بسم الله الرحمن الرحيم

STAT 105

Statistical Methods

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17 Dec 2023

CHAPTER 1

- Discrete Random Variables.
- Uniform Distribution.
- Binomial Distribution.
- Hyper geometric Distribution.
- Poisson Distribution.

Discrete Random Variables

- $0 \le f(x) \le 1$
- $\sum f(x) = 1$
- $f(x) = P(X = x)$
- $\mu_x = E(X) = \sum x\, f(x)$
- $\sigma_x^2 = Var(X) = E(X^2) - E(X)^2$
- $E(X^2) = \sum x^2 f(x)$
- $E(aX \pm b) = aE(X) \pm b$
- $Var(aX \pm b) = a^2 Var(X)$
- $F(x) = P(X \le x)$

Q1. Let the random variable X having the probability distribution (p.m.f) as:

x	-3	6	9
f(x)	1/6	1/2	1/3

(A)Find the probability that:

i. The random variable X assumes a non-negative value.

$$P(X = 6) + P(X = 9) = \frac{1}{2} + \frac{1}{3} = \frac{5}{6}$$

ii. The random variable X assumes a value less than 7.

$$P(X < 7) = P(X = -3) + P(X = 6) = \frac{1}{6} + \frac{1}{2} = \frac{4}{6}$$

iii. Calculate the cumulative distribution function (CDF)

$$F(X) = \begin{cases} 0 & x < -3 \\ 1/6 & -3 \le x < 6 \\ 2/3 & 6 \le x < 9 \\ 1 & x \ge 9 \end{cases}$$

iv. denoted by F, for $X = -4$, $X = -3$, $X = 0$, $X = 3$, $X = 6$, $X = 9$ and x more than 9

$F(-4) = 0$
$F(-3) = f(-3) = 1/6$
$F(0) \quad = f(X \le 0) = 1/6$
$F(3) \quad = f(X \le 3) = 1/6$
$F(6) \quad = f(X \le 6) = 1/6 + 1/2 = 2/3$
$F(9) \quad = f(X \le 9) = 1$
$P(X > 9) = 1 - P(X \le 9) = 1 - 1 = 0$

(B) Find μ_x and σ_x^2 and then deduce each of μ_y and σ_y^2; where $Y = 3X - 6$.

$$\mu_x = E(X) = \sum x\, f(x) = \left(-3 \times \frac{1}{6}\right) + \left(6 \times \frac{1}{2}\right) + \left(9 \times \frac{1}{3}\right) = \frac{11}{2}$$

$$\boxed{\sigma_x^2 = E(X^2) - E(X)^2}$$

$$E(X^2) = \sum x^2 f(x) = \left(-3^2 \times \frac{1}{6}\right) + \left(6^2 \times \frac{1}{2}\right) + \left(9^2 \times \frac{1}{3}\right) = \frac{93}{2}$$

$$\sigma_x^2 = E(X^2) - E(X)^2$$
$$= \frac{93}{2} - \left(\frac{11}{2}\right)^2 = \frac{65}{4}$$

$$\mu_y = E(Y) = E(3X - 6) = 3 \times E(X) - 6 = 3 \times \frac{11}{2} - 6 = \frac{21}{2}$$

$$\sigma_Y^2 = \mathrm{Var}(Y) = \mathrm{Var}(3X - 6) = 3^2 \mathrm{Var}(X) = 9 \times \frac{65}{4}$$

Q2. A large industrial firm purchases several word processors at the end of each year, the exact number depending of the frequency of repairs in the previous year. Suppose that the number of word processors, X that are purchased each year has the following probability distribution:

x	0	1	2	3
f(x)	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{4}{10}$	$\frac{2}{10}$

a. Find the cumulative distribution function F.

$$F(X) = \begin{cases} 0 & x < 0 \\ 0.1 & 0 \le x < 1 \\ 0.4 & 1 \le x < 2 \\ 0.8 & 2 \le x < 3 \\ 1 & x \ge 3 \end{cases}$$

b. Calculate

$F(-2) = 0$

$F(0) = 0.1$

$F(0.5) = 0.1$

$F(1) = 0.4$

$F(1.5) = 0.4$

$F(2) = 0.8$

$F(5) = 1$

c. Find μ_x and σ_x^2.

- $\mu_x = \left(0 \times \frac{1}{10}\right) + \left(1 \times \frac{3}{10}\right) + \left(2 \times \frac{4}{10}\right) + \left(3 \times \frac{2}{10}\right) = \frac{17}{10}$

$$\boxed{\sigma_x^2 = E(X^2) - E(X)^2}$$

- $E(X^2) = \sum x^2 f(x)$

$$= \left(0^2 \times \frac{1}{10}\right) + \left(1^2 \times \frac{3}{10}\right) + \left(2^2 \times \frac{4}{10}\right) + \left(3^2 \times \frac{2}{10}\right) = \frac{37}{10}$$

$$\sigma_x^2 = E(X^2) - E(X)^2 = \frac{37}{10} - \left(\frac{17}{10}\right)^2 = \frac{81}{100}$$

Q3. Let X be a discrete random variable with probability mass function:

$$f(x) = cx \quad ; \quad x = 1,2,3,4$$

What is the value of c?

x	1	2	3	4
f(x)	c	2c	3c	4c

$$c + 2c + 3c + 4c = 1 \quad \Longrightarrow c = \frac{1}{10}$$

Then probability mass(density) function is given by:

x	1	2	3	4
f(x)	$\frac{1}{10}$	$\frac{2}{10}$	$\frac{3}{10}$	$\frac{4}{10}$

And the cumulative probability distribution function is given by:

x	1	2	3	4
$F(x) = P(X \leq x)$	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{6}{10}$	1

Q4. Suppose that the number of cars X pass through a car wash between 4:00pm and 5:00pm on any sunny Friday has the following probability distribution:

x	4	5	6	7	8	9
f(x)	$\frac{1}{12}$	$\frac{1}{12}$	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{1}{6}$	$\frac{1}{6}$

Let $g(x) = 2X - 1$ represent the amount of dollars, paid to the attendant by the manger. Find the attendant's expected earnings (and its variance) for this particular time period.

$$\mu_x = \left(4 \times \frac{1}{12}\right) + \left(5 \times \frac{1}{12}\right) + \left(6 \times \frac{1}{4}\right) + \left(7 \times \frac{1}{4}\right) + \left(8 \times \frac{1}{6}\right) + \left(9 \times \frac{1}{6}\right) = \frac{41}{6}$$

$$\boxed{\sigma_x^2 = E(X^2) - E(X)^2}$$

$$E(X^2) = \left(4^2 \times \frac{1}{12}\right) + \left(5^2 \times \frac{1}{12}\right) + \cdots + \left(9^2 \times \frac{1}{6}\right) = \frac{293}{6}$$

$$\sigma_x^2 = E(X^2) - E(X)^2 = \frac{293}{6} - \left(\frac{41}{6}\right)^2 = \frac{77}{36}$$

$$\mu_y = E(Y) = E(2X - 1) = 2 \times E(X) - 1 = 2 \times \frac{41}{6} - 1 = \frac{38}{3}$$

$$\sigma_Y^2 = Var(Y) = Var(2X - 1) = 2^2 Var(X) = 4 \times \frac{77}{36}$$

Discrete Uniform Distribution

$$f(x) = {}^{1}/_{k}\ ; \quad x = x_1, x_2, \ldots, x_k$$

Q1: X have discrete uniform with parameter $k = 3, \quad x = 0, 1, 2$

x	0	1	2
f(x)	1/3	1/3	1/3

a. $P(X = 1) = 1/3$

b. $E(X) = \left(0 \times \frac{1}{3}\right) + \left(1 \times \frac{1}{3}\right) + \left(2 \times \frac{1}{3}\right) = 1$

c. $\boxed{Var(X) = E(X^2) - E(X)^2}$

$$E(X^2) = \left(0^2 \times \frac{1}{3}\right) + \left(1^2 \times \frac{1}{3}\right) + \left(2^2 \times \frac{1}{3}\right) = \frac{5}{3}$$

$$Var(X) = E(X^2) - E(X)^2 = \frac{5}{3} - (1)^2 = \frac{2}{3}$$

Binomial Distribution

$$P(X = x) = \binom{n}{x} p^x q^{n-x} \;;\quad x = 0,1 \ldots, n$$
$$* \; E(X) = np \qquad * \; Var(X) = npq$$
$$q = 1 - p$$

Q1. Suppose that 33% of the buildings in a certain city violate the building code. A building engineer randomly inspects a sample of 3 new buildings in the city.

(a) Find the (p.m.f) of the random variable X representing the number of buildings that violate the building code in the sample.

$$p = \frac{1}{3} \;,\; n = 3$$

$$f(x) = \binom{3}{x}\left(\frac{1}{3}\right)^x \left(\frac{2}{3}\right)^{3-x} \quad ; \quad x = 0,1,2,3$$

(b) Find the probability that:

(i) None of the buildings in the sample violating the building code.

$$P(X = 0) = f(0) = \binom{3}{0}\left(\frac{1}{3}\right)^0 \left(\frac{2}{3}\right)^3 = 0.296$$

X	0	1	2	3
P(X = x)	*			

(ii) One building in the sample violating the building code.

$$P(X = 1) = f(1) = \binom{3}{1}\left(\frac{1}{3}\right)^1 \left(\frac{2}{3}\right)^2 = 0.44$$

X	0	1	2	3
P(X = x)		*		

(iii) At least one building in the sample violating the building code.

$$P(X \geq 1) = 1 - P(X < 1) = 1 - f(0)$$
$$= 1 - 0.296 = 0.704$$

X	0	1	2	3
P(X = x)		*	*	*

(c) Find the expected number of buildings that violate the building code E(X).

$$E(X) = np = 3 \times \frac{1}{3} = 1$$

(d) Find Var(X).

$$Var(X) = npq = 3 \times \frac{1}{3} \times \frac{2}{3} = \frac{2}{3}$$

Q2. Suppose that the probability that a person dies when he or she contracts a certain disease is 0.4. A sample of 10 persons who contracted this disease is randomly chosen.

(1) What is the expected number of persons who will die in this sample?

$$p = 0.4 \quad , \quad n = 10$$

$$E(X) = np = 10 \times 0.4 = 4$$

(2) What is the variance of the number of persons who will die in this sample?

$$Var(X) = npq = 10 \times 0.4 \times 0.6 = 2.4$$

(3) What is the probability that exactly 4 persons will die among this sample?

$$f(x) = \binom{10}{x}(0.4)^x(0.6)^{10-x} \quad ; \quad x = 0,1,\ldots,10$$

$$P(X = 4) = f(4) = \binom{10}{4}(0.4)^4(0.6)^6 = 0.251$$

(4) What is the probability that less than 3 persons will die among this sample?

$$P(X < 3) = f(0) + f(1) + f(2)$$

$$= \binom{10}{0}(0.4)^0(0.6)^{10} + \binom{10}{1}(0.4)^1(0.6)^9 + \binom{10}{2}(0.4)^2(0.6)^8 = 0.167$$

(5) What is the probability that more than 8 persons will die among this sample?

$$P(X > 8) = f(9) + f(10)$$

$$= \binom{10}{9}(0.4)^9(0.6)^1 + \binom{10}{10}(0.4)^{10}(0.6)^0 = 0.0017$$

Q3. If X~ Binomial (n,p), E(X)=1, and Var(X)=0.75, find P(X=1).

$$X \sim \text{Binomial}(n, p) \quad \& \quad E(X) = 1 \quad \& \quad Var(X) = 0.75$$

$$\frac{Var(X)}{E(X)} = \frac{0.75}{1} \Rightarrow \frac{npq}{np} = \frac{0.75}{1} \Rightarrow q = 0.75 \Rightarrow p = 0.25.$$

$$E(X) = 1 \Rightarrow np = 1 \Rightarrow n \times 0.25 = 1 \Rightarrow n = 4.$$

$$f(x) = \binom{4}{x}(0.25)^x(0.75)^{4-x} \;\; ; \quad x = 0,1,2,3,4$$

$$P(X = 1) = f(1) = \binom{4}{1}(0.25)^1(0.75)^3 = 0.422$$

Q4. A traffic control engineer reports that 75% of the cars passing through a checkpoint are from Riyadh city. If at this checkpoint, five cars are selected at random.

(1) The probability that <u>none of them is from Riyadh</u> city equals to:

$$p = 0.75 \quad , \quad n = 5$$

$$f(x) = \binom{5}{x}(0.75)^x(0.25)^{5-x} \quad ; \quad x = 0,1,2,3,4,5$$

$$P(X = 0) = f(0) = \binom{5}{0}(0.75)^0(0.25)^5 = 0.00098$$

(2) The probability that <u>four of them are from Riyadh</u> city equals to:

$$P(X = 4) = f(4) = \binom{5}{4}(0.75)^4(0.25)^1 = 0.3955$$

(3) The probability that <u>at least four of them are from Riyadh</u> city equals to:

$$P(X \geq 4) = f(4) + f(5)$$

$$\binom{5}{4}(0.75)^4(0.25)^1 + \binom{5}{5}(0.75)^5(0.25)^0 = 0.6328$$

(4) The expected number of cars that are from Riyadh city equals to:

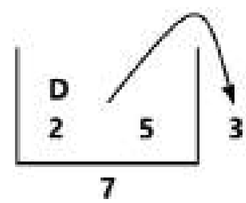
$$E(X) = np = 5 \times 0.75 = 3.75$$

Hypergeometric Distribution

$$f(x) = \frac{\binom{k}{x}\binom{N-k}{n-x}}{\binom{N}{n}} \quad ; \quad x = 0,1, \dots . \min(n,k)$$

$$* \; E(X) = n \times \frac{k}{N} \qquad * \; Var(X) = n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right)$$

Q1. A shipment of 7 television sets contains 2 defective sets. A hotel makes a random purchase of 3 of the sets.



Find the probability distribution function of the random variable X representing the number of defective sets purchased by the hotel.

$$N = 7 \;\; , \;\; n = 3 \;\; , \;\; k = 2$$

$$f(x) = \frac{\binom{2}{x}\binom{5}{3-x}}{\binom{7}{3}} \quad ; \quad x = 0,1,2$$

(i) Find the probability that the hotel purchased <u>no defective</u> television sets.

$$P(X = 0) = f(0) = \frac{\binom{2}{0}\binom{5}{3}}{\binom{7}{3}} = 0.29$$

(i) What is the expected number of defective television sets purchased by the hotel?

$$E(X) = n \times \frac{k}{N} = 3 \times \frac{2}{7} = \frac{6}{7}$$

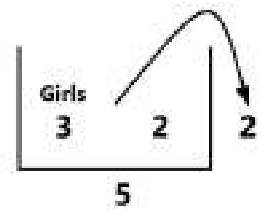
(ii) Find the variance of X.

$$Var(X) = n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right) = 3 \times \frac{2}{7}\left(\frac{5}{7}\right)\left(\frac{7-3}{7-1}\right) = 0.41$$

Q2. Suppose that a family has 5 children, 3 of them are girls and the rest are boys. A sample of 2 children is selected randomly and without replacement.

$$N = 5 \quad , \quad n = 2 \quad , \quad k = 3$$

$$f(x) = \frac{\binom{3}{x}\binom{2}{2-x}}{\binom{5}{2}} \quad ; \quad x = 0,1,2$$



a. The probability that no girls are selected is

$$P(X = 0) = f(0) = \frac{\binom{3}{0}\binom{2}{2}}{\binom{5}{2}} = 0.1$$

b. The probability that <u>at most one</u> girl are selected is

$$P(X \le 1) = f(0) + f(1) = \frac{\binom{3}{0}\binom{2}{2}}{\binom{5}{2}} + \frac{\binom{3}{1}\binom{2}{1}}{\binom{5}{2}} = 0.7$$

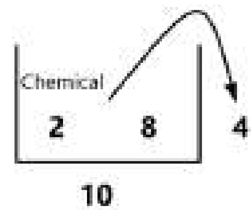
c. The expected number of girls in the sample is

$$E(X) = n \times \frac{k}{N} = 2 \times \frac{3}{5} = \frac{6}{5}$$

d. The variance of the number of girls in the sample is

$$Var(X) = n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right) = 2 \times \frac{3}{5}\left(1 - \frac{3}{5}\right)\left(\frac{5-2}{5-1}\right) = 0.36$$

Q3. A random committee of size 4 is selected from 2 chemical engineers and 8 industrial engineers.



(1) Write a formula for the probability distribution function of the random variable X representing the number of chemical engineers in the committee.

$$N = 10 \quad , \quad n = 4 \quad , \quad k = 2$$

$$f(x) = \frac{\binom{2}{x}\binom{8}{4-x}}{\binom{10}{4}} \quad ; \quad x = 0,1,2$$

(2) Find the probability that there will be no chemical engineers in the committee.

$$P(X = 0) = f(0) = \frac{\binom{2}{0}\binom{8}{4}}{\binom{10}{4}} = 0.33$$

(3) Find the probability that there will be at least one chemical engineer in the committee.

$$P(X \geq 1) = 1 - P(X < 1) = 1 - f(0) = 0.67$$

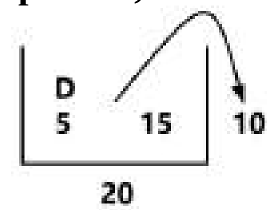
(4) What is the expected number of chemical engineers in the committee?

$$E(X) = n \times \frac{k}{N} = 4 \times \frac{2}{10} = 0.8$$

(5) What is the variance of the number of chemical engineers in the committee?

$$Var(X) = n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right) = 4 \times \frac{2}{10}\left(1 - \frac{2}{10}\right)\left(\frac{10-4}{10-1}\right) = 0.43$$

Q4. A shipment of 20 digital voice recorders contains 5 that are defective. If 10 of them are randomly chosen (without replacement) for inspection, then:



(1) The probability that 2 will be defective is:

$$N = 20 \quad , \quad n = 10 \quad , \quad k = 5$$

$$f(x) = \frac{\binom{5}{x}\binom{15}{10-x}}{\binom{20}{10}} \quad ; \quad x = 0,1,2,3,4,5$$

$$P(X = 2) = f(2) = \frac{\binom{5}{2}\binom{15}{8}}{\binom{20}{10}} = 0.35$$

(2) The probability that at most 1 will be defective is:

$$P(X \le 1) = f(0) + f(1) = \frac{\binom{5}{0}\binom{15}{10}}{\binom{20}{10}} + \frac{\binom{5}{1}\binom{15}{9}}{\binom{20}{10}} = 0.15$$

(3) The expected number of defective recorders in the sample is:

$$E(X) = n \times \frac{k}{N} = 10 \times \frac{5}{20} = \frac{5}{2}$$

(4) The variance of the number of defective recorders in the sample is:

$$Var(X) = n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right) = 10 \times \frac{5}{20}\left(\frac{15}{20}\right)\left(\frac{20-10}{20-1}\right) = 0.99$$

Poisson distribution

$$P(X = x) = \frac{e^{-\lambda}\ \lambda^{x}}{x!} \quad ; \quad x = 0,1,2,\dots$$

$$E(X) = Var(X) = \lambda$$

Q1. On average, a certain intersection results in 3 traffic accidents per day. Assuming Poisson distribution,

What is the probability that at this intersection:

(1) No accidents will occur in a given day?

$$\lambda_{\text{one day}} = 3$$

$$f(x) = \frac{e^{-3}(3)^{x}}{x!} \; ; \quad x = 0,1,2,\dots$$

$$P(X = 0) = f(0) = \frac{e^{-3}(3)^{0}}{0!} = 0.05$$

(2) More than 3 accidents will occur in a given day?

$$P(X > 3) = 1 - P(X \le 3) = 1 - [f(0) + f(1) + f(2) + f(3)]$$

$$= 1 - [\frac{e^{-3}(3)^{0}}{0!} + \frac{e^{-3}(3)^{1}}{1!} + \frac{e^{-3}(3)^{2}}{2!} + \frac{e^{-3}(3)^{3}}{3!}] = 0.35$$

(3) Exactly 5 accidents will occur in a period of two days?

$$\lambda_{\text{two days}} = 6$$

$$f(x) = \frac{e^{-6}(6)^{x}}{x!} \quad ; \quad x = 0, \dots \dots \infty$$

$$P(X = 5) = f(5) = \frac{e^{-6}(6)^{5}}{5!} = 0.16$$

(4) What is the average number of traffic accidents in a period of 4 days?

$$E(X) = \lambda_{4\text{ days}} = 4 \times 3 = 12$$

Q2. Suppose that the number of telephone calls received per day has a Poisson distribution with mean of 4 calls per day.

(a). The probability that 2 calls will be received in a given day is

$$\lambda_{\text{one day}} = 4$$

$$f(x) = \frac{e^{-4}(4)^x}{x!} \quad ; \quad x = 0,1,2,\ldots$$

$$P(X=2) = f(2) = \frac{e^{-4}(4)^2}{2!} = 0.15$$

(b). The expected number of telephone calls received in a given week is

$$E(X) = \lambda_{\text{week}} = 4 \times 7 = 28$$

(c). The probability that at least 2 calls will be received in a period of 12 hours is

$$\lambda_{\text{12hours}} = 2$$

$$f(x) = \frac{e^{-2}(2)^x}{x!} \quad ; \quad x = 0,1,2,\ldots$$

$$P(X \geq 2) = 1 - P(X < 2) = 1 - [f(0) + f(1)]$$

$$= 1 - \left[\frac{e^{-2}(2)^0}{0!} + \frac{e^{-2}(2)^1}{1!}\right] = 0.59$$

Q3. Suppose that X~ Binomial (1000,0.002). By using Poisson approximation, P(X=3) is approximately equal to

When we have a Binomial distribution with a small p and a large n, we can make an approximation to Poisson distribution:

$$\begin{matrix} E(X_{\text{Binomial}}) = np \\ E(X_{\text{Poisson}}) = \lambda \end{matrix} \Longrightarrow \boxed{\lambda = np} \Longrightarrow \boxed{\lambda = 1000 \times 0.002} \Longrightarrow \boxed{\lambda = 2}$$

$$P(X=3) = f(3) = \frac{e^{-2}2^3}{3!} = 0.18045$$

	f(x) or PDF		E(X)	Var(X)
Binomial	$\binom{n}{x} p^x q^{n-x}$	$x = 0, 1, 2, \ldots., n$	np	npq
Hyper geometric	$\frac{\binom{k}{x}\binom{N-k}{n-x}}{\binom{N}{n}}$	$x = 0,1, \ldots. \min(n, k)$	$n \times \frac{k}{N}$	$n \times \frac{k}{N}\left(1 - \frac{k}{N}\right)\left(\frac{N-n}{N-1}\right)$
Poisson	$\frac{\lambda^x e^{-\lambda}}{x!}$	$x = 0, 1, 2, \ldots.$	λ	λ

CHAPTER 2

- Continuous Random Variables.
- Uniform Distribution.
- Exponential Distribution
- Normal (N) Distribution.
- Student (T) Distribution.
- Chi-square (χ^2) Distribution.
- Fisher (F) Distribution.

Continuous Random Variables

- $0 \le f(x) \le 1$
- $\int_{-\infty}^{\infty} f(x)dx = 1$
- $P(a < X < b) = \int_a^b f(x)dx$
- $E(X) = \int_{-\infty}^{\infty} x\, f(x)dx$
- $Var(X) = E(X^2) - E(X)^2$
- $E(X^2) = \int_{-\infty}^{\infty} x^2\, f(x)dx$
- $E(aX \pm b) = aE(X) \pm b$
- $Var(aX \pm b) = a^2 Var(X)$
- $F(x) = P(X \le x)$

Q1. Suppose X is a continuous random variable what is $P(X = 16)$?

$$P(X = 16) = 0$$

Note, for any continuous random variable, $P(x = a) = 0$

Q2: Suppose we have the (p.d.f): $f(x) = k\sqrt{x}$, $0 < x < 1$

(1) Find the value of k

$$\int_{-\infty}^{\infty} f(x)dx = 1 \Rightarrow \int_0^1 k\sqrt{x}\, dx = 1$$

$$\Rightarrow \int_0^1 k\sqrt{x}\, dx = 1 \Rightarrow k\int_0^1 x^{0.5}\, dx = 1$$

$$\Rightarrow k\left[\frac{x^{1.5}}{1.5}\right]_0^1 = 1 \Rightarrow \frac{k}{1.5} = 1 \Rightarrow \mathbf{k = 1.5}$$

then, $\boxed{f(x) = 1.5\sqrt{x} \quad , \quad 0 < x < 1}$

(2) Find the probability $P(0.3 < X < 0.6)$:

$$P(0.3 < X < 0.6) = \int_{0.3}^{0.6} f(x)dx$$

$$= \int_{0.3}^{0.6} 1.5x^{0.5}dx$$

$$= 1.5\left[\frac{x^{1.5}}{1.5}\right]_{0.3}^{0.6}$$

$$= [x^{1.5}]_{0.3}^{0.6} = 0.6^{1.5} - 0.3^{1.5} = \mathbf{0.3004}$$

(3) Find the cumulative distribution function $F(x)$:

$$F(X) = \begin{cases} 0 & x \leq 0 \\ x^{1.5} & 0 \leq x \leq 1 \\ 1 & x \geq 0 \end{cases}$$

(4) Calculate:

$F(-2) = 0$

$F(-1) = 0$

$F(0) \quad = 0$

$F(0.5) = (0.5)^{1.5} = 0.35$

$F(0.8) = (0.8)^{1.5} = 0.72$

$F(1) \quad = 1$

$F(2) \quad = 1$

(5) Find $E(X)$:

$$E(X) = \int_{-\infty}^{\infty} x\, f(x) dx$$

$$= \int_0^1 1.5x^{1.5} dx$$

$$= 1.5 \left[\frac{x^{2.5}}{2.5}\right]_0^1$$

$$= \frac{1.5}{2.5} \left[x^{2.5}\right]_0^1 = \mathbf{0.6}$$

Q3: Suppose we have the (p.d.f):

$$\mathbf{f(x) = 3x^2 \quad , \quad 0 < x < 1}$$

1. Find the mean μ:

$$E(X) = \int_{-\infty}^{\infty} x\, f(x)dx = \int_0^1 3x^3 dx$$

$$= 3\int_0^1 x^3 dx$$

$$= \frac{3}{4}\,[x^4]_0^1 = \frac{3}{4}$$

2. $\mathbf{P(X > 0.5) =}$

$$= \int_{0.5}^{1} f(x)dx$$

$$= \int_{0.5}^{1} 3x^2\, dx$$

$$= 3\int_{0.5}^{1} x^2\, dx$$

$$= \frac{3}{3}\,[x^3]_{0.5}^{1} = 0.875$$

3. $\mathbf{P(0.4 < X < 0.6) =}$

$$= \int_{0.4}^{0.6} f(x)dx$$

$$= \int_{0.4}^{0.6} 3x^2\, dx$$

$$= 3\int_{0.4}^{0.6} x^2\, dx$$

$$= \frac{3}{3}\,[x^3]_{0.4}^{0.6} = 0.154$$

Q4: Suppose we have the (p.d.f):

$$\mathbf{f(x) = c(4 - x)} \quad , \qquad \mathbf{-2 < x < 2}$$

1. **Find the value of c:**

$$f(X) = \int_{-\infty}^{\infty} f(x)dx = 1$$

$$\Longrightarrow \int_{-2}^{2} c(4 - x)\, dx = 1$$

$$\Longrightarrow c\int_{-2}^{2} 4 - x\, dx = 1$$

$$\Longrightarrow c \quad \left[4x - \frac{x^2}{2}\right]_{-2}^{2} = 1$$

$$\Longrightarrow c \quad \left[\left(8 - \frac{4}{2}\right) - \left(-8 - \frac{4}{2}\right)\right] = 1$$

$$\Longrightarrow c \quad [16] = 1$$

$$\Longrightarrow c \quad = \frac{1}{16}$$

2. $\mathbf{P(X > 0)} = \int_{0}^{2} f(x)dx$

$$= \frac{1}{16}\int_{0}^{2} 4 - x\, dx = \frac{1}{16}\left[4x - \frac{x^2}{2}\right]_{0}^{2} = \frac{1}{16}\left[\left(8 - \frac{4}{2}\right)\right] = \frac{6}{16}$$

3. $\mathbf{P(X < \quad 3) = 1}$
4. $\mathbf{P(X < -3) = 0}$

5. **Find $\mathbf{E(X)}$**

$$= \int_{-2}^{2} x\, f(x)dx = \frac{1}{16}\int_{-2}^{2} 4x - x^2 \;\; dx = \frac{1}{16}\left[2\, x^2 - \frac{x^3}{3}\right]_{-2}^{2} = -\frac{1}{3}$$

The Uniform Distribution:

$$f(X) = \frac{1}{b-a} \quad ; \quad a < X < b$$

$$E(X) = \frac{a+b}{2} \quad ; \quad Var(X) = \frac{(b-a)^2}{12} \quad ; \quad F(X) = \frac{X-a}{b-a} \quad ; \quad a < X < b$$

Question 1:

If the random variable X has a uniform distribution on the interval (0,10), then

(1) $P(X < 6)$ equal to:

$$f(x) = \frac{1}{10} \quad , \; 0 < X < 10 \qquad \boxed{a = 0 \quad , \quad b = 10}$$

$$P(X < 6) = \int_0^6 \frac{1}{10} \; dx$$

$$= \frac{1}{10} [x]_0^6 = \frac{1}{10}(6 - 0) = 0.6$$

or $\quad F(6) = \frac{x-a}{b-a} = \frac{6-0}{10-0} = \frac{6}{10} = 0.6$

(2) The mean of X is

$$E(X) = \frac{a+b}{2} = \frac{10}{2} = 5$$

(3) The variance of X is

$$Var(X) = \frac{(b-a)^2}{12} = \frac{100}{12} = 8.33$$

Question 2:

Suppose that the random variable X has the following uniform distribution:

$$f(X) = 3 \quad ; \quad \frac{2}{3} < X < 1$$

1. $P(0.33 < X < 0.5) = 0$
2. $P(X > 1.25) = 0$
3. The variance of X is

$$Var(X) = \frac{(b-a)^2}{12} = \frac{\left(1 - {}^2/_3\right)^2}{12} = 0.00926$$

Question 3:

Suppose that the continuous random variable X has the density function:

$$f(x) = 0.2 \quad , 0 < X < 5$$

$$\boxed{a = 0 \quad , \quad b = 5}$$

(1) $P(X > 1)$ equal to

$$P(X > 1) = 1 - P(X < 1) = 1 - F(1) = 1 - \left(\frac{1-0}{5-0}\right) = 1 - \frac{1}{5} = 0.8$$

(2) $P(X \geq 1)$

$P(X \geq 1) = 0.8$, the same as part (1) because we had a continuous random variable, and hence, $P(X = 1) = 0$

(3) The mean $\mu = E(X)$ is equal to

$$E(X) = \frac{a+b}{2} = \frac{5}{2} = 2.5$$

(4) $E(X^2)$

$$Var(X) = E(X^2) - [E(X)]^2 \Rightarrow E(X^2) = Var(X) + [E(X)]^2$$

$$Var(X) = \frac{(b-a)^2}{12} = \frac{25}{12} = 2.0833$$

$$E(X^2) = Var(X) + [E(X)]^2 = 2.08 + 2.5^2 = 8.33$$

(5) $Var(X)$ equal to

$$Var(X) = \frac{(b-a)^2}{12} = \frac{25}{12} = 2.0833$$

(6) If $F(X)$ is the cumulative distribution function of X then $F(1)$ equal to

$$F(X) = \frac{X-a}{b-a} \quad ; \quad a < X < b$$

$$F(X) = \frac{X}{5} \quad ; \quad 0 < X < 5$$

$$F(1) = \frac{1}{5} = 0.2$$

The Exponential Distribution:

$$f(X) = \frac{1}{\beta} e^{-\frac{X}{\beta}} \quad ; \quad X > 0$$

$$E(X) = \beta \quad ; \quad Var(X) = \beta^2 \quad ; \quad F(X) = 1 - e^{-\frac{X}{\beta}} \ ; \ X > 0$$

Question 1:

If the random variable X has an exponential distribution with mean 4, then:

$$X \sim expo(4) \quad , \quad f(x) = \frac{1}{4} e^{-\frac{x}{4}} \quad ; \quad x > 0$$

(1) $P(X < 8)$ equal to

$$P(X < 8) = F(8) = 1 - e^{-\frac{8}{4}} = 1 - e^{-2} = 0.8647$$

(2) The variance of X is

$$Var(X) = \beta^2 = 4^2 = 16$$

Question 2:

The lifetime of a specific battery a random variable X with probability density function given by:

$$f(x) = \frac{1}{200} e^{-\frac{x}{200}} \quad ; \quad x > 0$$

(1) The mean lifetime of the battery equal to:

$$E(X) = \beta = 200$$

(2) $P(X > 100)$ equal to

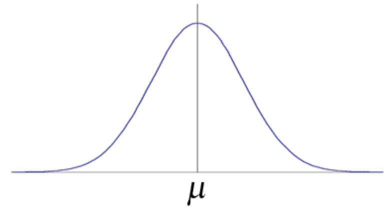
$$P(X > 100) = 1 - P(X < 100) = 1 - F(100)$$

$$1 - \left[1 - e^{-\frac{100}{200}}\right] = 1 - 1 + e^{-0.5} = e^{-0.5} = 0.6065$$

(3) $P(X = 200) = 0$

The Normal Distribution

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}; \quad -\infty < X < \infty$$



Mean = Median = Mode

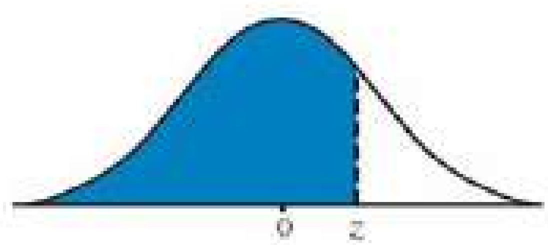
Normal distribution	$X \sim N(\mu, \sigma^2)$	
Standard normal	$Z \sim N(0, 1)$	$Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$

Q1. Suppose that Z is distributed according to the standard normal distribution.

1) The area under the curve to the left of is $Z = 1.43$:

$$P(Z < 1.43) = 0.9236$$

z	0.00	0.01	0.02	0.03	0.04
0.00	0.50000	0.50399	0.50798	0.5[illegible]197	0.51595
0.10	0.53983	0.54380	0.54776	0.5[illegible]172	0.55567
0.20	0.57926	0.58317	0.58706	0.5[illegible]095	0.59483
0.30	0.61791	0.62172	0.62552	0.6[illegible]930	0.63307
0.40	0.65542	0.65910	0.66276	0.6[illegible]640	0.67003
0.50	0.69146	0.69497	0.69847	0.7[illegible]194	0.70540
0.60	0.72575	0.72907	0.73237	0.7[illegible]565	0.73891
0.70	0.75804	0.76115	0.76424	0.7[illegible]730	0.77035
0.80	0.78814	0.79103	0.79389	0.7[illegible]673	0.79955
0.90	0.81594	0.81859	0.82121	0.8[illegible]381	0.82639
1.00	0.84134	0.84375	0.84614	0.8[illegible]849	0.85083
1.10	0.86433	0.86650	0.86864	0.8[illegible]076	0.87286
1.20	0.88493	0.88686	0.88877	0.8[illegible]065	0.89251
1.30	0.90320	0.90490	0.90658	0.9[illegible]824	0.90988
1.40	[illegible]	[illegible]	[illegible]	0.92364	0.92507
1.50	0.93319	0.93448	0.93574	0.93699	0.93822
1.60	0.94520	0.94630	0.94738	0.94845	0.94950
1.70	0.95543	0.95637	0.95728	0.95818	0.95907
1.80	0.96407	0.96485	0.96562	0.96638	0.96712



P (Z < أطراف الجدول) = داخل الجدول

2) The area under the curve to the left of is $Z = 1.39$:

$$P(Z < 1.39) = 0.9177$$

3) The area under the curve to the right of is $Z = -0.89$:

$$P(Z > -0.89) = 1 - P(Z < -0.89) = 1 - 0.1867 = 0.8133$$

4) The area under the curve between $Z = -2.16$ and $Z = -0.65$ is:

$$P(-2.16 < Z < -0.65)$$
$$= P(Z < -0.65) - P(Z < -2.16)$$
$$= 0.2578 - 0.0154 = 0.2424$$

5) The value of k such that is $P(0.93 < Z < k)$:

$$P(0.93 < Z < k) = 0.0427$$
$$P(Z < k) - P(Z < 0.93) = 0.0427$$
$$P(Z < k) - 0.8238 = 0.0427$$
$$P(Z < k) = 0.8665$$

P(Z < اطراف الجدول) = داخل الجدول $\Rightarrow k = 1.11$

Q2. The finished inside diameter of a piston ring is normally distributed with a mean of 12 centimeters (c.m) and a standard deviation of 0.03 centimeter. Then,

1) The proportion of rings that will have inside diameter less than 12.05 is:

$$X \sim N(\mu, \sigma)$$

$$X \sim N(12, 0.03)$$

$$P(X < 12.05) = P\left(Z < \frac{12.05-\mu}{\sigma}\right)$$

$$= P\left(Z < \frac{12.05-12}{0.03}\right)$$

$$= P(Z < 1.67) = 0.9525$$

2) The proportion of rings that will have inside diameter exceeding 11.97 is:

$$P(X > 11.97) = P\left(Z > \frac{11.97-\mu}{\sigma}\right)$$

$$= P\left(Z > \frac{11.97-12}{0.03}\right)$$

$$= P(Z > -1)$$

$$= 1 - P(Z < -1)$$

$$= 1 - 0.1587 = 0.8413$$

3) The probability that a piston ring will have an inside diameter between 11.95 and 12.05 is:

$$P(11.95 < X < 12.05)$$

$$= P\left(\frac{11.95-12}{0.03} < Z < \frac{12.05-12}{0.03}\right)$$

$$= P(-1.67 < Z < 1.67)$$

$$= P(Z < 1.67) - P(Z < -1.67)$$

$$= 0.9525 - 0.0475 = 0.905$$

Q3. The weight of a large number of fat persons is nicely modeled with a normal distribution with mean of 128 kg and a standard deviation of 9 kg.

$$X \sim N(\mu, \sigma)$$
$$X \sim N(128, 9)$$

(1): The percentage of fat persons with weights at most 110 kg is

$$P(X \le 110) = P\left(Z < \frac{110-128}{9}\right)$$
$$= P(Z < -2) = 0.0228$$

(2): The percentage of fat persons with weights more than 149 kg is

$$P(X > 149) = P\left(Z > \frac{149-128}{9}\right)$$
$$= 1 - P(Z < 2.33)$$
$$= 1 - 0.9901 = 0.0099$$

(3): The weight x above which 86% of those persons will be

$$P(X > x) = 0.86$$
$$1 - P(X < x) = 0.86$$
$$P(X < x) = 0.14$$
$$P\left(Z < \frac{x-128}{9}\right) = 0.14$$

By searching inside the table for 0.14, and transforming X to Z, we got:

$$\frac{x-128}{9} = -1.08 \Rightarrow x = 118.28$$

(4): The weight x below which 50% of those persons will be

$P(X < x) = 0.5$, by searching inside the table for 0.5, and transforming X to Z

$$\frac{x-128}{9} = 0 \Rightarrow x = 128$$

Q4. If the random variable X has a normal distribution with the mean μ and the variance σ^2, then $P(X < \mu + 2\sigma)$ equals to

$$P(X < \mu + 2\sigma) = P\left(Z < \frac{(\mu + 2\sigma) - \mu}{\sigma}\right) = P(Z < 2) = 0.9772$$

Q5. If the random variable X has a normal distribution with the mean μ and the variance 1, and if P(X<3)=0.877, then μ equals to

$$\text{Given that } \sigma = 1$$

$$P(X < 3) = 0.877 \Rightarrow P\left(Z < \frac{3 - \mu}{1}\right) = 0.877$$

$$3 - \mu = 1.16 \Rightarrow \mu = 1.84$$

Q6. Suppose that the marks of the students in a certain course are distributed according to a normal distribution with the mean 70 and the variance 25. If it is known that 33% of the student failed the exam, then the passing mark x is

$$X \sim N(70,5)$$

$$P(X < x) = 0.33 \Rightarrow P\left(Z < \frac{x - 70}{5}\right) = 0.33$$

By searching inside the table for 0.33, and transforming X to Z, we got:

$$\frac{x - 70}{5} = -0.44 \Rightarrow x = 67.8$$

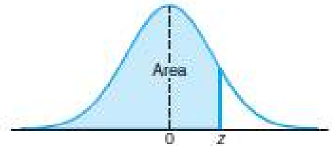


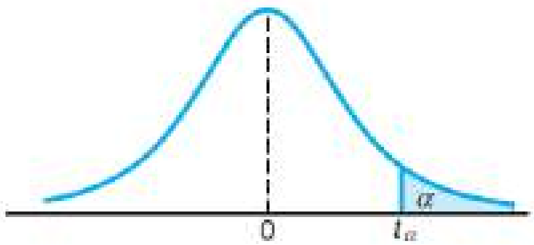
Table A.3 Areas under the Normal Curve

z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
−3.4	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0003	0.0002
−3.3	0.0005	0.0005	0.0005	0.0004	0.0004	0.0004	0.0004	0.0004	0.0004	0.0003
−3.2	0.0007	0.0007	0.0006	0.0006	0.0006	0.0006	0.0006	0.0005	0.0005	0.0005
−3.1	0.0010	0.0009	0.0009	0.0009	0.0008	0.0008	0.0008	0.0008	0.0007	0.0007
−3.0	0.0013	0.0013	0.0013	0.0012	0.0012	0.0011	0.0011	0.0011	0.0010	0.0010
−2.9	0.0019	0.0018	0.0018	0.0017	0.0016	0.0016	0.0015	0.0015	0.0014	0.0014
−2.8	0.0026	0.0025	0.0024	0.0023	0.0023	0.0022	0.0021	0.0021	0.0020	0.0019
−2.7	0.0035	0.0034	0.0033	0.0032	0.0031	0.0030	0.0029	0.0028	0.0027	0.0026
−2.6	0.0047	0.0045	0.0044	0.0043	0.0041	0.0040	0.0039	0.0038	0.0037	0.0036
−2.5	0.0062	0.0060	0.0059	0.0057	0.0055	0.0054	0.0052	0.0051	0.0049	0.0048
−2.4	0.0082	0.0080	0.0078	0.0075	0.0073	0.0071	0.0069	0.0068	0.0066	0.0064
−2.3	0.0107	0.0104	0.0102	0.0099	0.0096	0.0094	0.0091	0.0089	0.0087	0.0084
−2.2	0.0139	0.0136	0.0132	0.0129	0.0125	0.0122	0.0119	0.0116	0.0113	0.0110
−2.1	0.0179	0.0174	0.0170	0.0166	0.0162	0.0158	0.0154	0.0150	0.0146	0.0143
−2.0	0.0228	0.0222	0.0217	0.0212	0.0207	0.0202	0.0197	0.0192	0.0188	0.0183
−1.9	0.0287	0.0281	0.0274	0.0268	0.0262	0.0256	0.0250	0.0244	0.0239	0.0233
−1.8	0.0359	0.0351	0.0344	0.0336	0.0329	0.0322	0.0314	0.0307	0.0301	0.0294
−1.7	0.0446	0.0436	0.0427	0.0418	0.0409	0.0401	0.0392	0.0384	0.0375	0.0367
−1.6	0.0548	0.0537	0.0526	0.0516	0.0505	0.0495	0.0485	0.0475	0.0465	0.0455
−1.5	0.0668	0.0655	0.0643	0.0630	0.0618	0.0606	0.0594	0.0582	0.0571	0.0559
−1.4	0.0808	0.0793	0.0778	0.0764	0.0749	0.0735	0.0721	0.0708	0.0694	0.0681
−1.3	0.0968	0.0951	0.0934	0.0918	0.0901	0.0885	0.0869	0.0853	0.0838	0.0823
−1.2	0.1151	0.1131	0.1112	0.1093	0.1075	0.1056	0.1038	0.1020	0.1003	0.0985
−1.1	0.1357	0.1335	0.1314	0.1292	0.1271	0.1251	0.1230	0.1210	0.1190	0.1170
−1.0	0.1587	0.1562	0.1539	0.1515	0.1492	0.1469	0.1446	0.1423	0.1401	0.1379
−0.9	0.1841	0.1814	0.1788	0.1762	0.1736	0.1711	0.1685	0.1660	0.1635	0.1611
−0.8	0.2119	0.2090	0.2061	0.2033	0.2005	0.1977	0.1949	0.1922	0.1894	0.1867
−0.7	0.2420	0.2389	0.2358	0.2327	0.2296	0.2266	0.2236	0.2206	0.2177	0.2148
−0.6	0.2743	0.2709	0.2676	0.2643	0.2611	0.2578	0.2546	0.2514	0.2483	0.2451
−0.5	0.3085	0.3050	0.3015	0.2981	0.2946	0.2912	0.2877	0.2843	0.2810	0.2776
−0.4	0.3446	0.3409	0.3372	0.3336	0.3300	0.3264	0.3228	0.3192	0.3156	0.3121
−0.3	0.3821	0.3783	0.3745	0.3707	0.3669	0.3632	0.3594	0.3557	0.3520	0.3483
−0.2	0.4207	0.4168	0.4129	0.4090	0.4052	0.4013	0.3974	0.3936	0.3897	0.3859
−0.1	0.4602	0.4562	0.4522	0.4483	0.4443	0.4404	0.4364	0.4325	0.4286	0.4247
−0.0	0.5000	0.4960	0.4920	0.4880	0.4840	0.4801	0.4761	0.4721	0.4681	0.4641

Table A.3 (continued) Areas under the Normal Curve

z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9953	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9976	0.9977	0.9977	0.9978	0.9979	0.9979	0.9980	0.9981
2.9	0.9981	0.9982	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9989	0.9990	0.9990
3.1	0.9990	0.9991	0.9991	0.9991	0.9992	0.9992	0.9992	0.9992	0.9993	0.9993
3.2	0.9993	0.9993	0.9994	0.9994	0.9994	0.9994	0.9994	0.9995	0.9995	0.9995
3.3	0.9995	0.9995	0.9995	0.9996	0.9996	0.9996	0.9996	0.9996	0.9996	0.9997
3.4	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9997	0.9998

The student (t) Distribution



$$t_{v,\alpha} \Rightarrow P(T > t_v) = \alpha$$
$$\Rightarrow P(T < t_v) = 1 - \alpha$$
$$\Rightarrow P(T < -t_v) = \alpha \qquad ; v = n - 1$$

- If $P(T < t_{22}) = 0.99 \Rightarrow P(T > t_{22}) = 0.01$
  $\Rightarrow t_{22,0.01} = 2.508$
- If $P(T > t_{18,0.975}) = 0.975$
  $\Rightarrow t_{18,0.975} = -t_{18,0.025} = -2.101$
- If $P(T > t_{v,\alpha}) = \alpha$ where $v = 24, \alpha = 0.995$
  $\Rightarrow t_{24,0.995} = -t_{24,0.005} = -2.797$
- If $P(T > t_{v,\alpha}) = \alpha$ where $v = 7, \alpha = 0.975$
  $\Rightarrow t_{7,0.975} = -t_{7,0.025} = -2.365$

Q2. A random sample of size 15 selected from a normal distribution with <u>unknown variance</u>, and let $T = \frac{\bar{X}-\mu}{S/\sqrt{n}}$ and k such that $P(1.761 < T < k) = 0.045$,then the value of k is:

v	0.40	0.30	0.20	0.15	0.10	0.05	0.025
1	0.325	0.727	1.376	1.963	3.078	6.314	12.706
2	0.289	0.617	1.061	1.386	1.886	2.920	4.303
3	0.277	0.584	0.978	1.250	1.638	2.353	3.182
4	0.271	0.569	0.941	1.190	1.533	2.132	2.776
5	0.267	0.559	0.920	1.156	1.476	2.015	2.571
6	0.265	0.553	0.906	1.134	1.440	1.943	2.447
7	0.263	0.549	0.896	1.119	1.415	1.895	2.365
8	0.262	0.546	0.889	1.108	1.397	1.860	2.306
9	0.261	0.543	0.883	1.100	1.383	1.833	2.262
10	0.260	0.542	0.879	1.093	1.372	1.812	2.228
11	0.260	0.540	0.876	1.088	1.363	1.796	2.201
12	0.259	0.539	0.873	1.083	1.356	1.782	2.179
13	0.259	0.538	0.870	1.079	1.350	1.771	2.160
14	0.258	0.537	0.868	1.076	1.345	1.761	2.145
15	0.258	0.536	0.866	1.074	1.341	1.753	2.131

$t_{14,0.05} = 1.761$

$P(t_{14,0.05} < T < t_{14,\alpha}) = 0.045$

$\Rightarrow P(T < t_{14,\alpha}) - P(T < t_{14,0.05}) = 0.045$

$\Rightarrow P(T < t_{14,\alpha}) - P(T > t_{14,0.95}) = 0.045$

$\Rightarrow P(T < t_{14,\alpha}) - \quad 0.95 \quad = 0.045$

$\Rightarrow P(T < t_{14,\alpha}) = 0.995$

$\Rightarrow P(T > t_{14,\alpha}) = 0.005$

$\Rightarrow k = t_{14,\alpha} = 2.977$

Q3. A random variable T with degree 17 and $P(-1.333 < T < k) = 0.75$,then the value of k is:

	α						
v	0.40	0.30	0.20	0.15	0.10	0.05	0.025
1	0.325	0.727	1.376	1.963	3.078	6.314	12.706
2	0.289	0.617	1.061	1.386	1.886	2.920	4.303
3	0.277	0.584	0.978	1.250	1.638	2.353	3.182
4	0.271	0.569	0.941	1.190	1.533	2.132	2.776
5	0.267	0.559	0.920	1.156	1.476	2.015	2.571
6	0.265	0.553	0.906	1.134	1.440	1.943	2.447
7	0.263	0.549	0.896	1.119	1.415	1.895	2.365
8	0.262	0.546	0.889	1.108	1.397	1.860	2.306
9	0.261	0.543	0.883	1.100	1.383	1.833	2.262
10	0.260	0.542	0.879	1.093	1.372	1.812	2.228
11	0.260	0.540	0.876	1.088	1.363	1.796	2.201
12	0.259	0.539	0.873	1.083	1.356	1.782	2.179
13	0.259	0.538	0.870	1.079	1.350	1.771	2.160
14	0.258	0.537	0.868	1.076	1.345	1.761	2.145
15	0.258	0.536	0.866	1.074	1.341	1.753	2.131
16	0.258	0.535	0.865	1.071	1.337	1.746	2.120
17	0.257	0.534	0.863	1.069	1.333	1.740	2.110
18	0.257	0.534	0.862	1.067	1.330	1.734	2.101

$$t_{17,0.10} = 1.333 \Longrightarrow t_{17,0.90} = -1.333$$

$$P(t_{17,0.90} < T < t_{17,\alpha}) = 0.75$$

$$\Longrightarrow P(T < t_{17,\alpha}) - P(T < t_{17,0.90}) = 0.75$$

$$\Longrightarrow P(T < t_{17,\alpha}) - P(T > t_{17,0.10}) = 0.75$$

$$\Longrightarrow P(T < t_{17,\alpha}) - 0.10 = 0.75$$

$$\Longrightarrow P(T < t_{17,\alpha}) = 0.85$$

$$\Longrightarrow P(T > t_{17,\alpha}) = 0.15$$

$$\Longrightarrow k = t_{17,\alpha} = 1.069$$

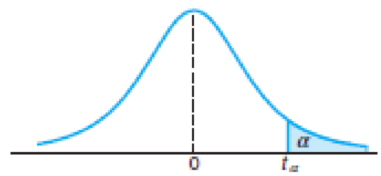


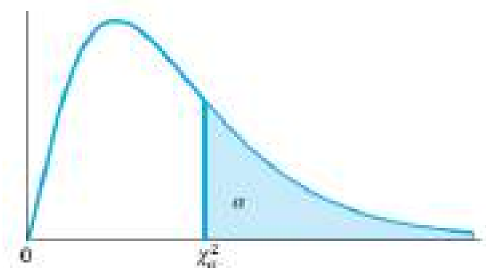
Table A.4 Critical Values of the t-Distribution

	α						
v	**0.40**	**0.30**	**0.20**	**0.15**	**0.10**	**0.05**	**0.025**
1	0.325	0.727	1.376	1.963	3.078	6.314	12.706
2	0.289	0.617	1.061	1.386	1.886	2.920	4.303
3	0.277	0.584	0.978	1.250	1.638	2.353	3.182
4	0.271	0.569	0.941	1.190	1.533	2.132	2.776
5	0.267	0.559	0.920	1.156	1.476	2.015	2.571
6	0.265	0.553	0.906	1.134	1.440	1.943	2.447
7	0.263	0.549	0.896	1.119	1.415	1.895	2.365
8	0.262	0.546	0.889	1.108	1.397	1.860	2.306
9	0.261	0.543	0.883	1.100	1.383	1.833	2.262
10	0.260	0.542	0.879	1.093	1.372	1.812	2.228
11	0.260	0.540	0.876	1.088	1.363	1.796	2.201
12	0.259	0.539	0.873	1.083	1.356	1.782	2.179
13	0.259	0.538	0.870	1.079	1.350	1.771	2.160
14	0.258	0.537	0.868	1.076	1.345	1.761	2.145
15	0.258	0.536	0.866	1.074	1.341	1.753	2.131
16	0.258	0.535	0.865	1.071	1.337	1.746	2.120
17	0.257	0.534	0.863	1.069	1.333	1.740	2.110
18	0.257	0.534	0.862	1.067	1.330	1.734	2.101
19	0.257	0.533	0.861	1.066	1.328	1.729	2.093
20	0.257	0.533	0.860	1.064	1.325	1.725	2.086
21	0.257	0.532	0.859	1.063	1.323	1.721	2.080
22	0.256	0.532	0.858	1.061	1.321	1.717	2.074
23	0.256	0.532	0.858	1.060	1.319	1.714	2.069
24	0.256	0.531	0.857	1.059	1.318	1.711	2.064
25	0.256	0.531	0.856	1.058	1.316	1.708	2.060
26	0.256	0.531	0.856	1.058	1.315	1.706	2.056
27	0.256	0.531	0.855	1.057	1.314	1.703	2.052
28	0.256	0.530	0.855	1.056	1.313	1.701	2.048
29	0.256	0.530	0.854	1.055	1.311	1.699	2.045
30	0.256	0.530	0.854	1.055	1.310	1.697	2.042
40	0.255	0.529	0.851	1.050	1.303	1.684	2.021
60	0.254	0.527	0.848	1.045	1.296	1.671	2.000
120	0.254	0.526	0.845	1.041	1.289	1.658	1.980
∞	0.253	0.524	0.842	1.036	1.282	1.645	1.960

Table A.4 (continued) Critical Values of the t-Distribution

	α						
v	**0.02**	**0.015**	**0.01**	**0.0075**	**0.005**	**0.0025**	**0.0005**
1	15.894	21.205	31.821	42.433	63.656	127.321	636.578
2	4.849	5.643	6.965	8.073	9.925	14.089	31.600
3	3.482	3.896	4.541	5.047	5.841	7.453	12.924
4	2.999	3.298	3.747	4.088	4.604	5.598	8.610
5	2.757	3.003	3.365	3.634	4.032	4.773	6.869
6	2.612	2.829	3.143	3.372	3.707	4.317	5.959
7	2.517	2.715	2.998	3.203	3.499	4.029	5.408
8	2.449	2.634	2.896	3.085	3.355	3.833	5.041
9	2.398	2.574	2.821	2.998	3.250	3.690	4.781
10	2.359	2.527	2.764	2.932	3.169	3.581	4.587
11	2.328	2.491	2.718	2.879	3.106	3.497	4.437
12	2.303	2.461	2.681	2.836	3.055	3.428	4.318
13	2.282	2.436	2.650	2.801	3.012	3.372	4.221
14	2.264	2.415	2.624	2.771	2.977	3.326	4.140
15	2.249	2.397	2.602	2.746	2.947	3.286	4.073
16	2.235	2.382	2.583	2.724	2.921	3.252	4.015
17	2.224	2.368	2.567	2.706	2.898	3.222	3.965
18	2.214	2.356	2.552	2.689	2.878	3.197	3.922
19	2.205	2.346	2.539	2.674	2.861	3.174	3.883
20	2.197	2.336	2.528	2.661	2.845	3.153	3.850
21	2.189	2.328	2.518	2.649	2.831	3.135	3.819
22	2.183	2.320	2.508	2.639	2.819	3.119	3.792
23	2.177	2.313	2.500	2.629	2.807	3.104	3.768
24	2.172	2.307	2.492	2.620	2.797	3.091	3.745
25	2.167	2.301	2.485	2.612	2.787	3.078	3.725
26	2.162	2.296	2.479	2.605	2.779	3.067	3.707
27	2.158	2.291	2.473	2.598	2.771	3.057	3.689
28	2.154	2.286	2.467	2.592	2.763	3.047	3.674
29	2.150	2.282	2.462	2.586	2.756	3.038	3.660
30	2.147	2.278	2.457	2.581	2.750	3.030	3.646
40	2.123	2.250	2.423	2.542	2.704	2.971	3.551
60	2.099	2.223	2.390	2.504	2.660	2.915	3.460
120	2.076	2.196	2.358	2.468	2.617	2.860	3.373
∞	2.054	2.170	2.326	2.432	2.576	2.807	3.290

<u>Chi-square Distribution</u>



$* \ E(\chi^2) = v \qquad * \ Var(\chi^2) = 2v$

$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{v=(n-1)} \qquad \chi^2_{v,\alpha} \Rightarrow P(\chi^2 > \chi^2_{v,\alpha}) = \alpha$

$\Rightarrow P(\chi^2 < \chi^2_{v,\alpha}) = 1 - \alpha$

Questions. For a chi-squared distribution:

(1) **the value χ^2_α such that $P(\chi^2 > \chi^2_\alpha) = 0.01$ when $v = 21$ is:**

$$\Rightarrow \chi^2_{v,\alpha} = \chi^2_{21,0.01} = 38.93$$

(2) **the value χ^2_α such that $P(\chi^2 > \chi^2_\alpha) = 0.95$ when $v = 14$ is:**

$$\Rightarrow \chi^2_{v,\alpha} = \chi^2_{14,0.95} = 6.571$$

(3) **the value χ^2_α such that $P(\chi^2_\alpha < X < 23.209) = 0.015$ when $v = 10$ is:**

Table A.5 (continued) Critical Values of the Chi-Squared Distribution

	α									
v	0.30	0.25	0.20	0.10	0.05	0.025	0.02	0.01	0.005	0.001
1	1.074	1.323	1.642	2.706	3.841	5.024	5.412	6.635	7.879	10.827
2	2.408	2.773	3.219	4.605	5.991	7.378	7.824	9.210	10.597	13.815
3	3.665	4.108	4.642	6.251	7.815	9.348	9.837	11.345	12.838	16.266
4	4.878	5.385	5.989	7.779	9.488	11.143	11.668	13.277	14.860	18.466
5	6.064	6.626	7.289	9.236	11.070	12.832	13.388	15.086	16.750	20.515
6	7.231	7.841	8.558	10.645	12.592	14.449	15.033	16.812	18.548	22.457
7	8.383	9.037	9.803	12.017	14.067	16.013	16.622	18.475	20.278	24.321
8	9.524	10.219	11.030	13.362	15.507	17.535	18.168	20.090	21.955	26.124
9	10.656	11.389	12.242	14.684	16.919	19.023	19.679	21.666	23.589	27.877
10	11.781	12.549	13.442	15.987	18.307	20.483	21.161	23.209	25.188	29.588
11	12.899	13.701	14.631	17.275	19.675	21.920	22.618	24.725	26.757	31.264

$$\chi^2_{0.01,10} = 23.209$$

$$\Longrightarrow P\left(\chi^2_{\alpha,10} < X < \chi^2_{0.01,10}\right) = 0.015$$

$$\Longrightarrow P\left(X < \chi^2_{0.01,10}\right) - P\left(X < \chi^2_{\alpha,10}\right) = 0.015$$

$$\Longrightarrow P\left(X > \chi^2_{0.99,10}\right) - P\left(X < \chi^2_{\alpha,10}\right) = 0.015$$

$$\Longrightarrow \quad 0.99 \quad - P\left(X < \chi^2_{\alpha,10}\right) = 0.015$$

$$\Longrightarrow P\left(X < \chi^2_{\alpha,10}\right) = 0.975$$

$$\Longrightarrow P\left(X > \chi^2_{\alpha,10}\right) = 0.025$$

$$\chi^2_{0.025,10} = 20.483$$

(4) **the value χ^2_α such that $P(6.57 < X < \chi^2_\alpha) = 0.85$ when $v = 14$ is:**

Table A.5 Critical Values of the Chi-Squared Distribution

v					α					
	0.995	0.99	0.98	0.975	0.95	0.90	0.80	0.75	0.70	0.50
1	0.0^4393	0.0^3157	0.0^3628	0.0^3982	0.00393	0.0158	0.0642	0.102	0.148	0.455
2	0.0100	0.0201	0.0404	0.0506	0.103	0.211	0.446	0.575	0.713	1.386
3	0.0717	0.115	0.185	0.216	0.352	0.584	1.005	1.213	1.424	2.366
4	0.207	0.297	0.429	0.484	0.711	1.064	1.649	1.923	2.195	3.357
5	0.412	0.554	0.752	0.831	1.145	1.610	2.343	2.675	3.000	4.351
6	0.676	0.872	1.134	1.237	1.635	2.204	3.070	3.455	3.828	5.348
7	0.989	1.239	1.564	1.690	2.167	2.833	3.822	4.255	4.671	6.346
8	1.344	1.647	2.032	2.180	2.733	3.490	4.594	5.071	5.527	7.344
9	1.735	2.088	2.532	2.700	3.325	4.168	5.380	5.899	6.393	8.343
10	2.156	2.558	3.059	3.247	3.940	4.865	6.179	6.737	7.267	9.342
11	2.603	3.053	3.609	3.816	4.575	5.578	6.989	7.584	8.148	10.341
12	3.074	3.571	4.178	4.404	5.226	6.304	7.807	8.438	9.034	11.340
13	3.565	4.107	4.765	5.009	5.892	7.041	8.634	9.299	9.926	12.340
14	4.075	4.660	5.368	5.629	6.571	7.790	9.467	10.165	10.821	13.339
15	4.601	5.229	5.985	6.262	7.261	8.547	10.307	11.037	11.721	14.339

$$\chi^2_{0.95,14} = 6.571$$

$$\Rightarrow P(\chi^2_{0.95,14} < X < \chi^2_{\alpha,14}) = 0.85$$

$$\Rightarrow P(X < \chi^2_{\alpha,14}) - P(X < \chi^2_{0.95,14}) = 0.85$$

$$\Rightarrow P(X < \chi^2_{\alpha,14}) - P(X > \chi^2_{0.05,14}) = 0.85$$

$$\Rightarrow P(X < \chi^2_{\alpha,14}) - \quad 0.05 \quad = 0.85$$

$$\Rightarrow P(X < \chi^2_{\alpha,14}) = 0.90$$

$$\Rightarrow P(X > \chi^2_{\alpha,14}) = 0.10$$

$$\Rightarrow \chi^2_{0.1,14} = 21.064$$

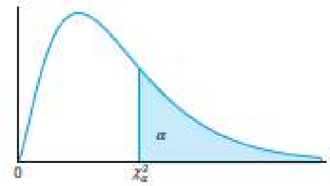


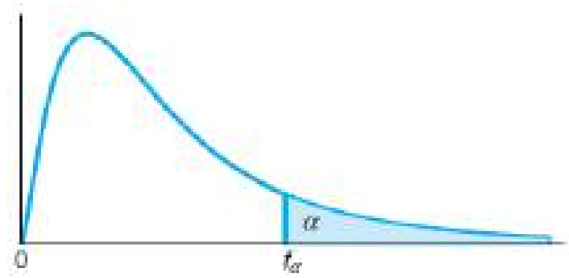
Table A.5 Critical Values of the Chi-Squared Distribution

	α									
v	**0.995**	**0.99**	**0.98**	**0.975**	**0.95**	**0.90**	**0.80**	**0.75**	**0.70**	**0.50**
1	$0.0^{4}393$	$0.0^{3}157$	$0.0^{3}628$	$0.0^{3}982$	0.00393	0.0158	0.0642	0.102	0.148	0.455
2	0.0100	0.0201	0.0404	0.0506	0.103	0.211	0.446	0.575	0.713	1.386
3	0.0717	0.115	0.185	0.216	0.352	0.584	1.005	1.213	1.424	2.366
4	0.207	0.297	0.429	0.484	0.711	1.064	1.649	1.923	2.195	3.357
5	0.412	0.554	0.752	0.831	1.145	1.610	2.343	2.675	3.000	4.351
6	0.676	0.872	1.134	1.237	1.635	2.204	3.070	3.455	3.828	5.348
7	0.989	1.239	1.564	1.690	2.167	2.833	3.822	4.255	4.671	6.346
8	1.344	1.647	2.032	2.180	2.733	3.490	4.594	5.071	5.527	7.344
9	1.735	2.088	2.532	2.700	3.325	4.168	5.380	5.899	6.393	8.343
10	2.156	2.558	3.059	3.247	3.940	4.865	6.179	6.737	7.267	9.342
11	2.603	3.053	3.609	3.816	4.575	5.578	6.989	7.584	8.148	10.341
12	3.074	3.571	4.178	4.404	5.226	6.304	7.807	8.438	9.034	11.340
13	3.565	4.107	4.765	5.009	5.892	7.041	8.634	9.299	9.926	12.340
14	4.075	4.660	5.368	5.629	6.571	7.790	9.467	10.165	10.821	13.339
15	4.601	5.229	5.985	6.262	7.261	8.547	10.307	11.037	11.721	14.339
16	5.142	5.812	6.614	6.908	7.962	9.312	11.152	11.912	12.624	15.338
17	5.697	6.408	7.255	7.564	8.672	10.085	12.002	12.792	13.531	16.338
18	6.265	7.015	7.906	8.231	9.390	10.865	12.857	13.675	14.440	17.338
19	6.844	7.633	8.567	8.907	10.117	11.651	13.716	14.562	15.352	18.338
20	7.434	8.260	9.237	9.591	10.851	12.443	14.578	15.452	16.266	19.337
21	8.034	8.897	9.915	10.283	11.591	13.240	15.445	16.344	17.182	20.337
22	8.643	9.542	10.600	10.982	12.338	14.041	16.314	17.240	18.101	21.337
23	9.260	10.196	11.293	11.689	13.091	14.848	17.187	18.137	19.021	22.337
24	9.886	10.856	11.992	12.401	13.848	15.659	18.062	19.037	19.943	23.337
25	10.520	11.524	12.697	13.120	14.611	16.473	18.940	19.939	20.867	24.337
26	11.160	12.198	13.409	13.844	15.379	17.292	19.820	20.843	21.792	25.336
27	11.808	12.878	14.125	14.573	16.151	18.114	20.703	21.749	22.719	26.336
28	12.461	13.565	14.847	15.308	16.928	18.939	21.588	22.657	23.647	27.336
29	13.121	14.256	15.574	16.047	17.708	19.768	22.475	23.567	24.577	28.336
30	13.787	14.953	16.306	16.791	18.493	20.599	23.364	24.478	25.508	29.336
40	20.707	22.164	23.838	24.433	26.509	29.051	32.345	33.66	34.872	39.335
50	27.991	29.707	31.664	32.357	34.764	37.689	41.449	42.942	44.313	49.335
60	35.534	37.485	39.699	40.482	43.188	46.459	50.641	52.294	53.809	59.335

Table A.5 (continued) Critical Values of the Chi-Squared Distribution

	α									
v	**0.30**	**0.25**	**0.20**	**0.10**	**0.05**	**0.025**	**0.02**	**0.01**	**0.005**	**0.001**
1	1.074	1.323	1.642	2.706	3.841	5.024	5.412	6.635	7.879	10.827
2	2.408	2.773	3.219	4.605	5.991	7.378	7.824	9.210	10.597	13.815
3	3.665	4.108	4.642	6.251	7.815	9.348	9.837	11.345	12.838	16.266
4	4.878	5.385	5.989	7.779	9.488	11.143	11.668	13.277	14.860	18.466
5	6.064	6.626	7.289	9.236	11.070	12.832	13.388	15.086	16.750	20.515
6	7.231	7.841	8.558	10.645	12.592	14.449	15.033	16.812	18.548	22.457
7	8.383	9.037	9.803	12.017	14.067	16.013	16.622	18.475	20.278	24.321
8	9.524	10.219	11.030	13.362	15.507	17.535	18.168	20.090	21.955	26.124
9	10.656	11.389	12.242	14.684	16.919	19.023	19.679	21.666	23.589	27.877
10	11.781	12.549	13.442	15.987	18.307	20.483	21.161	23.209	25.188	29.588
11	12.899	13.701	14.631	17.275	19.675	21.920	22.618	24.725	26.757	31.264
12	14.011	14.845	15.812	18.549	21.026	23.337	24.054	26.217	28.300	32.909
13	15.119	15.984	16.985	19.812	22.362	24.736	25.471	27.688	29.819	34.527
14	16.222	17.117	18.151	21.064	23.685	26.119	26.873	29.141	31.319	36.124
15	17.322	18.245	19.311	22.307	24.996	27.488	28.259	30.578	32.801	37.698
16	18.418	19.369	20.465	23.542	26.296	28.845	29.633	32.000	34.267	39.252
17	19.511	20.489	21.615	24.769	27.587	30.191	30.995	33.409	35.718	40.791
18	20.601	21.605	22.760	25.989	28.869	31.526	32.346	34.805	37.156	42.312
19	21.689	22.718	23.900	27.204	30.144	32.852	33.687	36.191	38.582	43.819
20	22.775	23.828	25.038	28.412	31.410	34.170	35.020	37.566	39.997	45.314
21	23.858	24.935	26.171	29.615	32.671	35.479	36.343	38.932	41.401	46.796
22	24.939	26.039	27.301	30.813	33.924	36.781	37.659	40.289	42.796	48.268
23	26.018	27.141	28.429	32.007	35.172	38.076	38.968	41.638	44.181	49.728
24	27.096	28.241	29.553	33.196	36.415	39.364	40.270	42.980	45.558	51.179
25	28.172	29.339	30.675	34.382	37.652	40.646	41.566	44.314	46.928	52.619
26	29.246	30.435	31.795	35.563	38.885	41.923	42.856	45.642	48.290	54.051
27	30.319	31.528	32.912	36.741	40.113	43.195	44.140	46.963	49.645	55.475
28	31.391	32.620	34.027	37.916	41.337	44.461	45.419	48.278	50.994	56.892
29	32.461	33.711	35.139	39.087	42.557	45.722	46.693	49.588	52.335	58.301
30	33.530	34.800	36.250	40.256	43.773	46.979	47.962	50.892	53.672	59.702
40	44.165	45.616	47.269	51.805	55.758	59.342	60.436	63.691	66.766	73.403
50	54.723	56.334	58.164	63.167	67.505	71.420	72.613	76.154	79.490	86.660
60	65.226	66.981	68.972	74.397	79.082	83.298	84.58	88.379	91.952	99.608

Fisher (F) Distribution



$$* \ E(F) = \frac{v_1}{v_2 - 2} \qquad * \ Var(F) = \frac{2v_2^2(v_1 + v_2 - 2)}{v_1(v_2 - 2)^2(v_2 - 4)} \ ; \ v_2 > 4$$

$$\frac{\chi^2_{v_1}}{\chi^2_{v_2}} \sim F_{v_1,v_2} \qquad F_{v_1,v_2,\alpha} = \frac{1}{F_{v_2,v_1,1-\alpha}} \qquad F_{v_1,v_2} \Rightarrow P\left(F > F_{v_1,v_2,\alpha}\right) = \alpha$$

$$\Rightarrow P\left(F < F_{v_1,v_2,\alpha}\right) = 1 - \alpha$$

Q.1. For the F distribution, then the value F such that $P\left(F > F_{8,6}\right) = 0.05$ is:

$$F_{8,6,0.05} = 4.15$$

Q.2. For the F distribution, then the value F such that $P\left(F > F_{8,6}\right) = 0.99$ is:

$$F_{8,6,0.99} = \frac{1}{F_{6,8,0.01}} = \frac{1}{6.37} = 0.157$$

Q.3.For the F distribution with degrees 8 and 6, then:

(a). the value a such that $P(F > a) = 0.95$

$$a = F_{8,6,0.95} = \frac{1}{F_{6,8,0.05}} = \frac{1}{3.58} = 0.2793$$

(b). the value b such that $P(F > b) = 0.01$

$$b = F_{8,6,0.01} = 8.10$$

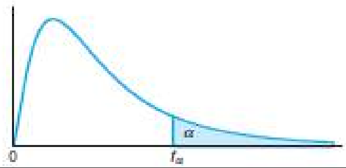


Table A.6 Critical Values of the F-Distribution

	$f_{0.05}(v_1, v_2)$								
	v_1								
v_2	**1**	**2**	**3**	**4**	**5**	**6**	**7**	**8**	**9**
1	161.45	199.50	215.71	224.58	230.16	233.99	236.77	238.88	240.54
2	18.51	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.38
3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81
4	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00
5	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77
6	5.99	5.14	4.76	4.53	4.39	4.28	4.21	4.15	4.10
7	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68
8	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39
9	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18
10	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02
11	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90
12	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80
13	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71
14	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65
15	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59
16	4.49	3.63	3.24	3.01	2.85	2.74	2.66	2.59	2.54
17	4.45	3.59	3.20	2.96	2.81	2.70	2.61	2.55	2.49
18	4.41	3.55	3.16	2.93	2.77	2.66	2.58	2.51	2.46
19	4.38	3.52	3.13	2.90	2.74	2.63	2.54	2.48	2.42
20	4.35	3.49	3.10	2.87	2.71	2.60	2.51	2.45	2.39
21	4.32	3.47	3.07	2.84	2.68	2.57	2.49	2.42	2.37
22	4.30	3.44	3.05	2.82	2.66	2.55	2.46	2.40	2.34
23	4.28	3.42	3.03	2.80	2.64	2.53	2.44	2.37	2.32
24	4.26	3.40	3.01	2.78	2.62	2.51	2.42	2.36	2.30
25	4.24	3.39	2.99	2.76	2.60	2.49	2.40	2.34	2.28
26	4.23	3.37	2.98	2.74	2.59	2.47	2.39	2.32	2.27
27	4.21	3.35	2.96	2.73	2.57	2.46	2.37	2.31	2.25
28	4.20	3.34	2.95	2.71	2.56	2.45	2.36	2.29	2.24
29	4.18	3.33	2.93	2.70	2.55	2.43	2.35	2.28	2.22
30	4.17	3.32	2.92	2.69	2.53	2.42	2.33	2.27	2.21
40	4.08	3.23	2.84	2.61	2.45	2.34	2.25	2.18	2.12
60	4.00	3.15	2.76	2.53	2.37	2.25	2.17	2.10	2.04
120	3.92	3.07	2.68	2.45	2.29	2.18	2.09	2.02	1.96
∞	3.84	3.00	2.60	2.37	2.21	2.10	2.01	1.94	1.88

Table A.6 (continued) Critical Values of the *F*-Distribution

	$f_{0.05}(v_1, v_2)$									
	v_1									
v_2	**10**	**12**	**15**	**20**	**24**	**30**	**40**	**60**	**120**	∞
1	241.88	243.91	245.95	248.01	249.05	250.10	251.14	252.20	253.25	254.31
2	19.40	19.41	19.43	19.45	19.45	19.46	19.47	19.48	19.49	19.50
3	8.79	8.74	8.70	8.66	8.64	8.62	8.59	8.57	8.55	8.53
4	5.96	5.91	5.86	5.80	5.77	5.75	5.72	5.69	5.66	5.63
5	4.74	4.68	4.62	4.56	4.53	4.50	4.46	4.43	4.40	4.36
6	4.06	4.00	3.94	3.87	3.84	3.81	3.77	3.74	3.70	3.67
7	3.64	3.57	3.51	3.44	3.41	3.38	3.34	3.30	3.27	3.23
8	3.35	3.28	3.22	3.15	3.12	3.08	3.04	3.01	2.97	2.93
9	3.14	3.07	3.01	2.94	2.90	2.86	2.83	2.79	2.75	2.71
10	2.98	2.91	2.85	2.77	2.74	2.70	2.66	2.62	2.58	2.54
11	2.85	2.79	2.72	2.65	2.61	2.57	2.53	2.49	2.45	2.40
12	2.75	2.69	2.62	2.54	2.51	2.47	2.43	2.38	2.34	2.30
13	2.67	2.60	2.53	2.46	2.42	2.38	2.34	2.30	2.25	2.21
14	2.60	2.53	2.46	2.39	2.35	2.31	2.27	2.22	2.18	2.13
15	2.54	2.48	2.40	2.33	2.29	2.25	2.20	2.16	2.11	2.07
16	2.49	2.42	2.35	2.28	2.24	2.19	2.15	2.11	2.06	2.01
17	2.45	2.38	2.31	2.23	2.19	2.15	2.10	2.06	2.01	1.96
18	2.41	2.34	2.27	2.19	2.15	2.11	2.06	2.02	1.97	1.92
19	2.38	2.31	2.23	2.16	2.11	2.07	2.03	1.98	1.93	1.88
20	2.35	2.28	2.20	2.12	2.08	2.04	1.99	1.95	1.90	1.84
21	2.32	2.25	2.18	2.10	2.05	2.01	1.96	1.92	1.87	1.81
22	2.30	2.23	2.15	2.07	2.03	1.98	1.94	1.89	1.84	1.78
23	2.27	2.20	2.13	2.05	2.01	1.96	1.91	1.86	1.81	1.76
24	2.25	2.18	2.11	2.03	1.98	1.94	1.89	1.84	1.79	1.73
25	2.24	2.16	2.09	2.01	1.96	1.92	1.87	1.82	1.77	1.71
26	2.22	2.15	2.07	1.99	1.95	1.90	1.85	1.80	1.75	1.69
27	2.20	2.13	2.06	1.97	1.93	1.88	1.84	1.79	1.73	1.67
28	2.19	2.12	2.04	1.96	1.91	1.87	1.82	1.77	1.71	1.65
29	2.18	2.10	2.03	1.94	1.90	1.85	1.81	1.75	1.70	1.64
30	2.16	2.09	2.01	1.93	1.89	1.84	1.79	1.74	1.68	1.62
40	2.08	2.00	1.92	1.84	1.79	1.74	1.69	1.64	1.58	1.51
60	1.99	1.92	1.84	1.75	1.70	1.65	1.59	1.53	1.47	1.39
120	1.91	1.83	1.75	1.66	1.61	1.55	1.50	1.43	1.35	1.25
∞	1.83	1.75	1.67	1.57	1.52	1.46	1.39	1.32	1.22	1.00

Table A.6 (continued) Critical Values of the F-Distribution

$f_{0.01}(v_1, v_2)$

v_2 \ v_1	1	2	3	4	5	6	7	8	9
1	4052.18	4999.50	5403.35	5624.58	5763.65	5858.99	5928.36	5981.07	6022.47
2	98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39
3	34.12	30.82	29.46	28.71	28.24	27.91	27.67	27.49	27.35
4	21.20	18.00	16.69	15.98	15.52	15.21	14.98	14.80	14.66
5	16.26	13.27	12.06	11.39	10.97	10.67	10.46	10.29	10.16
6	13.75	10.92	9.78	9.15	8.75	8.47	8.26	8.10	7.98
7	12.25	9.55	8.45	7.85	7.46	7.19	6.99	6.84	6.72
8	11.26	8.65	7.59	7.01	6.63	6.37	6.18	6.03	5.91
9	10.56	8.02	6.99	6.42	6.06	5.80	5.61	5.47	5.35
10	10.04	7.56	6.55	5.99	5.64	5.39	5.20	5.06	4.94
11	9.65	7.21	6.22	5.67	5.32	5.07	4.89	4.74	4.63
12	9.33	6.93	5.95	5.41	5.06	4.82	4.64	4.50	4.39
13	9.07	6.70	5.74	5.21	4.86	4.62	4.44	4.30	4.19
14	8.86	6.51	5.56	5.04	4.69	4.46	4.28	4.14	4.03
15	8.68	6.36	5.42	4.89	4.56	4.32	4.14	4.00	3.89
16	8.53	6.23	5.29	4.77	4.44	4.20	4.03	3.89	3.78
17	8.40	6.11	5.18	4.67	4.34	4.10	3.93	3.79	3.68
18	8.29	6.01	5.09	4.58	4.25	4.01	3.84	3.71	3.60
19	8.18	5.93	5.01	4.50	4.17	3.94	3.77	3.63	3.52
20	8.10	5.85	4.94	4.43	4.10	3.87	3.70	3.56	3.46
21	8.02	5.78	4.87	4.37	4.04	3.81	3.64	3.51	3.40
22	7.95	5.72	4.82	4.31	3.99	3.76	3.59	3.45	3.35
23	7.88	5.66	4.76	4.26	3.94	3.71	3.54	3.41	3.30
24	7.82	5.61	4.72	4.22	3.90	3.67	3.50	3.36	3.26
25	7.77	5.57	4.68	4.18	3.85	3.63	3.46	3.32	3.22
26	7.72	5.53	4.64	4.14	3.82	3.59	3.42	3.29	3.18
27	7.68	5.49	4.60	4.11	3.78	3.56	3.39	3.26	3.15
28	7.64	5.45	4.57	4.07	3.75	3.53	3.36	3.23	3.12
29	7.60	5.42	4.54	4.04	3.73	3.50	3.33	3.20	3.09
30	7.56	5.39	4.51	4.02	3.70	3.47	3.30	3.17	3.07
40	7.31	5.18	4.31	3.83	3.51	3.29	3.12	2.99	2.89
60	7.08	4.98	4.13	3.65	3.34	3.12	2.95	2.82	2.72
120	6.85	4.79	3.95	3.48	3.17	2.96	2.79	2.66	2.56
∞	6.63	4.61	3.78	3.32	3.02	2.80	2.64	2.51	2.41

Table A.6 (continued) Critical Values of the F-Distribution

	$f_{0.01}(v_1, v_2)$									
	v_1									
v_2	**10**	**12**	**15**	**20**	**24**	**30**	**40**	**60**	**120**	∞
1	6055.85	6106.32	6157.28	6208.73	6234.63	6260.65	6286.78	6313.03	6339.39	6365.86
2	99.40	99.42	99.43	99.45	99.46	99.47	99.47	99.48	99.49	99.50
3	27.23	27.05	26.87	26.69	26.60	26.50	26.41	26.32	26.22	26.13
4	14.55	14.37	14.20	14.02	13.93	13.84	13.75	13.65	13.56	13.46
5	10.05	9.89	9.72	9.55	9.47	9.38	9.29	9.20	9.11	9.02
6	7.87	7.72	7.56	7.40	7.31	7.23	7.14	7.06	6.97	6.88
7	6.62	6.47	6.31	6.16	6.07	5.99	5.91	5.82	5.74	5.65
8	5.81	5.67	5.52	5.36	5.28	5.20	5.12	5.03	4.95	4.86
9	5.26	5.11	4.96	4.81	4.73	4.65	4.57	4.48	4.40	4.31
10	4.85	4.71	4.56	4.41	4.33	4.25	4.17	4.08	4.00	3.91
11	4.54	4.40	4.25	4.10	4.02	3.94	3.86	3.78	3.69	3.60
12	4.30	4.16	4.01	3.86	3.78	3.70	3.62	3.54	3.45	3.36
13	4.10	3.96	3.82	3.66	3.59	3.51	3.43	3.34	3.25	3.17
14	3.94	3.80	3.66	3.51	3.43	3.35	3.27	3.18	3.09	3.00
15	3.80	3.67	3.52	3.37	3.29	3.21	3.13	3.05	2.96	2.87
16	3.69	3.55	3.41	3.26	3.18	3.10	3.02	2.93	2.84	2.75
17	3.59	3.46	3.31	3.16	3.08	3.00	2.92	2.83	2.75	2.65
18	3.51	3.37	3.23	3.08	3.00	2.92	2.84	2.75	2.66	2.57
19	3.43	3.30	3.15	3.00	2.92	2.84	2.76	2.67	2.58	2.49
20	3.37	3.23	3.09	2.94	2.86	2.78	2.69	2.61	2.52	2.42
21	3.31	3.17	3.03	2.88	2.80	2.72	2.64	2.55	2.46	2.36
22	3.26	3.12	2.98	2.83	2.75	2.67	2.58	2.50	2.40	2.31
23	3.21	3.07	2.93	2.78	2.70	2.62	2.54	2.45	2.35	2.26
24	3.17	3.03	2.89	2.74	2.66	2.58	2.49	2.40	2.31	2.21
25	3.13	2.99	2.85	2.70	2.62	2.54	2.45	2.36	2.27	2.17
26	3.09	2.96	2.81	2.66	2.58	2.50	2.42	2.33	2.23	2.13
27	3.06	2.93	2.78	2.63	2.55	2.47	2.38	2.29	2.20	2.10
28	3.03	2.90	2.75	2.60	2.52	2.44	2.35	2.26	2.17	2.06
29	3.00	2.87	2.73	2.57	2.49	2.41	2.33	2.23	2.14	2.03
30	2.98	2.84	2.70	2.55	2.47	2.39	2.30	2.21	2.11	2.01
40	2.80	2.66	2.52	2.37	2.29	2.20	2.11	2.02	1.92	1.80
60	2.63	2.50	2.35	2.20	2.12	2.03	1.94	1.84	1.73	1.60
120	2.47	2.34	2.19	2.03	1.95	1.86	1.76	1.66	1.53	1.38
∞	2.32	2.18	2.04	1.88	1.79	1.70	1.59	1.47	1.32	1.00

CHAPTER 3

- Sampling Distributions:
- Single mean.
- Two means.
- Single proportion.
- Two proportions.

Single Mean	Two Means
$\bar{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$ $E(\bar{X}) = \bar{X} = \mu$ $Var(\bar{X}) = \frac{\sigma^2}{n}$	$\bar{X}_1 - \bar{X}_2 \sim N\left(\mu_1 - \mu_2\ , \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$ $E(\bar{X}_1 - \bar{X}_2) = \mu_1 - \mu_2$ $Var(\bar{X}_1 - \bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$
Single Proportion	**Two Proportions**
For large sample size $(n \geq 30,\ np > 5,\ nq > 5)$ $\hat{p} \sim N\left(p, \sqrt{\frac{pq}{n}}\right)$ $E(\hat{p}) = p$ $Var(\hat{p}) = \frac{pq}{n}$	For large sample size $(n_1 \geq 30,\ n_1p_1 > 5,\ n_1q_1 > 5)$ $(n_2 \geq 30,\ n_2p_2 > 5,\ n_2q_2 > 5)$ $\hat{p}_1 - \hat{p}_2 \sim N\left(p_1 - p_2, \sqrt{\frac{p_1q_1}{n_1} + \frac{p_2q_2}{n_2}}\right)$ $E(\hat{p}_1 - \hat{p}_2) = p_1 - p_2$ $Var(\hat{p}_1 - \hat{p}_2) = \frac{p_1q_1}{n_1} + \frac{p_2q_2}{n_2}$

Sampling Distribution

Sampling Distribution: Single Mean

$$* \ \overline{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$$

$$* \ E(\overline{X}) = \overline{X} = \mu \qquad * \ Var(\overline{X}) = \frac{\sigma^2}{n}$$

Question 1: The average life of a certain battery is 5 years, with a standard deviation of 1 year. Assume that the live of the battery approximately follows a normal distribution.

1) The sample mean X of a random sample of 5 batteries selected from this product has a mean $E(\overline{X}) = \mu$ equal to:

$$\mu = 5 \ ; \ \sigma = 1 \ ; \ n = 5$$

$$E(\overline{X}) = \mu = 5$$

2) The variance $Var(\overline{X})$ of the sample mean X of a random sample of 5 batteries selected from this product is equal to:

$$Var(\overline{X}) = \frac{\sigma^2}{n} = \frac{1}{5} = 0.2$$

4) The probability that the average life of a random sample of <u>size 16</u> of such batteries will be between 4.5 and 5.4 years is:

$$n = 16 \rightarrow \frac{\sigma}{\sqrt{n}} = \frac{1}{4}$$

$$P(4.5 < \overline{X} < 5.4) = P\left(\frac{4.5-\mu}{\frac{\sigma}{\sqrt{n}}} < Z < \frac{5.4-\mu}{\frac{\sigma}{\sqrt{n}}}\right) = P\left(\frac{4.5-5}{\frac{1}{4}} < Z < \frac{5.4-5}{\frac{1}{4}}\right)$$

$$= P(-2 < Z < 1.6)$$

$$= P(Z < 1.6) - P(Z < -2)$$

$$= 0.9452 - 0.0228 = 0.9224$$

5) The probability that the average life of a random sample of size 16 of such batteries will be less than 5.5 years is:

$$P(\bar{X} < 5.5) = P\left(Z < \frac{5.5-\mu}{\frac{\sigma}{\sqrt{n}}}\right) = P\left(Z < \frac{5.5-5}{1/4}\right) = P(Z < 2) = 0.9772$$

6) If $P(\bar{X} > a) = 0.1492$ where X represents the sample mean for a random sample of size 9 of such batteries, then the numerical value of a is:

$P(\bar{X} > a) = 0.1492$; $n = 9$

$$P\left(Z > \frac{a-\mu}{\frac{\sigma}{\sqrt{n}}}\right) = 0.1492$$

$$\Rightarrow 1 - P\left(Z < \frac{a-5}{\frac{1}{3}}\right) = 0.1492$$

$$\Rightarrow P\left(Z < \frac{a-5}{\frac{1}{3}}\right) = 0.8508$$

$$\frac{a-5}{\frac{1}{3}} = 1.04$$

$$a = 5 + \frac{1.04}{3} = 5.347$$

Question 2: Suppose that you take a random sample of size $n = 64$ from a distribution with mean $\mu = 55$ and standard deviation $\sigma = 10$. Let X be the sample mean.

(a) What is the approximated sampling distribution of X?

$$\mu = 55 \quad ; \quad \sigma = 10 \quad ; \quad n = 64$$

$$\bar{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right) = \bar{X} \sim N\left(55, \frac{10}{8}\right)$$

(b) What is the mean of $\bar{X}$?

$$E(\bar{X}) = \mu = 55$$

(c) What is the standard error (standard deviation) of $\bar{X}$?

$$S.D(\bar{X}) = \frac{\sigma}{\sqrt{n}} = \frac{10}{\sqrt{64}} = \frac{10}{8}$$

(d) Find the probability that the sample mean x exceeds 52.

$$\text{(a) } P(\bar{X} > 52) = P\left(Z > \frac{52-55}{\frac{10}{8}}\right)$$

$$= P(Z > -2.4)$$

$$= 1 - P(Z < -2.4)$$

$$= 1 - 0.0082 = 0.9918$$

Sampling Distribution: Two Means

$$* \ \bar{X}_1 - \bar{X}_2 \sim N\left(\mu_1 - \mu_2 \ , \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$$

$$* \ E(\bar{X}_1 - \bar{X}_2) = \mu_1 - \mu_2 \qquad * \ Var(\bar{X}_1 - \bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

Question 3: A random sample of size $n_1 = 36$ is taken from a normal population with a mean $\mu_1 = 70$ and a standard deviation $\sigma_1 = 4$. A second independent random sample of size $n_2 = 49$ is taken from a normal population with a mean $\mu_2 = 85$ and a standard deviation $\sigma_2 = 5$. Let $\bar{X}_1$ and $\bar{X}_2$ be the averages of the first and second samples, respectively.

$$\boxed{\begin{array}{l} n_1 = 36, \mu_1 = 70, \sigma_1 = 4 \\ n_2 = 49, \mu_2 = 85, \sigma_2 = 5 \end{array}}$$

a. Find $E(\bar{X}_1 - \bar{X}_2)$:

$$E(\bar{X}_1 - \bar{X}_2) = \mu_1 - \mu_2 = 70 - 85 = -15$$

a. Find $Var(\bar{X}_1 - \bar{X}_2)$:

$$Var(\bar{X}_1 - \bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2} = \frac{16}{36} + \frac{25}{49} = 0.955$$

$$S.D(\bar{X}_1 - \bar{X}_2) = \sqrt{0.955}$$

b. Find $P(\bar{X}_1 - \bar{X}_2 > -16)$:

$$= \quad P\left(Z > \frac{-16-(-15)}{\sqrt{0.955}}\right)$$

$$= 1 - P\left(Z < \frac{-16-(-15)}{\sqrt{0.955}}\right)$$

$$= 1 - P(Z < -1.02) = 0.8461$$

Question 4: The distribution of heights of a certain breed of terrier has a mean of 72 centimeters and a standard deviation of 10 centimeters, whereas the distribution of heights of a certain breed of poodle has a mean of 28 centimeters with a standard deviation of 5 centimeters. Assuming that the sample means can be measured to any degree of accuracy, find the probability that the sample mean for a random sample of heights of 64 terriers exceeds the sample mean for a random sample of heights of 100 poodles by <u>less than 44.2</u> centimeter

$n_1 = 64 \ , \ \mu_1 = 72 \ , \ \sigma_1 = 10$
$n_2 = 100 \ , \ \mu_2 = 28 \ , \ \sigma_2 = 5$

$$E(\bar{X}_1 - \bar{X}_2) = \mu_1 - \mu_2 = 72 - 28 = 44$$

$$Var(\bar{X}_1 - \bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2} = \frac{100}{64} + \frac{25}{100} = 1.8125$$

$$P(\bar{X}_1 - \bar{X}_2 < 44.2) =$$

$$= P\left(Z < \frac{44.2 - (44)}{\sqrt{1.8125}}\right)$$

$$= P(Z < 0.15)$$

$$= 0.5596$$

Sampling Distribution: Single Proportion

$$* \ \hat{p} \sim N\left(p, \sqrt{\frac{pq}{n}}\right)$$

$$* \ E(\hat{p}) = p \qquad * \ Var(\hat{p}) = \frac{pq}{n}$$

Question 5: Suppose that you take a random sample of size n=100 from a population with proportion of diabetic equal to p =0.25. Let pˆ be the sample proportion of of diabetic.

(a) What is the mean and the variance of $\hat{p}$?

$$\boxed{p = 0.25 \ ; \quad n = 100}$$

$$E(\hat{p}) = p = 0.25$$

$$Var(\hat{p}) = \frac{pq}{n} = \frac{0.25 \times 0.75}{100} = 0.003$$

(b) What is the approximated sampling distribution of $\hat{p}$?

$$\hat{p} \sim N\left(0.25, \sqrt{0.003}\right)$$

(c) Find the probability that the sample proportion $\hat{p}$ is less than 0.2.

$$P(\hat{p} < 0.2) = P\left(Z < \frac{0.2 - 0.25}{\sqrt{0.003}}\right) = P(Z < -0.91) = 0.1814$$

Sampling Distribution: Two Proportions

$$* \ \hat{p}_1 - \hat{p}_2 \sim N\left(p_1 - p_2, \sqrt{\frac{p_1q_1}{n_1} + \frac{p_2q_2}{n_2}}\right)$$

$$* \ E(\hat{p}_1 - \hat{p}_2) = p_1 - p_2 \qquad * \ Var(\hat{p}_1 - \hat{p}_2) = \frac{p_1q_1}{n_1} + \frac{p_2q_2}{n_2}$$

Question 6: Suppose that 25% of the male students and 20% of the female students in a certain university smoke cigarette. A random sample of 5 male students is taken. Another random sample of 10 female students is independently taken from this university. Let and be the proportions of smokers in the two samples, respectively.

$$p_1 = 0.25 \ ; \quad n_1 = 5$$

$$p_2 = 0.20 \ ; \quad n_2 = 10$$

(1): $E(\hat{p}_1 - \hat{p}_2) = p_1 - p_2 = 0.25 - 0.2 = 0.05$

(2): $Var(\hat{p}_1 - \hat{p}_2) = \frac{p_1q_1}{n_1} + \frac{p_2q_2}{n_2} = \frac{0.25\times0.75}{5} + \frac{0.2\times0.8}{10} = 0.054$

(3): $\hat{p}_1 - \hat{p}_2 \sim N(0.05, \sqrt{0.054})$

(4): $P(0.1 < \hat{p}_1 - \hat{p}_2 < 0.2) = P\left(\frac{0.1-0.05}{\sqrt{0.054}} < Z < \frac{0.2-0.05}{\sqrt{0.054}}\right)$

$$= P(0.22 < Z < 0.65)$$

$$= P(Z < 0.65) - P(Z < 0.22)$$

$$= 0.7422 - 0.5871 = 0.1551$$

Sampling Distribution		
single Population	Mean	$\bar{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right)$
	Proportion	$\hat{p} \sim N\left(p, \sqrt{\frac{pq}{n}}\right)$
Two Populations	Means	$\bar{X}_1 - \bar{X}_2 \sim N\left(\mu_1 - \mu_2 , \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$
	Proportions	$\hat{p}_1 - \hat{p}_2 \sim N\left(p_1 - p_2, \sqrt{\frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}}\right)$

CHAPTER 4

- Estimation (point estimation and confidence interval):

- Single mean
- Two means
- Single proportion
- Two proportion
- Single variance.
- The ratio of two variances.

Estimation and Confidence Interval

Estimation and Confidence Interval: Single Mean:

- To find the confidence intervals for a single mean:

1- $\overline{X} \pm \left(Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$

2- $\overline{X} \pm \left(t_{\frac{\alpha}{2},n-1} \frac{S}{\sqrt{n}}\right)$

- $Z_{\frac{\alpha}{2}} = a \Longrightarrow P(Z > a) = \frac{\alpha}{2}$

- To estimate an error:

$$e = Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$$

- To estimate the sample size with particular error:

$$n = \left(\frac{Z_{\frac{\alpha}{2}} \sigma}{e}\right)^2$$

.

Q1. Suppose that we are interested in making some statistical inferences about the mean, μ, of a normal population with standard deviation $\sigma = 2$. Suppose that a random sample of size $n = 49$ from this population gave a sample mean $\overline{X} = 4.5$.

$$\boxed{\sigma = 2 \;\;,\;\; \overline{X} = 4.5 \;\;,\;\; n = 49}$$

(1) The distribution of $\overline{X}$:

$$\overline{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right) \Longrightarrow \overline{X} \sim N\left(\mu, \frac{2}{\sqrt{49}}\right) \Longrightarrow \overline{X} \sim N\left(\mu, \frac{2}{7}\right)$$

(2) A good point estimate of μ is:

$$\hat{\mu} = \overline{X} = 4.5$$

(3) The standard error of $\overline{X}$ is:

$$S.E\,(\overline{X}) = \frac{2}{7} = 0.2857$$

(4) A 95% confidence interval for μ is:

$$\overline{X} \pm \left(Z_{\frac{\alpha}{2}} \; \frac{\sigma}{\sqrt{n}}\right)$$

$$95\% \rightarrow \alpha = 0.05$$

$$Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96$$

$$4.5 \pm \left(1.96 \times \frac{2}{7}\right)$$

The 95% confidence interval is: $(3.94,\; 5.06)$

(5) If the upper confidence limit of a confidence interval is 5.2, then the lower confidence limit is

The upper limit $= 4.5 + e = 5.2$

and the lower limit $= 4.5 - e = ?$

$$4.5 + e = 5.2 \;\Rightarrow\; e = 5.2 - 4.5 \Rightarrow (e = 0.7)$$

Then, the lower limit $= 4.5 - e = 4.5 - 0.7 = 3.8$

(6) The confidence level of the confidence interval (3.88, 5.12) is

We have the interval:(3.88,5.12):

$$\text{The upper limit} = 5.12$$

$$\Rightarrow \overline{X} + \left(Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right) = 5.12$$

$$\Rightarrow 4.5 + \left(Z_{\frac{\alpha}{2}} \times \frac{2}{7}\right) = 5.12$$

$$\Rightarrow Z_{\frac{\alpha}{2}} = 2.17 \Rightarrow \frac{\alpha}{2} = 1 - 0.985$$

$$\Rightarrow \frac{\alpha}{2} = 0.015 \Rightarrow \alpha = 0.03$$

Hence, the confidence level is 97%.

Or, we can do the same thing with the lower limit:

$$\text{The lower limit} = 3.88$$

$$\Rightarrow \overline{X} - \left(Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right) = 3.88$$

$$\Rightarrow 4.5 - \left(Z_{\frac{\alpha}{2}} \times \frac{2}{7}\right) = 3.88$$

$$\Rightarrow Z_{\frac{\alpha}{2}} = 2.17 \Rightarrow \frac{\alpha}{2} = 1 - 0.985$$

$$\Rightarrow \frac{\alpha}{2} = 0.015 \Rightarrow \alpha = 0.03$$

Hence, the confidence level is 97%

(7) If we use $\overline{X}$ to estimate μ, then we are 95% confident that our estimation error will not exceed.

$$95\% \rightarrow \alpha = 0.05 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96$$

$$e = \frac{Z_{\frac{\alpha}{2}} \times \sigma}{\sqrt{n}} = \frac{1.96 \times 2}{7} = 0.56$$

(8) If we want to be 95% confident that the estimation error will not exceed e=0.1 when we use $\overline{X}$ to estimate μ, then the sample size n must be equal to

$$95\% \rightarrow \alpha = 0.05 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96 \;\; \& \;\; e = 0.1$$

$$n = \left(\frac{Z_{\frac{\alpha}{2}} \times \sigma}{e}\right)^2 = \left(\frac{1.96 \times 2}{0.1}\right)^2 = 1536.64 \simeq 1537$$

Q2. The following measurements were recorded for lifetime, in years, of certain type of machine: 3.4, 4.8, 3.6, 3.3, 5.6, 3.7, 4.4, 5.2, and 4.8. Assuming that the measurements represent a random sample from a normal population, then a 99% confidence interval for the mean life time of the machine is

$$99\% \rightarrow \alpha = 0.01 \rightarrow t_{\frac{\alpha}{2}, n-1} = t_{0.005, 8} = 3.355.$$

$$\overline{X} \pm \left(t_{\frac{\alpha}{2}, n-1} \times \frac{S}{\sqrt{n}}\right)$$

$$= 4.31 \pm \left(3.355 \times \frac{0.84}{3}\right)$$

$$\overline{X} = \frac{\sum x_i}{n} = 4.31$$

$$S^2 = \frac{\sum (x_i - \overline{X})^2}{n-1} = 0.71$$

$$S^2 = 0.71 \Rightarrow S = 0.84$$

The 99% confidence interval is: $(3.37, \; 5.25)$

Q3. A researcher wants to estimate the mean lifespan of a certain light bulbs. Suppose that the distribution is normal with standard deviation of 5 hours.

1. **Determine the sample size needed on order that the researcher will be 90% confident that the error will not exceed 2 hours when he uses the sample mean as a point estimate for the true mean.**

$$\sigma = 5 \quad \& \quad e = 2 \quad \& \quad \alpha = 0.10$$

$$\alpha = 0.10 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.05} = 1.645$$

$$n = \left(\frac{Z_{\frac{\alpha}{2}} \times \sigma}{e}\right)^2 = \left(\frac{1.645 \times 5}{2}\right)^2 = 16.9 \simeq 17$$

2. **Suppose that the researcher selected a random sample of 49 bulbs and found that the sample mean is 390 hours.**

(i) Find a good point estimate for the true mean μ.

$$\hat{\mu} = \overline{X} = 390$$

(ii) Find a 95% confidence interval for the true mean μ.

$$\sigma = 5 \quad \& \quad n = 49 \quad \& \quad \overline{X} = 390$$

$$\overline{X} \pm \left(Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$$

$$390 \pm \left(1.96 \frac{5}{\sqrt{49}}\right)$$

$$95\% \rightarrow \alpha = 0.05$$

$$Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96$$

The 95% confidence interval is: (388.6, 391.3)

Q4. The amount of time that customers using ATM (Automatic Teller Machine) is a random variable with a standard deviation of 1.4 minutes. If we wish to estimate the population mean μ by the sample mean, and if we want to be 96% confident that the sample mean will be within 0.3 minutes of the population mean, then the sample size needed is X

$$\boxed{\sigma = 1.4 \ , \ e = 0.3 \ , \ \alpha = 0.04}$$

$$\alpha = 0.04 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.02} = 2.055$$

$$n = \left(\frac{Z_{\frac{\alpha}{2}} \times \sigma}{e}\right)^2 = \left(\frac{2.055 \times 1.4}{0.3}\right)^2 = 91.9 \simeq 92$$

	population normal or not normal n large $(n \geq 30)$		population normal n small $(n < 30)$	
	σ known	σ unknown	σ known	σ unknown
Estimation	$\overline{X} \pm \left(Z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$	$\overline{X} \pm \left(Z_{1-\frac{\alpha}{2}} \frac{s}{\sqrt{n}}\right)$	$\overline{X} \pm \left(Z_{1-\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$	$\overline{X} \pm \left(t_{1-\frac{\alpha}{2},n-1} \times \frac{S}{\sqrt{n}}\right)$
testing	$Z = \frac{\overline{X} - \mu_0}{\sigma/\sqrt{n}}$	$Z = \frac{\overline{X} - \mu_0}{s/\sqrt{n}}$	$Z = \frac{\overline{X} - \mu_0}{\sigma/\sqrt{n}}$	$T = \frac{\overline{X} - \mu_0}{s/\sqrt{n}}$

<u>Estimation and Confidence Interval: Two Means</u>

To find the confidence intervals for two means:

1- $(\bar{X}_1 - \bar{X}_2) \pm \left(Z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right)$

2- $(\bar{X}_1 - \bar{X}_2) \pm \left(t_{\frac{\alpha}{2}, n_1+n_2-2}\, Sp \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \right)$

$$Sp^2 = \frac{S_1^2(n_1 - 1) + S_2^2(n_2 - 1)}{n_1 + n_2 - 2}$$

Q1. The tensile strength of type I thread is approximately normally distributed with standard deviation of 6.8 kilograms. A sample of 20 pieces of the thread has an average tensile strength of 72.8 kilograms. Then,

The tensile strength of type II thread is approximately normally distributed with standard deviation of 6.8 kilograms. A sample of 25 pieces of the thread has an average tensile strength of 64.4 kilograms. Then for the 98% confidence interval of the difference in tensile strength means between type I and type II, we have:

Theard 1 :	$n_1 = 20, \bar{X}_1 = 72.8, \sigma_1 = 6.8$
Theard 2 :	$n_2 = 25, \bar{X}_2 = 64.4, \sigma_2 = 6.8$

$$98\% \rightarrow \alpha = 0.02 \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.01} = 2.33$$

$$(\bar{X}_1 - \bar{X}_2) \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} \right)$$

$$(72.8 - 64.4) \pm \left(2.33 \times \sqrt{\frac{6.8^2}{20} + \frac{6.8^2}{25}} \right)$$

$$8.4 \pm (2.33)(2.04) = (\ 3.65\ , 13.15\ )$$

Q2. Two random samples were independently selected from two normal populations with equal variances. The results are summarized as follows.

	First sample	Second sample
Sample size (n)	12	14
Sample mean ($\overline{X}$)	10.5	10
Sample variance (S^2)	4	5

1. **Find a point estimate for $\mu_1 - \mu_2$.**

$$E(\overline{X}_1 - \overline{X}_2) = 10.5 - 10 = 0.5$$

2. **Find 95% confidence interval for $\mu_1 - \mu_2$.**

$$\alpha = 0.05 \rightarrow t_{\frac{\alpha}{2}, n_1+n_2-2} = t_{0.025, 24} = 2.064$$

$$(\overline{X}_1 - \overline{X}_2) \pm \left(t_{\frac{\alpha}{2}, n_1+n_2-2}\, Sp\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}\right)$$

$$(0.5) \pm \left(2.064 \times 2.13 \times \sqrt{\frac{1}{12} + \frac{1}{14}}\right)$$

$$(-1.23, 2.23)$$

$$Sp^2 = \frac{S_1^2(n_1 - 1) + S_2^2(n_2 - 1)}{n_1 + n_2 - 2}$$

$$= \frac{4(11)+5(13)}{24} = 4.54$$

$$Sp^2 = 4.54 \Rightarrow Sp = 2.13$$

<u>Estimation and Confidence Interval: Single Proportion</u>

* Point estimate for P is: $\frac{x}{n}$

* Interval estimate for P is: $\hat{p} \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}\hat{q}}{n}}\right)$

Q1: A random sample of 200 students from a certain school showed that 15 students smoke. Let p be the proportion of smokers in the school.

1. **Find a point Estimate for p.**

$$n = 200 \quad \& \quad x = 15$$

$$\hat{p} = \frac{x}{n} = \frac{15}{200} = 0.075 \rightarrow \hat{q} = 0.925$$

2. **Find 95% confidence interval for p.**

$95\% \rightarrow \alpha = 0.05 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96$

$$\hat{p} \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}\hat{q}}{n}}\right)$$

$$= 0.075 \pm \left(1.96 \times \sqrt{\frac{0.075 \times 0.925}{200}}\right)$$

The 95% confidence interval is: $(0.038,\ 0.112)$

Estimation and Confidence Interval: Two Proportions

$$* \text{ Point estimate for } P_1 - P_2 = \hat{p}_1 - \hat{p}_2 = \frac{x_1}{n_1} - \frac{x_2}{n_2}$$

$$* \text{ Interval estimate for } P_1 - P_2 \text{ is: } (\hat{p}_1 - \hat{p}_2) \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}_1\hat{q}_1}{n_1} + \frac{\hat{p}_2\hat{q}_2}{n_2}} \right)$$

Q1. A random sample of 100 students from school "A" showed that 15 students smoke. Another independent random sample of 200 students from school "B" showed that 20 students smoke. Let p1 be the proportion of smokers in school "A" and p2 is the proportion of smokers in school "B".

(1) Find a point Estimate for $p_1 - p_2$.

$$n_1 = 100 \quad x_1 = 15 \quad \rightarrow \hat{p}_1 = \frac{15}{100} = 0.15$$

$$n_2 = 200 \quad x_2 = 20 \rightarrow \hat{p}_2 = \frac{20}{200} = 0.10$$

$$\hat{p}_1 - \hat{p}_2 = 0.15 - 0.1 = 0.05$$

(2) Find 95% confidence interval for $p_1 - p_2$.

$$95\% \rightarrow \alpha = 0.05 \quad \rightarrow \quad Z_{\frac{\alpha}{2}} = Z_{0.025} = 1.96$$

$$(\hat{p}_1 - \hat{p}_2) \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}_1\hat{q}_1}{n_1} + \frac{\hat{p}_2\hat{q}_2}{n_2}} \right)$$

$$= (0.05) \pm \left(1.96 \times \sqrt{\frac{(0.15)(0.85)}{100} + \frac{(0.1)(0.9)}{200}} \right)$$

$$= 0.05 \pm \left(1.96 \times \sqrt{0.001725} \right)$$

The 95% confidence interval is: $(-0.031,\ 0.131)$

<u>Confidence Interval for the Population Variance and Standard Deviation:</u>

$$P\left(\frac{(n-1)S^2}{\chi^2_{(n-1),\frac{\alpha}{2}}} < \sigma^2 < \frac{(n-1)S^2}{\chi^2_{(n-1),1-\frac{\alpha}{2}}}\right) = 1-\alpha$$

$$P\left(\frac{S_1^2}{S_2^2} \times \frac{1}{F_{v_1,v_2,\frac{\alpha}{2}}} < \frac{\sigma_1^2}{\sigma_2^2} < \frac{S_1^2}{S_2^2} \times F_{v_2,v_1,\frac{\alpha}{2}}\right) = 1-\alpha$$

Q1. We are interested in the content of a soft-drink dispensing machine. A random sample of 25 drinks gave a variance of 2.03 deciliters^2. Assume that the contents are approximately normally distributed.

$$\boxed{n = 25 \quad , \quad S^2 = 2.03}$$

(1) The point estimate of the population variance of the contents is:

$$S^2 = 2.03$$

(2) The <u>lower bound</u> of the of the 90 % confidence interval of the population variance σ^2 is:

$$\frac{(n-1)S^2}{\chi^2_{(n-1),\frac{\alpha}{2}}} = \frac{(25-1)\times 2.03}{\chi^2_{(25-1),\frac{0.1}{2}}} = \frac{(24)\times 2.03}{\chi^2_{(24),0.05}} = \frac{48.72}{36.415} = 1.34$$

(3) The <u>upper bound</u> of the of the 90 % confidence interval of the population variance σ^2 is:

$$\frac{(n-1)S^2}{\chi^2_{(n-1),1-\frac{\alpha}{2}}} = \frac{(25-1)\times 2.03}{\chi^2_{(25-1),1-\frac{0.1}{2}}} = \frac{(24)\times 2.03}{\chi^2_{(24),095}} = \frac{48.72}{13.848} = 3.52$$

Q2. In a series of experiments to determine the absorption rate of certain pesticides into skin, measured amounts of two pesticides were applied to several skin specimens. For pesticide A, the variance of the amounts absorbed in $n_1 = 6$ specimens were $S_1^2 = 2.3$, while for pesticide B, the variance of the amounts absorbed in $n_2 = 10$ specimens was $S_2^2 = 0.6$. Assume that for each pesticide, the amounts absorbed are a simple random sample from a normal population.

$$n_1 = 6 \quad , \quad S_1^2 = 2.3$$
$$n_2 = 10 \, , \quad S_2^2 = 0.6$$

1. Find A point estimate of the ratio σ_1^2/σ_2^2 of the two population variances is:

$$\frac{S_1^2}{S_2^2} = \frac{2.3}{0.6} = 3.833$$

2. The <u>lower bound</u> of the of the 90 % confidence interval of the ratio σ_1^2/σ_2^2

$$\frac{S_1^2}{S_2^2} \times \frac{1}{F_{v_1,v_2,\frac{\alpha}{2}}} = \frac{2.3}{0.6} \times \frac{1}{F_{5,9,\frac{0.1}{2}}} = 3.833 \times \frac{1}{F_{5,9,0.05}} = 3.833 \times \frac{1}{3.48} = 1.1$$

3. The <u>upper bound</u> of the of the 90 % confidence interval of the ratio σ_1^2/σ_2^2

$$\frac{S_1^2}{S_2^2} \times F_{v_2,v_1,\frac{\alpha}{2}} = \frac{2.3}{0.6} \times F_{9,5,\frac{0.1}{2}} = 3.833 \times F_{9,5,0.05} = 3.833 \times 4.77 = 18.28$$

Estimation and Confident intervals		
single Population	Mean	$\bar{X} \pm \left(Z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right)$ σ known
		$\bar{X} \pm \left(t_{\frac{\alpha}{2},n-1} \frac{S}{\sqrt{n}}\right)$ σ unknown
	Proportion	$\hat{p} \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}\hat{q}}{n}}\right)$
	Variance	$\frac{(n-1)S^2}{\chi^2_{(n-1),\frac{\alpha}{2}}} < \sigma^2 < \frac{(n-1)S^2}{\chi^2_{(n-1),1-\frac{\alpha}{2}}}$
Two Populations	Means	$(\bar{X}_1 - \bar{X}_2) \pm \left(Z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}\right)$ σ_1 and σ_2 known
		$(\bar{X}_1 - \bar{X}_2) \pm \left(t_{\frac{\alpha}{2},n_1+n_2-2}\ Sp \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}\right)$ σ_1 and σ_2 unknown $S_P^2 = \frac{S_1^2(n_1-1) + S_2^2(n_2-1)}{n_1+n_2-2}$
	Proportions	$(\hat{p}_1 - \hat{p}_2) \pm \left(Z_{\frac{\alpha}{2}} \times \sqrt{\frac{\hat{p}_1\hat{q}_1}{n_1} + \frac{\hat{p}_2\hat{q}_2}{n_2}}\right)$
	Ratio of two Variance	$\frac{S_1^2}{S_2^2} \times \frac{1}{F_{v_1,v_2,\frac{\alpha}{2}}} < \frac{\sigma_1^2}{\sigma_2^2} < \frac{S_1^2}{S_2^2} \times F_{v_2,v_1,\frac{\alpha}{2}}$

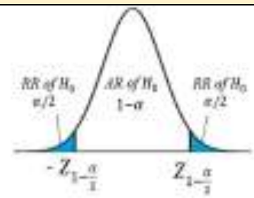
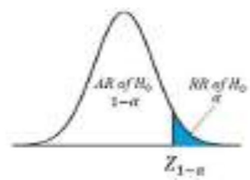
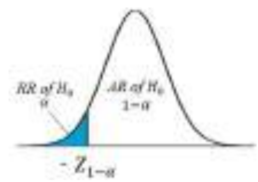
CHAPTER 5

- Testing of The Hypothesis for:
  - Single mean
  - Two means
  - Single proportion
  - Two proportion
  - Single variance.
  - Two variances.
  - Two Samples Paired Observation

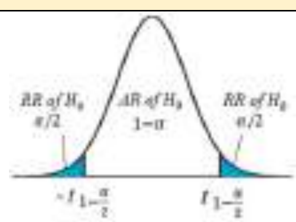
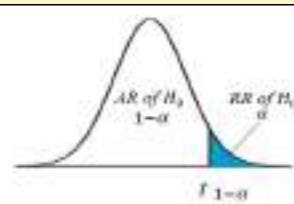
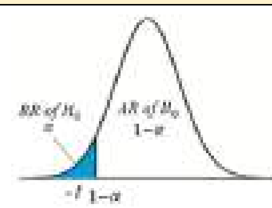
Hypotheses Testing

Single Mean

(if σ known):

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: \mu = \mu_o$ $H_A: \mu \neq \mu_o$	$H_0: \mu = \mu_o$ $H_A: \mu > \mu_o$	$H_0: \mu = \mu_o$ $H_A: \mu < \mu_o$
Test Statistics (TS)	$Z = \dfrac{\overline{X} - \mu_0}{\sigma/\sqrt{n}} \sim N(0,1)$		
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$Z < -Z_{1-(\alpha/2)}$ or $Z > Z_{1-(\alpha/2)}$ Two sides test	$Z > Z_{1-\alpha}$ One side test	$Z < -Z_{1-\alpha}$ One side test

(if σ unknown):

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: \mu = \mu_o$ $H_A: \mu \neq \mu_o$	$H_0: \mu = \mu_o$ $H_A: \mu > \mu_o$	$H_0: \mu = \mu_o$ $H_A: \mu < \mu_o$
Test Statistics (TS)	$t = \dfrac{\overline{X} - \mu_0}{S/\sqrt{n}} \sim t_{n-1}$		
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$t < -t_{1-(\alpha/2)}$ or $t > t_{1-(\alpha/2)}$ Two sides test	$t > t_{1-\alpha}$ One side test	$t < -t_{1-\alpha}$ One side test

Question 1:

Suppose that we are interested in making some statistical inferences about the mean μ , of a normal population with standard deviation σ =2. Suppose that a random sample of size n=49 from this population gave a sample mean $\bar{X} = 4.5$.

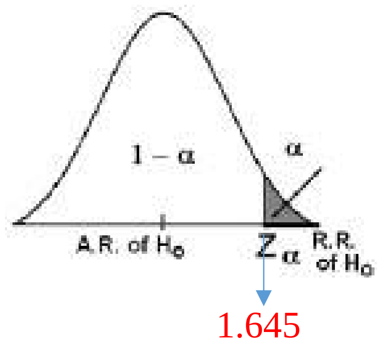
$$\boxed{\sigma = 2 \quad , \quad n = 49 \quad , \quad \bar{X} = 4.5}$$

1. **If we want to test $H_0: \mu = 5$ vs $H_1: \mu > 5$, then the test statistic equals to**

$$Z = \frac{\bar{X} - \mu_o}{\sigma/\sqrt{n}} = \frac{4.5 - 5}{2/7} = -1.75$$

2. **If we want to test, $H_0: \mu = 5$ vs $H_1: \mu > 5$, the Rejection Region of H_0**

$$\alpha = 0.05 \rightarrow Z_\alpha = Z_{0.05} = 1.645$$



The Rejection Region (R.R) is $(1.645, \infty)$

3. **If we want to test $H_0: \mu = 5$ vs $H_1: \mu > 5$at, $\alpha = 0.05$ then we**

$$Z = -1.75 \notin R.R = (1.645, \infty)$$

Then we accept H_0.

$$\boxed{P - value = P(Z > -1.75) = 1 - P(Z < -1.75) = 1 - 0.0401 = 0.9599 > \alpha}$$

<u>Question 2:</u>

An electrical firm manufactures light bulbs that have a length of life that is normally distributed with a standard deviation of 30 hours. A sample of 50 bulbs were selected randomly and found to have an average of 750 hours. Let μ be the population mean of life of all bulbs manufactured by this firm.

Test $H_0: \mu = 740$ vs $H_1: \mu < 740$? Use a 0.05 level of significance.

$$\boxed{\sigma = 30 \quad , \quad n = 50 \quad , \quad \overline{X} = 750}$$

- Hypotheses:

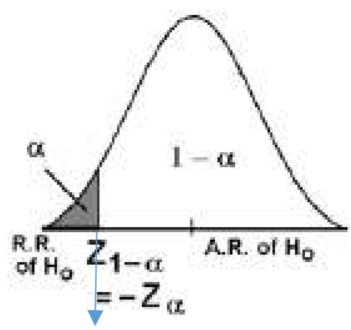
$$H_0: \mu = 740 \quad \text{vs} \quad H_1: \mu < 740$$

- Test Statistic (T.S):

$$Z = \frac{\overline{X} - \mu_o}{\sigma/\sqrt{n}} = \frac{750 - 740}{30/\sqrt{50}} = 2.37$$

- Rejection Region (R.R):

$$Z_\alpha = Z_{0.05} = 1.645$$



- Decision: -1.645

$$Z = 2.37 \notin \text{R.R, then we accept } H_0$$

$$\boxed{P - \text{value} = P(Z < 2.37) = 0.9911 > \alpha}$$

Question 3:

A random sample of size $\mathbf{n = 36}$ from a normal quantitative population produced a mean $\bar{X} = 15.2$ and a variance $S^2 = 9$.

Test $H_0: \mu = 15$ vs $H_1: \mu \neq 15$, use $\alpha = 0.05$.

$$\boxed{s = 3 \ , \ n = 36 \ , \ \bar{X} = 15.2 \ , \ \alpha = 0.05}$$

- Hypotheses:

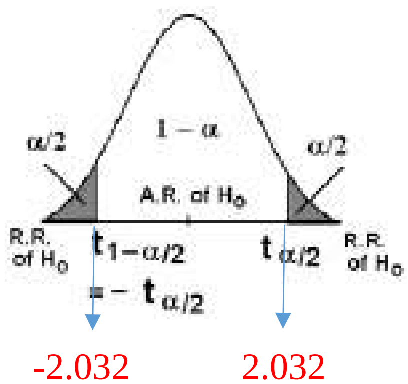
$$H_0: \mu = 15 \quad \text{vs} \quad H_1: \mu \neq 15$$

- Test statistic (T.S):

$$t = \frac{\bar{X} - \mu_o}{S/\sqrt{n}} = \frac{15.2 - 15}{3/\sqrt{36}} = 0.4$$

- Rejection Region (R.R):

$$t_{\frac{\alpha}{2}, n-1} = t_{0.025, 35} = 2.032$$

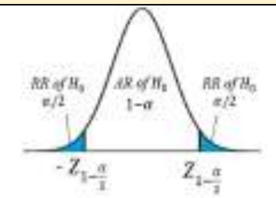
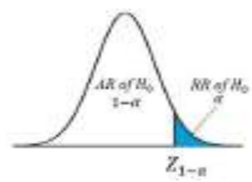
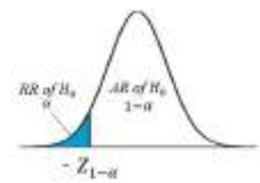


- Decision:

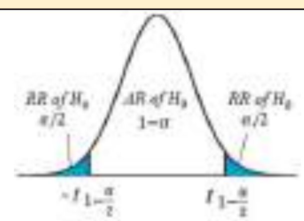
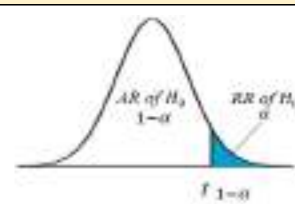
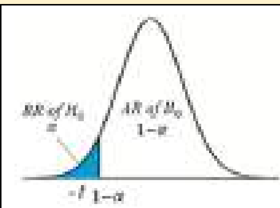
$$t = 0.4 \notin R.R, \text{then we accept } H_0$$

Two Means:

(if σ_1 and σ_2 known):

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 \neq d$	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 > d$	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 < d$
Test Statistics (TS)	$Z = \frac{(\bar{X}_1 - \bar{X}_2) - d}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0,1)$		
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$Z < -Z_{1-(\alpha/2)}$ or $Z > Z_{1-(\alpha/2)}$ Two sides test	$Z > Z_{1-\alpha}$ One side test	$Z < -Z_{1-\alpha}$ One side test

(if σ_1 and σ_2 unknown):

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 \neq d$	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 > d$	$H_0: \mu_1 - \mu_2 = d$ $H_A: \mu_1 - \mu_2 < d$
Test Statistics (TS)	$t = \frac{(\bar{X}_1 - \bar{X}_2) - d}{\sqrt{\frac{S_p^2}{n_1} + \frac{S_p^2}{n_2}}} = \frac{(\bar{X}_1 - \bar{X}_2) - d}{S_p\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1+n_2-2}$		$S_p^2 = \frac{S_1^2(n_1-1) + S_2^2(n_2-1)}{n_1+n_2-2}$
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$t < -t_{1-(\alpha/2)}$ or $t > t_{1-(\alpha/2)}$ Two sides test	$t > t_{1-\alpha}$ One side test	$t < -t_{1-\alpha}$ One side test

Question 1:

A standardized chemistry test was given to 50 girls and 75 boys. The girls made an average of 84, while the boys made an average grade of 82. Assume the population standard deviations are 6 and 8 for girls and boys respectively. To test the null hypothesis

$$H_0: \mu_1 - \mu_2 \leq 0 \ \text{ vs } H_A: \mu_1 - \mu_2 > 0 \quad \text{use } \alpha = 0.05$$

(1) The standard error of $(\overline{X}_1 - \overline{X}_2)$ is:

girls: $n_1 = 50, \ \overline{X}_1 = 84, \ \sigma_1 = 6$

boys: $n_2 = 75, \ \overline{X}_2 = 82, \ \sigma_2 = 8$

$$S.E(\overline{X}_1 - \overline{X}_2) = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} = \sqrt{\frac{6^2}{50} + \frac{8^2}{75}} = 1.2543$$

(2) The value of the test statistic is:

$$Z = \frac{(\overline{X}_1 - \overline{X}_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} = \frac{(84-82)}{\sqrt{\frac{6^2}{50} + \frac{8^2}{75}}} = \frac{2}{1.2543} = \boxed{1.5945}$$

(3) The rejection region (RR) of H_0 is:

$$Z_{1-\alpha} = Z_{1-0.05} = Z_{0.95} = 1.645$$

1.645

A	$(1.645, \infty)$	B	$(-\infty, -1.645)$	C	$(1.96, \infty)$	D	$(-\infty, -1.96)$

(4) The decision is:

A	Reject H_0
B	Do not reject (Accept) H_0.
C	Accept both H_0 and H_A.
D	Reject both H_0 and H_A.

$$\boxed{P - \text{value} = P(Z > 1.59) = 1 - P(Z < 1.59) = 0.056 > 0.05}$$

Question 2:

Two random samples were independently selected from two normal populations with equal variances. The results are summarized as follows:

	First sample	Second sample
Sample size (n)	**12**	**14**
Sample mean $(\bar{X})$	**10.5**	**10**
Sample variance (S^2)	**4**	**5**

Let μ_1 and μ_2 be the true means of the first and second populations, respectively. Test $H_0: \mu_1 = \mu_2$ vs $H_1: \mu_1 \neq \mu_2$.(use $\alpha = 0.05$)

- Hypotheses:

$$H_0: \mu_1 = \mu_2 \quad \text{or} \quad H_0: \mu_1 - \mu_2 = 0$$

$$H_1: \mu_1 \neq \mu_2 \qquad\quad H_1: \mu_1 - \mu_2 \neq 0$$

- Test statistic (T.S):

$$t = \frac{(\bar{X}_1 - \bar{X}_2) - d}{Sp\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

$$= \frac{(10.5 - 10) - 0}{2.13\sqrt{\frac{1}{12} + \frac{1}{14}}} = 0.597$$

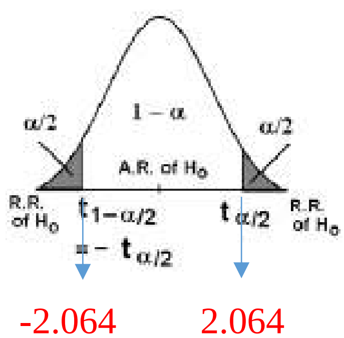
$$Sp^2 = \frac{S_1^2(n_1 - 1) + S_2^2(n_2 - 1)}{n_1 + n_2 - 2}$$

$$= \frac{4(11) + 5(13)}{24} = 4.54$$

$$Sp^2 = 4.54 \Rightarrow Sp = 2.13$$

- Rejection Region (R.R):

$$\alpha = 0.05 \rightarrow t_{\frac{\alpha}{2}, n_1 + n_2 - 2} = t_{0.025, 24} = 2.064$$

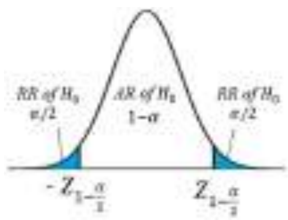
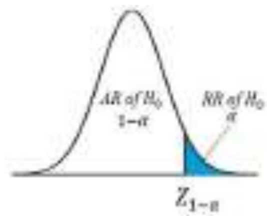
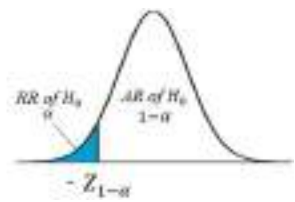


- Decision:

$$t = 0.597 \notin R.R \text{ Then we accept } H_0.$$

$$\boxed{P - \text{value} = 2 \times P(t_{24} < 0.597) \approx 2 \times 0.3 = 0.6 > \alpha}$$

<u>Single Proportion:</u>

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: p = p_o$ $H_A: p \neq p_o$	$H_0: p = p_o$ $H_A: p > p_o$	$H_0: p = p_o$ $H_A: p < p_o$
Test Statistics (TS)	$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}} \sim N(0,1)$		
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$Z < -Z_{1-(\alpha/2)}$ or $Z > Z_{1-(\alpha/2)}$ Two sides test	$Z > Z_{1-\alpha}$ One side test	$Z < -Z_{1-\alpha}$ One side test

Question 1:

A researcher was interested in making some statistical inferences about the proportion of smokers (p) among the students of a certain university. A random sample of 500 students showed that 150 students smoke.

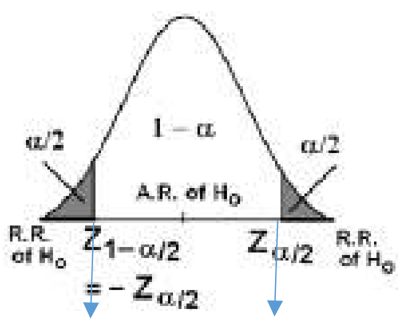
$$n = 500 \quad , \quad x = 150 \ \rightarrow \hat{p} = \frac{150}{500} = 0.3$$

(1) If we want to test $H_0: p = 0.25$ vs $H_1: p \neq 0.25$ then the test statistic:

$$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}} = \frac{0.3 - 0.25}{\sqrt{\frac{0.25 \times 0.75}{500}}} = 2.58$$

(2) If we want to test $H_0: p = 0.25$ vs $H_1: p \neq 0.25$ at $\alpha = 0.1$, then the Acceptance Region of H_0 is

$$\alpha = 0.1 \rightarrow Z_{\frac{\alpha}{2}} = Z_{0.05} = 1.645 \rightarrow A.R = (-1.645, 1.645).$$



-1.645 1.645

(3) If we want to test $H_0: p = 0.25$ vs $H_1: p \neq 0.25$ at $\alpha = 0.1$, then we:

$$Z = 2.58 \notin A.R, \text{ then we reject } H_0.$$

$$P - value = 2 \times P(Z > 2.58) = 2 \times 0.0049 = 0.0098 < \alpha$$

<u>Question 2:</u>

In a random sample of 500 homes in a certain city, it is found that 114 are heated by oil. Let p be the proportion of homes in this city that are heated by oil. A builder claims that less than 20% of the homes in this city are heated by oil. Would you agree with this claim? Use a 0.02 level of significance.

$$\boxed{n = 500 \quad , \quad x = 114 \; \rightarrow \hat{p} = \frac{114}{500} = 0.23}$$

- Hypotheses:

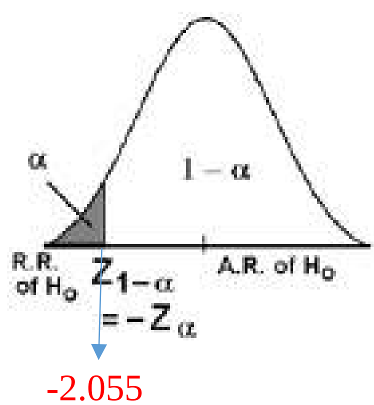
$$H_0: p = 0.2 \quad \text{vs} \quad H_1: p < 0.2$$

- Test statistic (T.S):

$$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}} = \frac{0.23 - 0.2}{\sqrt{\frac{0.2 \times 0.8}{500}}} = 1.67$$

- Rejection Region (R.R):

$$\alpha = 0.02 \; \rightarrow Z_\alpha = Z_{0.02} = 2.055$$



- Decision:

$$Z = 1.67 \; \notin \text{R.R, then we accept } H_0.$$

$$\boxed{P - \text{value} = P(Z > 1.67) = 1 - P(Z < 1.67) = 1 - 0.9525 = 0.0475 > \alpha}$$

- **Two Samples Test for Paired Observation**

Q1: The following contains the calcium levels of eleven test subjects at zero hours and three hours after taking a multi-vitamin containing calcium.

Pair	0 hour (X_i)	3 hours (Y_i)	Difference $D_i = X_i - Y_i$
1	17.0	17.0	0.0
2	13.2	12.9	0.3
3	35.3	35.4	-0.1
4	13.6	13.2	0.4
5	32.7	32.5	0.2
6	18.4	18.1	0.3
7	22.5	22.5	0.0
8	26.8	26.7	0.1
9	15.1	15.0	0.1

The sample mean and sample standard deviation of the differences D are 0.144 and 0.167, respectively. To test whether the data provide sufficient evidence to indicate a difference in mean calcium levels ($H_0: \mu_1 = \mu_2$ against$H_1: \mu_1 \neq \mu_2$) with $\alpha = 0.10$ we have: $\boxed{\bar{D} = 0.144 \;,\quad S_d = 0.167 \quad,\quad n = 9}$

[1]. the reliability coefficient (the tabulated value) is:

$$t_{\frac{\alpha}{2},n-1} = t_{\frac{0.1}{2},9-1} = t_{0.05,8} = \boxed{1.860}$$

[2]. the value of the test statistic is:

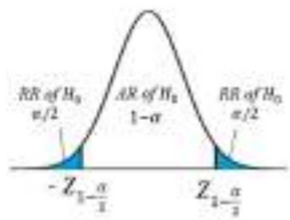
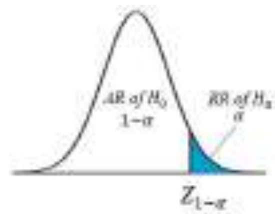
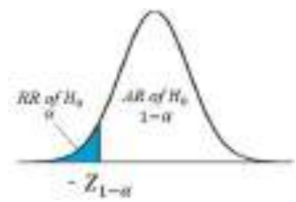
$$\boxed{\begin{array}{l} H_0: \mu_1 = \mu_2 \\ H_1: \mu_1 \neq \mu_2 \end{array}} \Rightarrow \boxed{\begin{array}{l} H_0: \mu_1 - \mu_2 = 0 \\ H_1: \mu_1 - \mu_2 \neq 0 \end{array}} \Rightarrow \boxed{\begin{array}{l} H_0: \mu_D = 0 \\ H_1: \mu_D \neq 0 \end{array}}$$

$$T = \frac{\bar{D} - \mu_D}{S_d/\sqrt{n}} = \frac{0.144 - 0}{0.167/\sqrt{9}} = \boxed{2.5868}$$

[3]. the decision is:

$$T = 2.5868 \in R.R, \text{then we } \boxed{\text{Reject } H_0}$$

Two Proportions:

Hypothesis Null H_0 Alternative (Research) H_A	$H_0: p_1 - p_2 = d$ $H_A: p_1 - p_2 \neq d$	$H_0: p_1 - p_2 = d$ $H_A: p_1 - p_2 > d$	$H_0: p_1 - p_2 = d$ $H_A: p_1 - p_2 < d$
Test Statistics (TS)	$Z = \frac{(\hat{p}_1 - \hat{p}_2) - d}{\sqrt{\frac{\hat{p}\hat{q}}{n_1} + \frac{\hat{p}\hat{q}}{n_2}}} = \frac{(\hat{p}_1 - \hat{p}_2) - d}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \sim N(0,1)$		
Rejection Region (RR) of H_0 Acceptance Region (AR) of H_0			
Decision	We reject H_0 at the significance level α if		
	$Z < -Z_{1-(\alpha/2)}$ or $Z > Z_{1-(\alpha/2)}$ Two sides test	$Z > Z_{1-\alpha}$ One side test	$Z < -Z_{1-\alpha}$ One side test

$$\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$$

<u>Question 1:</u>

A random sample of 100 students from school "A" showed that 15 students smoke. Another independent random sample of 200 students from school "B" showed that 20 students smoke. Let p_1 be the proportion of smokers in school "A" and p_2 is the proportion of smokers in school "B". Test $H_0: p_1 = p_2$ vs $H_1: p_1 > p_2$. (use α=0.05)

$$n_1 = 100 \quad x_1 = 15 \rightarrow \hat{p}_1 = \frac{x_1}{n_1} = \frac{15}{100} = 0.15$$

$$n_2 = 200 \quad x_2 = 20 \rightarrow \hat{p}_2 = \frac{x_2}{n_2} = \frac{20}{200} = 0.10$$

- Hypotheses:

$$H_0: p_1 = p_2 \qquad \text{or} \qquad H_0: p_1 - p_2 = 0$$

$$H_1: p_1 > p_2 \qquad\qquad H_1: p_1 - p_2 > 0$$

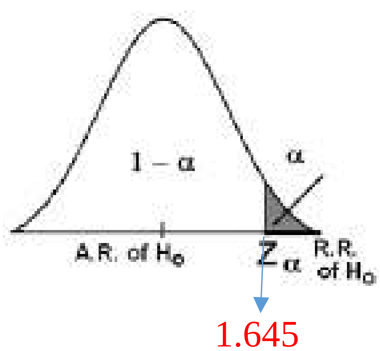
- Test statistic (T.S):

$$Z = \frac{(\hat{p}_1 - \hat{p}_2)}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_1}+\frac{1}{n_2}\right)}} = \frac{(0.15-0.10)}{\sqrt{(0.12)(0.88)\left(\frac{1}{100}+\frac{1}{200}\right)}} = 1.26$$

$$\hat{p} = \frac{x_1 + x_2}{n_1 + n_2} = \frac{15+20}{100+200} = 0.12$$

- Rejection Region (R.R):

$$\alpha = 0.05 \rightarrow Z_\alpha = Z_{0.05} = 1.645$$



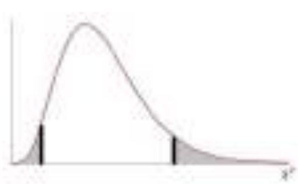
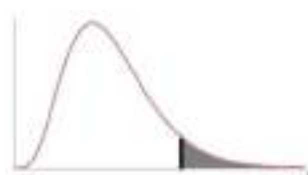
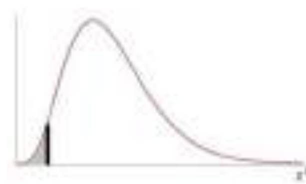
- Decision:

$$Z = 1.26 \notin R.R = (1.645, \infty), \text{ Then we accept } H_0.$$

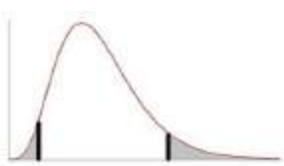
$$P - value = P(Z > 1.26) = 1 - P(Z < 1.26) = 1 - 0.8962 = 0.1038 > \alpha$$

<u>Testing the Population Variance</u>

- One sample variance:

Hypotheses	$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 \neq \sigma_0^2$	$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 > \sigma_0^2$	$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 < \sigma_0^2$
Test Statistics (T.S)	$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2} \sim \chi_{n-1}^2$		
R.R and A.R of H_0	 $\chi^2_{1-\frac{\alpha}{2}}$ $\chi^2_{\frac{\alpha}{2}}$	 χ^2_{α}	 $\chi^2_{1-\alpha}$
Decision	Reject H_0 at the significance level α		

- Two-sample variances:

Hypotheses	$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 \neq \sigma_2^2$	$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 > \sigma_2^2$	$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 < \sigma_2^2$
Test Statistics (T.S)	$F = \frac{S_1^2}{S_2^2} \sim F_{n_1-1, n_2-1}$		
R.R and A.R of H_0	 $F_{1-\frac{\alpha}{2}}$ $F_{\frac{\alpha}{2}}$	 F_{α}	 $F_{1-\alpha}$
Decision	Reject H_0 at the significance level α		

Question 1:

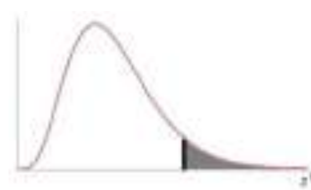
We are interested in the content of a soft-drink dispensing machine. A random sample of 25 drinks gave a variance of 2.03 deciliters^2. Assume that the contents are approximately normally distributed. The soft-drink machine is said to be out of control if the variance exceeds 1.15 deciliters^2. To test $H_0: \sigma^2 = 1.15$ vs $H_1: \sigma^2 > 1.15$ **with 0.05 level of significance, then:**

(1) The test statistic used is:

$$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2} = \frac{(25-1)2.03}{1.15} = 42.365$$

(A)	7.500	(B)	0.833	(C)	42.365	(D)	10.823

(2) The critical region is:



$$\chi^2_{\alpha,n-1} = \chi^2_{0.05,24} = 36.415$$

(A)	$(36.415, \infty)$	(B)	$(33.196, \infty)$	(C)	$(1.96, \infty)$	(D)	$(1.645, \infty)$

(3) The decision is:

(A)	Not reject H_0	(B)	Reject H_0

Question 2:

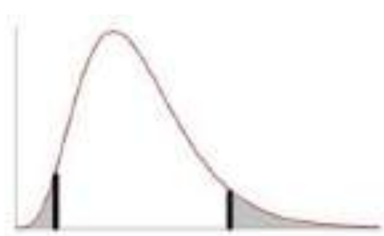
In a series of experiments to determine the absorption rate of certain pesticides into skin, measured amounts of two pesticides were applied to several skin specimens. For pesticide A, the variance of the amounts absorbed in $n_1 = 6$ specimens was $S_1^2 = 2.3$, while for pesticide B, the variance of the amounts absorbed in $n_2 = 10$ specimens was $S_2^2 = 0.6$. Assume that for each pesticide, the amounts absorbed are a simple random sample from a normal population. To test the claim that the variance in the amount absorbed is greater for pesticide A than for pesticide B, that is $H_0: \sigma_1^2 = \sigma_2^2$ vs $H_1: \sigma_1^2 \neq \sigma_2^2$ with 0.10 level of significance, then:

1. **The value of test statistic is:**

$$F = \frac{S_1^2}{S_2^2} = \frac{2.3}{0.6} = 3.833$$

(A)	3.833	(B)	0.1	(C)	0.057	(D)	2.35

2. **The non-rejection region is:**



$$F_{1-\frac{\alpha}{2}, n_1-1, n_2-1} = F_{0.95,5,9} = \frac{1}{F_{0.05,9,5}} = \frac{1}{4.77} = 0.21$$

$$F_{\frac{\alpha}{2}, n_1-1, n_2-1} = F_{0.05,5,9} = 3.48$$

(A)	$(-\infty, -1.96)$	(B)	$(-1.96, \infty)$	(C)	$(-1.96, 1.96)$	(D)	$(0.21, 3.48)$

3. **The decision is:**

(A)	Not reject H_0	(B)	Reject H_0

<u>Question 3</u>:

The following data represent the sizes and variances of two normal samples:

Sample1 $n_1 = 10\ ,\ S_1^2 = 1.8$

Sample2 $n_2 = 13\ ,\ S_2^2 = 1.5$

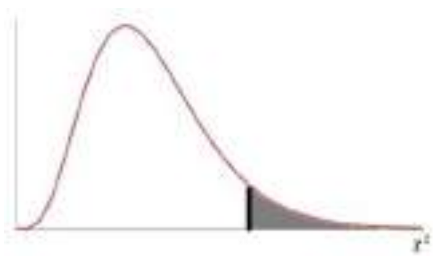
To test $\mathbf{H_0: \sigma_1^2 = 1.5}$ **vs** $\mathbf{H_1: \sigma_1^2 > 1.5}$ **with level 0.05, then:**

a. The test statistic used is:

$$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2} = \frac{(10-1)1.8}{1.5} = 10.8$$

(A)	7.5	(B)	0.833	(C)	1.2	(D)	10.8

b. The rejection region is:



$$\chi^2_{\alpha,n-1} = \chi^2_{0.05,9} = 16.919$$

(A)	$(-\infty,-16.92)$	(B)	$(-16.92,\infty)$	(C)	$(16.92,\infty)$	(D)	$(1.645,\infty)$

c. The decision is:

(A)	Accept H_0	(B)	Reject H_0

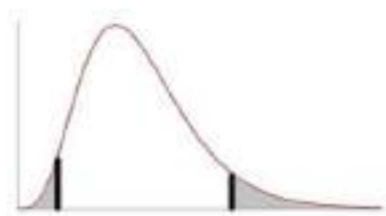
- **To test claim with level 0.10 that both variances <u>are equal</u>, then:**

d. The test statistic used is:

$$F = \frac{S_1^2}{S_2^2} = \frac{1.8}{1.5} = 1.2$$

(A)	7.5	(B)	0.833	(C)	1.2	(D)	10.8

e. The acceptance region is:



$$F_{1-\frac{\alpha}{2},n_1-1,n_2-1} = F_{0.95,9,12} = \frac{1}{F_{0.05,12,9}} = \frac{1}{3.07} = 0.326$$

$$F_{\frac{\alpha}{2},n_1-1,n_2-1} = F_{0.05,9,12} = 2.8$$

(A)	(0.326 , 2.8)	(B)	$(-2.8, 2.8)$	(C)	$(2.8, \infty)$	(D)	$(1.645, \infty)$

f. The decision is:

<u>(A)</u>	Accept H_0	(B)	Reject H_0

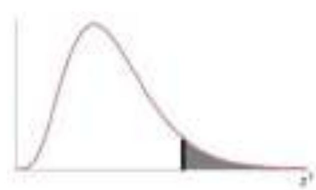
Testing		
single Population	Mean	$Z = \frac{\bar{X} - \mu_o}{\sigma/\sqrt{n}}$ σ known
		$t = \frac{\bar{X} - \mu_o}{S/\sqrt{n}}$ σ unknown
	Proportion	$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 q_0}{n}}}$
	Variance	$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2}$
Two Populations	Means	$Z = \frac{(\bar{X}_1 - \bar{X}_2) - d}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$ σ_1 & σ_2 known
		$t = \frac{(\bar{X}_1 - \bar{X}_2) - d}{Sp\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ σ_1 & σ_2 unknown
	Proportions	$Z = \frac{(\hat{p}_1 - \hat{p}_2)}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$ $\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$
	Variances	$F = \frac{S_1^2}{S_2^2}$

Two samples T test for paired observation $T = \frac{\bar{D} - \mu_D}{S_d/\sqrt{n}}$

CHAPTER 6

- Goodness of Fit Test.
- Test for Independency and Homogeneity.

Goodness of Fit Test

Hypotheses	$H_0: The\ data\ follow\ (\dots)\ distribution$ $H_1: The\ data\ do\ not\ follow\ (\dots)\ distribution$
Test Statistics (T.S)	$\chi^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i}$
R.R and A.R of H_0	 χ^2_α
Decision	If $\chi^2 > \chi^2_{\alpha,k-m-1} \Longrightarrow$ Reject H_0 k: number of groups. m: number of parameters to be estimated.

<u>Q1</u>. A doctor believes that the proportions of births in this country on <u>each day of the week are equal</u>. A simple random sample of 700 births from a recent year is selected, and the results are below.

Day	Sunday	Monday	Tuesday	Wednesday	Thursday	Friday	Saturday
Frequency	65	103	114	116	115	112	75
Expected E_i	$E_1 = 100$	$E_2 = \cdots$	$E_3 = \cdots$	$E_4 = 100$	$E_5 = 100$	$E_6 = 100$	$E_7 = 100$

At a significance level of α =0.01, we want to test the hypothesis if there is enough evidence to support the doctor's claim.

1. The expected frequency E_2 is

$$E_2 = 100$$

2. The degree of freedom of the χ^2 test statistic is

$$\text{v} = k - m - 1$$
$$= 7 - 1 - 1 = \boxed{5}$$

3. The value of the χ^2 test statistic is

$$\chi^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i}$$

$$= \frac{(65-100)^2}{100} + \frac{(103-100)^2}{100} + \frac{(114-100)^2}{100} + \frac{(116-100)^2}{100} + \frac{(115-100)^2}{100} + \frac{(112-100)^2}{100} + \frac{(75-100)^2}{100} = \boxed{26.8}$$

4. The critical value is

$$\chi^2_{\alpha,k-m-1} = \chi^2_{0.01,5} = \boxed{15.086}$$

5. The decision about the doctor's claim is

$$26.8 \in (15.086\,,\infty) \Longrightarrow \boxed{\text{Reject } H_0}$$

<u>Q2:</u> We test the balance of a dice. If the results of tossing the dice 180 times are:

X	1	2	3	4	5	6	
Frequency	15	40	36	25	45	19	180
E_i	30	30	30	30	30	30	

1. The expected frequency of each face is

(A)	40	(B)	50	(C)	30	(D)	60

2. The test statistic used to perfume this test is distributed as:

(A)	T	(B)	F	(C)	Normal	(D)	Chi squares

2. The degree of freedom of the value of the test statistic is:

$$v = k - m - 1$$
$$= 6 - 1 - 1 = \boxed{4}$$

(A)	6	(B)	4	(C)	7	(D)	8

3. The value of the test statistic is:

$$\chi^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i} = \frac{(15-30)^2}{30} + \frac{(40-30)^2}{30} + \cdots + \frac{(19-30)^2}{30} = \boxed{24.4}$$

(A)	8.94	(B)	22.4	(C)	24.4	(D)	44.7

4. The table value, at the 0.05 significance level, is

$$\chi^2_{\alpha,k-m-1} = \chi^2_{0.05,4} = \boxed{9.488}$$

(A)	22.5	(B)	17.6	(C)	18.45	(D)	11.070

5. The conclusion is

$$24.4 \in (9.488\,, \infty) \Longrightarrow \boxed{\text{Reject } H_0}$$

(A)	Accept the balance	(B)	Reject the balance	(C)	No Decision

<u>Q3</u>: We investigate the number of errors in electrical circles panel printers. A random sample of 200 panel printers is selected and we recorded X= Number of errors in each printer. The table is the observed and the expected of each value of X in the sample:

Value of X	$X_1 = 0$	$X_2 = 1$	$X_3 = 2$	$X_4 = 3$	$X_5 = 4$
O_i	$O_1 = 75$	$O_2 = 66$	$O_3 = 42$	$O_4 = 13$	$O_5 = 4$
E_i	$E_1 = 71.76$	$E_2 = \cdots$	$E_3 = 37.7$	$E_4 = \cdots$	$E_5 = 3.3$

Test at <u>$\alpha = 0.10$</u> the inspector claim that X follows <u>Poisson distribution</u>

1) The estimate of Poisson parameter λ is:

$$\lambda = \frac{(0\times75)+(1\times66)+(2\times42)+(3\times13)+(4\times4)}{200} = \boxed{1.025}$$

2) The expected number of errors E_2 is:

$$E_2 = \frac{\lambda^{x_2} e^{-\lambda}}{x_2!} \times 200 = \frac{(1.025)^1 e^{-(1.025)}}{1!} \times 200 = \boxed{73.55}$$

3) The expected number of errors E_4 is:

$$E_4 = \frac{\lambda^{x_4} e^{-\lambda}}{x_4!} \times 200 = \frac{(1.025)^3 e^{-(1.025)}}{3!} \times 200 = \boxed{12.88}$$

4) The degree of freedom of the χ^2test statistic is:

Value of X	$X_1 = 0$	$X_2 = 1$	$X_3 = 2$	$X_4 = 3$
O_i	$O_1 = 75$	$O_2 = 66$	$O_3 = 42$	$O_4 = 13 + 4 = 17$
E_i	$E_1 = 71.76$	$E_2 = 73.55$	$E_3 = 37.7$	$E_4 = 12.88 + 3.3$ $E_4 = 16.18$

$$\begin{aligned} v &= k - m - 1 \\ &= 4 - 1 - 1 = \boxed{2} \end{aligned}$$

5) The value of the χ^2 test statistic is:

$$X^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i}$$

$$= \frac{(75-71.76)^2}{71.76} + \frac{(66-73.55)^2}{73.55} + \frac{(42-37.7)^2}{37.7} + \frac{(17-16.18)^2}{16.18} = \boxed{1.46}$$

6) The rejection region R.R is

$$\chi^2_{\alpha,k-m-1} = \chi^2_{0.1,2} = 4.605 \implies R.R \in \boxed{(4.605\,,\infty)}$$

7) The conclusion is

$$1.46 \notin (4.605\,,\infty) \implies \boxed{\text{Accept the claim}}$$

Q4. 1000 bags of oranges, each containing 10 oranges. Some of the oranges are rotten. The observed numbers of rotten oranges per bag are tabulated as follows:

No. of rotten oranges	0	1	2	3	4	5	6
Observed counts (O_i)	334	369	191	63	22	12	9
Expected counts (E_i)	$E_1 = 297.4$	$E_2 = 383.4$	$E_3 = \cdots$	$E_4 = \cdots$	$E_5 = 17.3$	$E_6 = 2.7$	$E_7 = 0.3$

Suppose that we want to test at level of significance <u>0.05</u>, that this data suggests that the distribution of rotten oranges in the individual bags follow a binomial distribution with parameters <u>n=10</u> and unknown p <u>(Bin (10, p))</u>. Then:

(1) The null hypothesis H_0 is:

A	The counts of rotten oranges follow Poisson distribution.
B	The counts of rotten oranges do not follow a binomial distribution (Bin (10, p) for some p).
C	The counts of rotten oranges follow a binomial distribution (Bin (10, p) for some p).
D	The counts of rotten oranges follow a binomial distribution (Bin (n, 10).

(2) The estimate of the unknown parameter p is:

$$E(X) = n \times p = \frac{(0\times334)+(1\times369)+\cdots+(6\times9)}{1000} = 1.142$$

$$n \times p = 1.142 \Longrightarrow 10 \times p = 1.142 \Longrightarrow \boxed{p = 0.1142}$$

(3) The expected number of errors E_3 is:

$$p(X = 2) = \binom{10}{2}(0.1142)^2(1 - 0.1142)^8 = 0.2224$$

$$E_3 = 1000 \times p(X = 2) = 1000 \times 0.2224 = \boxed{222.4}$$

(4) The expected number of errors E_4 is:

$$p(X = 3) = \binom{10}{3}(0.1142)^3(1 - 0.1142)^7 = 0.0765$$

$$E_4 = 1000 \times p(X = 3) = 1000 \times 0.0765 = \boxed{76.5}$$

(5) The degree of freedom of the χ^2 test statistic is:

No. of rotten oranges	0	1	2	3	4	5	6
Observed counts (O_i)	334	369	191	63	22	12	9
Expected counts (E_i)	$E_1 = 297.4$	$E_2 = 383.4$	$E_3 = 222.4$	$E_4 = 76.5$	$E_5 = 17.3$	$E_6 = 2.7$	$E_7 = 0.3$

No. of rotten oranges	0	1	2	3	4
Observed counts (O_i)	334	369	191	63	22 + 12 + 9 = 43
Expected counts (E_i)	$E_1 = 297.4$	$E_2 = 383.4$	$E_3 = 222.4$	$E_4 = 76.5$	$E_5 = 17.3 + 2.7 + 0.3$ $E_5 = 20.3$

$$v = k - m - 1$$
$$= 5 - 1 - 1 = \boxed{3}$$

(6) The value of the χ^2 test statistic is:

$$X^2 = \sum_{i=1}^{k} \frac{(O_i - E_i)^2}{E_i}$$
$$= \frac{(334-297.4)^2}{297.4} + \frac{(369-383.4)^2}{383.4} + \frac{(191-222.4)^2}{222.4} + \frac{(63-76.5)^2}{76.5} + \frac{(43-20.3)^2}{20.3} = \boxed{37.24}$$

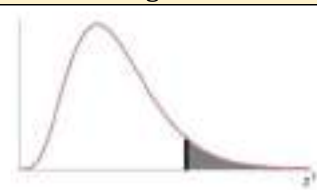
(7) The rejection region (R.R) of H_0 is:

$$\chi^2_{\alpha,k-m-1} = \chi^2_{0.05,3} = 7.815 \implies R.R \in \boxed{(7.815\,,\infty)}$$

(8) The decision is:

$$37.24 \in (7.815\,,\infty) \implies \boxed{\text{Reject } H_0}$$

Test for Independence (categorical data)

Hypotheses	H_0: The two random variable are independent H_1: The two random variable are not independent
Test Statistics (T.S)	$\chi^2 = \sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij}-E_{ij})^2}{E_{ij}}$ $E_{ij} = \frac{(j^{th}\text{ column total})\times(i^{th}\text{ row total})}{\text{grand total}}$
R.R and A.R of H_0	 $v = (c-1)\times(r-1)$
Decision	If $\chi^2 > \chi^2_{\alpha} \implies$ Reject H_0

<u>Q1:</u> We test the hypothesis at level of significance 0.05, that the size of a family is independent of the level of education attained by the father. A random sample of 400 married men, all retired, were classified according to education and number of children. The table shows the observed O_{ij} and the expected E_{ij} values:

			Number of Children				
			0-1	2-3	4-5	over 5	
Education	Elementary	O	25	48	44	50	
		E	32.57	53.86	38.83	41.75	
	Secondary		30	53	28	32	143
			27.89	E_{22}	E_{23}	35.75	
	Collage		23	28	21	18	
			17.55	29.03	20.93	22.5	
				129	93		

1) The test statistic used to perfume this test is distributed as:

(A)	T	(B)	F	(C)	Normal	<u>(D)</u>	<u>Chi squares</u>

2) The mathematic expression of the test statistic is:

A	B	C	D
$\sum_{j=1}^{c}\frac{(O_j-E_j)^2}{O_j}$	$\sum_{j=1}^{c}\frac{(O_j-E_j)^2}{E_j}$	$\sum_{i=1}^{r}\sum_{j=1}^{c}\frac{(O_{ij}-E_{ij})^2}{O_{ij}}$	$\sum_{i=1}^{r}\sum_{j=1}^{c}\frac{(O_{ij}-E_{ij})^2}{E_{ij}}$

3) The expectation E_{23} is:

$$E_{23}=\frac{143\times 93}{400}=\boxed{33.25}$$

4) The expectation E_{22} is:

$$E_{22}=\frac{143\times 129}{400}=\boxed{46.12}$$

5) The value of the test statistic is:

$$\chi^2=\sum_{i=1}^{r}\sum_{j=1}^{c}\frac{(O_{ij}-E_{ij})^2}{E_{ij}}=\frac{(25-32.57)^2}{32.57}+\frac{(48-53.86)^2}{53.86}+\cdots+\frac{(18-22.5)^2}{22.5}=\boxed{9.75}$$

6) The degree of freedom of the test statistic is:

$$v=(c-1)\times(r-1)=(4-1)\times(3-1)=3\times 2=\boxed{6}$$

7) The rejection region R.R is:

$$\chi^2_{0.05,6}=12.592\Longrightarrow\boxed{(12.592\,,\infty)}$$

8) The decision about the independence is

$$9.75\notin(12.592\,,\infty)\Longrightarrow\boxed{\text{Accept } H_0}$$

Q2: In an experiment to study the dependence of hypertension on smoking habits, the following data were taken on 180 individuals:

	Non-smoker (NS)	Moderate smokers (MS)	Heavy smoker (HS)	Total
Hypertension	21 $(E_{11} = 33.35)$	36 $(E_{12} = 29.97)$	30 $(E_{13} = 32.68)$	
No hypertension	48 $(E_{21} = 35.65)$	26 $(E_{22} = 32.03)$	19 $(E_{23} = 25.32)$	93
Total		62	49	180

Suppose that we want to test the hypothesis at level of significance 0.05, that the presence or absence of hypertension and the smoking habits are independent.

1. The test statistic used to perfume this test is distributed as:

(A)	T	(B)	F	(C)	Normal	(D)	Chi squares

2. The mathematic expression of the test statistic is:

A	B	C	D
$\sum_{j=1}^{c} \frac{(O_j - E_j)^2}{O_j}$	$\sum_{j=1}^{c} \frac{(O_j - E_j)^2}{E_j}$	$\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{O_{ij}}$	$\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{E_{ij}}$

3. The value of the expectation E_{22} is:

$$E_{22} = \frac{93 \times 62}{180} = \boxed{32.03}$$

4. The value of the expectation E_{23} is:

$$E_{23} = \frac{93 \times 49}{180} = \boxed{25.32}$$

5. The value of the test statistic is:

$$\chi^2 = \sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{E_{ij}} = \frac{(21-33.35)^2}{33.35} + \frac{(36-29.97)^2}{29.97} + \cdots + \frac{(19-25.32)^2}{22.5} = \boxed{14.46}$$

6. The degree of freedom of the test statistic is:

$$v = (c-1) \times (r-1) = (3-1) \times (2-1) = 2 \times 1 = \boxed{2}$$

7. The rejection region (R.R) of H_0 is:

$$\chi^2_{0.05,2} = 5.991 \Longrightarrow \boxed{(5.991, \infty)}$$

8. The decision about the independence is:

$$14.46 \in (5.991, \infty) \Longrightarrow \boxed{\text{Reject } H_0}$$

Q3. Results of a random sample of children with pain from musculoskeletal injuries treated with acetaminophen, ibuprofen, or codeine are shown in the table. At $\alpha = 0.10$, we want to test the hypothesis that the treatment and result are independent

	Acetaminophen	Ibuprofen	Codeine	Total
Significant Improvement	58 ($E_{11} = 66.7$)	81 ($E_{12} = \ldots$)	61 ($E_{13} = 66.6$)	200
Slight Improvement	42 ($E_{21} = \ldots$)	19 ($E_{22} = 33.3$)	39 ($E_{23} = 33.4$)	100
Total	100	100		300

1. The distribution of the test statistic is

(A)	T	(B)	Binomial	(C)	Chi squares	(D)	Normal

2. The value of the expectation E_{12} is

$$E_{12} = \frac{200 \times 100}{300} = \boxed{66.7}$$

3. The value of the expectation E_{21} is

$$E_{21} = \frac{100 \times 100}{300} = \boxed{33.3}$$

4. The mathematical expression of the test statistic is

A	B	C	D
$\sum_{j=1}^{c} \frac{(O_j - E_j)^2}{O_j}$	$\sum_{j=1}^{c} \frac{(O_j - E_j)^2}{E_j}$	$\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{E_{ij}}$	$\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{O_{ij}}$

5. The value of the χ^2 test statistic is

$$\chi^2 = \sum_{i=1}^{r} \sum_{j=1}^{c} \frac{(O_{ij} - E_{ij})^2}{E_{ij}} = \frac{(58-66.7)^2}{66.7} + \frac{(81-66.7)^2}{66.7} + \cdots + \frac{(39-33.4)^2}{33.4} = \boxed{14.02}$$

6. The critical value is

$$v = (c-1) \times (r-1) = (3-1) \times (2-1) = 2 \times 1 = 2$$

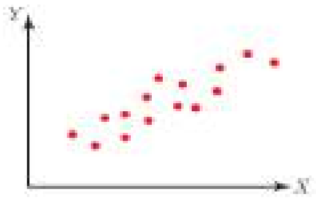
$$\chi^2_{0.10,2} = \boxed{4.605}$$

7. The decision about the independence is

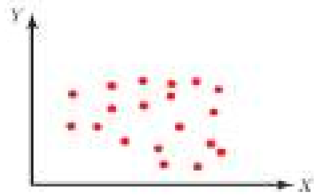
$$14.02 \in (4.605, \infty) \Longrightarrow \boxed{\text{Reject } H_0}$$

CHAPTER 7
Regression analysis

- Correlation:



Strong positive correlation
($\mathbf{r} = 0.80$)



Very Weak correlation
($\mathbf{r} = 0.25$)

Weak negative correlation
($\mathbf{r} = -0.45$)

Strong correlation & outlier
($\mathbf{r} = 0.7$)

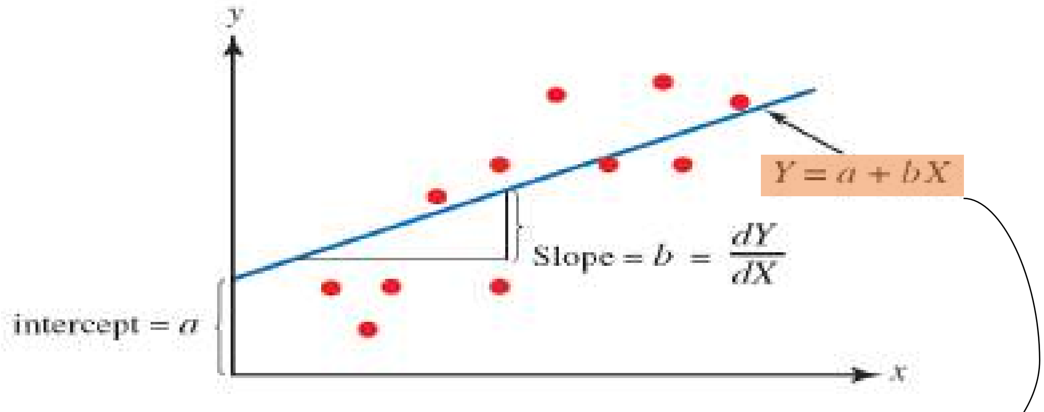
$$\mathrm{Corr}(X, Y) = r_{XY} = \frac{S_{XY}}{\sqrt{S_{XX}\ S_{YY}}}$$

$$S_{XX} = \sum (X_i - \overline{X})^2 = \sum X_i^2 - n\overline{X}^2$$

$$S_{YY} = \sum (Y_i - \overline{Y})^2 = \sum Y_i^2 - n\overline{Y}^2$$

$$S_{XY} = \sum (X_i - \overline{X})(Y_i - \overline{Y}) = \sum X_i Y_i - n\overline{XY}$$

- Simple Linear Regression:



$$\hat{y} = \mathbf{b_0} + \mathbf{b_1 X}$$

$$b_0 = \bar{y} - b\bar{x} \qquad b_1 = \frac{S_{XY}}{S_{XX}}$$

- Testing β_1:

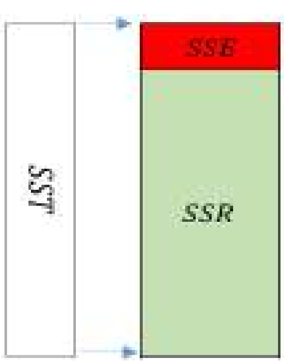
$$t = \frac{b_1 - \beta_{10}}{S(b_1)}$$

- Estimation for β_1:

$$\beta_1 \in \left(b_1 \pm t_{1-\frac{\alpha}{2},n-2}\, S(b_1)\right)$$

$$S(b_1) = \frac{\hat{\sigma}}{\sqrt{S_{XX}}} \quad , \quad \hat{\sigma}^2 = \frac{SSE}{n-2} = \frac{\sum(y_i - \hat{y}_i)^2}{n-2}$$

Coefficient of determination: $R^2 = \frac{SSR}{SST} = 1 - \frac{SSE}{SST} = r^2_{XY}$



$$SSR = \Sigma(\hat{y}_i - \bar{y})^2 = \frac{S^2_{XY}}{S_{XX}}$$

$$SSE = \Sigma(y_i - \hat{y}_i)^2 = S_{YY} - \frac{S^2_{XY}}{S_{XX}} = \sum e_i^2 = \Sigma y_i^2 - b_o \Sigma y_i - b_1 \Sigma x_i y_i$$

$$SST = \Sigma(y_i - \bar{y})^2 = S_{YY}$$

SST = SSE + SSR

Q1: The grades of a class of 9 students on a midterm report (X) and on the final examination (Y) are as follows:

X	82	66	78	80	85	85	99	99	68	$\sum_i x_i = 742$, $\sum_i y_i = 707$, $\sum_i x_i y_i = 59648$
Y	77	50	71	72	81	94	96	99	67	$\sum_i x_i^2 = 62240$, $\sum_i y_i^2 = 57557$

The value of Pearson Correlation coefficient is:

$$S_{XX} = \sum(X_i - \overline{X})^2 = \sum X_i^2 - n\overline{X}^2 = 62240 - 9 \times \left(\frac{742}{9}\right)^2 = \boxed{1066.222}$$

$$S_{YY} = \sum(Y_i - \overline{Y})^2 = \sum Y_i^2 - n\overline{Y}^2 = 57557 - 9 \times \left(\frac{707}{9}\right)^2 = \boxed{2018.222}$$

$$S_{XY} = \sum(X_i - \overline{X})(Y_i - \overline{Y}) = \sum X_i Y_i - n\overline{X}\overline{Y} = (59648) - 9 \times \left(\frac{742}{9}\right)\left(\frac{707}{9}\right) = \boxed{1359.778}$$

$$r = \frac{S_{XY}}{\sqrt{S_{XX}\ S_{YY}}} = \frac{1359.778}{\sqrt{1066.222 \times 2018.222}} = 0.927$$

If the estimate of the linear regression line is $\hat{y} = b_0 + b_1 X$, then

1. **The value of b_1 is:**

$$b_1 = \frac{S_{XY}}{S_{XX}} = \frac{1359.222}{1066.222} = 1.275$$

9. **The value of b_0 is:**

$$b_0 = \bar{y} - b_1\bar{x}$$
$$= \left(\frac{707}{9}\right) - (1.275) \times \left(\frac{742}{9}\right) = -26.56$$

10. **A student got 85 on the midterm, then the estimate of the final grade is:**

$$\hat{y} = b_0 + b_1 X = -26.56 + 1.275X$$
$$= -26.56 + 1.275 \times (85) = 81.82$$

Q2. A study was made by a retail merchant to determine the relation between weekly advertising expenditures and sales.

costs $ (X)	40	20	25	20	30	50	40	20	50	40	25	50
Sales $ (Y)	385	400	395	365	475	440	490	420	560	525	480	510

$$\sum x_i = 410, \sum y_i = 5445, \sum x_i^2 = 15650, \sum y_i^2 = 2512925, \sum x_i y_i = 191325,$$

1. The Pearson correlation coefficient of sales and advertising costs is:

$$S_{XX} = \sum (X_i - \overline{X})^2 = \sum X_i^2 - n\overline{X}^2 = 15650 - 12 \times \left(\frac{410}{12}\right)^2 = \boxed{1641.67}$$

$$S_{YY} = \sum (Y_i - \overline{Y})^2 = \sum Y_i^2 - n\overline{Y}^2 = 2512925 - 12 \times \left(\frac{5445}{12}\right)^2 = \boxed{42256.25}$$

$$S_{XY} = \sum (X_i - \overline{X})(Y_i - \overline{Y}) = \sum X_i Y_i - n\overline{X}\overline{Y} = (191325) - 12 \times \left(\frac{410}{12}\right)\left(\frac{5445}{12}\right) = \boxed{5287.5}$$

$$\boxed{r = \frac{S_{XY}}{\sqrt{S_{XX}\ S_{YY}}} = \frac{5287.5}{\sqrt{1641.67 \times 42256.25}} = 0.63}$$

- If the estimate of the linear regression line is $\hat{y} = b_0 + b_1 X$, then:

2. The value of b_1 is:

$$b_1 = \frac{S_{XY}}{S_{XX}} = \frac{5287.5}{1641.67} = 3.22$$

3. The value of b_0 is:

$$b_0 = \bar{y} - b_1 \bar{x}$$
$$= \left(\frac{5445}{12}\right) - (3.22) \times \left(\frac{410}{12}\right) = 343.71$$

4. If the advertising costs is $30, then the weekly sales are:

$$\hat{y} = b_0 + b_1 X = 343.71 + 3.22X$$
$$= 343.71 + 3.22 \times (30) = 440.31$$

- We want to test the hypothesis that $\beta_1 = 0$ against the alternative that $\beta_1 \neq 0$ at the 0.05 level of significance.

1. **The residual sum of squares SSR is:**

$$SSR = \frac{S_{XY}^2}{S_{XX}} = \frac{(5287.5)^2}{1641.67} = 17030.04$$

2. **The total sum of squares SST is:**

$$SST = S_{YY} = 42256.25$$

3. **The value of the statistic test is:**

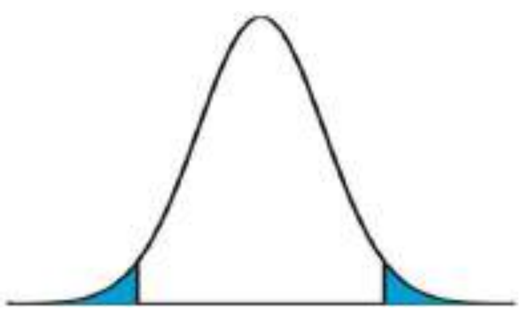
$$t = \frac{b_1 - \beta_{10}}{S(b_1)} = \frac{3.22 - 0}{1.24} = \boxed{2.6}$$

$$\hat{\sigma}^2 = \frac{SSE}{n-2} = \frac{SST-SSR}{n-2} = \frac{42256.25-17030.04}{12-2} = 2522.625 \implies \hat{\sigma} = 50.2257$$

$$S(b_1) = \frac{\hat{\sigma}}{\sqrt{S_{XX}}} = \frac{50.2257}{\sqrt{1641.67}} = 1.24$$

4. **The decision is:**

$$t_{1-\frac{\alpha}{2}, n-2} = t_{0.975,10} = 2.228$$



$$t = 2.6 \notin (-2.228, 2.228) \implies \text{Reject } H_0$$

Q3. The shear resistance of soil, Y, is determined by measurements as a function of the normal stress, X. We assume that the errors ε_i are normally distributed. The data are as shown below:

x_i	10	11	12	13	14	15	16	17	18	19	20	21
y_i	14.08	15.57	16.94	17.68	18.49	19.55	20.68	21.72	22.8	23.84	24.79	25.67

We have $\sum_i x_i = 186$, $\sum_i y_i = 241.81$, $\sum_i x_i^2 = 3026$, $\sum_i y_i^2 = 5025.399$, $\sum_i x_i y_i = 3895.65$

1. **The coefficient S_{XX} is**

$$S_{XX} = \sum (X_i - \bar{X})^2 = \sum X_i^2 - n\bar{X}^2 = 3026 - 12 \times \left(\frac{186}{12}\right)^2 = \boxed{143}$$

2. **The coefficient S_{YY} is**

$$S_{YY} = \sum (Y_i - \bar{Y})^2 = \sum Y_i^2 - n\bar{Y}^2 = 5025.399 - 12 \times \left(\frac{241.81}{12}\right)^2 = \boxed{152.726}$$

3. **The coefficient S_{XY} is**

$$S_{XY} = \sum (X_i - \bar{X})(Y_i - \bar{Y}) = \sum X_i Y_i - n\bar{X}\bar{Y} = (3895.65) - 12 \times \left(\frac{186}{12}\right)\left(\frac{241.81}{12}\right) = \boxed{147.595}$$

4. **The sample linear correlation coefficient r is**

$$r = \frac{S_{XY}}{\sqrt{S_{XX}\ S_{YY}}} = \frac{147.595}{\sqrt{143 \times 152.726}} = 0.9987$$

- If the estimate of the linear regression line is $\hat{y} = b_0 + b_1 X$, then:

5. **The value of b_1 is:**

$$b_1 = \frac{S_{XY}}{S_{XX}} = \frac{147.595}{143} = 1.032$$

6. **The value of b_0 is:**

$$b_0 = \bar{y} - b_1 \bar{x}$$
$$= \left(\frac{241.81}{12}\right) - (1.032) \times \left(\frac{186}{12}\right) = 4.15$$

-We want to test the hypothesis that $\beta_1 = 1$ against the alternative that $\beta_1 > 1$ at the 0.05 level of significance. The residuals e_i are
-0.394 0.064 0.402 0.109 -0.113 -0.085 0.013 0.021 0.069 0.077 -0.005 -0.158

7. Deduce that the value of SSE is

$$SSE = \sum e_i^2 = (-0.394)^2 + (0.064)^2 + (0.402)^2 + \cdots + (-0.158)^2 = \boxed{0.389}$$

8. The unbiased estimate of σ^2 is

$$\hat{\sigma}^2 = \frac{SSE}{n-2} = \frac{0.389}{12-2} = \boxed{0.0389} \implies \hat{\sigma} = 0.197$$

9. The value of the test statistic is

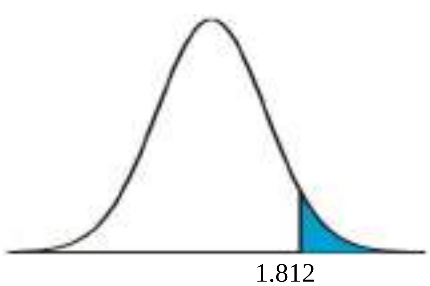
$$t = \frac{b_1 - \beta_{10}}{S(b_1)} = \frac{1.032 - 1}{0.0165} = \boxed{1.94}$$

$$S(b_1) = \frac{\hat{\sigma}}{\sqrt{S_{XX}}} = \frac{0.197}{\sqrt{143}} = 0.0165$$

10. The critical value is

$$t_{1-\alpha, n-2} = t_{0.95,10} = \boxed{1.812}$$

11. The decision is



$$t = 1.94 > 1.812 \implies \boxed{\text{Reject } H_0}$$

12. The coefficient of determination R^2 is

$$R^2 = \frac{SSR}{SST} = r_{XY}^2 = (0.9987)^2 = 0.9974$$

13. Determine the 90% confidence interval for the parameter β_1 (2 marks).

$$\beta_1 \in \left(b_1 \pm t_{1-\frac{\alpha}{2},n-2} \times S(b_1)\right)$$

$$\beta_1 \in \left(1.032 \pm t_{0.95,10} \times (0.0165)\right)$$

$$\beta_1 \in \left(1.032 \pm (1.812) \times (0.0165)\right)$$

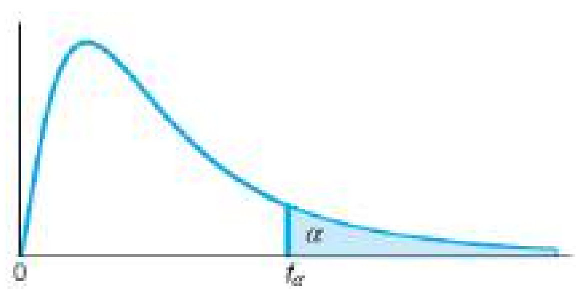
$$\beta_1 \in (1.002 \ , 1.062)$$

CHAPTER 8

On way ANOVA – Completely Randomized Designs (CRD):

$H_0: \mu_1 = \mu_2 = \cdots = \mu_k$
H_1: at least two of the means are not equal.

Source of Variation	Sum Square	Degrees of freedom	Mean Square	F
Treatment (Between)	$SSA = SS_{trt} = \sum_{i=1}^{k} n_i(\bar{y}_{i.} - \bar{y}_{..})^2$	$df_{trt} = k - 1$	$MSA = MS_{trt} = \frac{SSA}{k-1}$	$F = \frac{MSA}{MSE}$
Error (Within)	$SSE = SS_{er} = \sum_{i=1}^{k}\sum_{j=1}^{n}(y_{ij} - \bar{y}_{i.})^2$	$df_{er} = k(n-1) = kn - k = N - k$	$MSE = MS_{er} = \frac{SSE}{N-k}$	
Total	$SST = SS_{tot} = \sum_{i=1}^{k}\sum_{j=1}^{n}(y_{ij} - \bar{y}_{..})^2$ $\boxed{SST = SSA + SSE}$	$df_{tot} = kn - 1 = N - 1$		



Reject H_0 if $F > F_{\alpha,k-1,k(n-1)}$

Q1. **The statistics classroom is divided into three rows: front, middle, and back. The instructor noticed that the further the students were from him, the more likely they were to miss class or use an instant messenger during class. He wanted to see at level of significance 0.05: Are the students further away did worse on the exams? For this end, a random sample of the students in each row was taken. The score for those students on the second exam was recorded:**

–	Front:	82,	83,	97,	93,	55,	67,	53	
–	Middle:	83,	78,	68,	61,	77,	54,	69,	51,	63
–	Back:	38,	59,	55,	66,	45,	52,	52,	61	

Let the one-way ANOVA tabulated as follows:

Source of variation	**Sum of squares**	**Degrees of freedom**	**Mean Squares**	**Test Statistics**
Treatments	SS_{trt}	df_{trt}	MS_{trt}	
Errors	SS_{er}	df_{er}	MS_{er}	F_0
Total	$SS_{tot} = 5287.83$	df_{tot}		

										n_i	Sum	Mean	
Front	82	83	97	93	55	67	53			7	530	75.71	
Middle	83	78	68	61	77	54	69	51	63	9	604	67.11	
Back	38	59	55	66	45	52	52	61		8	428	53.50	
										24	1562	65.08	Total

1. **The value of SS_{trt} is:**

$$SS_{trt} = \sum_{i=1}^{k} n_i(\bar{y}_{i.} - \bar{y}_{..})^2$$

$$= 7 \times (75.71 - 65.08)^2 + 9 \times (67.11 - 65.08)^2 + 8 \times (53.50 - 65.08)^2 = 1901.52$$

2. The value of SS_{er} is:

$$SS_{er} = SS_{tot} - SS_{trt} = 5287.83 - 1901.52 = 3386.32$$

3. The degrees of freedom of the treatments (df_{trt}) is:

$$k - 1 = 3 - 1 = 2$$

4. The degrees of freedom of the errors (df_{er}) is:

$$N - k = 24 - 3 = 21$$

5. The degrees of freedom of the total (df_{tot}) is:

$$N - 1 = 24 - 1 = 23$$

6. The Mean Squares of the treatments (MS_{trt}) is:

$$MS_{trt} = \frac{SS_{trt}}{k - 1} = \frac{1901.52}{3 - 1} = 950.76$$

7. The Mean Squares of the errors (MS_{er}) is:

$$MS_{er} = \frac{SS_{er}}{N - k} = \frac{3386.32}{24 - 3} = 161.25$$

8. The value of the test statistic F_0 is:

$$F = \frac{MS_{trt}}{MS_{er}} = \frac{950.76}{161.25} = 5.896$$

9. The rejection region (R.R) of H_0 is:

$$F_{\alpha,(k-1),(N-k)} = F_{0.05,2,21} = 3.47 \Longrightarrow \boxed{\text{RR is } (3.47, \infty)}$$

10. The decision is:

$$F = 5.896 > 3.47 \Longrightarrow \boxed{\text{Reject } H_0}$$

<u>Q2.</u> Three types of medium sized cars assembled in New Zealand have been test driven by a motoring magazine and compared on a variety of criteria. In the area of fuel efficiency performance, five cars of each brand were each test driven 1000 km; the km per liter data are obtained as follows:

	Kilometres per liter					Total	Mean
Brand A	7.6	8.4	8	7.6	8.4	40	8
Brand B	7.8	8	9.1	8.5	9.6	43	8.6
Brand C	9.6	10.4	9.2	9.7	10.6	49.5	9.9

Let the one way ANOVA tabulated as follows:

Source of variation	**Sum of squares**	**Degrees of freedom**	**Mean Squares**	**Test Statistics**
Treatments	SSA	df_{trt}	MSA	
Errors	SSE	df_{er}	MSE	f
Total	SST	df_{tot}		

At a significance level of α =0.05, we want to compare the means of the three groups.

1. **Write the hypotheses H_0 and H_1. Explain (2 marks).**

$H_0: \mu_1 = \mu_2 = \mu_3$ vs H_1: at least two of the means are not equal.

Are the three bands have the same fuel consumption or not?

2. **The grand mean $\bar{y}_{..}$ is**

(A) 40+43+49.5)/3	(B) (40+43+49.5)/5	(C) (40+43+49.5)/15=8.83333

3. **The value of SSA is**

$$SS_{trt} = \sum_{i=1}^{k} n_i(\bar{y}_{i.} - \bar{y}_{..})^2$$
$$= 5 \times (8 - 8.83)^2 + 5 \times (8.6 - 8.83)^2 + 5 \times (9.9 - 8.83)^2 = \boxed{9.43}$$

4. $\sum\sum\left(\mathbf{y_{ij}}-\bar{\mathbf{y}}_{..}\right)^2=\sum\sum\mathbf{y_{ij}^2}-\mathbf{15}\bar{\mathbf{y}}_{..}^2$ **and** $\sum\sum\mathbf{y_{ij}^2}=\mathbf{1184.11}$ **. Then SST is:**

$$\begin{aligned}\text{SST} &= \sum\sum y_{ij}^2 - 15\bar{y}_{..}^2 \\ &= 1184.11 - 15(8.8333)^2 = 13.69\end{aligned}$$

5. The value of SSE is

$$\text{SSE} = \text{SST} - \text{SSA} = 13.69 - 9.43 = \boxed{4.26}$$

6. The degrees of freedom of the treatments ($\mathbf{df_{trt}}$) is

$$k - 1 = 3 - 1 = 2$$

7. The degrees of freedom of the error ($\mathbf{df_{er}}$) is

$$N - k = 15 - 3 = 12$$

8. The degrees of freedom of the total ($\mathbf{df_{tot}}$) is

$$N - 1 = 15 - 1 = 14$$

9. The Mean Squares of the treatments (MSA) is

$$\text{MSA} = \frac{\text{SSA}}{k-1} = \frac{9.43}{3-1} = 4.72$$

10. The Mean Squares of the errors (MSE) is

$$\text{MSE} = \frac{\text{SSE}}{N-k} = \frac{4.26}{15-3} = 0.355$$

11. The value of the test statistic F is

$$F = \frac{\text{MSA}}{\text{MSE}} = \frac{4.72}{0.355} = 13.296$$

12. The rejection region (R.R) of $\mathbf{H_0}$ is

$$F_{\alpha,(k-1),(N-k)} = F_{0.05,2,12} = 3.89 \Longrightarrow \boxed{\text{RR is } (3.89, \infty)}$$

13. The decision is

$$F = 13.296 > 3.89 \Longrightarrow \boxed{\text{Reject } H_0}$$

⏩ A pharmaceutical company conducts an experiment to test the effect of a new cholesterol medication. The company selects 15 subjects randomly from a larger population. Each subject is randomly assigned to one of three treatment groups. Within each treatment group, subjects receive a different dose of the new medication. In Group 1, subjects receive 0 mg/day; in Group 2, 50 mg/day; and in Group 3, 100 mg/day. The treatment levels represent all the levels of interest to the experimenter, so this experiment used a fixed-effects model to select treatment levels for study.

After 30 days, doctors measure the cholesterol level of each subject. The results for all 15 subjects appear in the table below:

Dosage		
Group 1, 0 mg	Group 2, 50 mg	Group 3, 100 mg
210	210	180
240	240	210
270	240	210
270	270	210
300	270	240

$\sum y_{i.}$ 1290 1230 1050

$\bar{y}_{i.}$ $\frac{1290}{5} = 258$ $\frac{1230}{5} = 246$ $\frac{1050}{5} = 210$

$\bar{y}_{..} = \frac{1290+1230+1050}{15} = 238$

At $\alpha = 0.05$, does dosage level have a significant effect on cholesterol level? The analysis of variance (ANOVA) Table:

Source of Variation	SS	df	MS	F
Treatment	"A"	2	"B"	"E"
Error	"C"	12	750	
Total	15240	14		

[1]. The value of the "sum of squares treatment (SSA)" denoted by "A" in the ANOVA table is:

$$\sum_{i=1}^{k} n_i(\bar{y}_{i.} - \bar{y}_{..})^2 = 5(258 - 238)^2 + 5(246 - 238)^2 + 5(210 - 238)^2 = 6240$$

A	B	C	D
4580	5147	6240	3697

[2]. The "sum squares error (SSE)" denoted by "C" in the ANOVA table is:

$$SSE = 15240 - 6240 = 9000$$

A	B	C	D
4578	3978	4256	9000

[3]. The value of the "mean squares treatment (MSA)" denoted by "B" in the ANOVA table is:

$$MSA = \frac{6240}{2} = 3120$$

A	B	C	D
3120	4512	5689	1647

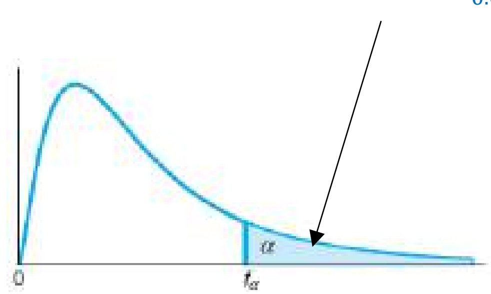
[4]. The value of the calculated F test, denoted by "E" in the ANOVA table is:

$$F = \frac{3120}{750} = 4.16$$

A	B	C	D
3.28	4.16	2.15	6.20

[5]. The decision is:

$$F = 4.16 > F_{0.05,2,12} = 3.88$$



3.88

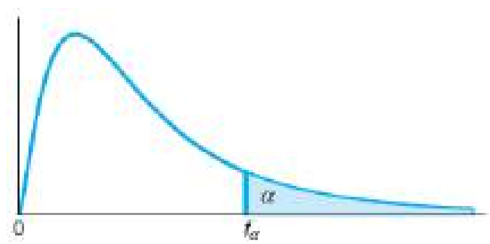
A	B	C
Reject H_0	Don't Reject H_0	Decision is not possible

Two-way ANOVA – Randomized Complete Block Design (RCBD):

$H_0^{trt}: \mu_1 = \mu_2 = \cdots = \mu_k$
H_1^{trt}: at least two $\mu_i \neq \mu_j$.

$H_0^{Block}: \mu_1 = \mu_2 = \cdots = \mu_b$
H_1^{Block}: at least two $\mu_i \neq \mu_j$.

Source of Variation	Sum Square	Degrees of freedom	Mean Square	F
Treatment	$SSA = \sum_{i=1}^{k} b(\bar{y}_{i.} - \bar{y}_{..})^2 = \frac{1}{b}\sum_{i=1}^{k} y_{i.}^2 - \frac{y_{..}^2}{N}$	$k-1$	$MSA = \frac{SSA}{k-1}$	$F_{trt} = \frac{MSA}{MSE}$
Block	$SSB = \sum_{j=1}^{b} k(\bar{y}_{.j} - \bar{y}_{..})^2 = \frac{1}{k}\sum_{j=1}^{b} y_{.j}^2 - \frac{y_{..}^2}{N}$	$b-1$	$MSB = \frac{SSB}{b-1}$	$F_{Block} = \frac{MSB}{MSE}$
Error	$SSE = \sum_{i=1}^{k}\sum_{j=1}^{b}(y_{ij} - \bar{y}_{i.} - \bar{y}_{.j} - \bar{y}_{..})^2$	$(k-1)(b-1)$	$MSE = \frac{SSE}{(k-1)(b-1)}$	
Total	$SST = \sum_{i=1}^{k}\sum_{j=1}^{b} b(\bar{y}_{ij} - \bar{y}_{..})^2 = \sum_{i=1}^{k}\sum_{j=1}^{b} y_{ij}^2 - \frac{y_{..}^2}{N}$ $\boxed{SST = SSA + SSB + SSE}$	$= bk - 1 = N - 1$		



Reject H_0^{trt} if $F_{trt} > F_{\alpha,k-1,(k-1)(b-1)}$

Reject H_0^{Block} if $F_{Block} > F_{\alpha,b-1,(k-1)(b-1)}$

An Experiment to test the difference between 5 treatments, using $\alpha = 0.02$ gave the following data:

Treatment	Blocks			Total
1	11	15	15	41
2	6	11	12	29
3	7	5	13	25
4	8	10	11	29
5	4	6	10	20
Total	36	47	61	144

The Analysis of Variance (ANOVA) Table:

Source of Variation	SS	df	MS	F
Treatment	SSA		MSA	F
Blocks	SSB			
Error	SSE		MSE	
Total				

1) SSA =

$$SSA = \frac{1}{b}\sum_{i=1}^{k} y_{i.}^2 - \frac{y_{..}^2}{N} = \frac{1}{3}\left[(41)^2 + (29)^2 + (25)^2 + (29)^2 + (20)^2\right] - \frac{(144)^2}{15}$$

$$= \frac{1}{3}[4388] - \frac{(144)^2}{15} = 80.2666$$

A) 20.5	B) 35.6	C) 28.3	D) 80.3

2) SSB =

$$SSB = \frac{1}{k}\sum_{j=1}^{b} y_{.j}^2 - \frac{y_{..}^2}{N} = \frac{1}{5}[(36)^2 + (47)^2 + (61)^2] - \frac{(144)^2}{15}$$

$$= \frac{1}{5}[7226] - \frac{(144)^2}{15} = 62.8$$

A) 36.9	B) 82.0	C) 62.8	D) 55.7

3) SSE =

$$SST = \sum_{i=1}^{k}\sum_{j=1}^{b} y_{ij}^2 - \frac{y_{..}^2}{N}$$

$$= \begin{bmatrix} (11)^2 + (6)^2 + (7)^2 + (8)^2 + (4)^2 \\ (15)^2 + (11)^2 + (5)^2 + (10)^2 + (6)^2 \\ (15)^2 + (12)^2 + (13)^2 + (11)^2 + (10)^2 \end{bmatrix} - \frac{(144)^2}{15}$$

$$= [1552] - \frac{(144)^2}{15} = 169.6$$

$SST = SSA + SSB + SSE$

$169.6 = 80.3 + 62.8 + SSE$

$SSE = 26.5$

A) 26.5	B) 15.0	C) 21.8	D) 17.7

4) MSA =

$$MSA = \frac{SSA}{k-1} = \frac{80.3}{4} = 20.075$$

A) 11.5	B) 15.3	C) 24.3	D) 20.1

5) MSE =

$$MSE = \frac{SSE}{(k-1)(b-1)} = \frac{26.5}{(5-1)(3-1)} = 3.3125$$

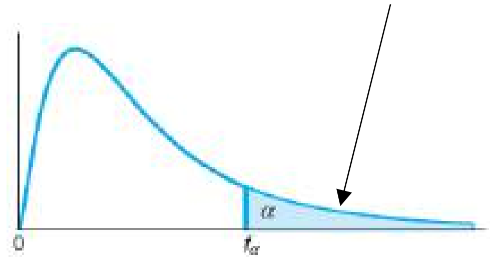
A) 11.7	B) 6.1	C) 9.8	D) 13.2

6) The decision is

$H_0: \mu_1 = \mu_2 = \mu_3 = \mu_4 = \mu_5$
H_1: at least two $\mu_i \neq \mu_j$.

Source of Variation	SS	df	MS	F
Treatment	80.3	4	$\frac{80.3}{4} = 20.1$	$\frac{20.1}{3.3125} = 6.1$
Blocks	62.8	2		
Error	26.5	8	$\frac{26.5}{8} = 3.3125$	
Total				

$F = 6.1 > F_{0.05,4,8} = 3.84$



3.84

Then we Reject H_0

A) Reject H_0	B) Accept H_0	C) No Decision

An Experiment to test the difference between 5 treatments, using $\alpha = 0.05$ gave the following data:

Treatment	Blocks			Total
1	17	10	18	45
2	16	9	14	39
3	15	8	9	32
4	16	9	14	39
5	17	9	13	39
Total	81	45	68	194

The Analysis of Variance (ANOVA) Table:

Source of Variation	SS	df	MS	F
Treatment	a		b	c
Blocks	d			
Error	e			
Total				

1) a =

$$SSA = \frac{1}{b}\sum_{i=1}^{k} y_{i.}^2 - \frac{y_{..}^2}{N} = \frac{1}{3}\left[(45)^2 + (39)^2 + (32)^2 + (39)^2 + (39)^2\right] - \frac{(194)^2}{15}$$

$$= \frac{1}{3}[7612] - \frac{(194)^2}{15} = 28.2666$$

$$SSA = b\sum_{i=1}^{k}(\bar{y}_{i.} - \bar{y}_{..})^2 =$$

$$3 \times \left[(15 - 12.933)^2 + (13 - 12.933)^2 + (10.67 - 12.933)^2 + (13 - 12.933)^2 + (13 - 12.933)^2\right]$$

A) 20.5	B) 35.6	C) 28.3	D) 10.5

2) b =

$$MSA = \frac{SSA}{k-1} = \frac{28.3}{4} = 7.075$$

A) 11.5	B) 7.1	C) 4.3	D) 15.5

3) c =

$$MSE = \frac{SSE}{(k-1)(b-1)} = \frac{17.7}{(5-1)(3-1)} = 2.2125$$

A) 6.7	B) 8.0	C) 1.8	D) 3.2

4) d =

$$SSB = \frac{1}{k}\sum_{j=1}^{b} y_{.j}^2 - \frac{y_{..}^2}{N} = \frac{1}{5}[(81)^2 + (45)^2 + (68)^2] - \frac{(194)^2}{15}$$

$$= \frac{1}{5}[13210] - \frac{(194)^2}{15} = 132.9333$$

A) 132.9	B) 182.0	C) 112.8	D) 100.7

5) e =

$$SST = \sum_{i=1}^{k}\sum_{j=1}^{b} y_{ij}^2 - \frac{y_{..}^2}{N}$$

$$= \begin{bmatrix} (17)^2 + (16)^2 + (15)^2 + (16)^2 + (17)^2 \\ (10)^2 + (9)^2 + (8)^2 + (9)^2 + (9)^2 \\ (18)^2 + (14)^2 + (9)^2 + (14)^2 + (13)^2 \end{bmatrix} - \frac{(194)^2}{15}$$

$$= [2688] - \frac{(194)^2}{15} = 178.933$$

$$SST = SSA + SSB + SSE$$

$$178.9 = 28.3 + 132.9 + SSE$$

$$SSE = 17.733$$

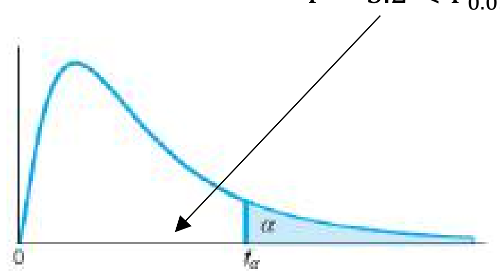
A) 11.5	B) 15.0	C) 21.8	D) 17.7

6) The decision is

$H_0: \mu_1 = \mu_2 = \mu_3 = \mu_4 = \mu_5$
$H_1:$ at least two $\mu_i \neq \mu_j$.

Source of Variation	SS	df	MS	F
Treatment	28.3	4	$\frac{28.3}{4} = 7.1$	$\frac{7.1}{2.2125} = 3.2$
Blocks	7.1	2		
Error	17.7	8	$\frac{17.7}{8} = 2.2125$	
Total				

$F = 3.2 < F_{0.05,4,8} = 3.84$



3.84

Then we Accept H_0

A) Reject H_0	B) Accept H_0	C) No Decision

■ **TABLE 4.3**

Randomized Complete Block Design for the Vascular Graft Experiment

Extrusion Pressure (PSI)	Batch of Resin (Block) 1	2	3	4	5	6	Treatment Total
8500	90.3	89.2	98.2	93.9	87.4	97.9	556.9
8700	92.5	89.5	90.6	94.7	87.0	95.8	550.1
8900	85.5	90.8	89.6	86.2	88.0	93.4	533.5
9100	82.5	89.5	85.6	87.4	78.9	90.7	514.6
Block Totals	350.8	359.0	364.0	362.2	341.3	377.8	$y_{..} = 2155.1$

To perform the analysis of variance, we need the following sums of squares:

$$SS_T = \sum_{i=1}^{4}\sum_{j=1}^{6} y_{ij}^2 - \frac{y_{..}^2}{N} = 193{,}999.31 - \frac{(2155.1)^2}{24} = 480.31$$

$$SS_{\text{Treatments}} = \frac{1}{b}\sum_{i=1}^{4} y_{i.}^2 - \frac{y_{..}^2}{N} = \frac{1}{6}[(556.9)^2 + (550.1)^2 + (533.5)^2 + (514.6)^2] - \frac{(2155.1)^2}{24} = 178.17$$

$$SS_{\text{Blocks}} = \frac{1}{a}\sum_{j=1}^{6} y_{.j}^2 - \frac{y_{..}^2}{N} = \frac{1}{4}[(350.8)^2 + (359.0)^2 + \cdots + (377.8)^2] - \frac{(2155.1)^2}{24} = 192.25$$

$$SS_E = SS_T - SS_{\text{Treatments}} - SS_{\text{Blocks}} = 480.31 - 178.17 - 192.25 = 109.89$$

The ANOVA is shown in Table 4.4. Using $\alpha = 0.05$, the critical value of F is $F_{0.05,3,15} = 3.29$. Because $8.11 > 3.29$, we conclude that extrusion pressure affects the mean yield. The P-value for the test is also quite small. Also, the resin batches (blocks) seem to differ significantly, because the mean square for blocks is large relative to error.

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Ratio
Pressure	3	178.17125	59.3904	8.1071
Batch	5	192.25208	38.4504	5.2487
Error	15	109.88625	7.3257	
C.Total	23	480.30958		

$F_{\alpha,k-1,(k-1)(b-1)} = F_{0.05,3,15} = 3.29 \Longrightarrow F_{trt} = 8.10 > 3.29$ then we reject H_0^{trt}

$F_{\alpha,b-1,(k-1)(b-1)} = F_{0.05,5,15} = 2.9 \Longrightarrow F_{Block} = 5.25 > 2.9$ then we reject H_0^{Block}