

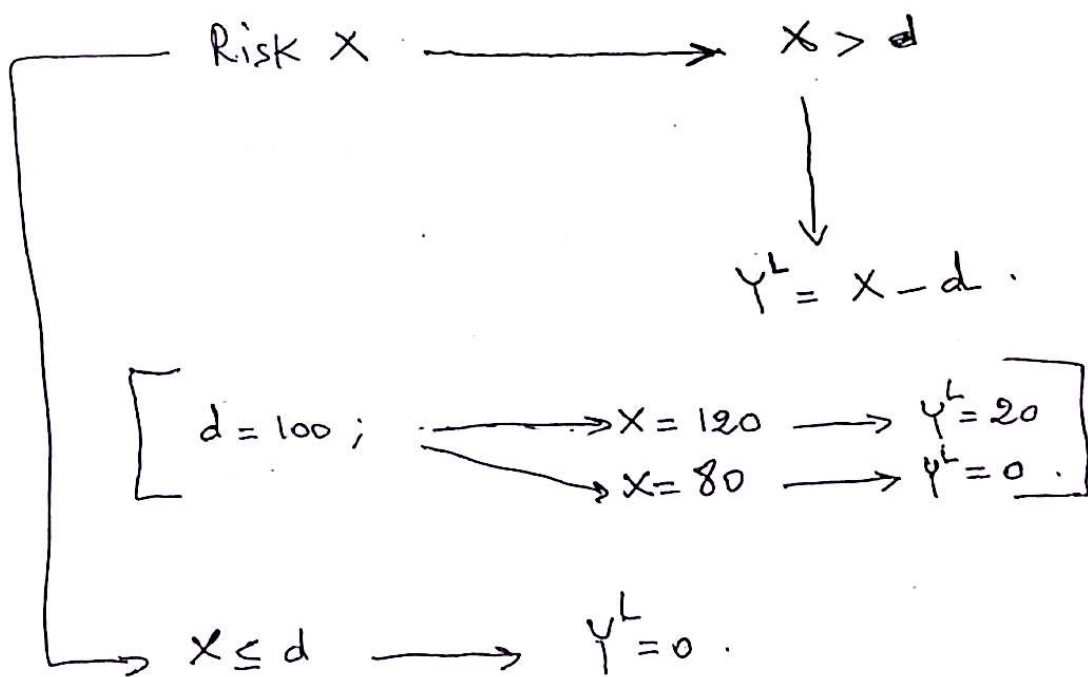
Chapter ③: Modifications of loss distributions.

insurance policy \longrightarrow risk X

\downarrow
payment Y

in some cases, Y is different from X .

①: Ordinary policy deductible:



Y^L called cost-per loss is given:

$$Y^L = (X - d)_+$$

Y^L is called also left censored and shifted variable.

- cdf of Y^L : -

$$\begin{aligned}
 F_{Y^L}(y) &= P(Y^L \leq y) \\
 &= P((X-d)_+ \leq y) \\
 &= P((X-d)_+ \leq y; X \leq d) + P((X-d)_+ \leq y; X > d) \\
 &= P(0 \leq y; X \leq d) + P(X-d \leq y; X > d)
 \end{aligned}$$

$y \geq 0$:

$$\begin{aligned}
 F_{Y^L}(y) &= F_X(d) + F_X(y+d) - F_X(d) \\
 F_{Y^L}(y) &= \begin{cases} F_X(y+d) & ; y \geq 0 \\ 0 & ; y < 0 \end{cases}
 \end{aligned}$$

- Remark: • $F_{Y^L}(0) = F_X(d)$, $F_{Y^L}(0^-) = 0$.
- Y^L is a mixed distribution.

$$f_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ F_X(d) & y = 0 \\ f_X(y+d) & y > 0 \end{cases}$$

- mgf of Y^L : (X continuous)

$$\begin{aligned}
 M_{Y^L}(t) &= E e^{tY} = E e^{t(X-d)_+} \\
 &= \int_{-\infty}^d e^{t(x-d)_+} f_X(x) dx + \int_d^{\infty} e^{t(x-d)_+} f_X(x) dx \\
 &= F_X(d) + e^{-td} \left(\int_d^{\infty} e^{tx} f_X(x) dx \right) \\
 &= F_X(d) + e^{-td} \left(M_X(t) - \int_0^d e^{tx} f_X(x) dx \right)
 \end{aligned}$$

(2)

Example: Determine the pdf, cdf and sdf for Y_L if the ground-up loss amount function has an exponential distribution with mean $\frac{1}{\theta}$ and an ordinary deductible of d .

$$X \sim \text{Exp}(\theta); \quad f_X(x) = \theta e^{-\theta x}; \quad x \geq 0,$$

$$F_X(x) = \begin{cases} 1 - e^{-\theta x}; & x \geq 0 \\ 0; & x < 0 \end{cases}$$

$$S_X(x) = \begin{cases} e^{-\theta x}; & x \geq 0 \\ 1; & x < 0 \end{cases}$$

$$\bullet \quad F_{Y_L}(y) = F_X(y+d) = \begin{cases} 0 & y < 0 \\ 1 - e^{-\theta(y+d)} & y \geq 0 \end{cases};$$

$$F_{Y_L}(y) = \begin{cases} 0 & y < 0 \\ F_X(y+d) & y \geq 0 \end{cases}$$

$$= \begin{cases} 0 & y < 0 \\ 1 - e^{-\theta(y+d)} & y \geq 0 \end{cases}$$

$$f_{Y_L}(y) = \begin{cases} 0 & y < 0 \\ \theta e^{-\theta(y+d)} & y \geq 0 \end{cases}$$

$$S_{Y_L}(y) = \begin{cases} 1 & y < 0 \\ \frac{e^{-\theta(y+d)}}{e^{-\theta d}} & y \geq 0 \end{cases}$$

$$\bullet \quad E(Y_L) = E(X-d)_+ = \int_0^{\infty} (x-d)_+ f_X(x) dx$$

$$= \int_d^{\infty} (x-d) f_X(x) dx = \int_d^{\infty} (x-d) \theta e^{-\theta x} dx$$

$$= -\int_d^{\infty} (x-d) e^{-\theta x} dx + \int_d^{\infty} e^{-\theta x} dx = \frac{1}{\theta} e^{-\theta d} = \frac{1}{\theta} e^{-\theta d}$$

$$\begin{aligned}
 m_{Y^L}(t) &= E e^{tY^L} = E e^{t(x-d)_+} \\
 &= \int_0^\infty e^{t(x-d)_+} \theta e^{-\theta x} dx = \int_0^d \theta e^{-\theta x} dx + \int_d^\infty e^{t(x-d)} \theta e^{-\theta x} dx \\
 &= 1 - e^{-\theta d} + e^{-dt} \int_d^\infty \theta e^{(t-\theta)x} dx \\
 &= 1 - e^{-\theta d} + e^{-dt} \frac{\theta}{t-\theta} e^{(t-\theta)x} \Big|_d^\infty \\
 &= 1 - e^{-\theta d} + e^{-dt} \frac{\theta}{t-\theta} (-e^{-(t-\theta)d}) \\
 &= \underbrace{1 - e^{-\theta d}}_{1-q} + \frac{\theta}{\theta-t} e^{-\theta d}
 \end{aligned}$$

$t < \theta$

• Now we define the cost-per payment:

$$Y^L = (Y^L | X > d) = (Y^L | Y^L > 0)$$

A random variable X , want $B \equiv (X > c)$

$$Y = (X | X > c)$$

$$\begin{aligned}
 F_Y(y) &= P(Y \leq y | X > c) = \frac{P(c < X \leq y)}{P(X > c)} \\
 &= P(X \leq y | X > c) = \frac{P(c < X \leq y)}{P(X > c)}
 \end{aligned}$$

$$= \begin{cases} 0 & y < c \\ \frac{F_X(y) - F_X(c)}{1 - F_X(c)} & y \geq c \end{cases}$$

$$f_Y(y) = \begin{cases} 0 & y < c \\ \frac{f_X(y)}{1 - F_X(c)} & y \geq c \end{cases}$$

$$E(Y) = \int_c^\infty \frac{y f_X(y)}{S_X(c)} dx =$$

(4)

• CDF of Y^L :

$$F_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{F_Y(y) - F_Y(0)}{1 - F_Y(0)} & y \geq 0 \end{cases}$$

$$= \begin{cases} 0 & y < 0 \\ \frac{F_X(y+d) - F_X(d)}{1 - F_X(d)} & y \geq 0 \end{cases}$$

• f_{Y^L} :

$$f_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{f_X(y+d)}{1 - F_X(d)} & y \geq 0 \end{cases}$$

• Mean of Y^L :

$$E(Y^L) = \int_0^{\infty} y \frac{f_X(y+d)}{1 - F_X(d)} dy$$

$$E(Y^L) = \int_0^{\infty} y dF_{Y^L}(y) = 0 \cdot f_{Y^L}(0) + \int_0^{\infty} y \cdot f_X(y+d) dy$$

$$E(Y^L) = \frac{E(Y^L)}{1 - F_X(d)}$$

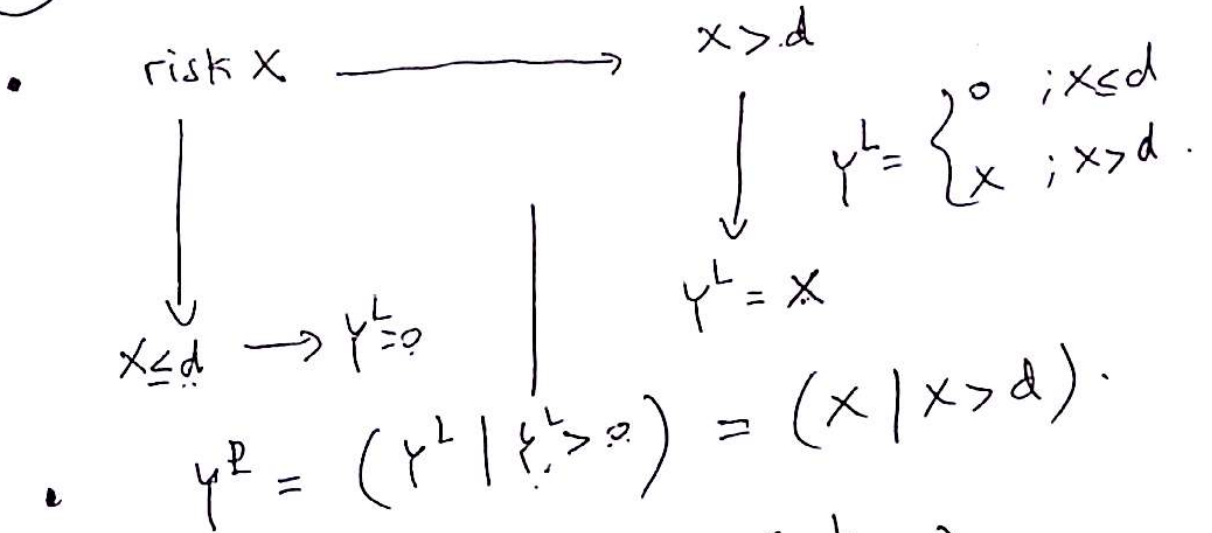
• $E(X \wedge d) = ?$

$$Y^L = \begin{cases} 0 & X < d \\ X - d & X > d \end{cases}$$

$$E Y^L = E(X) - E(X \wedge d)$$

②

Franchise deductible:



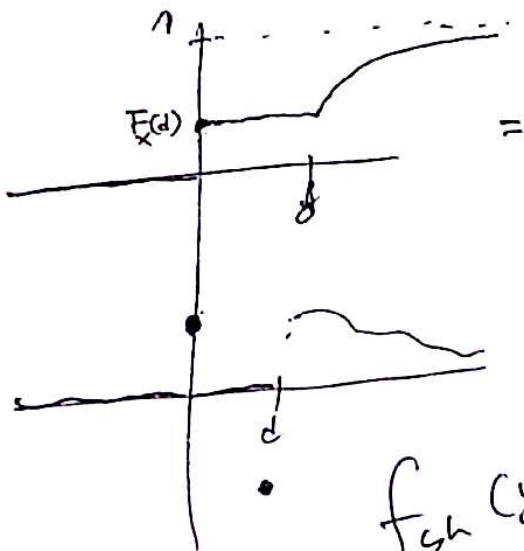
• CDF of Y^L : $F_{Y^L}(y) = P(Y^L \leq y)$

$= P(Y^L \leq y; X \leq d) + P(Y^L \leq y; X > d)$

$= P(0 \leq y, X \leq d) + P(X \leq y; X > d)$

$= F_X(d) + \begin{cases} 0 & y < 0 \\ F_X(y) - F_X(d) & 0 \leq y < d \\ F_X(y) & y \geq d \end{cases}$

$y \geq 0$:



• $f_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ F_X(d) & 0 \leq y < d \\ 0 & y = 0 \\ f_X(y) & y \geq d \end{cases}$

⑥

• CDF of Y^L :

$$F_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{F_{Y^L}(y) - F_{Y^L}(0)}{1 - F_{Y^L}(0)} & y \geq 0 \end{cases}$$

$$= \begin{cases} 0 & 0 \leq y < d \\ \frac{F_X(y) - F_X(d)}{1 - F_X(d)} & y \geq d \end{cases}$$

$$= \begin{cases} 0 & y < d \\ \frac{F_X(y) - F_X(d)}{1 - F_X(d)} & y \geq d \end{cases}$$

$$f_{Y^L}(y) = \begin{cases} 0 & y < d \\ \frac{f_X(y)}{1 - F_X(d)} & y \geq d \end{cases}$$

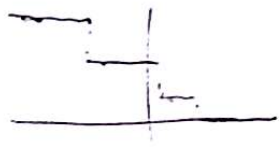
Example: $X \sim \text{Exp}(\theta)$. Insurance policies are subject to a franchise deductible of d .

Compute the cdf, pdf, sdf, hazard rate function of Y^L and Y^I .

Y^L :

$$F_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ 1 - e^{-\theta y} & 0 \leq y < d \\ 1 - e^{-\theta y} & y \geq d \end{cases}$$

$$f_{Y^L}(y) = \begin{cases} 0 & y < d, y \neq 0 \\ \theta e^{-\theta y} & y = 0 \\ \theta e^{-\theta y} & y \geq d \end{cases}$$



$$S_{Y^L}(y) = \begin{cases} 1 & y < 0 \\ e^{-\theta y} & 0 \leq y < d \\ e^{-\theta y} & y \geq d \end{cases}$$

$$h_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{1 - e^{-\theta d}}{e^{-\theta y}} & 0 \leq y < d \\ 0 & y \geq d \end{cases}$$

• Y^L :

$$F_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{e^{-\theta y} - e^{-\theta d}}{1 - e^{-\theta d}} & 0 \leq y < d \\ e^{-\theta(y-d)} & y \geq d \end{cases}$$

$$f_{Y^L}(y) = \begin{cases} 0 & y < 0 \\ \frac{\theta e^{-\theta y}}{1 - e^{-\theta d}} & 0 \leq y < d \\ \theta e^{-\theta(y-d)} & y \geq d \end{cases}$$

$$S_{Y^L}(x) = \begin{cases} 1 & x < 0 \\ e^{-\theta(x-d)} & x \geq 0 \end{cases}$$

$$h_{Y^L}(x) = \begin{cases} 0 & x < 0 \\ \theta e^{-\theta(x-d)} & x \geq 0 \end{cases}$$

• Mean of Y^L and Y^R :

$$Y^L = X - X \wedge d + d \mathbb{1}_{\{X > d\}} = \begin{cases} X & X \leq d \\ d & X > d \end{cases}$$

$$E(Y^L) = E(X) - E(X \wedge d) + d S_X(d)$$

$$E(Y^L) = \frac{E(Y^L)}{1 - F_X(d)}$$

Ex: Let $X \sim \text{Exp}(\theta)$ with franchise deductible of d .

Compute $E(Y^L)$ and $E(Y^F)$.

$$E(X \wedge d) = \int_0^d f_X(x) dx$$

$$E(Y^L) = \int_{-\infty}^0 + \int_0^0 + \int_0^d + \int_d^{\infty}$$

$$= 0 + 0 + 0 + \int_d^{\infty} y \theta e^{-\theta y} dy$$

$$\rightarrow E(X \wedge d) = \int_0^d e^{-\theta x} dx = -\frac{1}{\theta} e^{-\theta x} \Big|_0^d$$

$$= \frac{1}{\theta} (1 - e^{-\theta d})$$

$$E(Y^L) = E(X) - E(X \wedge d) + \int_d^{\infty} f_X(x) dx$$

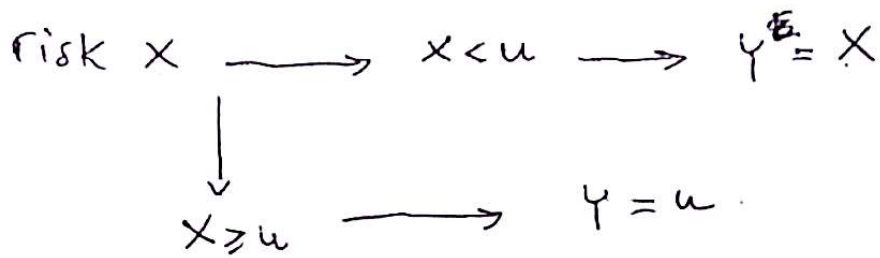
$$= \frac{1}{\theta} - \frac{1}{\theta} (1 - e^{-\theta d}) + d e^{-\theta d} = \left(\frac{1}{\theta} + d\right) e^{-\theta d}$$

$$E(Y^F) = \frac{E(Y^L)}{f_X(d)} = \frac{\left(\frac{1}{\theta} + d\right) e^{-\theta d}}{e^{-\theta d}}$$

$$= \frac{1}{\theta} + d$$

(5)

3 Policy limit:



The limited loss random variable

$$Y = X \wedge u$$

• CDF of Y :

$$\begin{aligned}
 F_Y(y) &= P(Y \leq y) = P(X \wedge u \leq y) \\
 &= P(X \wedge u \leq y; X \leq u) + P(X \wedge u \leq y; X > u) \\
 &= P(X \leq y; X \leq u) + P(u \leq y; X > u) \\
 &= F_X(y \wedge u) + \begin{cases} 0 & u > y \\ 1 - F_X(u) & u \leq y \end{cases} \\
 &= \begin{cases} F_X(y) & y < u \\ 1 & y \geq u \end{cases}
 \end{aligned}$$

• PDF of Y :

$$f_Y(y) = \begin{cases} f_X(y) & y < u \\ 1 - F_X(u) & y = u \\ 0 & y > u \end{cases}$$

$$E(Y) = E(X \wedge u) = \int_0^u S_X(x) dx$$

Example: let $X \sim \text{Gamma}(1, \beta)$.

$$F_Y(y) = \begin{cases} 0 & y < 0 \\ 1 - e^{-\beta y} & 0 \leq y < u \\ 1 & y \geq u \end{cases}$$

$$f_Y(y) = \begin{cases} 0 & y < 0 \\ \beta e^{-\beta y} & 0 \leq y < u \\ e^{-\beta u} & y = u \\ 0 & y > u \end{cases}$$

$$E(Y) = \int_0^u e^{-\beta x} dx = -\frac{1}{\beta} e^{-\beta x} \Big|_0^u = \frac{1}{\beta} (1 - e^{-\beta u})$$

④ Coinsurance:

An insurance policy is defined with a coinsurance factor $0 < \alpha < 1$.

$$\text{risk } X \longrightarrow Y = \alpha X$$

$$F_Y(y) = F_X\left(\frac{y}{\alpha}\right)$$

$$f_Y(y) = \frac{1}{\alpha} f_X\left(\frac{y}{\alpha}\right)$$

⑤ Inflation effect:

we suppose that r is the inflation coefficient. Then a payment f becomes $(1+r)f$.

Suppose now an ordinary policy deductible with risk X .

$$X \longrightarrow (1+r)X$$

$$\downarrow$$

$$Y^L = ((1+r)X - d)_+ = (1+r) \left(X - \frac{d}{1+r} \right)_+$$

(6) Combination :

we suppose we have an ordinary deductible of d , limit of u , with coinsurance factor α and inflation coefficient r .

For a risk X , we have :

$$Y^L = \begin{cases} 0 & X < \frac{d}{1+r} \\ \alpha \cdot ((1+r)X - d) & \frac{d}{1+r} \leq X < \frac{u}{1+r} \\ \alpha(u-d) & X \geq \frac{u}{1+r} \end{cases}$$

we write $Y^L = \alpha \left[(1+r)X \wedge u - (1+r)X \wedge d \right]$

Ex : $r = 0.01$; $\alpha = 0.8$; $d = 200$; $u = 500$;

$x = 100$	$\xrightarrow{x=101}$	$Y^L = 0$
$x = 300$	$\xrightarrow{x=303}$	$Y^L = 82.4$
$x = 600$	$\xrightarrow{x=606}$	$Y^L = 240$

• CDF of Y^L :
* $\alpha = 1$; $r = 0$.

$$F_{Y^L}(y) = P(X \wedge u - X \wedge d \leq y)$$

$$= P\left(X \wedge u - X \wedge d \leq y ; \begin{matrix} x < d \\ d \leq x < u \\ x \geq u \end{matrix} \right)$$

$$= P(0 \leq y ; x < d)$$

$$+ P(x \leq d+y ; d \leq x \leq u)$$

$$+ P(u-d \leq y ; x \geq u)$$

$$= F_X(d) + F_X((d+y) \wedge u) - F_X(d) + \begin{cases} 0, & y < 0 \\ 1 - F_X(u), & y \geq u-d \end{cases}$$

$$= \begin{cases} F_X(d+y) & 0 \leq y < u-d \\ 1 & y \geq u-d \end{cases}$$

1

In general:

$$F_{y^L}(y) = \begin{cases} F_x\left(\frac{d+y/\alpha}{1+r}\right) & ; 0 \leq y < \alpha(u-d) \\ 1 & ; y \geq \alpha(u-d) \\ 0 & ; y < 0 \end{cases}$$

• $y^R = y^L \mid \{y^L > 0 : y < 0$

$$F_{y^R}(y) = \begin{cases} \frac{F_{y^L}(y) - F_{y^L}(0)}{1 - F_{y^L}(0)} & ; y \geq 0 \\ 0 & ; y < 0 \end{cases}$$

$$= \begin{cases} \frac{F_x\left(\frac{d+y/\alpha}{1+r}\right) - F_x\left(\frac{d}{1+r}\right)}{1 - F_x\left(\frac{d}{1+r}\right)} & ; y < \alpha(u-d) \\ 1 & ; y \geq \alpha(u-d) \end{cases}$$

$$f_{y^L}(y) = \begin{cases} F_x\left(\frac{d}{1+r}\right) & ; y = 0 \\ \frac{1}{\alpha(1-r)} f_x\left(\frac{d+y/\alpha}{1+r}\right) & ; 0 < y < \alpha(u-d) \\ 1 - F_x\left(\frac{u}{1+r}\right) & ; y = \alpha(u-d) \\ 0 & ; y > \alpha(u-d) \end{cases}$$

$$f_{y^R}(y) = \begin{cases} 0 & ; y < 0, y > \alpha(u-d) \\ \frac{\frac{1}{\alpha(1-r)} f_x\left(\frac{d+y/\alpha}{1+r}\right)}{1 - F_x\left(\frac{d}{1+r}\right)} & ; 0 \leq y \leq \alpha(u-d) \\ \cancel{1 - F_x\left(\frac{u}{1+r}\right)} & ; \cancel{\alpha(u-d)} \end{cases}$$

Ex: A health insurance has a deductible of 200, a policy limit of 5,000, and a coinsurance factor of 80%. Calculate the expected claim amount per loss and the expected claim amount per payment if losses follow an exponential distribution with mean 1,000.

$$Y^L = \alpha (X \wedge u - X \wedge d)$$

$$E(X \wedge u) = \int_0^{\infty} (x \wedge u) \lambda e^{-\lambda x} dx$$

$$= \int_0^u x \lambda e^{-\lambda x} dx + \int_u^{\infty} u \lambda e^{-\lambda x} dx$$

$$= -x e^{-\lambda x} \Big|_0^u + \int_0^u e^{-\lambda x} dx + u S_X(u)$$

$$= -u e^{-\lambda u} + 0 + \frac{1}{\lambda} (1 - e^{-\lambda u}) + u e^{-\lambda u}$$

$$= -\cancel{u e^{-\lambda u}} + \frac{1}{\lambda} (1 - e^{-\lambda u}) + \cancel{u e^{-\lambda u}}$$

$$= \left(-\cancel{2u} + \frac{1}{\lambda} \right) e^{-\lambda u} + \frac{1}{\lambda} + u$$

$$E Y^L = \alpha (E(X \wedge u) - E(X \wedge d))$$

$$= \alpha \left(0 + \frac{1}{\lambda} (1 - e^{-\lambda u}) - d - \frac{1}{\lambda} (1 - e^{-\lambda d}) \right)$$

$$= \frac{\alpha}{\lambda} (e^{-\lambda d} - e^{-\lambda u})$$

$$\lambda = 0.001$$

$$u = 5,000$$

$$d = 200, \alpha = 0.8$$

$$E Y^L = 649.59$$

$$Y^L = Y^P | Y^L > 0$$

$$E Y^P = \frac{E Y^L}{S_{Y^L}(0)} = \frac{E Y^L}{S_X(d)}$$

$$= \frac{649.59}{e^{-0.2}} = 793.41$$

(14)

(7)

loss elimination ratio:

This ratio calculates the difference between losses with and without a combination of deductible, limit, coinsurance and inflation.

$$LER = \frac{E(X) - E(Y^L)}{E(X)}$$

Ex: loss X follows an exponential distribution with mean ~~100~~ 1,000. We suppose that the policy is subject to a deductible of 500. Calculate the loss elimination ratio.

$$\bullet \quad E(X) = \frac{1}{\lambda} = 1,000.$$

$$\bullet \quad Y^L = X - X \wedge d = (X - d)_+$$

$$E(Y^L) = E(X) - E(X \wedge d)$$

$$= \frac{1}{\lambda} - \frac{1}{\lambda} (1 - e^{-\lambda d})$$

$$= \frac{1}{\lambda} e^{-\lambda d}$$

$$\bullet \quad LER = \frac{\frac{1}{\lambda} - \frac{1}{\lambda} e^{-\lambda d}}{\frac{1}{\lambda}} = 1 - e^{-\lambda d} = 1 - e^{-0.5} = 0.393 = 39.3\%$$

[For example if insurance company pays 2000 ~~for~~ without deductible, but with deductible 60.7% of 2000 = 1214]

(15)

⑧ Impact of deductible on the number of payments:

Collective risk model:

* N number of insurance policies in a period of time.

* X_1, X_2, \dots iid.

* $S = X_1 + \dots + X_N$

we suppose that each X_i is subject to a deductible d .

N^L = number of cost-per loss.

N^P = number of cost-per payment.

$N^L = N$

we suppose: $x_1 = 100, x_2 = 150, x_3 = 120, x_4 = 130$.
 $d = 125: y_1^L = 0, y_2^L = 25, y_3^L = 0, y_4^L = 5$.

$N^P = \sum_{i=1}^N \mathbb{1}_{(X_i \geq d)} = \sum_{i=1}^N z_i$

$m_{N^P}(t) = m_N(\ln(m_z(t)))$, $z = \mathbb{1}_{(x \geq d)}$
 $z \sim \text{Bernoulli}(p)$, $p = P(X \geq d) = S_X(d)$

$m_{N^P}(t) = m_N(\ln(1-p + pe^t))$

Ex: $N \sim \text{Poi}(\lambda) \rightarrow N^P \sim \text{Poi}(\lambda p)$

$m_{N^P}(t) = e^{\lambda(1-p+pe^t-1)} = e^{\lambda p(e^t-1)}$