

Q1

a) $X \sim \text{uniform}(0, 4000)$

$$E[X] = \int_0^{4000} x \frac{1}{4000} dx = \frac{1}{4000} \frac{x^2}{2} \Big|_{x=0}^{4000} \\ = 2000$$

$$b) E[X^2] = \int_0^{4000} x^2 \frac{1}{4000} dx \\ = \frac{1}{4000} \frac{x^3}{3} \Big|_{x=0}^{4000} = \frac{1}{4000} \frac{4000^3}{3} = \frac{4000^2}{3} \\ = 5,333,333.33$$

$$\Rightarrow \text{Var}[X] = E[X^2] - (E[X])^2 \\ = 5,333,333.33 - 2000^2 = 1,333,333.33$$

$$\Rightarrow \text{SD}[X] = 1,154.701$$

$$c) E[Y] = \int_d^{4000} (x-d) f_X(x) dx \\ = \int_{400}^{4000} (x-400) \frac{1}{4000} dx \\ = \frac{1}{4000} \frac{(x-400)^2}{2} \Big|_{400}^{4000} \\ = \frac{1}{4000} \frac{(3600)^2}{2} = 1,620$$

d) If K is the amount of loss not covered by the policy, then

$$K = \begin{cases} X & \text{if } 0 \leq X \leq 400 \\ 400 & \text{if } 400 \leq X \leq 4000 \end{cases}$$

$$\Rightarrow E[K] = \int_0^{400} x \cdot \frac{1}{4000} dx + \int_{400}^{4000} 400 \cdot \frac{1}{4000} dx$$

$$= \frac{1}{4000} \cdot \frac{x^2}{2} \Big|_{x=0}^{400} + 400 \Pr(400 \leq X \leq 4000)$$

$$= \frac{400^2}{2 \times 4000} + 400 \times \frac{9}{10}$$

$$= 20 + 360 = 380$$

Q2

$$a) P_{GG} = Pr \{X_{n+1} = G | X_n = G\} = 0.4$$

$$P_{GD} = Pr \{X_{n+1} = D | X_n = G\} = 0.6$$

$$P_{DD} = Pr \{X_{n+1} = D | X_n = D\} = 0.3$$

$$P_{DG} = Pr \{X_{n+1} = G | X_n = D\} = 0.7$$

$$Pr \{X_2 = G, X_3 = G, X_4 = G, X_5 = G, X_6 = D | X_1 = G\} = ?$$

$$= Pr \{X_6 = D | X_2 = G, X_3 = G, X_4 = G, X_5 = G\} \cdot Pr \{X_2 = G, X_3 = G, X_4 = G, X_5 = G | X_1 = G\}$$

$$= Pr \{X_6 = D | X_5 = G\} \cdot Pr \{X_2 = G, X_3 = G, X_4 = G, X_5 = G | X_1 = G\}$$

$$= P_{GD}$$

repeat the step.

$$Pr \{X_5 = G | X_2 = G, X_3 = G, X_4 = G, X_1 = G\} \cdot Pr \left\{ \begin{matrix} X_2 = G \\ X_3 = G \\ X_4 = G \end{matrix} | X_1 = G \right\}$$

$$Pr \{X_5 = G | X_4 = G\}$$

$$= P_{GD}$$

$$P_{GG}$$

and so on continue.

$$= P_{GD} \cdot P_{GG}^4 = 0.6 (0.4)^4 = 0.01536$$

we get:

Q2

$$b) P_r(X_5=2 | X_3=1)$$

$$= P_{12}^2 = 0.72$$

$$P^2 = \begin{bmatrix} 0.1 & 0.9 \\ 0.3 & 0.7 \end{bmatrix} \cdot \begin{bmatrix} 0.1 & 0.9 \\ 0.3 & 0.7 \end{bmatrix}$$

$$= \begin{matrix} 1 & 2 \\ 2 & \end{matrix} \begin{bmatrix} & 0.72 \\ & \end{bmatrix}$$

Q3

a)

	(S,S)	(S,C)	(C,S)	(C,C)
(S,S)	0.7	0.3	0	0
(S,C)	0	0	0.4	0.6
(C,S)	0.5	0.5	0	0
(C,C)	0	0	0.2	0.8

$\downarrow \pi_0$ $\downarrow \pi_1$ $\downarrow \pi_2$ $\downarrow \pi_3$

Q3 b) limiting distribution is $\pi = (\pi_0, \pi_1, \pi_2, \pi_3)$

$$\begin{cases} 0.7\pi_0 + 0.5\pi_2 = \pi_0 \Rightarrow \pi_2 = 0.6\pi_0 \\ 0.3\pi_0 + 0.5\pi_2 = \pi_1 \Rightarrow \pi_1 = 0.6\pi_0 \\ 0.6\pi_1 + 0.8\pi_3 = \pi_3 \Rightarrow \pi_3 = 1.8\pi_0 \\ \text{f } \pi_0 + \pi_1 + \pi_2 + \pi_3 = 1 \end{cases}$$

$$\Rightarrow \pi_0 = \frac{1}{4} = 0.25$$

$$\Rightarrow \pi = (0.25, 0.15, 0.15, 0.45)$$

\therefore The long run fraction of days in which it is cloudy is $\pi_2 + \pi_3 = 0.15 + 0.45 = 0.60$

f The long run fraction of days in which it is sunny is $\pi_0 + \pi_1 = 0.25 + 0.15 = 0.40$

Q4

a) $\lambda = 2$, $X(t) = \#$ messages that arrives at the telegraph office at anytime.

$$P\left\{ \frac{X(12) - X(8)}{4} = \frac{1}{k} \right\} = \frac{(2 \cdot 4)^1 e^{-2 \cdot 4}}{1!}$$

$$= 8 \cdot e^{-8}$$

$$\boxed{Q_4} \quad b) \Pr\{X(2) = 5\}$$

$$\Pr\{X(2) - X(0) = 5\}$$

$$\mu = \int_0^2 \lambda(u) du$$

$$= \int_0^1 (2t+1) dt + \int_1^2 3 dt$$

$$= \left. \frac{2t^2}{2} + t \right|_0^1 + \left. 3t \right|_1^2$$

$$= 2 + 3 = 5$$

$$\therefore \Pr\{X(2) = 5\} = \Pr\{X(2) - X(0) = 5\}$$

$$= \frac{\mu^k e^{-\mu}}{k!} = \frac{5^5 e^{-5}}{5!}$$

$$\begin{aligned}
 \boxed{Q5} \quad a) (i) P_r \{X_2=1\} &= P_r \{X_2=1 | X_0=1\} \cdot P_r \{X_0=1\} \\
 &= P_{11}^2 \cdot P_1 \quad (P_1 = P_r \{X_0=1\} = 1) \\
 &= P_{11}^2 = 0.22
 \end{aligned}$$

$$P^2 = \begin{bmatrix} 0.49 & 0.28 & 0.23 \\ 0.43 & 0.22 & 0.35 \\ 0.47 & 0.20 & 0.33 \end{bmatrix}$$

(ii) To find $\bar{\pi} = (\pi_0, \pi_1, \pi_2)$

Solve: ~~Find π_0, π_1, π_2~~

$$\begin{cases} 4\pi_0 - 3\pi_1 - 4\pi_2 = 0 \\ 3\pi_0 - 7\pi_1 + \pi_2 = 0 \\ \pi_0 + \pi_1 + \pi_2 = 1 \end{cases}$$

by Cramer's Rule

$$\Delta = -66, \quad \Delta_0 = -31, \quad \Delta_1 = -16, \quad \Delta_2 = -19$$

$$\therefore \pi_0 = \frac{\Delta_0}{\Delta} = \frac{31}{66}, \quad \pi_1 = \frac{\Delta_1}{\Delta} = \frac{16}{66}, \quad \pi_2 = \frac{19}{66}$$

$$(iii) \pi_2 = \frac{19}{66} \approx 0.2879$$

Q6

$$a) \Pr \{B(4) > a \mid B(0) = 1\} = 0.1587$$

$$\Rightarrow 1 - \Pr \{B(4) \leq a \mid B(0) = 1\} = 0.1587$$

$$\Rightarrow 1 - \Phi \left(\frac{a-1}{\sqrt{4}} \right) = 0.1587$$

$$\Rightarrow \Phi \left(\frac{a-1}{2} \right) = 0.8413$$

$$\Rightarrow \frac{a-1}{2} = 1$$

using
Normal
tables

$$\Rightarrow a-1 = 2$$

$$\Rightarrow \boxed{a=3}$$

b) $X(t) = cB\left(\frac{t}{c^2}\right)$, $c > 0$, is a BM.

[1] show $X(t) - X(s) \sim N(0, t-s)$:

$$\begin{aligned} X(t) - X(s) &= cB\left(\frac{t}{c^2}\right) - cB\left(\frac{s}{c^2}\right) \\ &= c\left[B\left(\frac{t}{c^2}\right) - B\left(\frac{s}{c^2}\right)\right] \\ &\sim cN\left(0, \frac{t}{c^2} - \frac{s}{c^2}\right) \end{aligned}$$

$$\sim N\left(c \cdot 0, e^2 \cdot \frac{t-s}{c^2}\right) \quad \text{option BM}$$

$$\Rightarrow X(t) - X(s) \sim N(0, t-s) \quad \checkmark$$

[2] show $\text{Cov}[X(t), X(s)] = \min(s, t)$:

Case (1) $0 \leq s \leq t$

Since $B(t)$ is BM, $\text{Cov}[B(s), B(t)] = \min(s, t)$
 $\forall s, t > 0$

as $0 \leq s \leq t$

$$\Rightarrow 0 \leq \frac{s}{c^2} \leq \frac{t}{c^2}$$

$$\text{So } \text{Cov}\left[B\left(\frac{s}{c^2}\right), B\left(\frac{t}{c^2}\right)\right] = \frac{s}{c^2}$$

$$\begin{aligned} \left. \begin{array}{l} \text{from } \text{Cov}(aX, bY) \\ = a \text{Cov}(X, Y) \end{array} \right\} &\Rightarrow c^2 \cdot \text{Cov}\left[B\left(\frac{s}{c^2}\right), B\left(\frac{t}{c^2}\right)\right] = s \\ &\Rightarrow \text{Cov}\left[c \cdot B\left(\frac{s}{c^2}\right), B\left(\frac{t}{c^2}\right)\right] = s \end{aligned}$$

$$= \text{Cov}[X(s), X(t)] = s = \min(s, t)$$

Similarly, Case (2) $0 \leq t \leq s$.

$$[3] E[X(t)] = E\left[cB\left(\frac{t}{c^2}\right)\right] = 0$$

as $\checkmark N\left(0, \frac{t}{c^2}\right)$

$$[1] \wedge [2] \wedge [3] \Rightarrow X(t) \text{ is a BM.}$$

c) $X(t) = e^{uB(t) - \frac{u^2}{2}t}$ is a martingale.

□ $E |X(t)| < \infty$?

$$E |e^{uB(t) - \frac{u^2}{2}t}| = E \left[e^{uB(t) - \frac{u^2}{2}t} \right]$$

$$= e^{-\frac{u^2}{2}t} E \left[e^{uB(t)} \right]$$

as $B(t) \sim N(0, t) \Rightarrow uB(t) \sim N(0, u^2 t)$

$$E \left[e^{uB(t)} \right] = e^{0 + \frac{u^2 t}{2}} = e^{\frac{u^2}{2}t}$$

$$B(t) - B(s) \sim N(0, t-s)$$

$$\Rightarrow E \left[e^{u(B(t)-B(s))} \right] = e^{\frac{u^2}{2}(t-s)}$$

$$\therefore E |X(t)| = e^{-\frac{u^2}{2}t} \cdot e^{\frac{u^2}{2}t} = e^0 = 1 < \infty$$

□ For $0 \leq s \leq t$, $E[X(t) | \mathcal{F}_s] = X(s)$?

$$\text{L.H.S.} = E \left[e^{uB(t) - \frac{u^2}{2}t} \mid \mathcal{F}_s \right]$$

$$= e^{-\frac{u^2}{2}t} E \left[e^{uB(t)} \mid \mathcal{F}_s \right]$$

$$= e^{-\frac{u^2}{2}t} E \left[e^{u(B(t)-B(s)) + uB(s)} \mid \mathcal{F}_s \right]$$

$$= e^{-\frac{u^2}{2}t} E \left[e^{u(B(t)-B(s))} \cdot e^{uB(s)} \mid \mathcal{F}_s \right]$$

$$= e^{uB(s) - \frac{u^2}{2}t} E \left[e^{u(B(t)-B(s))} \mid \mathcal{F}_s \right]$$

determined by \mathcal{F}_s

$$= e^{uB(s) - \frac{u^2}{2}t} E \left[e^{u(B(t)-B(s))} \right]$$

independent of \mathcal{F}_s

$$= e^{uB(s) - \frac{u^2}{2}t} \frac{e^{\frac{u^2}{2}(t-s)}}{e^{\frac{u^2}{2}(t-s)}} = e^{uB(s) - \frac{u^2}{2}s} = X(s) = \text{R.H.S.}$$

□ $\&$ □ $\Rightarrow X(t)$ is a martingale.