

## 3.5.6 Scheffé's Method for Comparing All Contrasts

In many situations, experimenters may not know in advance which contrasts they wish to compare, or they may be interested in more than  $(a - 1)$  possible comparisons. In many exploratory experiments, the comparisons of interest are discovered only after preliminary examination of the data. Scheffé (1953) has proposed a method for comparing any and all possible contrasts between treatment means. In the Scheffé method, the type I error is at most for any of the possible comparisons

## ■ The Steps for conducting Scheffé method.

Suppose that a set of  $m$  contrasts in the treatment means

1- Form the contrast number  $u$ ,  $u = 1, 2, \dots, m$

$$\Gamma_u = c_{1u}\mu_1 + c_{2u}\mu_2 + \dots + c_{au}\mu_a = \sum_{i=1}^a c_{iu}\mu_i$$

2- Compute the point estimate to contrast  $\Gamma_u$  .

$$\hat{\Gamma}_u = C_u = \sum_{i=1}^a c_{iu}\bar{y}_i.$$

where  $C_u$  is unbiased estimate to  $\Gamma_u$

3- The standard error of this contrast is

$$S_{C_u} = \sqrt{\frac{MSE}{n} \sum_{i=1}^a c_{iu}^2} \quad \text{Balance} \qquad S_{C_u} = \sqrt{MSE \sum_{i=1}^a \left(\frac{c_{iu}^2}{n_i}\right)} \quad \text{Unbalance}$$

4- The **critical value** against which  $C_u$  should be compared is

$$S_{\alpha,u} = S_{C_u} \sqrt{(a-1)F_{\alpha,a-1,N-a}}$$

5- If  $|C_u| > S_{\alpha,u}$  the null hypothesis  $H_0: \Gamma_u = 0$  is rejected

□ **Example 1** : The response time in milliseconds was determined for three different types of circuits that could be used in an automatic valve shutoff mechanism. The results from a completely randomized experiment are shown in the following table.

Circuit Type		Response Time			
1	9	12	10	8	15
2	20	21	23	17	30
3	6	5	8	16	7

- Use Scheffé method to compare type 2 with the other two.
- Verify the results using the R program.

# The Answer of Example 1 (a)

**The contrast is:**  $\mu_2 = \frac{\mu_1 + \mu_3}{2} \rightarrow \Gamma = -\mu_1 + 2\mu_2 - \mu_3 = 0, \quad c_i = (-1), (+2), (-1)$

- **Create ANOVA table**

<b>Circuite Type</b>	response Time					$y_{i\cdot}$	$\bar{y}_{i\cdot}$	$c_i$
<b>1</b>	<b>9</b>	<b>12</b>	<b>10</b>	<b>8</b>	<b>15</b>	<b>54</b>	<b>10.8</b>	<b>-1</b>
<b>2</b>	<b>20</b>	<b>21</b>	<b>23</b>	<b>17</b>	<b>30</b>	<b>111</b>	<b>22.2</b>	<b>+2</b>
<b>3</b>	<b>6</b>	<b>5</b>	<b>8</b>	<b>16</b>	<b>7</b>	<b>42</b>	<b>8.4</b>	<b>-1</b>

- **Hypothesis  $H_0, H_1$**

$$H_0: \Gamma = 0 \quad vs \quad H_1: \Gamma \neq 0$$

<i>S.O.V</i>	<i>SS</i>	<i>df</i>	<i>MS</i>	<i>F</i>	<i>F crit</i>	<i>P-value</i>
<b>Circuite Type</b>	543.6	2	271.8	16.0828	3.8853	0.0004
<b>Error</b>	202.8	12	16.9			
<b>Total</b>	746.4	14				

- **Point estimate of Contrast  $\Gamma$**

$$C = \sum_{i=1}^3 c_i \bar{y}_{i\cdot} = (-1)\bar{y}_{1\cdot} + (2)\bar{y}_{2\cdot} + (-1)\bar{y}_{3\cdot} = (-1)(10.8) + (2)(22.2) + (-1)(8.4) = 25.2.$$

- **The standard error of  $C$ :**  $S_C = \sqrt{\frac{MSE}{n} \sum_{i=1}^3 c_i^2} = \sqrt{\frac{16.9}{5} (6)} = \sqrt{20.2} = 4.494$

- **Critical value:**  $S_\alpha = S_C \sqrt{(a-1)F_{\alpha, a-1, N-a}} = 4.494 \sqrt{(2)(3.8853)} = 12.527$

- **Decision :**  $C = 25.2 > S_\alpha = 12.527$ , **The null hypothesis  $H_0: \Gamma = 0$  is rejected**

## The Answer of Example 1 (b)

# Input data

```
y=c(9,12,10,8,15,20,21,23,17,30,6,5,8,16,7)
Type = factor(rep(c("Type1","Type2","Type3"),each=5))
data1 = data.frame(y,Type)
```

#Run Model

```
Model = aov(y~Type, data = data1); summary(Model)
```

# Scheffé's Method to contrast (mean of Type2- mean of Type1 & Type3)

```
install.packages("DescTools")
library(DescTools)
```

```
# Contrast: (mean of Type2- mean of Type1 & Type3)
```

```
# Vector: -Type1 + 2Type2 -Type3
```

```
c.i <- matrix(c(-1,2,-1), ncol = 1)
Cont1 = ScheffeTest(Model, contrasts = c.i)
print(Cont1)
```

## Output

```
> Model = aov(y~Type, data = data1); summary(Model)
              Df Sum Sq Mean Sq F value    Pr(>F)
Type           2  543.6   271.8   16.08 0.000402 ***
Residuals     12  202.8    16.9
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
>
> Cont1 = ScheffeTest(Model, contrasts = c.i)
>
> print(Cont1)

Posthoc multiple comparisons of means: Scheffe Test
 95% family-wise confidence level

$Type
              diff   lwr.ci   upr.ci   pval
Type2-Type1,Type3 25.2 12.64661 37.75339 0.00045 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

We note that  $p \text{ value} = 0.00045 < 0.05$  we reject null hypothesis  $H_0: \Gamma = 0$

## 3.5.7 Comparing Pairs of Treatment Means

In many practical situations, we will wish to compare only *pairs of means* ( $\mu_i - \mu_j$ ). Frequently, we can determine which means differ by testing the differences between all pairs  $\left[\frac{a(a-1)}{2}\right]$  of treatment means. Suppose that we are interested in comparing all pairs of a treatment means and that the null hypotheses that we wish to test are  $H_0: \mu_i = \mu_j$ , for all  $i \neq j$ . We now present two popular methods for making such comparisons.

### ▪ Tukey's Test.

When the null hypothesis of equal treatment means  $H_0: \mu_1 = \mu_2 = \dots = \mu_a$ , is rejected, we wish to test all pairwise mean comparisons  $H_0: \mu_i = \mu_j$  vs  $H_1: \mu_i \neq \mu_j$

The steps of **Tukey** procedure as follows

1. At significant level  $\alpha$ , determine critical value  $q_\alpha(a, f)$  from **Appendix Table VII**, where  $a = \text{number of treatments}$ ,  $f = \text{error degrees of freedom}$
2. Compute  $T_\alpha$ , which given by

$$T_\alpha = q_\alpha(a, f) \sqrt{\frac{MSE}{n}} \quad (3.35)$$

3. If the absolute difference  $|\bar{y}_i. - \bar{y}_j.| > T_\alpha$  null hypothesis  $H_0: \mu_i = \mu_j$  is rejected and it concluded that existing significant difference between mean  $\mu_i$  and mean  $\mu_j$

□ **Example 2:** Use the data in example 1, to conduct comparing pairs of treatment means by Tukey's test.

a) **Manually**

b) **R program**

## The Answer

<i>S.O.V</i>	<i>SS</i>	<i>df</i>	<i>MS</i>	<i>F</i>	<i>F crit</i>	<i>P-value</i>
<b>Circuite Type</b>	543.6	2	271.8	16.0828	3.8853	0.0004
<b>Error</b>	202.8	12	16.9			
<b>Total</b>	746.4	14				

a) **Manually**

- Compute  $T_\alpha$ , which given by:

$$T_\alpha = q_\alpha(a, f) \sqrt{\frac{MSE}{n}} = q_{0.05}(3, 12) \sqrt{\frac{16.9}{5}} = 3.77(1.8385) = 6.93$$

$\bar{y}_i$	$\bar{y}_j$	$ \bar{y}_i - \bar{y}_j $	$T_\alpha$	$H_0$	Decision
10.8	22.2	11.4	6.93	$H_0: \mu_1 = \mu_2$	$H_0$ is rejected
10.8	8.4	2.4	6.93	$H_0: \mu_1 = \mu_3$	$H_0$ is not rejected
22.2	8.4	13.8	6.93	$H_0: \mu_2 = 3$	$H_0$ is rejected

## b) R program

```
# Input libraries
```

```
library(car); library(multcomp); library(DescTools)
```

```
# Input data
```

```
y=c(9,12,10,8,15,20,21,23,17,30,6,5,8,16,7)
```

```
Type = factor(rep(c("Type1", "Type2", "Type3"), each=5))
```

```
data1 = data.frame(y, Type)
```

```
#Run Model
```

```
Model = aov(y~Type, data = data1); summary(Model)
```

```
# TukeyHSD Method to pairwise comparisons
```

```
TukeyHSD(Model)
```

```
> TukeyHSD(Model)
```

```
Tukey multiple comparisons of means
```

```
95% family-wise confidence level
```

```
Fit: aov(formula = y ~ Type, data = data1)
```

```
$Type
```

	diff	lwr	upr	p adj
Type2-Type1	11.4	4.463555	18.336445	0.0023656
Type3-Type1	-2.4	-9.336445	4.536445	0.6367043
Type3-Type2	-13.8	-20.736445	-6.863555	0.0005042

▪ **The Fisher Least Significant Difference (*LSD*) Method.**

The steps of Least Significant Difference (*LSD*) for testing all pairwise mean comparisons  $H_0: \mu_i = \mu_j$  vs  $H_1: \mu_i \neq \mu_j$  as follows

1. At significant level  $\alpha$ , compute the value (*LSD*), which given by

$$LSD = t_{\frac{\alpha}{2}, N-a} \sqrt{\frac{2MSE}{n}} \quad \text{or} \quad LSD = t_{\frac{\alpha}{2}, N-a} \sqrt{MSE \left( \frac{1}{n_i} + \frac{1}{n_j} \right)}$$

2. If the absolute difference  $|\bar{y}_i. - \bar{y}_j.| > LSD$  null hypothesis  $H_0: \mu_i = \mu_j$  is rejected and it concluded that existing significant difference between mean  $\mu_i$  and mean  $\mu_j$

□ **Example 3:** Use the data in example 1, to conduct comparing pairs of treatment means by *LSD*.

- a) **Manually**                      b) **R program**

=====

## The Answer

### a) Manually

Compute *LSD*, which given by:

$$LSD = t_{\frac{\alpha}{2}, N-a} \sqrt{\frac{2MSE}{n}} = t_{\frac{0.05}{2}, 12} \sqrt{\frac{2 \times 16.9}{5}} = 2.179(2.6) = 5.665$$

$\bar{y}_i$	$\bar{y}_j$	$ \bar{y}_i - \bar{y}_j $	<i>LSD</i>	$H_0$	Decision
10.8	22.2	11.4	5.665	$H_0: \mu_1 = \mu_2$	$H_0$ is rejected
10.8	8.4	2.4	5.665	$H_0: \mu_1 = \mu_3$	$H_0$ is not rejected
22.2	8.4	13.8	5.665	$H_0: \mu_2 = 3$	$H_0$ is rejected

### b) R program

```
library(car); library(multcomp); library(DescTools); library(agricolae)
y=c(9,12,10,8,15,20,21,23,17,30,6,5,8,16,7)
Type = factor(rep(c("Type1", "Type2", "Type3"), each=5))
data1 = data.frame(y, Type)
Model = aov(y~Type, data = data1); summary(Model)
print(LSD.test(Model, trt = "Type"))
```

## The output

```
> print(LSD.test(Model, trt = "Type"))
$statistics
  MSerror Df Mean      CV  t.value      LSD
  16.9  12  13.8 29.78957 2.178813 5.664913

$parameters
      test p.adjusted name.t ntr alpha
Fisher-LSD      none   Type   3  0.05

$means
      y      std r      se      LCL      UCL Min Max Q25 Q50 Q75
Type1 10.8 2.774887 5 1.838478  6.794301 14.8057   8  15   9  10  12
Type2 22.2 4.868265 5 1.838478 18.194301 26.2057  17  30  20  21  23
Type3  8.4 4.393177 5 1.838478  4.394301 12.4057   5  16   6   7   8

$comparison
NULL

$groups
      y groups
Type2 22.2    a
Type1 10.8    b
Type3  8.4    b
```

Means with the same letters are not significantly different; therefore, there is a significant difference between the mean of type 2 and both types 1 and 3, while there is no significant difference between the means of type 1 and 3.

## 3.5.8 Comparing Treatment Means with a Control

In many experiments, one of the treatments is a control, and the analyst is interested in comparing each of the other  $a - 1$  treatment means with the control. A procedure for making these comparisons has been developed by **Dunnnett** (1964). Suppose that treatment  $a$  is the control and we wish to test the hypotheses

$$H_0: \mu_i = \mu_a \text{ vs } H_1: \mu_i \neq \mu_a$$

The null hypothesis  $H_0: \mu_i = \mu_a$  is rejected using a type I error rate  $\alpha$  if

$$|\bar{y}_{i\cdot} - \bar{y}_{a\cdot}| > d_\alpha(a - 1, f) \sqrt{MSE \left( \frac{1}{n_i} + \frac{1}{n_a} \right)}$$

where the constant  $d_\alpha(a - 1, f)$  is given in **Appendix Table VIII**. (Both two- and one-sided tests are possible). Note that  $\alpha$  is the joint significance level associated with all  $a - 1$  tests.

□ **Example 4:** Use the data in example 1, to conduct comparing mean of type 3 as a control to mean of each type 1 and type 2

a) **Manually**

b) **R program**

**The Answer**

a) **Manually**

From Appendix Table VIII:  $d_{\alpha}(a - 1, f) = d_{0.05}(2, 12) = 2.5$

$$n_1 = n_2 = n_3 = 5, \sqrt{MSE \left( \frac{1}{n_i} + \frac{1}{n_3} \right)} = \sqrt{16.9 \left( \frac{1}{5} + \frac{1}{5} \right)} = 2.6$$

$$D_i = 2.5 \times 2.6 = 6.5$$

$\bar{y}_i$	$\bar{y}_i$	$ \bar{y}_i - \bar{y}_j $	$D_i$	$H_0$	Decision
10.8	$\bar{y}_3$				
22.2	8.4	2.4	6.5	$H_0: \mu_1 = \mu_3$	$H_0$ is not rejected
8.4	8.4	13.8	6.5	$H_0: \mu_2 = \mu_3$	$H_0$ is rejected

## b) R program

```
y <- c(9, 12, 10, 8, 15, 20, 21, 23, 17, 30, 6, 5, 8, 16, 7)
  Type <- factor(rep(c("Type1", "Type2", "Type3"), each = 5))
  data1 <- data.frame(y, Type)
# Run the ANOVA model using aov()
  Model <- aov(y ~ Type, data = data1); summary(Model)
#Dunnett test
  # Make Type3 control
  data1$Type <- relevel(data1$Type, ref = "Type3")
  library(multcomp)
  Result_MC <- glht(Model, mcp(Type = "Dunnett")); summary(Result_MC )
```

```
> summary(Result_MC )
```

### Simultaneous Tests for General Linear Hypotheses

#### Multiple Comparisons of Means: Dunnett Contrasts

```
Fit: aov(formula = y ~ Type, data = data1)
```

#### Linear Hypotheses:

	Estimate	Std. Error	t value	Pr(> t )
Type1 - Type3 == 0	2.4	2.6	0.923	0.569159
Type2 - Type3 == 0	13.8	2.6	5.308	0.000354 ***

---

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Adjusted p values reported -- single-step method)
```