

chapter 3

Ex. 3.1:

(a)

$$\begin{aligned}
 y &= \text{weight (Response)} \\
 x_1 &= \text{age} \\
 x_2 &= \text{sex} \\
 x_3 &= \text{height} \\
 x_4 &= \text{mean daily food intake} \\
 x_5 &= \text{mean daily energy expenditure} \\
 y_i &\sim N(\mu_i, \sigma^2) \\
 E(y_i) &= \mu_i
 \end{aligned}
 \quad \left\| \begin{array}{l} \text{explanatory variables} \end{array} \right.$$

Model:

$$\begin{aligned}
 \mu_i &= \beta_0 + \beta_1 x_{i1} + \dots + \beta_5 x_{i5} \\
 g(\mu_i) &= \mu_i \quad (\text{link function} = \text{identity function}) \\
 \eta(\underline{x}_i) &= \sum_{j=0}^5 \beta_j x_{ij} \quad (x_0 \equiv 1) \quad (\text{linear component})
 \end{aligned}$$

(b)

$$\begin{aligned}
 y &= \text{number of mice infected (in each level)} \\
 x &= \text{exposure level of bacteria, with 5 levels: } x_1, x_2, \dots, x_5 \\
 n &= 20 \text{ mice for each level.}
 \end{aligned}$$

$y_i \sim \text{Bino}(n, \pi_i)$, $n = 20$, $i = 1, 2, \dots, 5$ because 'infection' is a binary outcome (but the plausibility of the assumption of independence of infection for mice depends on the experimental conditions).

$$\Rightarrow E(y_i) = \mu_i = n\pi_i$$

Let $P_i = \frac{y_i}{n}$: proportion of infected mice in the i -th level

$$E(P_i) = \pi_i$$

Model:

$$\begin{aligned}
 g(\pi_i) &= \beta_0 + \beta_1 x_i \quad \text{for some link function such as } g(\pi) = \ln\left(\frac{\pi}{1-\pi}\right) \\
 \eta(\underline{x}_i) &= \beta_0 + \beta_1 x_i \leftarrow \text{Linear Predictor}
 \end{aligned}$$

or $g(\pi_i) = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} + \beta_4 x_{i4} + \beta_5 x_{i5}$ with the β_k 's subject to a corner point or sum-to-zero constraint.

(c)

$$\begin{aligned}
 y &= \text{number of trips per week (Response)} \\
 x_1 &= x_{i1} = \text{number of people in the household.} \\
 x_2 &= x_{i2} = \text{household income} \\
 x_3 &= x_{i3} = \text{distance to supermarket} \\
 y_i &\sim P_0(\mu_i) \cong \text{Poisson}(\lambda_i) \\
 \Rightarrow E(y_i) &= \mu_i
 \end{aligned}
 \quad \left\| \begin{array}{l} \text{explanatory variables} \end{array} \right.$$

model:

$$\begin{aligned}
 g(\mu_i) &= \log \lambda_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} \\
 &= \eta(\underline{x}_i)
 \end{aligned}$$

for some link function such as $g(\mu_i) = \ln(\mu_i)$.

$$\eta(\underline{x}_i) = \sum_{j=0}^3 \beta_j x_{ij} \quad (x_{i0} = 1) \leftarrow \text{Linear Predictor}$$

Ex 3.2:

$y \sim \text{Gamma}(\alpha, \beta)$ ($y > 0, \alpha > 0, \beta > 0$)

- α is assumed to be known:

$$\begin{aligned}
 f(y; \beta) &= \frac{\beta^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-y\beta} \\
 &= \beta^\alpha [\Gamma(\alpha)]^{-1} y^{\alpha-1} e^{-y\beta} \\
 &= e^{\ln(\beta^\alpha)} e^{\ln[\Gamma(\alpha)]^{-1}} e^{\ln(y^{\alpha-1})} e^{-y\beta} \\
 &= e^{\alpha \ln(\beta)} e^{-\ln(\Gamma(\alpha))} e^{(\alpha-1)\ln(y)} e^{-y\beta} \\
 &= e^{-y\beta + \alpha \ln(\beta) - \ln(\Gamma(\alpha)) + (\alpha-1)\ln(y)} \\
 &= e^{y(-\beta) + \alpha \ln(\beta) - \ln[\Gamma(\alpha)] + (\alpha-1)\ln(y)}
 \end{aligned}$$

This is the canonical form of EF:

$$\begin{aligned}
 a(y) &= y \\
 b(\beta) &= -\beta \text{ (natural parameter)} \\
 c(\beta) &= \alpha \ln(\beta) \\
 d(y) &= -\ln[\Gamma(\alpha)] + (\alpha-1)\ln(y)
 \end{aligned}$$

- Natural parameter is $b(\beta) = -\beta$

$$\begin{aligned}
 E(y) &= -\frac{c'(\beta)}{b'(\beta)} = -\frac{(\alpha/\beta)}{(-1)} = \frac{\alpha}{\beta} \\
 \text{Var}(y) &= \frac{b''(\beta)c'(\beta) - c''(\beta)b'(\beta)}{[b'(\beta)]^3} = \frac{0 - (-\alpha/\beta^2)(-1)}{(-1)^3} = \frac{\alpha}{\beta^2}
 \end{aligned}$$

where:

$$\begin{aligned}
 b'(\beta) &= \frac{d}{d\beta} b(\beta) = \frac{d}{d\beta} (-\beta) = -1 \\
 c'(\beta) &= \frac{d}{d\beta} c(\beta) = \frac{d}{d\beta} (\alpha \ln(\beta)) = \frac{\alpha}{\beta} \\
 b''(\beta) &= 0 \\
 c''(\beta) &= -\alpha/\beta^2
 \end{aligned}$$

Ex 3.3:

(a)

$$f(y; \theta) = \theta y^{-\theta-1} = e^{\ln(\theta)} e^{\ln[y^{-\theta-1}]} = e^{\ln(\theta)} e^{(-\theta-1)\ln(y)} = e^{\ln(y)(-\theta-1) + \ln(\theta)}$$

$$\begin{aligned}
 \Rightarrow a(y) &= \ln(y) \\
 b(\theta) &= -\theta - 1 \\
 c(\theta) &= \ln(\theta) \\
 d(y) &= 0
 \end{aligned}$$

(b)

$$\begin{aligned}
 f(y; \theta) &= \theta e^{-y\theta} = e^{\ln(\theta)} e^{-y\theta} = e^{y(-\theta) + \ln(\theta)} \\
 \Rightarrow a(y) &= y \\
 b(\theta) &= -\theta \\
 c(\theta) &= \ln(\theta) \\
 d(y) &= 0
 \end{aligned}$$

(c)

$$\begin{aligned}
 f(y; \theta) &= \binom{y+r-1}{r-1} \theta^r (1-\theta); \text{ } r \text{ is known} \\
 &= e^{\ln \binom{y+r-1}{r-1}} e^{\ln(\theta^r)} e^{\ln[(1-\theta)^y]} = e^{\ln \binom{y+r-1}{r-1}} e^{r \ln(\theta)} e^{y \ln(1-\theta)} = e^{y \ln(1-\theta) + r \ln(\theta) + \ln \binom{y+r-1}{r-1}}
 \end{aligned}$$

$$\begin{aligned}
&\Rightarrow a(y) = y \\
&b(\theta) = \ln(1 - \theta) \\
&c(\theta) = r \ln(\theta) \\
&d(y) = \ln \binom{y+r-1}{r-1}
\end{aligned}$$

Ex 3.4:

For the distribution belonging to EF , we have:

$$\begin{aligned}
E[a(y)] &= -\frac{c'(\theta)}{b'(\theta)} \\
\text{Var}[a(y)] &= \frac{b''(\theta)c'(\theta) - c''(\theta)b'(\theta)}{[b'(\theta)]^3}
\end{aligned}$$

(a) For $y \sim \text{Po}(\theta)$:

$$\begin{aligned}
&\Rightarrow a(y) = y \\
&b(\theta) = \ln(\theta) \\
&c(\theta) = -\theta \\
&b(y) = -\ln(y!) \\
&\Rightarrow b'(\theta) = \frac{1}{\theta} \\
&b''(\theta) = -\frac{1}{\theta^2} \\
&c'(\theta) = -1 \\
&c''(\theta) = 0 \\
&\Rightarrow E(y) = \frac{-(-1)}{\left(\frac{1}{\theta}\right)} = \theta \\
&\text{Var}(y) = \frac{\left(-\frac{1}{\theta^2}\right)(-1) - 0}{\left(\frac{1}{\theta}\right)^3} = \frac{1/\theta^2}{1/\theta^3} = \frac{1}{\left(\frac{1}{\theta}\right)} = \theta
\end{aligned}$$

(b) For $y \sim N(\mu, \sigma^2)$: it is assumed that σ^2 is fixed,

$$\begin{aligned}
&\Rightarrow a(y) = y \\
&b(\mu) = \frac{\mu}{\sigma^2} \\
&c(\mu) = -\frac{\mu^2}{2\sigma^2} \\
&d(y) = -\frac{y^2}{2\sigma^2} - \frac{1}{2} \ln(2\pi\sigma^2) \\
&\Rightarrow \begin{cases} b'(\mu) = \frac{1}{\sigma^2} \\ b''(\mu) = 0 \\ c'(\mu) = -\frac{\mu}{\sigma^2} \\ c''(\mu) = -\frac{1}{\sigma^2} \end{cases} \\
&\Rightarrow E(y) = \frac{-(-\mu/\sigma^2)}{(1/\sigma^2)} = \mu \\
&\text{Var}(y) = \frac{0 - (-1/\sigma^2)(1/\sigma^2)}{(1/\sigma^2)^3} = \frac{1/\sigma^4}{1/\sigma^6} = \sigma^2
\end{aligned}$$

(c) For $y \sim \text{Bino}(n, \pi)$:

$$\begin{aligned}
&\Rightarrow a(y) = y \\
&b(\pi) = \ln \left(\frac{\pi}{1-\pi} \right) \\
&c(\pi) = n \ln(1-\pi) \\
&d(y) = \ln \binom{n}{y} \\
&\Rightarrow E(y) = - \frac{\left(-\frac{n}{1-\pi} \right)}{\frac{1}{\pi(1-\pi)}} = n\pi \\
\text{Var}(y) &= \frac{\frac{2\pi-1}{\pi^2(1-\pi)^2} \left(\frac{-n}{1-\pi} \right) - \left(-\frac{n}{(1-\pi)^2} \right) \left(\frac{1}{\pi(1-\pi)} \right)}{\left[\frac{1}{\pi(1-\pi)} \right]^3} = \frac{\frac{n(1-2\pi)}{\pi^2(1-\pi)^3} + \frac{n}{\pi(1-\pi)^3}}{\frac{1}{\pi^3(1-\pi)^3}} \\
&= \frac{\frac{n(1-2\pi) + n\pi}{\pi^2(1-\pi)^3}}{\frac{1}{\pi^3(1-\pi)^3}} = [n(1-2\pi) + n\pi]\pi = (n - 2n\pi + n\pi)\pi = (n - n\pi)\pi = n(1-\pi)\pi \\
&= n\pi(1-\pi)
\end{aligned}$$

where

$$\begin{aligned}
b'(\pi) &= \frac{1-\pi}{\pi} \left[\frac{(1-\pi) - \pi(-1)}{(1-\pi)^2} \right] = \frac{1-\pi}{\pi} \frac{1}{(1-\pi)^2} = \frac{1}{\pi(1-\pi)} \\
b''(\pi) &= \frac{0 - [(1-\pi) + \pi(-1)]}{[\pi(1-\pi)]^2} = \frac{2\pi-1}{\pi^2(1-\pi)^2} \\
c'(\pi) &= \frac{n}{1-\pi} (-1) = \frac{-n}{1-\pi} \\
c''(\pi) &= -n \left[\frac{-(-1)}{(1-\pi)^2} \right] = -\frac{n}{(1-\pi)^2}
\end{aligned}$$

Ex 3.5:

(a) For $y \sim NBin(r, \theta)$ (see Ex 3.3 part(c)),

$$\begin{aligned}
&\Rightarrow a(y) = y \\
&b(\theta) = \ln(1-\theta) \\
&c(\theta) = r \ln(\theta) \\
&d(y) = \ln \binom{y+r-1}{r-1} \\
&\Rightarrow b'(\theta) = \frac{-1}{1-\theta} \\
b''(\theta) &= -\frac{(1-\theta)(0) - (-1)}{(1-\theta)^2} = -\frac{1}{(1-\theta)^2} \\
c'(\theta) &= \frac{r}{\theta} \\
c''(\theta) &= -\frac{r}{\theta^2} \\
\Rightarrow E(y) &= -\frac{\left(\frac{r}{\theta} \right)}{\left(-\frac{1}{1-\theta} \right)} = \frac{r(1-\theta)}{\theta}
\end{aligned}$$

$$\begin{aligned}\text{Var}(y) &= \frac{-\frac{1}{(1-\theta)^2} \left(\frac{r}{\theta}\right) - \left(-\frac{r}{\theta^2}\right) \left(-\frac{1}{1-\theta}\right)}{\left(-\frac{1}{1-\theta}\right)^3} = \frac{-\frac{r}{\theta(1-\theta)^2} - \frac{r}{\theta^2(1-\theta)}}{-\frac{1}{(1-\theta)^3}} = \frac{-\left[\frac{\theta r + (1-\theta)r}{\theta^2(1-\theta)^2}\right]}{-\frac{1}{(1-\theta)^3}} \\ &= \frac{\frac{r}{\theta^2(1-\theta)^2}}{\frac{1}{(1-\theta)^3}} = \frac{\frac{r}{\theta^2}}{\frac{1}{(1-\theta)}} = \frac{r(1-\theta)}{\theta^2}\end{aligned}$$

Ex. 3.6:

Refare to Example 3.5.3 (Mortality Rote) Model:

$$\begin{aligned}E(\mu_i) &= \mu_i = n_i e^{\theta i} \\ \Leftrightarrow g(\mu_i) &= \ln(\mu_i) = \ln(n_i) + \theta i \\ \Leftrightarrow \ln(\mu_i) - \ln(n_i) &= \theta i \\ \Leftrightarrow \ln\left(\frac{\mu_i}{n_i}\right) &= \theta i \\ \Leftrightarrow \mu_i^* &= \theta i \text{ whene } \mu_i^* = \ln\left(\frac{\mu_i}{n_i}\right)\end{aligned}$$

age x_i	y_i	n_i	$\frac{y_i}{n_i}$	$\ln\left(\frac{y_i}{n_i}\right)$
30	1	17742	$5.64 * 10^{-5}$	-9.7837
35	5	16554	$3.02 * 10^{-4}$	-8.1049
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$
65	65	9933	6.54×10^{-3}	-5.0292

Let $y_i^* = \ln\left(\frac{y_i}{n_i}\right)$ and $x_i^* = \ln(x_i)$

Regnen y_i^* on i using the model

$$y_i^* = \beta_0 + \beta_1 x_i^* \text{ (Simple linear Regression)}$$

$$\hat{\beta}_0 = -13.182 \text{ and } \hat{\beta}_1 = 0.1331 \text{ [by using computer]}$$

Fitted value for y_i^* : $\hat{y}_i^* = \hat{\beta}_0 + \hat{\beta}_1 x_i^* = -13.182 + 0.1331 x_i^*$

Fitted values for y_i : $\hat{y}_i = n_i e^{\hat{y}_i^*}$

$$R^2 = 0.9685 \leftarrow (\text{good fit}) \text{ [see computer output]}$$

Ex. 3.7:

- $y_i \sim \text{Bernoulli}(\pi_i)$
- $y_1, y_2, \dots, y_N$ are midependent.

$$f(y_i, \pi_i) = \pi_i^{y_i} (1 - \pi_i)^{1-y_i}; y_i = 0, 1$$

(a)

$$\begin{aligned}f(y, \pi) &= \pi^y (1 - \pi)^{1-y} = e^{\ln(\pi^y)} e^{\ln[(1-\pi)^{1-y}]} = e^{y \ln(\pi)} e^{(1-y) \ln(1-\pi)} = e^{y \ln(\pi) + (1-y) \ln(1-\pi)} \\ &= e^{y \ln(\pi) + \ln(1-\pi) - y \ln(1-\pi)} = e^{y [\ln(\pi) - \ln(1-\pi)] + \ln(1-\pi)} = e^{y \ln\left(\frac{\pi}{1-\pi}\right) + \ln(1-\pi)} \in EF \\ &\Rightarrow a(y) = y\end{aligned}$$

$$\begin{aligned}b(\pi) &= \ln\left(\frac{\pi}{1-\pi}\right), c(\pi) = \ln(1-\pi) \\ d(y) &= 0\end{aligned}$$

(b) Natural parameter is

$$b(\pi) = \ln\left(\frac{\pi}{1-\pi}\right)$$

(c)

$$E(y) = -\frac{c'(\pi)}{b'(\pi)} = -\frac{\left(-\frac{1}{1-\pi}\right)}{\left(\frac{1}{\pi(1-\pi)}\right)} = \pi$$

where (see Ex. 3.4 Part (c))

$$c'(\pi) = -\frac{1}{1-\pi}$$

$$b'(\pi) = \frac{1}{\pi(1-\pi)}$$

Note: $\frac{\pi}{1-\pi}$ is called odds (معامل الترجيح)

(d) Link function $g(\pi) = \ln\left(\frac{\pi}{1-\pi}\right) = \mathbf{x}'\boldsymbol{\beta}$.

$$\ln\left(\frac{\pi}{1-\pi}\right) = \mathbf{x}'\boldsymbol{\beta}$$

$$\Leftrightarrow e^{\ln\left(\frac{\pi}{1-\pi}\right)} = e^{\mathbf{x}'\boldsymbol{\beta}}, (e^{\ln x} = x)$$

$$\Leftrightarrow \frac{\pi}{1-\pi} = e^{\mathbf{x}'\boldsymbol{\beta}}$$

$$\Leftrightarrow \pi = (1-\pi)e^{\mathbf{x}'\boldsymbol{\beta}} = e^{\mathbf{x}'\boldsymbol{\beta}} - \pi e^{\mathbf{x}'\boldsymbol{\beta}}$$

$$\Leftrightarrow \pi + \pi e^{\mathbf{x}'\boldsymbol{\beta}} = e^{\mathbf{x}'\boldsymbol{\beta}}$$

$$\Leftrightarrow \pi(1 + e^{\mathbf{x}'\boldsymbol{\beta}}) = e^{\mathbf{x}'\boldsymbol{\beta}}$$

$$\Leftrightarrow \pi = \frac{e^{\mathbf{x}'\boldsymbol{\beta}}}{1 + e^{\mathbf{x}'\boldsymbol{\beta}}} \quad (\text{logistic Regression model})$$

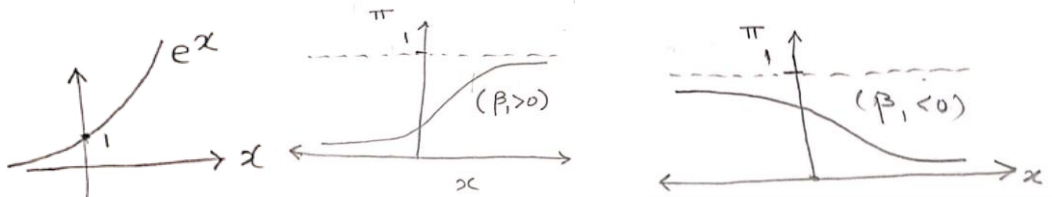
(e) For $\mathbf{x}'\boldsymbol{\beta} = \beta_1 + \beta_2 x$, we have

$$\pi = \frac{e^{\beta_1 + \beta_2 x}}{1 + e^{\beta_1 + \beta_2 x}} \quad (\text{logistic function})$$

note: $e^{\beta_1 + \beta_2 x} > 0$

$$0 < \frac{e^{\beta_1 + \beta_2 x}}{1 + e^{\beta_1 + \beta_2 x}} < 1$$

$$0 < \pi < 1$$



Ex. 3.8:

WLOG, let $\phi = 1$ (ϕ is nuisance parameter) and it is considered as fixed

$$f(y; \theta) = e^{(y-\theta)-e^{(y-\theta)}} = e^{y-\theta-e^{(y-\theta)}} = e^{y-\theta-e^y e^{-\theta}} = e^{e^y(-e^{-\theta})-\theta+y} = e^{a(y)b(\theta)+c(\theta)+d(y)}$$

$$a(y) = e^y; b(\theta) = -e^{-\theta}, c(\theta) = -\theta, d(y) = y$$

Ex. 3.9:

$y_1, y_2, \dots, y_n$ are independent.

$$y_i \sim \text{Pareto}(\theta) \therefore f(y; \theta) = \theta y^{-\theta-1} \in EF$$

$$\Rightarrow a(y) = \ln(y), b(\theta) = -\theta - 1, c(\theta) = \ln(\theta), d(y) = 0 \quad (\text{See Ex 3.3(a)}).$$

$$\begin{aligned}\Rightarrow E(y_i) &= \mu_i = (\beta_0 + \beta_1 x_i)^2 \\ \Leftrightarrow \sqrt{\mu_i} &= \beta_0 + \beta_1 x_i \\ g(\mu_i) &= \sqrt{\mu_i} \text{ link function}\end{aligned}$$

For Pareto, $a(y) = \ln(y) \neq y$

one of the conditions of Generalized linear model is that $a(y)$ (Canonical form)

Therefore, the model is not a generalized linear Model.

Ex. 3.10:

$y_1, y_2, \dots, y_N$ are independent

$y_i \sim N(\mu_i, \sigma^2) \in EF$ in canonical form.

$$\begin{aligned}E(y_i) &= \mu_i = \beta_0 + \ln(\beta_1 + \beta_2 x_i) \\ \mu_i &= \beta_0 + \ln(\beta_1 + \beta_2 x_i) \\ \Leftrightarrow e^{\mu_i} &= e^{\beta_0 + \ln(\beta_1 + \beta_2 x_i)} = e^{\beta_0} e^{\ln(\beta_1 + \beta_2 x_i)} = e^{\beta_0} (\beta_1 + \beta_2 x_i) = \beta_1 e^{\beta_0} + \beta_2 x_i e^{\beta_0} = \underbrace{\beta_1^* + \beta_2^* x_i}_{\text{linear component}}\end{aligned}$$

where $\beta_1^* = \beta_1 e^{\beta_0}$ and $\beta_2^* = \beta_2 e^{\beta_0}$. The model is a generalized linear model.

Ex. 3.11:

$y \sim \text{Pareto}(\theta)$.

$$f(y; \theta) = \theta y^{-\theta-1} \text{ (for one observation)}$$

Let $y_1, y_2, \dots, y_n$ are independent, and $y_i \sim \text{Pareto}(\theta)$ for $i = 1, 2, \dots, n$

Likelihood function is:

$$L(\theta; \underline{y}) = f(\underline{y}; \theta) = \prod_{i=1}^n f(y_i; \theta)$$

log-likelihood function is:

$$\begin{aligned}l(\theta; \underline{y}) &= \ln [L(\theta; \underline{y})] = \ln \left[\prod_{i=1}^n f(y_i; \theta) \right] = \sum_{i=1}^n \ln [f(y_i, \theta)] = \sum_{i=1}^n \ln [\theta y_i^{-\theta-1}] \\ &= \sum_{i=1}^n \{ \ln(\theta) + \ln[y_i^{-\theta-1}] \} = \sum_{i=1}^n \{ \ln(\theta) + (-\theta - 1) \ln(y_i) \} \\ &= n \ln(\theta) - (\theta + 1) \sum_{i=1}^n \ln(y_i)\end{aligned}$$

The Score Statistic is:

$$U(\theta, \underline{y}) = \frac{\partial l(\theta; \underline{y})}{\partial \theta} = \frac{n}{\theta} - \sum_{i=1}^n \ln(y_i)$$

Recall: $E[U(\theta, \underline{y})] = 0$

Information:

$$J = \text{Var}(u) = E(u^2) = -E(u') = -E \left[\frac{\partial}{\partial \theta} u(\theta; \underline{y}) \right] = -E \left\{ \frac{\partial}{\partial \theta} \left(\frac{n}{\theta} - \sum_{i=1}^n \ln(y_i) \right) \right\} = -E \left[-\frac{n}{\theta^2} \right] = \frac{n}{\theta^2}$$

2nd Solution:

For Pareto distr. , we have [sec Ex. 3.3(a)]

$$a(y) = \ln(y)$$

$$b(\theta) = -\theta - 1$$

$$c(\theta) = \ln(\theta)$$

$$\Rightarrow b'(\theta) = -1; b''(\theta) = 0$$

$$c'(\theta) = \frac{1}{\theta}; c''(\theta) = -\frac{1}{\theta^2}$$

We know that

$$E[a(y)] = -\frac{c'(\theta)}{b'(\theta)}$$

$$\therefore E[\ln(y)] = -\frac{1/\theta}{(-1)} = \frac{1}{\theta}$$

$$\text{Note: } E(u) = E\left[\frac{n}{\theta} - \sum \ln(y_i)\right] = \frac{n}{\theta} - \sum E[\ln(y_i)] = \frac{n}{\theta} - \sum \frac{1}{\theta} = \frac{n}{\theta} - \frac{n}{\theta} = 0$$

$$\text{Var}[a(y)] = \frac{b''(\theta)c'(\theta) - c''(\theta)b'(\theta)}{[b'(\theta)]^3}$$

$$\therefore \text{Var}[\ln(y)] = \frac{0 - (-1/\theta^2)(-1)}{(-1)^3} = \frac{-1/\theta^2}{-1} = \frac{1}{\theta^2}$$

$$\Rightarrow E\{[\ln(y)]^2\} = \text{Var}[\ln(y)] + \{E[\ln(y)]\}^2 = \frac{1}{\theta^2} + \left(\frac{1}{\theta}\right)^2 = \frac{2}{\theta^2}$$

Now,

$$J = \text{Var}(u) = E(u^2) = E\left\{\left[\frac{n}{\theta} - \sum \ln(y_i)\right]^2\right\} = E\left\{\left(\frac{n}{\theta}\right)^2 + \left[\sum \ln(y_i)\right]^2 - 2\frac{n}{\theta} \sum \ln(y_i)\right\}$$

$$= \left(\frac{n}{\theta}\right)^2 + E\left\{\left[\sum \ln(y_i)\right]^2\right\} - 2\frac{n}{\theta} E\left[\sum \ln(y_i)\right]$$

$$= \left(\frac{n}{\theta}\right)^2 + \left[\text{Var}\left[\sum \ln(y_i)\right] + \left\{E\left[\sum \ln(y_i)\right]\right\}^2\right] - 2\frac{n}{\theta} E\left[\sum \ln(y_i)\right]$$

$$= \left(\frac{n}{\theta}\right)^2 + \sum_{i=1}^n \text{Var}(\ln(y_i)) + \left[\sum_{i=1}^n E(\ln(y_i))\right]^2 - 2\frac{n}{\theta} \sum E[\ln(y_i)]$$

$$= \left(\frac{n}{\theta}\right)^2 + \left[\sum_{i=1}^n \frac{1}{\theta^2} + \left(\sum_{i=1}^n \frac{1}{\theta}\right)^2\right] - 2\frac{n}{\theta} \sum_{i=1}^n \frac{1}{\theta} = \left(\frac{n}{\theta}\right)^2 + \frac{n}{\theta^2} + \left(\frac{n}{\theta}\right)^2 - 2\frac{n}{\theta} \left(\frac{n}{\theta}\right)$$

$$= \frac{n^2}{\theta^2} + \frac{n}{\theta^2} + \frac{n^2}{\theta^2} - 2\frac{n^2}{\theta^2} = \frac{n}{\theta^2}$$

3rd Solution:

We know that

$$\text{Var}(u) = \frac{b''(\theta)c'(\theta)}{b'(\theta)} - c''(\theta) \quad (\text{for one observation})$$

$$\Rightarrow \text{Var}(u) = \frac{(0)\left(\frac{1}{\theta}\right)}{(-1)} - \left(-\frac{1}{\theta^2}\right) = \frac{1}{\theta^2}$$

$$\therefore J = \text{Var}(u) \quad (\text{for all observations})$$

$$\Rightarrow J = \text{Var}(u) = \sum_{i=1}^n \frac{1}{\theta^2} = \frac{n}{\theta^2}$$

$$E(u) = 0:$$

$$u = \frac{n}{\theta} - \sum_{i=1}^n \ln(y_i)$$

$$\Rightarrow E(u) = E\left[\frac{n}{\theta} - \sum_{i=1}^n \ln(y_i)\right] = \frac{n}{\theta} - E\left[\sum_{i=1}^n \ln(y_i)\right] = \frac{n}{\theta} - \sum_{i=1}^n E[\ln(y_i)] = \frac{n}{\theta} - \sum_{i=1}^n \frac{1}{\theta} = \frac{n}{\theta} - \frac{n}{\theta} = 0$$

where $E(\ln(y_i)) = \frac{1}{\theta}$.