | 3.69 | 3.60 |
| 12 | 4.30 | 4.16 | 4.01 | 3.86 | 3.78 | 3.70 | 3.62 | 3.54 | 3.45 | 3.36 |
| 13 | 4.10 | 3.96 | 3.82 | 3.66 | 3.59 | 3.51 | 3.43 | 3.34 | 3.25 | 3.17 |
| 14 | 3.94 | 3.80 | 3.66 | 3.51 | 3.43 | 3.35 | 3.27 | 3.18 | 3.09 | 3.00 |
| 15 | 3.80 | 3.67 | 3.52 | 3.37 | 3.29 | 3.21 | 3.13 | 3.05 | 2.96 | 2.87 |
| 16 | 3.69 | 3.55 | 3.41 | 3.26 | 3.18 | 3.10 | 3.02 | 2.93 | 2.84 | 2.75 |
| 17 | 3.59 | 3.46 | 3.31 | 3.16 | 3.08 | 3.00 | 2.92 | 2.83 | 2.75 | 2.65 |
| 18 | 3.51 | 3.37 | 3.23 | 3.08 | 3.00 | 2.92 | 2.84 | 2.75 | 2.66 | 2.57 |
| 19 | 3.43 | 3.30 | 3.15 | 3.00 | 2.92 | 2.84 | 2.76 | 2.67 | 2.58 | 2.49 |
| 20 | 3.37 | 3.23 | 3.09 | 2.94 | 2.86 | 2.78 | 2.69 | 2.61 | 2.52 | 2.42 |
| 21 | 3.31 | 3.17 | 3.03 | 2.88 | 2.80 | 2.72 | 2.64 | 2.55 | 2.46 | 2.36 |
| 22 | 3.26 | 3.12 | 2.98 | 2.83 | 2.75 | 2.67 | 2.58 | 2.50 | 2.40 | 2.31 |
| 23 | 3.21 | 3.07 | 2.93 | 2.78 | 2.70 | 2.62 | 2.54 | 2.45 | 2.35 | 2.26 |
| 24 | 3.17 | 3.03 | 2.89 | 2.74 | 2.66 | 2.58 | 2.49 | 2.40 | 2.31 | 2.21 |
| 25 | 3.13 | 2.99 | 2.85 | 2.70 | 2.62 | 2.54 | 2.45 | 2.36 | 2.27 | 2.17 |
| 26 | 3.09 | 2.96 | 2.81 | 2.66 | 2.58 | 2.50 | 2.42 | 2.33 | 2.23 | 2.13 |
| 27 | 3.06 | 2.93 | 2.78 | 2.63 | 2.55 | 2.47 | 2.38 | 2.29 | 2.20 | 2.10 |
| 28 | 3.03 | 2.90 | 2.75 | 2.60 | 2.52 | 2.44 | 2.35 | 2.26 | 2.17 | 2.06 |
| 29 | 3.00 | 2.87 | 2.73 | 2.57 | 2.49 | 2.41 | 2.33 | 2.23 | 2.14 | 2.03 |
| 30 | 2.98 | 2.84 | 2.70 | 2.55 | 2.47 | 2.39 | 2.30 | 2.21 | 2.11 | 2.01 |
| 40 | 2.80 | 2.66 | 2.52 | 2.37 | 2.29 | 2.20 | 2.11 | 2.02 | 1.92 | 1.80 |
| 60 | 2.63 | 2.50 | 2.35 | 2.20 | 2.12 | 2.03 | 1.94 | 1.84 | 1.73 | 1.60 |
| 120 | 2.47 | 2.34 | 2.19 | 2.03 | 1.95 | 1.86 | 1.76 | 1.66 | 1.53 | 1.38 |
| $\infty$ | 2.32 | 2.18 | 2.04 | 1.88 | 1.79 | 1.70 | 1.59 | 1.47 | 1.32 | 1.00 |

## Theorem

Writing $f_{\alpha}\left(\nu_{1}, \nu_{2}\right)$ for $f_{\alpha}$ with $\nu_{1}$ and $\nu_{2}$ degrees of freedom, we have

$$
f_{1-\alpha}\left(\nu_{1}, \nu_{2}\right)=\frac{1}{f_{\alpha}\left(\nu_{2}, \nu_{1}\right)}
$$

Thus, the $f$-value with 6 and 10 degrees of freedom, leaving an area of 0.95 to the right, is $f_{0.95}(6,10)=\frac{1}{f_{0.05}(10,6)}=\frac{1}{4.06}=0.246$.

## Plan

(1) Random sampling and statistics

2 Sampling Distribution of Means and the Central Limit Theorem
(3) Sampling Distribution of the Difference between Two Means
a Sampling Distribution of the Variance
(5) The Student's Distribution
(8) The Fisher Distribution
(7) The Fisher with Two Sample Variances
(8) Sampling Distribution of Proportions and the Central Limit
(9) Sampling Distribution of the Difference between Two Proportions

Suppose that random samples of size $n_{1}$ and $n_{2}$ are selected from two normal populations with variances $\sigma_{1}^{2}$ and $\sigma_{2}^{2}$, respectively. From Theorem 23, we know that

$$
\chi_{1}^{2}=\frac{\left(n_{1}-1\right) S_{1}^{2}}{\sigma_{1}^{2}} \text { and } \chi_{2}^{2}=\frac{\left(n_{2}-1\right) S_{2}^{2}}{\sigma_{2}^{2}}
$$

are random variables having chi-squared distributions with $\nu_{1}=n_{1}-1$ and $\nu_{2}=n_{2}-1$ degrees of freedom. Furthermore, since the samples are selected at random, we are dealing with independent random variables. Then, using Theorem 31 with $\chi_{1}^{2}=U$ and $\chi_{2}^{2}=V$, we obtain the following result.

## Theorem

If $S_{1}^{2}$ and $S_{2}^{2}$ are the variances of independent random samples of size $n_{1}$ and $n_{2}$ taken from normal populations with variances $\sigma_{1}^{2}$ and $\sigma_{2}^{2}$, respectively, then

$$
F=\frac{S_{1}^{2} / \sigma_{1}^{2}}{S_{2}^{2} / \sigma_{2}^{2}}
$$

has an F-distribution with $\nu_{1}=n_{1}-1$ and $\nu_{2}=n_{2}-1$ degrees of freedom.

## Example

For an $F$-distribution, find
(a) $f_{0.05}$ with $\nu_{1}=7$ and $\nu_{2}=15$;
(b) $f_{0.05}$ with $\nu_{1}=15$ and $\nu_{2}=7$ :
(c) $f_{0.01}$ with $\nu_{1}=24$ and $\nu_{2}=19$;
(d) $f_{0.95}$ with $\nu_{1}=19$ and $\nu_{2}=24$;
(e) $f_{0.99}$ with $\nu_{1}=28$ and $\nu_{2}=12$.

Solution
(a) 2.71 .
(d) $1 / 2.11=0.47$.
(b) 3.51 .
(c) 2.92 .
(e) $1 / 2.90=0.34$.

## Plan

(1) Random sampling and statistics
(2) Sampling Distribution of Means and the Central Limit Theorem
(3) Sampling Distribution of the Difference between Two Means
a Sampling Distribution of the Variance
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(7) The Fisher with Two Sample Variances
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(9) Sampling Distribution of the Difference between Two Proportions

In many situations the use of the sample proportion is easier and more reliable because, unlike the mean, the proportion does not depend on the population variance, which is usually an unknown quantity. We will represent the sample proportion by $\widehat{P}$ and the population proportion by $p$. Construction of the sampling distribution of the sample proportion is done in a manner similar to that of the mean. One has $\widehat{P}=X / n$ where $X$ is a number of success for a sample of size $n$. It is clear that $X$ is a binomial distribution $B(n, p)$. Its mean $\mu_{X}=n p$ and its variance $\sigma_{X}^{2}=n p(1-p)$.

## Theorem

The mean $\mu_{\widehat{p}}$ of the sample distribution $\widehat{P}$ is equal to the true population proportion $p$, and its variance $\sigma_{\widehat{p}}^{2}$ is equal to $p(1-p) / n$.

## Theorem

If $n p \geq 5$ and $n(1-p) \geq 5$, then the random variable $\widehat{P}$ is approximation a normal distribution with mean $\mu_{\widehat{p}}=p$ and standard deviation (or standard error) $\sigma_{\widehat{p}}=\sqrt{p(1-p) / n}$. Hence

$$
Z=\frac{\widehat{P}-p}{\sqrt{p(1-p) / n}}
$$

is approximately a standard normal distribution.

## Example

In the mid seventies, according to a report by the National Center for Health Statistics, 19.4 percent of the adult U.S. male population was obese. What is the probability that in a simple random sample of size 150 from this population fewer than 15 percent will be obese?

## Solution

Here $n=150, p=0.194$. Since $n p \geq 5$ and $n(1-p) \geq 5$, hence

$$
Z=\frac{\widehat{P}-0.194}{\sqrt{0.194(1-0.194) / 150}}=\frac{\widehat{P}-0.194}{0.032}
$$

is approximately a standard normal distribution.
$\operatorname{Pr}(\widehat{P} \leq 0.15)=\operatorname{Pr}\left(\frac{\widehat{P}-0.194}{0.03} \leq \frac{0.15-0.194}{0.03}\right) \simeq \operatorname{Pr}(Z \leq-1.37)=$ 0.0853.

## Plan

(1) Random sampling and statistics
(2) Sampling Distribution of Means and the Central Limit Theorem
(3) Sampling Distribution of the Difference between Two Means
a Sampling Distribution of the Variance
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## Theorem

The mean $\mu_{\widehat{p}_{1}-\widehat{p}_{2}}$ of the sample distribution of the difference between two sample proportions $\widehat{P}_{1}-\widehat{P}_{2}$ is equal to the difference $p_{1}-p_{2}$ between the true population proportions, and its variance $\sigma_{\hat{p}_{1}-\hat{p}_{2}}^{2}$ will be equal to $p_{1}\left(1-p_{1}\right) / n_{1}+p_{1}\left(1-p_{2}\right) / n_{2}$.

## Theorem

If $n_{1} p_{1} \geq 5, n_{1}\left(1-p_{1}\right) \geq 5, n_{2} p_{2} \geq 5, n_{2}\left(1-p_{2}\right)$, then the random variable $\widehat{P}_{1}-\widehat{P}_{2}$ is approximation a normal distribution with mean $\mu_{\widehat{p}_{1}-\widehat{p}_{2}}=p_{1}-p_{2}$ and standard deviation (or standard error) $\sigma_{\widehat{p}_{1}-\widehat{p_{2}}}=\sqrt{p_{1}\left(1-p_{1}\right) / n_{1}+p_{1}\left(1-p_{2}\right) / n_{2}}$. Hence

$$
Z=\frac{\left(\widehat{P}_{1}-\widehat{P}_{2}\right)-\left(p_{1}-p_{2}\right)}{\sqrt{\frac{p_{1}\left(1-p_{1}\right)}{n_{1}}+\frac{p_{2}\left(1-p_{2}\right)}{n_{2}}}}
$$

is approximately a standard normal distribution.

## Example

Suppose that there are two large high schools, each with more than 2000 students, in a certain town. At School 1, 70\% of students did their homework last night. Only 50\% of the students at School 2 did their homework last night. The counselor at School 1 takes a sample random sample of 100 students and records the proportion that did homework. School 2's counselor takes a sample random sample of 200 students and records the proportion that did homework. Find the probability of getting a difference in sample proportion $\widehat{P}_{1}-\widehat{P}_{2}$ of 0.10 or less from the two surveys.

Solution
Here $p_{1}=0.7, p_{2}=0.5, n_{1}=100$ and $n_{2}=200$. It is clear that $n_{1} p_{1} \geq 5, n_{1}\left(1-p_{1}\right) \geq 5, n_{2} p_{2} \geq 5, n_{2}\left(1-p_{2}\right)$. Also

$$
\mu_{\widehat{p}_{1}-\widehat{p_{2}}}=p_{1}-p_{2}=0.2
$$

and

$$
\sigma_{\widehat{p}_{1}-\widehat{p}_{2}}=\sqrt{p_{1}\left(1-p_{1}\right) / n_{1}+p_{1}\left(1-p_{2}\right) / n_{2}}=0.058
$$

Hence,

$$
Z=\frac{\widehat{P}_{1}-\widehat{P}_{2}-p_{1}-p_{2}}{\sqrt{\frac{p_{1}\left(1-p_{1}\right)}{n_{1}}+\frac{p_{2}\left(1-p_{2}\right)}{n_{2}}}}=\frac{\widehat{P}_{1}-\widehat{P}_{2}-0.2}{0.058}
$$

is approximately a standard normal.

$$
\begin{aligned}
\operatorname{Pr}\left(\widehat{P}_{1}-\widehat{P}_{2} \leq 0.10\right) & =\operatorname{Pr}\left(\frac{\widehat{P}_{1}-\widehat{P}_{2}-0.2}{0.058} \leq \frac{0.10-0.2}{0.058}\right) \\
& \simeq \operatorname{Pr}(Z \leq-1.72)=0.0427
\end{aligned}
$$

