King Saud University<br>College of Sciences<br>Department of Mathematics

Final Exam, S2-1442H
ACTU 475
Credibility Theory and Loss Distributions.
Time: 3 hours - Marks: 20

## Answer the following questions:

(Note that SND Table is attached in page 3)

## Q1: [6+3]

(a) If the random variable $X$ has probability density function $f(x)=\left(1+2 x^{2}\right) e^{-2 x}, x \geq 0$.
(i) Determine the survival function.
(ii) Determine the hazard rate function.
(iii) Determine the mean excess loss function.
(b) The cdf of a random variable is $F(x)=1-x^{-2}, x \geq 1$. Determine the mean, median, and mode of this random variable.

Q2: [4+2]
(a) Let $X \mid \Lambda$ have a Poisson distribution with parameter $\Lambda$. Let $\Lambda$ have a Gamma distribution with parameters $\alpha$ and $\theta$. Determine the expectation of $\Lambda$ and then the unconditional expectation of $X$.
(b) Consider a frailty model with frailty random variable $\Lambda$, such that $a(x)=\frac{1}{x+1}, x>0$.

Find the conditional survival function of $X$.
Q3: $[5+5]$
(a) The average claim size for a group of insureds is 1500, with a standard deviation of 7500 . Assume that claim counts have the Poisson distribution. Determine the expected number of claims so that the total loss will be within $5 \%$ of the expected total loss with probability 0.90 .
(b) A group of insureds had 6000 claims and a total loss of 15,600,000. The prior estimate of the total loss was $16,500,000$. Determine the limited fluctuation credibility estimate of the total loss for the group. Use the standard for full credibility determined in (a).

Q4: [5+5]
(a) There are two types of drivers. Good drivers make up $70 \%$ of the population and in one year have zero claims with probability 0.6 , one claim with probability 0.3 , and two claims with probability 0.1 . Bad drivers
make up the other $30 \%$ of the population and have zero, one, or two claims with probabilities $0.4,0.3$, and 0.3 , respectively.
(i) Describe this process by using the concept of the risk parameter $\Theta$.
(ii) For a particular policyholder, suppose that we have observed $x_{1}=0$ and $x_{2}=1$ for past claims. Determine the posterior distribution of $\Theta \mid X_{1}=0, X_{2}=1$ and the predictive distribution of $X_{3} \mid X_{1}=0, \mathrm{X}_{2}=1$.
(b)

Risk 1 produces claims of amounts $100,1,000$, and 20,000 with probabilities $0.5,0.3$, and 0.2 , respectively. For risk 2, the probabilities are $0.7,0.2$ and 0.1 . Risk 1 is twice as likely as risk 2 of being observed. A claim of 100 is observed, but the observed risk is unknown. Determine the Buhlmann credibility estimate of the expected value of the second claim amount from the same risk.

Q5: [5]
Polices have a deductible of 7 . Six losses are observed, with value $10,13,14,15,17,19$. Ground-up losses have an exponential distribution with parameter $\theta$. Determine the maximum likelihood estimate of $\theta$.

Standard Normal Cumulative Probability Table

Cumulative probabilltles for POSITIVE z-values are shown in the following table:


| z | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 | 0.06 | 0.07 | 0.08 | 0.09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 | 0.5239 | 0.5279 | 0.5319 | 0.5359 |
| 0.1 | 0.5396 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 | 0.5636 | 0.5675 | 0.5714 | 0.5753 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 | 0.6026 | 0.6064 | 0.6103 | 0.6141 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 | 0.6406 | 0.6443 | 0.6480 | 0.6517 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 | 0.6772 | 0.6808 | 0.6844 | 0.6879 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 | 0.7123 | 0.7157 | 0.7190 | 0.7224 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 | 0.7454 | 0.7486 | 0.7517 | 0.7549 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7704 | 0.7734 | 0.7764 | 0.7794 | 0.7823 | 0.7852 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 | 0.8051 | 0.8078 | 0.8106 | 0.8133 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 | 0.8315 | 0.8340 | 0.8365 | 0.8389 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 | 0.8554 | 0.8577 | 0.8599 | 0.8621 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 | 0.8770 | 0.8790 | 0.8810 | 0.8830 |
| 1.2 | 0.8849 | 0.8069 | 0.8888 | 0.8907 | 0.8925 | 0.8944 | 0.8962 | 0.8960 | 0.8997 | 0.9015 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 | 0.9115 | 0.9131 | 0.9147 | 0.9162 | 0.9177 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 | 0.9265 | 0.9279 | 0.9292 | 0.9306 | 0.9319 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 | 0.9394 | 0.9406 | 0.9418 | 0.9429 | 0.9441 |
| 1.6 | 0.9452 | 0.9453 | 0.9474 | 0.9484 | 0.9495 | 0.9505 | 0.9515 | 0.9525 | 0.9535 | 0.9545 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 | 0.9599 | 0.9608 | 0.9616 | 0.9625 | 0.9633 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 | 0.9678 | 0.9686 | 0.9693 | 0.9699 | 0.9706 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 | 0.9744 | 0.9750 | 0.9756 | 0.9761 | 0.9767 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 | 0.9798 | 0.9803 | 0.9808 | 0.9812 | 0.9817 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 | 0.9842 | 0.9846 | 0.9850 | 0.9854 | 0.9857 |
| 2.2 | 0.9861 | 0.9864 | 0.9668 | 0.9871 | 0.9875 | 0.9878 | 0.9881 | 0.9884 | 0.9887 | 0.9890 |
| 2.3 | 0.9893 | 0.9896 | 0.9698 | 0.9901 | 0.9904 | 0.9906 | 0.9909 | 0.9911 | 0.9913 | 0.9916 |
| 2.4 | 0.9918 | 0.9920 | 0.9922 | 0.9925 | 0.9927 | 0.9929 | 0.9931 | 0.9932 | 0.9934 | 0.9936 |
| 2.5 | 0.9938 | 0.9940 | 0.9941 | 0.9943 | 0.9945 | 0.9946 | 0.9948 | 0.9949 | 0.9951 | 0.9952 |
| 2.6 | 0.9953 | 0.9955 | 0.9956 | 0.9957 | 0.9959 | 0.9960 | 0.9961 | 0.9962 | 0.9963 | 0.9964 |
| 2.7 | 0.9965 | 0.9966 | 0.9967 | 0.9968 | 0.9969 | 0.9970 | 0.9971 | 0.9972 | 0.9973 | 0.9974 |
| 2.8 | 0.9974 | 0.9975 | 0.9976 | 0.9977 | 0.9977 | 0.9978 | 0.9979 | 0.9979 | 0.9980 | 0.9981 |
| 2.9 | 0.9981 | 0.9982 | 0.9982 | 0.9983 | 0.9984 | 0.9984 | 0.9985 | 0.9985 | 0.9986 | 0.9986 |
| 3.0 | 0.9987 | 0.9987 | 0.9987 | 0.9988 | 0.9988 | 0.9989 | 0.9989 | 0.9969 | 0.9990 | 0.9990 |
| 3.1 | 0.9990 | 0.9991 | 0.9991 | 0.9991 | 0.9992 | 0.9992 | 0.9992 | 0.9992 | 0.9993 | 0.9983 |
| 3.2 | 0.9993 | 0.9993 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9994 | 0.9995 | 0.9995 | 0.9995 |
| 3.3 | 0.9995 | 0.9995 | 0.9995 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9996 | 0.9997 |
| 3.4 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9997 | 0.9998 |

## The Model Answer

Q1: [6+3]
(a)
(i)

The survival function is

$$
\begin{aligned}
& S(x)=\int_{x}^{\infty}\left(1+2 t^{2}\right) e^{-2 t} d t \\
& \quad=-\frac{1}{2} e^{-2 t}+2 I, \text { where } I=\int_{x}^{\infty} t^{2} e^{-2 t} d t \\
& I=\int t^{2} e^{-2 t} d t \\
& \quad=-\frac{1}{4} e^{-2 t}-\frac{t}{2} e^{-2 t}-\frac{t^{2}}{2} e^{-2 t} \\
& \therefore S(x)=-\left.\left(1+t+t^{2}\right) e^{-2 t}\right|_{x} ^{\infty} \\
& \quad=\left(1+x+x^{2}\right) e^{-2 x}, x \geq 0
\end{aligned}
$$

(ii)
$\because$ The hazard rate function is
$h(x)=-\frac{d}{d x}[\ln S(x)]$
and $\because S(x)=\left(1+x+x^{2}\right) e^{-2 x}$
$\ln S(x)=-2 x+\ln \left(1+x+x^{2}\right)$
$\therefore h(x)=2-\frac{1+2 x}{1+x+x^{2}}$
or simply,
$h(x)=\frac{f(x)}{S(x)}=\frac{1+2 x^{2}}{1+x+x^{2}}$
(iii)

The mean excess loss function is
$e_{X}(x)=\frac{\int_{x}^{\infty} S(t) d t}{S(x)}$
From (i) $S(x)=\left(1+x+x^{2}\right) e^{-2 x}$

We can deduce that,

$$
\begin{align*}
\int_{x}^{\infty} S(t) d t & =\int_{x}^{\infty}\left(1+t+t^{2}\right) e^{-2 t} d t \\
& =-\left.\left(1+t+\frac{1}{2} t^{2}\right) e^{-2 t}\right|_{x} ^{\infty} \\
& =\left(1+x+\frac{1}{2} x^{2}\right) e^{-2 x} \tag{3}
\end{align*}
$$

Where $I=\int t^{2} e^{-2 t} d t=-\frac{1}{4} e^{-2 t}-\frac{t}{2} e^{-2 t}-\frac{t^{2}}{2} e^{-2 t}$, $\int t e^{-2 t} d t=-\frac{1}{4} e^{-2 t}-\frac{t}{2} e^{-2 t}$ and $\int e^{-2 t} d t=-\frac{1}{2} e^{-2 t}$
$\therefore$ By substituting (2) and (3) in (1), we get
$e_{X}(x)=\frac{1+x+\frac{1}{2} x^{2}}{1+x+x^{2}}$
(b)

The pdf is $f(x)=2 x^{-3}, x \geq 1$.

The mean is

$$
\begin{aligned}
E(X) & =\int_{1}^{\infty} 2 x^{-2} d x \\
& =2
\end{aligned}
$$

To get the median, solve the equation $F(x)=1-x^{-2}=0.5$
$\Rightarrow$ The median $\simeq 1.4142$

The mode is the value at which the pdf is highest, so to get the mode,
$\because f(x)=2 x^{-3}, x \geq 1$ is a decreasing function and its highest value at $x=1$
$\therefore$ The mode $=1$
Q2: [4+2]
(a)
$\because \Lambda \sim \operatorname{Gamma}(\alpha, \theta)$
$\therefore f_{\Lambda}(\lambda)=\frac{(\lambda / \theta)^{\alpha} e^{-\lambda / \theta}}{\lambda \Gamma(\alpha)}$
$\therefore E(\Lambda)=\alpha \theta$
$\because X \mid \Lambda \sim \operatorname{Poisson}(\Lambda)$
$\therefore E(X \mid \Lambda)=\Lambda$
$\because E(X)=E[E(X \mid \Lambda)]=E(\Lambda)$
$\therefore E(X)=\alpha \theta=E(\Lambda)$
Note: Also, you can obtain this result by using the formula

$$
f_{X}(x)=\int_{0}^{\infty} f_{X \mid \Lambda}(x \mid \lambda) f_{\Lambda}(\lambda) d \lambda \text { where } f_{X \mid \Lambda}(x \mid \lambda)=\frac{e^{-\lambda} \lambda^{x}}{x!}
$$

(b)

We first find $A(x)$

$$
\begin{aligned}
A(x) & =\int_{0}^{x} a(t) d t \\
& =\int_{0}^{x} \frac{d t}{1+t}=\ln (1+x)
\end{aligned}
$$

Thus,

$$
\begin{aligned}
S_{X \mid \Lambda}(x \mid \lambda) & =e^{-\lambda A(x)} \\
& =e^{-\lambda \ln (1+x)} \\
& =\frac{1}{(1+x)^{\lambda}}
\end{aligned}
$$

## Q3: $[5+5]$

(a)
at $p=0.90, \Phi\left(y_{p}\right)=(1+p) / 2=0.95$
$\Rightarrow y_{p}=1.645$ (by using SND table)
$\Rightarrow \lambda_{0}=\left(y_{p} / r\right)^{2}=(1.645 / 0.05)^{2}=1082.41$

To get the expected number of claims, use the following formula:

$$
n \lambda=\lambda_{0}\left[1+\left(\frac{\sigma}{\theta}\right)^{2}\right]
$$

where $\sigma^{2}=7500^{2}, \theta=1500$
$\therefore$ The expected \# of claims $=1082.41\left[1+\left(\frac{7500}{1500}\right)^{2}\right]$

$$
=28142.66
$$

(b)

The credibility factor is $Z=\sqrt{\frac{n}{\lambda_{0} \sigma^{2} / \xi^{2}}}$

$$
=\sqrt{\frac{6000}{28142.66}}=0.461735
$$

The partial credibility through premium is

$$
\begin{aligned}
P_{c}= & Z \bar{X}+(1-Z) M \\
& =0.461735(15600000)+(1-0.461735)(16500000) \\
& =16084438.5
\end{aligned}
$$

Q4: $[5+5]$
(a)
(i)

| $x$ | $\operatorname{Pr}(X=x \mid \Theta=G)$ | $\operatorname{Pr}(X=x \mid \Theta=B)$ | $\theta$ | $\operatorname{Pr}(\Theta=\theta)$ |
| :---: | :---: | :---: | :---: | :---: |
| 0 | 0.6 | 0.4 | $G$ | 0.7 |
| 1 | 0.3 | 0.3 | $B$ | 0.3 |
| 2 | 0.1 | 0.3 |  |  |

(ii)

For the posterior distribution, the posterior probabilities are given by

$$
\begin{aligned}
& \pi(G \mid 0,1)=\frac{f(0 \mid G) f(1 \mid G) \pi(G)}{f_{X}(0,1)} \\
& \text { where } f_{X}(0,1)=\sum_{\theta} f_{X_{1} \mid \Theta}(0 \mid \theta) f_{X_{2} \mid \Theta}(1 \mid \theta) \pi(\theta) \\
& f_{X}(0,1)=0.6(0.3)(0.7)+0.4(0.3)(0.3) \\
& \quad=0.162
\end{aligned} \quad \begin{aligned}
\pi(G \mid 0,1) & =\frac{0.6(0.3)(0.7)}{0.162}=0.7778 \\
\pi(B \mid 0,1) & =\frac{0.4(0.3)(0.3)}{0.162}=0.2222
\end{aligned}
$$

For the predictive distribution, the predictive probabilities are given by

$$
\begin{aligned}
f_{X_{3} \mid X}(0 \mid 0,1) & =\sum_{\theta} f(0 \mid \theta) \pi(\theta \mid 0,1) \\
& =f(0 \mid G) \pi(G \mid 0,1)+f(0 \mid B) \pi(B \mid 0,1) \\
& =0.6(0.7778)+0.4(0.2222) \\
& =0.55556 \\
f_{X_{3} \mid X}(1 \mid 0,1) & =\sum_{\theta} f(1 \mid \theta) \pi(\theta \mid 0,1) \\
& =f(1 \mid G) \pi(G \mid 0,1)+f(1 \mid B) \pi(B \mid 0,1) \\
& =0.3(0.7778)+0.3(0.2222) \\
& =0.3
\end{aligned}
$$

and $f_{X_{3} \mid X}(2 \mid 0,1)=\sum_{\theta} f(2 \mid \theta) \pi(\theta \mid 0,1)$

$$
\begin{aligned}
& =f(2 \mid G) \pi(G \mid 0,1)+f(2 \mid B) \pi(B \mid 0,1) \\
& =0.1(0.7778)+0.3(0.2222) \\
& =0.14444 .
\end{aligned}
$$

(b)

The required calculations are given in the following table.

| Risk | 100 | 1,000 | 20,000 | $\mu(\Theta)$ | $v(\Theta)$ | $\operatorname{Pr}(\Theta=\theta)$ |
| :--- | :--- | :--- | :--- | :---: | :---: | :---: |
| 1 | 0.5 | 0.3 | 0.2 | 4,350 | $61,382,500$ | $2 / 3$ |
| 2 | 0.7 | 0.2 | 0.1 | 2,270 | $35,054,100$ | $1 / 3$ |

To determine the Buhlmann credibility estimate of the expected value of the second claim amount from the same risk, we should find the following quantities.

$$
\begin{aligned}
\mu & =E[\mu(\Theta)] \\
& =(2 / 3)(4350)+(1 / 3)(2270)=3,656.667, \\
v & =E[v(\Theta)] \\
& =(2 / 3)(61,382,500)+(1 / 3)(35,054,100)=52,606,366.67, \\
a & =\operatorname{var}[\mu(\Theta)] \\
& =(2 / 3)(4350)^{2}+(1 / 3)(2270)^{2}-3,656.667^{2}=961,419.7845, \\
k & =\frac{v}{a} \\
& =54.71737, \\
Z & =\frac{n}{n+k} \\
& =\frac{1}{1+54.71737} \\
\therefore Z & =\frac{1}{55.71737}=0.0179477 .
\end{aligned}
$$

The Buhlmann estimate is

$$
\begin{aligned}
E\left(X_{2} \mid 100\right) & =P_{c}=Z \bar{X}+(1-Z) \mu \\
& =0.0179477(100)+(1-0.0179477)(3,656.667) \\
& =3,592.83 .
\end{aligned}
$$

where $X_{1}=100$
Q5: [5]

The likelihood function is

$$
\begin{aligned}
L(\theta) & =\prod_{j=1}^{6} \frac{f\left(x_{j} \mid \theta\right)}{1-F(7 \mid \theta)} \\
& =\frac{f(10 \mid \theta) f(13 \mid \theta) f(14 \mid \theta) f(15 \mid \theta) f(17 \mid \theta) f(19 \mid \theta)}{[1-F(7 \mid \theta)]^{6}} \\
& =\frac{\frac{1}{\theta^{6}} e^{-1 / \theta(10+13+14+15+17+19)}}{\left[e^{-7 / \theta}\right]^{6}} \\
\therefore \quad L(\theta) & =\frac{1}{\theta^{6}} e^{-46 / \theta}
\end{aligned}
$$

The loglikelihood function is
$l(\theta)=-\frac{46}{\theta}-6 \ln \theta$
To get $\hat{\theta}$, set $l^{\prime}(\theta)=0$
$\Rightarrow l^{\prime}(\theta)=\frac{46}{\theta^{2}}-\frac{6}{\theta}=0$
$\therefore \hat{\theta}=\frac{46}{6}$

$$
\simeq 7.6667
$$

Another Solution
$L(\theta)=\prod_{j=1}^{6} \frac{f\left(x_{j} \mid \theta\right)}{1-F(7 \mid \theta)}=\frac{\prod_{j=1}^{6} \frac{1}{\theta} e^{-x_{j} / \theta}}{\left(e^{-7 / \theta}\right)^{6}}$
The loglikelihood function is
$l(\theta)=-6 \ln \theta-\frac{1}{\theta} \sum_{j=1}^{6} x_{j}+6\left(\frac{7}{\theta}\right)$
To get $\hat{\theta}$, set $l^{\prime}(\theta)=0$
$\Rightarrow l^{\prime}(\theta)=-\frac{6}{\theta}+\frac{88}{\theta^{2}}-\frac{42}{\theta^{2}}=0$ where $\sum_{j=1}^{6} x_{j}=88$
$l^{\prime}(\theta)=-\frac{6}{\theta}+\frac{46}{\theta^{2}}=0$
$\therefore \hat{\theta}=\frac{46}{6}$
$\simeq 7.6667$

