

1 // Lecture (11)⁺⁺

- Ch 2 Sec 2.5 p. 71: Martingales

Definition: A stochastic process $\{X_n : n = 0, 1, \dots\}$ is a martingale if for $n = 0, 1, \dots$

$$(a) E[|X_n|] < \infty$$

$$(b) E[X_{n+1} | X_0, \dots, X_n] = X_n$$

* See also some properties for martingale p. 72 and revise all properties p. 50 Textbook.

Note that:

- ① The law of total probability for conditional expectation can be written as

$$E[g(X)] = \sum_y E[g(X) | Y=y] P_Y(y)$$

$$\therefore E[g(X)] = E[E[g(X) | Y]]$$

For $g(X) = X$, we have $E(X) = \sum_y E[X | Y=y] P_Y(y)$

$$\therefore E(X) = E[E[X | Y]]$$

- ② $E[g(X) | Y=y] = E[g(X)]$ if X and Y are independent

- ③ $E[g(X)h(Y) | Y=y] = h(y)E[g(X) | Y=y]$, h is determined by y .

Remarks (i) If $g(X) = 1$ then $E[h(Y) | Y=y] = h(y)$

(ii) If Z is determined by X_0, \dots, X_n , then $E[Z | X_0, \dots, X_n] = Z$

$$\begin{aligned} \textcircled{4} E[g(X)h(Y)] &= \sum_y E[g(X)h(Y)|Y=y] P_Y(y) \\ &= \sum_y h(y) E[g(X)|Y=y] P_Y(y) \end{aligned}$$

... by using $\textcircled{1}$
... by using $\textcircled{3}$

$$\therefore E[g(X)h(Y)] = E[h(Y)E[g(X)|Y]]$$

* Pb 2.5.1 p. 77

Use the law of total probability for Conditional Expectations

$$E[E\{X|Y, Z\}|Z] = E[X|Z] \text{ to show}$$

$$E[X_{n+2}|X_0, \dots, X_n] = E[E\{X_{n+2}|X_0, \dots, X_{n+1}\}|X_0, \dots, X_n] \textcircled{2}$$

Conclude that when X_n is a martingale,

$$E[X_{n+2}|X_0, \dots, X_n] = X_n \textcircled{3}$$

Answer:

Let $Z = X_0, \dots, X_n$ and $Y = X_{n+1}$ in R.H.S of $\textcircled{2}$

$$\therefore E[E\{X_{n+2}|X_0, \dots, X_n, X_{n+1}\}|X_0, \dots, X_n] \text{ R.H.S of } \textcircled{2}$$

$$= E[E\{X_{n+2}|Z, Y\}|Z] = E[X_{n+2}|Z] \text{ ... by using } \textcircled{1}$$

$$= E[X_{n+2}|X_0, \dots, X_n] \text{ L.H.S of } \textcircled{2}$$

Thus formula $\textcircled{2}$ is already proved.

Now, if X_n is a martingale, then for $n \geq 0$

$$E[X_{n+1}|X_0, \dots, X_n] = X_n \textcircled{4}, \text{ so } E[X_{n+2}|X_0, \dots, X_{n+1}] = X_{n+1} \textcircled{5}$$

Substitute $\textcircled{5}$ in the R.H.S of $\textcircled{2}$, we get

$$E[X_{n+2}|X_0, \dots, X_n] = E[X_{n+1}|X_0, \dots, X_n] = X_n \text{ ... by using } \textcircled{4}$$

Thus formula $\textcircled{3}$ is also proved. #

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* Pb 2.5.2 p.78

Let U_1, U_2, \dots be independent random variables each uniformly distributed over the interval $(0, 1]$. Show that $X_0 = 1$ and $X_n = 2^n U_1 \dots U_n$ for $n = 1, 2, \dots$ defines a martingale.

Answer:

To prove that X_n is a martingale, we have to prove that:

1 $E[|X_n|] < \infty$

2 $E[X_{n+1} | X_0, \dots, X_n] = X_n$

For condition 1, as X_n is a non-negative random variable

$$E[|X_n|] = E[X_n]$$

$$= E[2^n U_1 U_2 \dots U_n]$$

$$= 2^n E[U_1] E[U_2] \dots E[U_n]$$

where $U_i, i=1, 2, \dots, n$ are indep. r.v.s

$$= 2^n \cdot \underbrace{\frac{1}{2} \cdot \frac{1}{2} \dots \frac{1}{2}}_{n \text{ times}}$$

$$\therefore E[|X_n|] = \frac{2^n}{2^n} = 1 < \infty \quad (1)$$

For condition 2,

$$E[X_{n+1} | X_0, \dots, X_n]$$

$$= E[2^{n+1} U_1 \dots U_n U_{n+1} | X_0, \dots, X_n]$$

$$= 2^n U_1 \dots U_n E[2 U_{n+1} | X_0, \dots, X_n]$$

this product is out because it's determined by X_0, \dots, X_n .

Remember
If $U \sim \text{uniform}(a, b)$
then $E(U) = \frac{1}{2}(a+b)$
i.e. $E(U_i) = \frac{1}{2}(0+1) = \frac{1}{2}$
for $U_i \sim \text{uniform}(0, 1)$

$$\begin{aligned}
 & \therefore E[X_{n+1} | X_0, \dots, X_n] \\
 &= 2^n U_1 \dots U_n \cdot 2 E[U_{n+1}] \\
 &= 2^n U_1 \dots U_n \cdot 2 \cdot \frac{1}{2} \quad \text{because } U_{n+1} \text{ is independent of } X_0, \dots, X_n \\
 &= X_n, \quad n = 1, 2, \dots \quad \text{where } E[U_i] = \frac{1}{2}, i = 1, 2, \dots, n \quad \textcircled{2}
 \end{aligned}$$

From ① and ② we can deduce that $X_n, n = 0, 1, 2, \dots$ where $X_0 = 1$ defines a martingale. $\#$

Note that:

We can deduce from pb 2.5.1 that the martingale equality (b) (see the definition of a martingale) can be extended to future times to be in the form

$$E[X_m | X_0, \dots, X_n] = X_n \text{ for } m \geq n.$$