

Chapter 2

Ex. 2.1:

$$y_{jk} \sim N(\mu_j, \sigma^2); j = 1, 2$$

$$\text{i.e. } y_{1k} \sim N(\mu_1, \sigma^2); k = 1, \dots, n_1 \text{ and } y_{2k} \sim N(\mu_2, \sigma^2); k = 1, 2, \dots, n_2$$

y_{jk} are independent.

$$H_0: \mu_1 = \mu_2 = \mu$$

$$H_1: \mu_1 \neq \mu_2$$

(a) use Minitab.

Group	
Trtreatment	control
$\bar{y}_1 = 4.86$	$\bar{y}_2 = 4.726$
$s_1^2 = 0.626$	$s_2^2 = 0.746$
$n_1 = 20$	$n_2 = 20$

$$N = n_1 + n_2 = 40$$

$$\bar{y}_\cdot = 4.793$$

$$s^2 = 0.673$$

$$s_p^2 = 0.686$$

$$s_p = 0.82825$$

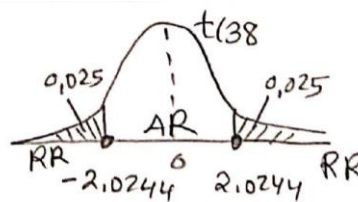
$$df = n_1 + n_2 - 2 = 38$$

(b) T-test:

$$t = \frac{\bar{y}_1 - \bar{y}_2}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{4.86 - 4.726}{0.82825 \sqrt{\frac{1}{20} + \frac{1}{20}}} = \frac{0.134}{0.2619156} = 0.5116$$

$$t_{\frac{\alpha}{2}(n_1+n_2-2)} = t_{0.025}(38) = 2.0244$$

$$|t| = 0.5116 < t_{0.025}(38)$$



$\Rightarrow$ Do not reject to $H_0: \mu_1 = \mu_2$ at significance Level $\alpha = 0,05$

(c)

(i) MLE: (Maximum Likelihood Estimativ):

First: For Model under $H_0: \mu_1 = \mu_2 = \mu_0$.

$$f(y_{jk}; \mu) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_{jk}-\mu)^2}$$

$$\Rightarrow \ln f(y_{jk}; \mu) = \ln \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] - \frac{1}{2\sigma^2} (y_{jk} - \mu)^2$$

$$\begin{aligned}
l(\mu) &= \sum_{j=1}^J \sum_{k=1}^K \ln f(y_{jk}; \mu) = \sum_j \sum_k \left\{ \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right) - \frac{1}{2\sigma^2} (y_{jk} - \mu)^2 \right\} \\
\Rightarrow \frac{\partial l(\mu)}{\partial \mu} &= -\frac{1}{2\sigma^2} \sum \sum 2(y_{jk} - \mu)(-1) = \frac{1}{\sigma^2} \sum \sum (y_{jk} - \mu) \\
\frac{\partial l(\mu)}{\partial \mu} \stackrel{\text{set}}{=} 0 &\Leftrightarrow \frac{1}{\sigma^2} \sum \sum (y_{jk} - \mu) = 0 \Leftrightarrow \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu) = 0 \Leftrightarrow \sum_{j=1}^J \sum_{k=1}^K y_{jk} - JK\mu = 0 \\
&\Leftrightarrow \sum \sum y_{jk} = JK\mu \Leftrightarrow \hat{\mu} = \frac{\sum \sum y_{jk}}{JK} = \frac{\sum \sum y_{jk}}{N} = \bar{y}_{..}.
\end{aligned}$$

$$J = 2; k = 20 \Rightarrow N = JK = (2)(20) = 40$$

$\therefore$ MLE of μ (under H_0) is:

$$\hat{\mu} = \bar{y}_{..}$$

Second: For Model under $H_1: \mu_1 \neq \mu_2$:

$$\begin{aligned}
f(y_{jk}; \mu_j) &= \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_{jk}-\mu_j)^2}; j = 1, 2, k = 1, \dots, 20 \\
\Rightarrow \ln f(y_{jk}; \mu_j) &= \ln \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] - \frac{1}{2\sigma^2} (y_{jk} - \mu_j)^2 \\
\ell(\mu_1, \mu_2) &= \sum_{j=1}^J \sum_{k=1}^K \ln(y_{jk}; \mu_j) = \sum \sum \left\{ \ln \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] - \frac{1}{2\sigma^2} (y_{jk} - \mu_j)^2 \right\} \\
&= \sum_{j=1}^J \sum_{k=1}^K \ln \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] - \frac{1}{2\sigma^2} \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu_j)^2 \\
&= \sum \sum \left[\frac{1}{\sqrt{2\pi\sigma^2}} \right] - \frac{1}{2\sigma^2} \left[\sum_{k=1}^K (y_{1k} - \mu_1)^2 + \sum_{k=1}^K (y_{2k} - \mu_2)^2 \right] \\
\Rightarrow \frac{\partial \ell(\mu_1, \mu_2)}{\partial \mu_1} &= \frac{1}{\sigma^2} \sum_{k=1}^K (y_{1k} - \mu_1) \text{ and } \frac{\partial \ell(\mu_1, \mu_2)}{\partial \mu_2} = \frac{1}{\sigma^2} \sum_{k=1}^K (y_{2k} - \mu_2) \\
\frac{\partial \ell(\mu_1, \mu_2)}{\partial \mu_1} \stackrel{\text{set}}{=} 0 &\text{ and } \frac{\partial \ell(\mu_1, \mu_2)}{\partial \mu_2} \stackrel{\text{set}}{=} 0 \\
\Rightarrow \frac{1}{\sigma^2} \sum_{k=1}^K (y_{1k} - \mu_1) &= 0 \text{ and } \frac{1}{\sigma^2} \sum_{k=1}^K (y_{2k} - \mu_2) = 0 \\
\Rightarrow \sum_{k=1}^K (y_{1k} - \mu_1) &= 0 \text{ and } \sum_{k=1}^K (y_{2k} - \mu_2) = 0 \\
\Rightarrow \sum_{k=1}^K y_{1k} - K\mu_1 &= 0 \text{ and } \sum_{k=1}^K y_{2k} - K\mu_2 = 0 \\
\Rightarrow \sum_{k=1}^K y_{1k} &= K\mu_1 \text{ and } \sum_{k=1}^K y_{2k} = K\mu_2 \\
\hat{\mu}_1 = \frac{\sum_{k=1}^K y_{1k}}{K} &= \bar{y}_{1.} \text{ and } \hat{\mu}_2 = \frac{\sum_{k=1}^K y_{2k}}{K} = \bar{y}_{2.}
\end{aligned}$$

$\therefore$ The MLE of μ_1 and μ_2 under H_1 are

$$\hat{\mu}_1 = \bar{y}_{1.}$$

$$\hat{\mu}_2 = \bar{y}_{2.}$$

(ii) LSE (Least Square Estimation):

First: For Model under $H_0: \mu_1 = \mu_2 = \mu$:

$$\begin{aligned}
S_0 &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu)^2 \\
&\Rightarrow \frac{\partial S_0}{\partial \mu} = -2 \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu) \\
\frac{\partial S_0}{\partial \mu} \stackrel{\text{set}}{=} 0 &\Leftrightarrow -2 \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu) = 0 \Leftrightarrow \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu) = 0 \Leftrightarrow \sum_{j=1}^J \sum_{k=1}^K y_{jk} - JK\mu = 0 \\
&\Leftrightarrow \mu = \frac{\sum \sum y_{jk}}{JK} = \frac{\sum \sum y_{jk}}{N} = \bar{y}_{..} \quad (\text{as } N = JK \text{ where } J = 2 \text{ and } K = 20)
\end{aligned}$$

$\therefore$ The LSE of μ under H_0 is:

$$\hat{\mu} = \bar{y}_{..} \quad (\text{same as MLE})$$

Second: For Model under $H_1: \mu_1 \neq \mu_2$:

$$\begin{aligned}
S_1 &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu_j)^2 = \sum_{k=1}^K (y_{1k} - \mu_1)^2 + \sum_{k=1}^K (y_{2k} - \mu_2)^2 \\
&\Rightarrow \frac{\partial S_1}{\partial \mu_1} = -2 \sum_{k=1}^K (y_{1k} - \mu_1) \text{ and } \frac{\partial S_1}{\partial \mu_2} = -2 \sum_{k=1}^K (y_{2k} - \mu_2) \\
&\quad \frac{\partial S_1}{\partial \mu_1} \stackrel{\text{set}}{=} 0 \text{ and } \frac{\partial S_1}{\partial \mu_2} \stackrel{\text{set}}{=} 0 \\
&\quad \Rightarrow -2 \sum_{k=1}^K (y_{1k} - \mu_1) = 0 \text{ and } -2 \sum_{k=1}^K (y_{2k} - \mu_2) = 0 \\
&\quad \Rightarrow \sum_{k=1}^K (y_{1k} - \mu_1) = 0 \text{ and } \sum_{k=1}^K (y_{2k} - \mu_2) = 0 \\
&\quad \Rightarrow \sum_{k=1}^K y_{1k} - K\mu_1 = 0 \text{ and } \sum_{k=1}^K y_{2k} - K\mu_2 = 0 \\
&\quad \Rightarrow \sum_{k=1}^K y_{1k} = K\mu_1 \text{ and } \sum_{k=1}^K y_{2k} = K\mu_2 \\
&\quad \Rightarrow \mu_1 = \frac{\sum_{k=1}^K y_{1k}}{K} = \bar{y}_{1.} \text{ and } \mu_2 = \frac{\sum_{k=1}^K y_{2k}}{K} = \bar{y}_{2.}
\end{aligned}$$

$\therefore$ The LSE of μ_1 and μ_2 under H_1 are:

$$\hat{\mu}_1 = \bar{y}_{1.} \text{ and } \hat{\mu}_2 = \bar{y}_{2.}$$

which is the same as MLE

(d) Recall:

- The LSE of μ under ($H_0: \mu_1 = \mu_2 = \mu$) is $\hat{\mu} = \bar{y}_{..}$.
- The LSE of μ_1 and μ_2 under ($H_1: \mu_1 \neq \mu_2$) are:
 $\hat{\mu}_1 = \bar{y}_{1.}$ and $\hat{\mu}_2 = \bar{y}_{2.}$

(i) The minimum value of $S_0 = \sum \sum (y_{jk} - \mu)^2$ is:

$$\hat{S}_0 = \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \hat{\mu})^2 = \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{..})^2 ; \bar{y}_{..} = \frac{\sum \sum y_{jk}}{JK} \text{ (grand average)}$$

(ii) The minimum value of $S_1 = \sum \sum (y_{jk} - \mu_j)^2$ is:

$$\begin{aligned}
\hat{S}_1 &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \hat{\mu}_j)^2 = \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})^2 ; \bar{y}_{j.} = \frac{\sum_{k=1}^K y_{jk}}{K} \\
&= \sum_{k=1}^K (y_{1k} - \bar{y}_{1.})^2 + \sum_{k=1}^K (y_{2k} - \bar{y}_{2.})^2
\end{aligned}$$

(e)

(1)

$$\begin{aligned}
S_1 &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu_j)^2 ; (\text{under } H_1: \mu_1 \neq \mu_2) \\
&= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.} + \bar{y}_{j.} - \mu_j)^2 ; \left(\bar{y}_{j.} = \frac{\sum_{k=1}^K y_{jk}}{K}; j = 1, 2 \right) \\
&= \sum_{j=1}^J \sum_{k=1}^K [(y_{jk} - \bar{y}_{j.}) + (\bar{y}_{j.} - \mu_j)]^2 \\
&= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})^2 + \sum_{j=1}^J \sum_{k=1}^K (\bar{y}_{j.} - \mu_j)^2 + 2 \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})(\bar{y}_{j.} - \mu_j) \\
&= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})^2 + K \sum_{j=1}^J (\bar{y}_{j.} - \mu_j)^2 + 2 \sum_{j=1}^J (\bar{y}_{j.} - \mu_j) \underbrace{\sum_{k=1}^K (y_{jk} - \bar{y}_{j.})}_{=0}
\end{aligned}$$

Note that:

$$\begin{aligned}
&\sum_{k=1}^K (y_{jk} - \bar{y}_{j.}) = 0 \text{ for } j = 1, \dots, J \\
\Rightarrow S_1 &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})^2 + K \sum_{j=1}^J (\bar{y}_{j.} - \mu_j)^2
\end{aligned}$$

or

$$\begin{aligned}
&\sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \mu_j)^2 = \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{j.})^2 + K \sum_{j=1}^J (\bar{y}_{j.} - \mu_j)^2 \\
\Leftrightarrow \underbrace{\sum_j \sum_k (y_{jk} - \bar{y}_{j.})^2}_{\hat{S}_1} &= \sum_j \sum_k (y_{jk} - \mu_j)^2 - K \sum_j (\bar{y}_{j.} - \mu_j)^2 \\
\Leftrightarrow \hat{S}_1 &= \sum_j \sum_k (y_{jk} - \mu_j)^2 - K \sum_j (\bar{y}_{j.} - \mu_j)^2 \\
\Leftrightarrow \frac{\hat{S}_1}{\sigma^2} &= \frac{1}{\sigma^2} \sum_j \sum_k (y_{jk} - \mu_j)^2 - \frac{K}{\sigma^2} \sum_j (\bar{y}_{j.} - \mu_j)^2
\end{aligned}$$

Note:

LSE of μ_j is $\hat{\mu}_j = \bar{y}_{j.}$ under H_1 :

$$\hat{S}_1 = \sum_j \sum_k (y_{jk} - \hat{\mu}_j)^2 = \sum_j \sum_k (y_{jk} - \bar{y}_{j.})^2$$

Distribution of $\frac{\hat{S}_1^2}{\sigma^2}$ (under H_1): We found that:

$$\frac{1}{\sigma^2} \hat{S}_1 = \frac{1}{\sigma^2} \sum_j \sum_k (y_{jk} - \mu_j)^2 - \frac{K}{\sigma^2} \sum_j (\bar{y}_{j.} - \mu_j)^2 = \sum_{j=1}^J \sum_{k=1}^K \left(\frac{y_{jk} - \mu_j}{\sigma} \right)^2 - \sum_{j=1}^J \left(\frac{\bar{y}_{j.} - \mu_j}{\sigma/\sqrt{K}} \right)^2$$

Recall:

If $y_1, y_2, \dots, y_n$ iid $N(\mu, \sigma^2)$ then $\bar{y} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$ since y_{jk} are independent and $y_{jk} \sim N(\mu_j, \sigma^2)$,

(1) We have $\bar{y}_{j.} \sim N\left(\mu_j, \frac{\sigma^2}{K}\right); j = 1, 2 \Rightarrow \frac{\bar{y}_{j.} - \mu_j}{\frac{\sigma}{\sqrt{K}}} \sim N(0, 1) \left(\frac{\bar{y}_{j.} - \mu_j}{\frac{\sigma}{\sqrt{K}}} \right)^2 = \frac{K}{\sigma^2} (\bar{y}_{j.} - \mu_j)^2 \sim \chi_{(1)}^2$

$$\Rightarrow \sum_{j=1}^2 \left(\frac{\bar{y}_{j\cdot} - \mu_j}{\sigma/\sqrt{K}} \right)^2 \sim \chi_{(2)}^2$$

$$(2) y_{jk} \sim N(\mu_j, \sigma^2) \Rightarrow \frac{y_{jk} - \mu_j}{\sigma} \sim N(0,1) \Rightarrow \left(\frac{y_{jk} - \mu_j}{\sigma} \right)^2 \sim \chi_{(1)}^2 \Rightarrow \sum_{j=1}^J \sum_{k=1}^K \left(\frac{y_{jk} - \mu_j}{\sigma} \right)^2 \sim \chi_{(JK)}^2$$

Therefore:

$$\frac{1}{\sigma^2} \hat{S}_1 = \chi_{(JK)}^2 - \chi_{(2)}^2 \Rightarrow \chi_{(JK)}^2 = \frac{\hat{S}_1}{\sigma^2} + \chi_{(2)}^2 \Rightarrow \frac{\hat{S}_1}{\sigma^2} \sim \chi_{(JK-2)}^2$$

(2)

$$\begin{aligned} S_0 &= \sum \sum (y_{jk} - \mu)^2; (\text{ under } H_0: \mu_1 = \mu_2 = \mu) \\ &= \sum \sum (y_{jk} - \bar{y}_{..} + \bar{y}_{..} - \mu)^2 = \sum \sum [(y_{jk} - \bar{y}_{..}) + (\bar{y}_{..} - \mu)]^2; \left(\bar{y}_{..} = \frac{\sum \sum y_{jk}}{JK} \right) \\ &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{..})^2 + \sum_{j=1}^J \sum_{k=1}^K (\bar{y}_{..} - \mu)^2 + 2 \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{..})(\bar{y}_{..} - \mu) \\ &= \sum_{j=1}^J \sum_{k=1}^K (y_{jk} - \bar{y}_{..})^2 + JK(\bar{y}_{..} - \mu)^2 + 2(\bar{y}_{..} - \mu) \underbrace{\sum \sum (y_{jk} - \bar{y}_{..})}_{=0} \\ &\Rightarrow S_0 = \sum \sum (y_{jk} - \bar{y}_{..})^2 + JK(\bar{y}_{..} - \mu)^2 \end{aligned}$$

or

$$\begin{aligned} \sum \sum (y_{jk} - \mu)^2 &= \sum \sum (y_{jk} - \bar{y}_{..})^2 + JK(\bar{y}_{..} - \mu)^2 \\ \Leftrightarrow \underbrace{\sum \sum (y_{jk} - \bar{y}_{..})^2}_{\hat{S}_0} &= \sum \sum (y_{jk} - \mu)^2 - JK(\bar{y}_{..} - \mu)^2 \\ \hat{S}_0 &= \sum \sum (y_{jk} - \mu)^2 - JK(\bar{y}_{..} - \mu)^2 \\ \Leftrightarrow \frac{\hat{S}_0}{\sigma^2} &= \frac{1}{\sigma^2} \sum \sum (y_{jk} - \mu)^2 - \frac{JK}{\sigma^2} (\bar{y}_{..} - \mu)^2 \end{aligned}$$

Note:

The LSE of μ is $\hat{\mu} = \bar{y}_{..}$ under H_0 :

$$\hat{S}_0 = \sum \sum (y_{jk} - \hat{\mu})^2 = \sum \sum (y_{jk} - \bar{y}_{..})^2$$

Distribution of $\frac{\hat{S}_0}{\sigma^2}$ under H_0 :

We have found that

$$\begin{aligned} \frac{\hat{S}_0}{\sigma^2} &= \frac{1}{\sigma^2} \sum \sum (y_{jk} - \mu)^2 - \frac{JK}{\sigma^2} (\bar{y}_{..} - \mu)^2 \\ &= \underbrace{\sum_{j=1}^J \sum_{k=1}^K \left(\frac{y_{jk} - \mu}{\sigma} \right)^2}_{\chi_{(JK)}^2} - \underbrace{\left(\frac{\bar{y}_{..} - \mu}{\sigma/\sqrt{JK}} \right)^2}_{\chi_{(1)}^2} \\ &= \chi_{(JK)}^2 - \chi_{(1)}^2 \\ \Leftrightarrow \chi_{(JK)}^2 &= \frac{\hat{S}_0}{\sigma^2} + \chi_{(1)}^2 \Rightarrow \frac{\hat{S}_0}{\sigma^2} \sim \chi_{(JK-1)}^2 \end{aligned}$$

(f) We have found that

$$\frac{\hat{S}_0}{\sigma^2} \sim \chi^2_{(JK-1)} \quad (\text{under } H_0)$$

$$\frac{\hat{S}_1}{\sigma^2} \sim \chi^2_{(JK-2)} \quad (\text{under } H_1) \rightarrow *$$

Now, if H_0 is true,

$$\frac{\hat{S}_0 - \hat{S}_1}{\sigma^2} \sim \chi^2_{(1)} \quad (\text{as } (JK-1) - (JK-2) = 1) \rightarrow **$$

From * and **:

$$F = \frac{\frac{\hat{S}_0 - \hat{S}_1}{\sigma^2} / (1)}{\frac{\hat{S}_1}{\sigma^2} / (JK-2)} = \frac{\hat{S}_0 - \hat{S}_1}{\hat{S}_1 / (JK-2)} \sim F(1, JK-2)$$

(g)

$$H_0: \mu_1 = \mu_2 = \mu$$

$$H_1: \mu_1 \neq \mu_2$$

$$\sum \sum y_{jk} = 191.73, N = JK = (2)(20) = 40$$

$$\begin{aligned} \hat{S}_0 &= \sum \sum (y_{jk} - \hat{\mu})^2 = \sum \sum (y_{jk} - \bar{y}_{..})^2 = \sum \sum y_{jk}^2 - JK(\bar{y}_{..})^2 \\ &= \sum \sum y_{jk}^2 - (2)(20)(\bar{y}_{..})^2 = 945.241 - (40)(4.793)^2 = 26.32704 \end{aligned}$$

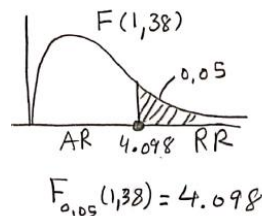
$$\begin{aligned} \hat{S}_1 &= \sum_{j=1}^2 \sum_{k=1}^{20} (y_{jk} - \bar{y}_{j.})^2 = \sum_{k=1}^{20} (y_{1k} - \bar{y}_{1.})^2 + \sum_{k=1}^{20} (y_{2k} - \bar{y}_{2.})^2 \\ &= \left[\sum_{k=1}^{20} y_{1k}^2 - K(\bar{y}_{1.})^2 \right] + \left[\sum_{k=1}^{20} y_{2k}^2 - K(\bar{y}_{2.})^2 \right] \\ &= [484.278 - (20)(4.86)^2] + [460.964 - (20)(4.726)^2] \end{aligned}$$

$$F_0 = \frac{\hat{S}_0 - \hat{S}_1}{\hat{S}_1 / (JK-2)} = \frac{26.32704 - 26.14848}{26.14848 / 38} = 0.2595$$

$$F_0 = 0.2595 < F_{0.05}(1, 38) = 4.098$$

$$F_0 \in AR$$

$\therefore$ We do not reject H_0



(h)

$$F_0 = 0.2595 \quad (\sim F(1, 38))$$

$$t_0 = 0.5116 \quad (\sim t(38))$$

$$t_0^2 = (0.5116)^2 = 0.2617$$

$$t_0^2 = F_0 \quad (\text{The difference is because of rounding errors})$$

(i) Calculating the residuals: $e_{jk} = y_{jk} - \bar{y}_{..}$.

[Minitab For Model H_0].

The residuals are consistent with the assumptions of independence, equal variances and Normality.

Ex. 2.2

(a)

$$H_0: \mu_1 = \mu_2$$

$$H_1: \mu_1 \neq \mu_2$$

Unpaired t-test:

Before	After
$\bar{y}_1 = 103.25$	$\bar{y}_2 = 100.60$
$s_1^2 = 182.72$	$s_2^2 = 155.61$
$n_1 = 20$	$n_2 = 20$

$$N = n_1 + n_2 = 40$$

$$\bar{y}_{..} = 4.793$$

$$s^2 = 0.673$$

$$s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2} = 169.165$$

$$s_p = 13.006$$

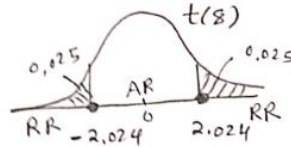
$$df = n_1 + n_2 - 2 = 38$$

$$t_0 = \frac{\bar{y}_1 - \bar{y}_2}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{103.25 - 100.60}{13.006 \sqrt{\frac{1}{20} + \frac{1}{20}}} = 0.644$$

$$|t_0| < t_{0.025, (38)} = 2.024$$

$$t_0 \in AR$$

$\therefore$ We do not reject $H_0: \mu_1 = \mu_2$ at significance level $\alpha = 0.05$.



(b) $\mu_D = \mu_1 - \mu_2$; $E(D) = E(y_1 - y_2) = E(y_1) - E(y_2) = \mu_1 - \mu_2$

$$H_0: \mu_D = 0$$

$$H_1: \mu_D \neq 0$$

paired t-test

Man	Before y_{1k}	After y_{2k}	Difference $D_k = y_{1k} - y_{2k}$
1	1008.8	977.0	38.0
2	102.0	107.5	-5.5
$\vdots$	$\vdots$	$\vdots$	$\vdots$
20	93.0	93.0	0.0

$$t_0 = \frac{\bar{D}}{S_D/\sqrt{n}} = \frac{\bar{D}}{S_{\bar{D}}} = \frac{2.645}{0.9206} = 2.873$$

$$\bar{D} = 2.645$$

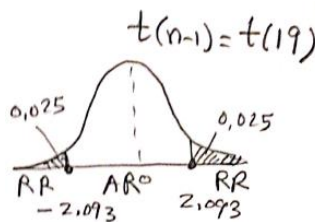
$$S_{\bar{D}} = \frac{S_D}{\sqrt{n}} = \frac{4.117}{\sqrt{20}} = 0.9206$$

$$S_D^2 = 16.947 \Rightarrow S_D = 4.117$$

$$|t_0| > t_{0,025}(19) = 2.093$$

$$t_0 \in RR$$

$\therefore$ we reject $H_0: \mu_1 = \mu_2$ at significance level $\alpha = 0.05$.



(c) The conclusions are:

(1) For t -test: $\begin{cases} H_0: \mu_1 = \mu_2 \text{ is not rejected} \\ \text{at } \alpha = 0.05 \end{cases}$

(2) For paired t -test: $\begin{cases} H_0: \mu_1 = \mu_2 \text{ is rejected} \\ \text{at } \alpha = 0.05 \end{cases}$

(d)

- We reached different conclusions.
- The paired t -test is more appropriate than t -test.
- Assumptions:

(1) For t -test: $\begin{cases} 1. y_{jk} \sim N(\mu_j, \sigma_j^2) \\ 2. y_{jk} \text{ are independent} \\ 3. \sigma_1^2 = \sigma_2^2 \end{cases}$

i.e. For (a) it is assumed that the y_{jk} 's are independent and $y_{jk} \sim N(\mu_j, \sigma_j^2)$ for all j and for all k .

(2) For paired t -test $\begin{cases} 1. y_{jk} \sim N(\mu_j, \sigma_j^2) \\ 2. y_{11}, \dots, y_{1k} \text{ are independent} \\ 3. y_{21}, \dots, y_{2k} \text{ are independent} \end{cases}$

but y_{1k} and y_{2k} are not independent.

i.e. For (b) it is assumed that the D_k 's are independent with $D_k \sim N(\mu_D, \sigma_D^2)$. The analysis in (b) does not involve assuming that y_{1k} and y_{2k} (i.e., 'before' and 'after' weights for the same man) are independent, so it is more appropriate.

Ex 2.4:

$$\begin{aligned}
 E(y) &= \ln (\beta_0 + \beta_1 x + \beta_2 x^2) \\
 \Rightarrow e^{E(y)} &= \beta_0 + \beta_1 x + \beta_2 x^2 \\
 \Rightarrow g[E(y)] &= \beta_0 + \beta_1 x + \beta_2 x^2; (g(x) = e^x) \\
 \Rightarrow g[E(\underline{y})] &= \underline{x} \underline{\beta}
 \end{aligned}$$

$$\underline{y} = \begin{bmatrix} 3.15 \\ 4.85 \\ 6.50 \\ 7.20 \\ 8.25 \\ 16.50 \end{bmatrix}, \underline{x} = \begin{bmatrix} \beta_0 & \beta_1 & \beta_2 \\ 1 & 1.0 & 1.0 \\ 1 & 1.2 & 1.44 \\ 1 & 1.4 & 1.96 \\ 1 & 1.6 & 2.56 \\ 1 & 1.8 & 3.24 \\ 1 & 2.0 & 4.00 \\ & x & x^2 \end{bmatrix}, \underline{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \end{bmatrix}$$

Ex 2.5:

$$\begin{aligned}
 E(y_{jk}) = \mu_{jk} &= \mu + \alpha_j + \beta_k \begin{cases} y_{jk} \sim N(\mu_{jk}, \sigma^2); j = 1, 2 \text{ and } k = 1, 2, 3 \\ \text{independent} \end{cases} \\
 \Rightarrow E(\underline{y}) &= \underline{x} \underline{\beta}
 \end{aligned}$$

constraints:

$$\begin{aligned}
 \sum \alpha_j &= 0 \Leftrightarrow \alpha_2 = -\alpha_1 \\
 \sum \beta_k &= 0 \Leftrightarrow \beta_3 = -\beta_1 - \beta_2
 \end{aligned}$$

$$\underline{y} = \begin{bmatrix} y_{11} \\ y_{12} \\ y_{13} \\ y_{21} \\ y_{22} \\ y_{23} \end{bmatrix}, \underline{x} = \begin{bmatrix} \mu & \alpha_1 & \beta_1 & \beta_2 \\ 1 & 1 & 1 & 0 \\ 1 & 1 & 0 & 1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & 1 & 0 \\ 1 & -1 & 0 & 1 \\ 1 & -1 & -1 & -1 \end{bmatrix} \text{ and } \underline{\beta} = \begin{bmatrix} \mu \\ \alpha_1 \\ \beta_1 \\ \beta_2 \end{bmatrix}.$$