

Tutorials

Problem 2.5.2 p.78

← To be Solved in Lecture class

* * * First, note that from the properties of conditional expectation
 $E[h(y) | Y=y] = h(y)$
 ↓
 determined by Y

Similarly, if Z is determined by X_0, \dots, X_n , then
 $E[Z | X_0, \dots, X_n] = Z$

Solution for 2.5.2

For $i \geq 1$, U_i independent identically distributed
 with $U_i \sim \text{Uniform}(0,1)$

$X_0 = 1$, $X_n = 2^n \cdot U_1 \dots U_n$, $n = 1, 2, \dots$

To prove X_n is a martingale, we have to prove:

$$\text{[1]} \quad E[|X_n|] < \infty$$

and

$$\text{[2]} \quad E[X_{n+1} | X_0, \dots, X_n] = X_n$$

For [1], as X_n is non-negative random variable,

$$\begin{aligned} E[|X_n|] &= E[X_n] = E[2^n \cdot U_1 \dots U_n] \\ &= 2^n \cdot E[U_1] \dots E[U_n] \end{aligned}$$

As U_i 's are independent, $E[U_1 \dots U_n] = E[U_1] \dots E[U_n]$

Remember if $U \sim \text{uniform}(a,b)$ then $E(U) = \frac{1}{2}(a+b)$
 So, $E[U_i] = \frac{1}{2}(0+1) = \frac{1}{2}$

$$\begin{aligned} &\rightarrow = 2^n \cdot \underbrace{\frac{1}{2} \cdot \frac{1}{2} \dots \frac{1}{2}}_{n\text{-times}} \\ &= \frac{2^n}{2^n} = 1 \\ \therefore E[|X_n|] &= 1 < \infty \quad (\text{finite}). \end{aligned}$$

$$\text{[2]} \quad E[X_{n+1} | X_0, \dots, X_n]$$

$$= E[2^{n+1} U_1 \dots U_n \cdot U_{n+1} | X_0, \dots, X_n]$$

$$= E[2^n \cdot U_1 \dots U_n \cdot 2 \cdot U_{n+1} | X_0, \dots, X_n]$$

$$= E[2^n \cdot U_1 \dots U_n \cdot 2 \cdot U_{n+1} \mid X_0, \dots, X_n]$$

$$= \underbrace{2^n \cdot U_1 \dots U_n}_{\substack{\text{is out because they are determined by} \\ X_0, \dots, X_n}} \cdot E[2 \cdot U_{n+1} \mid X_0, \dots, X_n]$$

$$= 2^n \cdot U_1 \dots U_n \cdot 2 E[U_{n+1} \mid X_0, \dots, X_n]$$

$$= 2^n \cdot U_1 \dots U_n \cdot 2 \underbrace{E[U_{n+1}]}$$

because U_{n+1} is independent of U_1, \dots, U_n and so independent of X_0, \dots, X_n .
Remember $E[X \mid Y] = E[X]$ if X is indep. of Y .

$$= 2^n \cdot U_1 \dots U_n \cdot 2 \cdot \frac{1}{2}$$

as $E[U_i] = \frac{1}{2}$

$$= 2^n \cdot U_1 \dots U_n$$

$$= X_n$$

So, we proved the second condition, and hence X_n is a martingale.

Problem 2.5.3

← To be solved in tutorials class.

$S_0 = 0$, $S_n = \xi_1 + \dots + \xi_n$, $n \geq 1$
where ξ_i are independent identically distributed with $\xi_i \sim \exp(1)$ (since $E[\xi_i] = 1$)

* Remember when $X \sim \exp(\lambda)$, then $E[X] = \frac{1}{\lambda}$.
we should prove that $X_n = 2^n e^{-S_n}$, $n \geq 0$ is a martingale.

□ X_n is a non-negative r.v., so

$$E[|X_n|] = E[X_n]$$

$$= E[2^n e^{-S_n}]$$

cause if $X \sim \exp(\lambda)$
 $E[X] = \int_0^\infty x \cdot \lambda e^{-\lambda x} dx$

because if $X \sim \text{Exp}(1)$,
 $E[e^{-X}] = \int_0^{\infty} e^{-x} \cdot 1 \cdot e^{-x} dx$

$$* E[e^{-\xi_i}] = \int_0^{\infty} e^{-x} \cdot e^{-x} dx$$

$$= \int_0^{\infty} e^{-2x} dx$$

$$= -\frac{1}{2} e^{-2x} \Big|_0^{\infty}$$

$$= \frac{1}{2}$$

$$= E[2^n e^{-S_n}]$$

$$= 2^n E[e^{-\xi_1} \cdot e^{-\xi_2} \dots e^{-\xi_n}]$$

$$= 2^n \underbrace{E[e^{-\xi_1}] \dots E[e^{-\xi_n}]}_{\text{since } \xi_i \text{ are independent}}$$

So, $E[|X_n|] = 1 < \infty$.

[2] Prove that $E[X_{n+1} | X_0, \dots, X_n] = X_n$

$$E[2^{n+1} e^{-S_{n+1}} | X_0, \dots, X_n]$$

$$= E[2^n \cdot e^{-S_n} \cdot 2 \cdot e^{-\xi_{n+1}} | X_0, \dots, X_n]$$

$$= 2^n \cdot e^{-S_n} E[2e^{-\xi_{n+1}} | X_0, \dots, X_n]$$

↑ because $2^n \cdot e^{-S_n}$ is determined by X_0, \dots, X_n

$$= 2^n \cdot e^{-S_n} \cdot 2 \underbrace{E[e^{-\xi_{n+1}}]}$$

since ξ_{n+1} is independent of X_0, \dots, X_n .

$$= 2^n \cdot e^{-S_n} \cdot 2 \cdot \frac{1}{2}$$

← as $E[e^{-\xi_i}] = \frac{1}{2}$

$$= 2^n \cdot e^{-S_n}$$

$$= X_n$$

So, X_n is a martingale.

Problem 2.5.4

← To be solved in tutorials class.
 Only Prove X_n is martingale.

ξ_i are independent identically distributed

ξ_i are independent identically distributed with $\xi_i \sim \text{Bernolli}(P)$ $0 < P < 1$

$$X_0 = 1, \quad X_n = P^{-n} \cdot \xi_1 \cdots \xi_n, \quad n=1,2,\dots$$

$$\textcircled{1} E[|X_n|] = E[X_n] = E[P^{-n} \cdot \xi_1 \cdots \xi_n]$$

$$= P^{-n} E[\xi_1] \cdots E[\xi_n]$$

as ξ_i s are independent

As if $X \sim \text{Bernolli}(P)$
then $E[X] = P$

$$= P^{-n} \cdot \underbrace{P \cdots P}_{n\text{-times}}$$

$$= P^{-n} \cdot P^n$$

$$= P^0 = 1 < \infty$$

$$\textcircled{2} E[X_{n+1} | X_0, \dots, X_n]$$

$$= E[\bar{P}^{(n+1)} \xi_1 \cdots \xi_{n+1} | X_0, \dots, X_n]$$

$$= E[P^{-n} \xi_1 \cdots \xi_n \cdot \bar{P}^1 \cdot \xi_{n+1} | X_0, \dots, X_n]$$

$$= \underbrace{P^{-n} \xi_1 \cdots \xi_n}_{\text{is out because it is determined by } X_0, \dots, X_n} E[\bar{P}^1 \cdot \xi_{n+1} | X_0, \dots, X_n]$$

is out because it is determined by X_0, \dots, X_n .

$$= P^{-n} \xi_1 \cdots \xi_n \cdot \bar{P}^1 E[\xi_{n+1} | X_0, \dots, X_n]$$

$$= P^{-n} \xi_1 \cdots \xi_n \cdot \bar{P}^1 \cdot \underbrace{E[\xi_{n+1}]}$$

as ξ_{n+1} is independent of X_0, \dots, X_n

$$= P^{-n} \xi_1 \cdots \xi_n \cdot \bar{P}^1 \cdot P$$

$$= P^{-n} \xi_1 \cdots \xi_n$$

$$= X_n$$

So, X_n is a martingale.