

Probability Distributions

Lupton, Ch. 2-4

MacKay, Ch. 23

Presentation Outline

- A motivating example
- Basic definitions (p.d.f, mean, standard deviation, etc)
- Characteristic functions
- Some standard distributions
 - Gaussian
 - Multivariate Gaussian
 - Log-Normal
 - Poisson
 - Binomial / Multinomial
 - Cauchy (Lorentzian)
 - Beta
- Distributions related to the Gaussian
 - χ^2
 - Student's t
 - F
- Additional distributions you may come across
 - Exponential, biexponential, gamma, Von Mises, Gaussian with wrap-around, entropic
- Some laws / theorems
 - Weak Law of Large Numbers
 - Central Limit Theorem
- Summary

A Motivating Example

- (Taken from **Problem 1** in Robert Lupton's book)
- Studying stellar rotation. Have cluster of stars – are their spin axes correlated?
- Want to test the hypothesis of random orientations
- Therefore, need to know what random orientations will look like in data

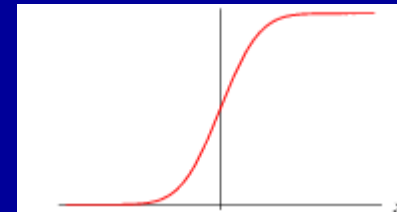
A Motivating Example

- What you see: $v_r \sin(\theta)$
- Uncorrelated distribution means uniform distribution in solid angle Ω
- Therefore need to:
 - Convert between Ω and θ
 - Determine distribution in terms of the observed $\sin(\theta)$

Probability Density Function (p.d.f)

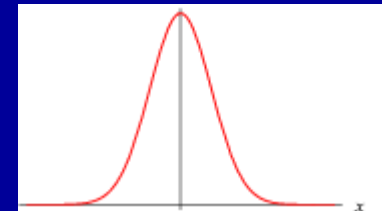
Define $F(x_0)$ as probability that random variable x is less than x_0 .

$F(-\infty)=0$ and $F(\infty)=1$



Define *probability density function (p.d.f)*:

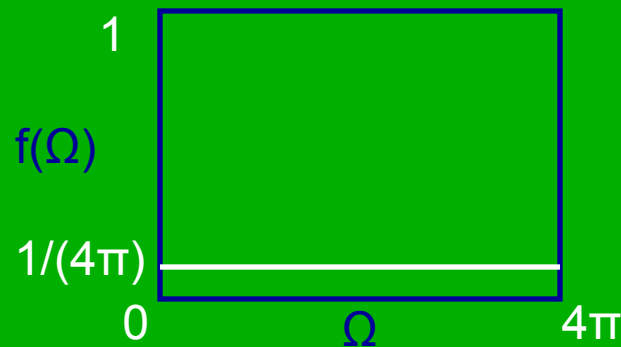
$$f(x) = \frac{dF}{dx}$$



$$Pr(x \in x, x + dx) = F(x + dx) - F(x) = f(x)dx$$

A Motivating Example

- Can calculate the p.d.f. of solid angle Ω
- $f(\Omega) = dF/d\Omega = \text{constant}$ (for uniform distribution)
- Normalize to 1: $\int_{-\infty}^{\infty} f(x)dx = 1$
- Therefore, $f(\Omega) = dF/d\Omega = 1/(4\pi)$



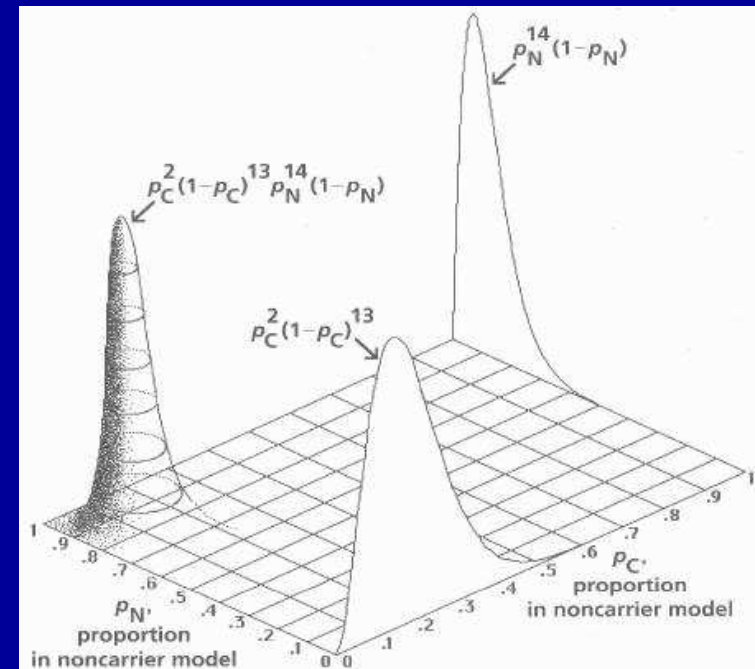
Marginal Distribution

Two random variables, x
and y

$\phi(x)$ = *marginal distribution*
of x :

$$\phi(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

If x and y independent,
 $\phi(x) = f(x)$



P.D.F. of a Function of x

Have p.d.f. $g(x)$ but want p.d.f. $\gamma(\xi)$ of $\xi(x)$.

$$\gamma(\xi) = g(x) \frac{dx}{d\xi}$$

analogous to changing variables in integral

A Motivating Example

- Have p.d.f. of Ω , want p.d.f. of $v_r \sin(\theta)$
- Convert between Ω and θ

$$f(\Omega) = 1/(4\pi)$$

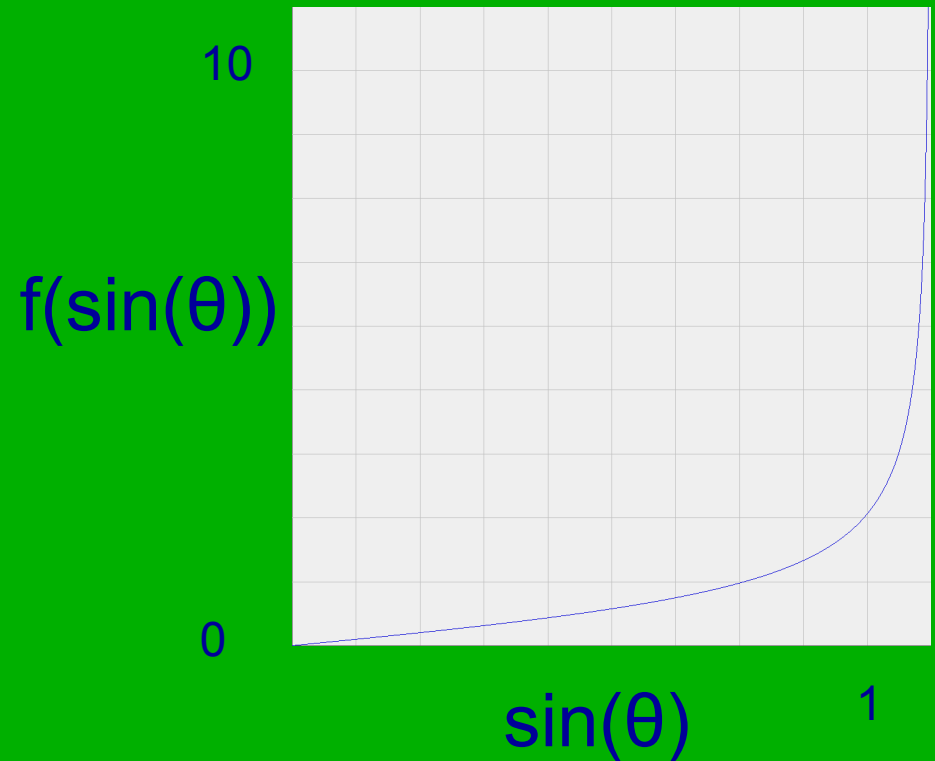
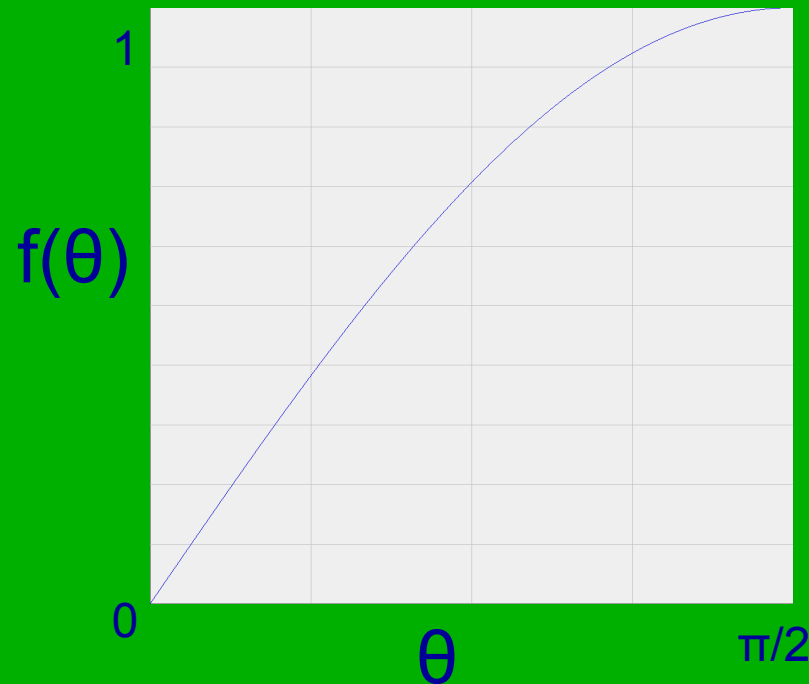
$$d\Omega = 2\pi \sin(\theta) d\theta$$

$$dF(\theta)/d\theta = f(\Omega) * (d\Omega/d\theta)$$

$$\Rightarrow f(\theta) = (1/2)\sin(\theta), \text{ for } \theta=(0,\pi)$$

Since we can't distinguish between θ and $\pi - \theta$, we use $f(\theta) = \sin(\theta)$ for $\theta=(0,\pi/2)$

A Motivating Example



- p.d.f.s of $f(\theta)$ and $f(\sin(\theta))$
- compare data with distribution of $v_r \sin(\theta)$

Center of Distribution

- *(arithmetic) mean μ :*

$$\mu = \int_{-\infty}^{\infty} x f(x) dx$$

- *median:*

$$\frac{1}{2} = \int_{-\infty}^{\text{median}} f(x) dx$$

- *mode:*

$$0 = \left. \frac{df}{dx} \right|_{\text{mode}}$$

- *(geometric) mean u :*

$$\ln(u) = \int_{-\infty}^{\infty} \ln x f(x) dx$$

- $\text{median} \approx (2 \times \text{mean} + \text{mode}) / 3$

Width of Distribution

- variance V :

$$V(x) \equiv \sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

- standard deviation σ : $\sigma^2 = V$

- absolute deviation (about d):

$$\delta = \int_{-\infty}^{\infty} |x - d| f(x) dx$$

- minimized if taken about median

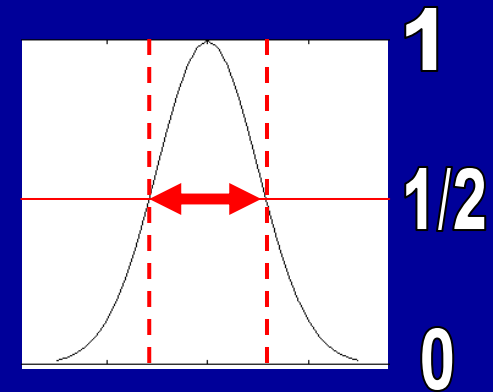
- mean deviation $\bar{\delta}$:

$$\bar{\delta} = \int_{-\infty}^{\infty} |x - \mu| f(x) dx$$

- note: above not guaranteed to exist

Width of Distribution

- **full width at half maximum (FWHM)**: distance between two points where $f(x)$ first falls to half its maximum value



- **interquantile ranges**
 - example: interquartile range:

$$\frac{3/4}{1/4} = \int_{\text{lower quartile}}^{\text{upper quartile}} f(x) dx$$

- **expectation value:** $\langle y \rangle = \int_{-\infty}^{\infty} y f(x) dx$

A Motivating Example

- Find mean and variance of $v_r \sin(\theta)$

$$\mu = \langle v_r \sin \theta \rangle = \int_{-\infty}^{\infty} (v_r \sin \theta f(\theta)) d\theta$$

$$= v_r \int_0^{\pi/2} \sin^2 \theta d(\theta) = v_r \frac{\pi}{4}$$

$$V(\sin \theta) = \int_{-\infty}^{\infty} (\sin \theta - \mu)^2 f(\sin \theta) d(\sin \theta)$$

$$= \frac{2}{3} - \left(\frac{\pi}{4}\right)^2$$

$$\Rightarrow V(v_r \sin \theta) = v_r \left(\frac{2}{3} - \left(\frac{\pi}{4}\right)^2 \right)$$

Moments

- Generalize mean and variance to higher orders:

$$\mu'_r = \langle x^r \rangle$$

$$\mu = \mu'_1, \quad \sigma^2 = \mu_2 = \langle x^2 \rangle - \mu^2$$

$$\mu_r = \langle (x - \mu'_1)^r \rangle$$

- $\mu_3 / \mu_2^{3/2} = \text{skewness}$

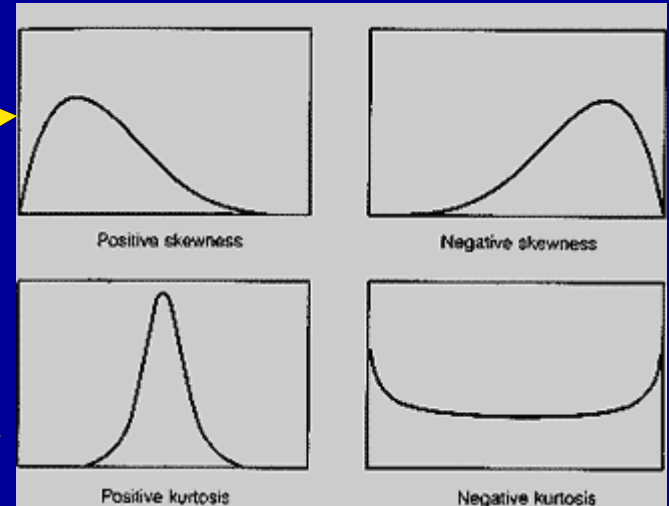
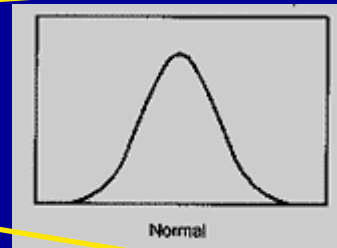
- $\mu_4 / \mu_2^2 - 3 = \text{kurtosis}$

- mixed moments
 - covariance

$$\sigma_{xy} = \langle (x - \langle x \rangle)(y - \langle y \rangle) \rangle$$

- product-moment correlation coefficient
 - (always within $[-1, 1]$)

$$\rho_{xy} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$



Multiple-Variable Distributions

Example: the distribution of $g(x,y) = xy$ given the distributions of x and y is calculated as:

$$F_g = \int_{-\infty}^0 f_x(x) dx \int_{g/x}^{\infty} f_y(y) dy + \int_0^{\infty} f_x(x) dx \int_{-\infty}^{g/x} f_y(y) dy$$

$$f_g = - \int_{-\infty}^0 f_x(x) f_y(g/x) \frac{dx}{x} + \int_0^{\infty} f_x(x) f_y(g/x) \frac{dx}{x}$$

Characteristic Functions

- Use an inverse Fourier transform of f to ease calculation of moments, etc.
- Definition:

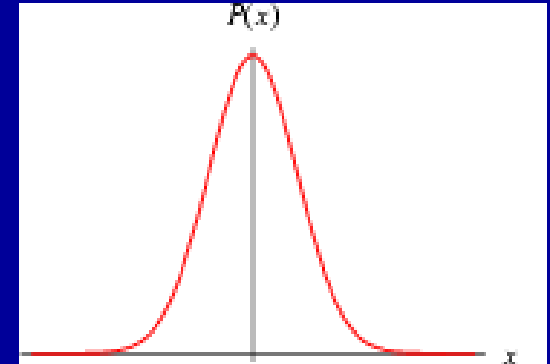
$$\phi(t) = \langle e^{ixt} \rangle = 1 + (it)\mu'_1 + \frac{(it)^2}{2!}\mu'_2 + \dots$$

$$f(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-ixt} \phi(t) dt$$

Useful Properties of Characteristic Functions

1. Derive moments:
$$\mu'_r = \left. \frac{d^r \phi}{d(it)^r} \right|_{t=0}$$
2. One-to-one mapping between f and ϕ
3. Easy to determine distribution of sum of n distributions
4. If $z=x/n$, then $\phi_z(t) = \langle e^{itx/n} \rangle = \phi(t/n)$
5. If $z=(x-\mu)/\sigma$, then $\phi_z(t) = e^{-it\mu/\sigma} \phi(t/\sigma)$
6. Can always get moments directly from ϕ even if you can't invert ϕ to get f

Gaussian Distribution



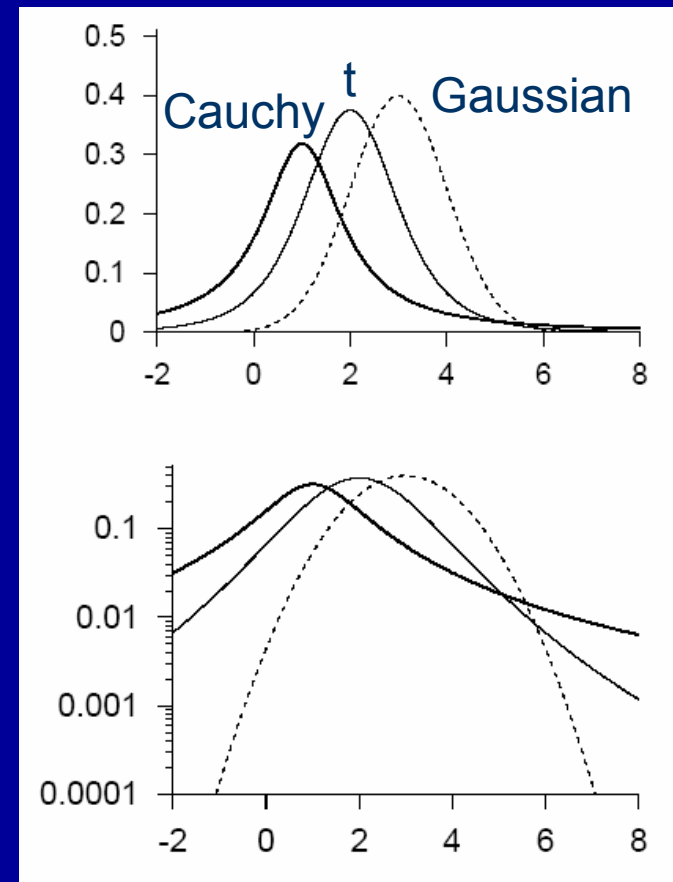
- Definition:

$$dF(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

- A.k.a. Normal Distribution, written $N(\mu, \sigma^2)$
- Used for:
 - description of noise in detectors
 - CMB anisotropies and primordial density perturbations
 - emission lines
 - beam profiles
 - and nearly everything else...

Cautionary Note on Gaussians

- May be overused: MacKay points out that Gaussians have very light tails
- Consider using another unimodal distribution (such as Student's t-distribution or Cauchy)



Characteristic Function Example

- Characteristic function of a Gaussian

$$\phi(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2} + itx} dx$$

$$= e^{it\mu - t^2\sigma^2/2}$$

$$\Rightarrow \mu'_1 = \mu, \quad \mu_2 = \sigma^2, \quad \mu_{2n+1} = 0, \quad \mu_4 = 3\sigma^2$$

- Both skewness and kurtosis equal to zero

Multivariate Gaussian Distribution

- n variables, each individually following Gaussian distribution
- p.d.f.:

$$f(\vec{x}) = \frac{1}{(2\pi)^{n/2} |V|^{1/2}} \exp(-(\vec{x} - \vec{\mu})^T V^{-1} (\vec{x} - \vec{\mu}) / 2)$$

- Covariance matrix V (symmetric, positive definite), elements:

$$V_{ij} = \langle (x_i - \bar{x}_i)(x_j - \bar{x}_j) \rangle$$

Covariance Matrix

- Leading diagonal consists of variances
- Other terms are covariances
- Changing variables to diagonalize V , p.d.f. becomes product of n independent Gaussian variables

Log-Normal Distribution

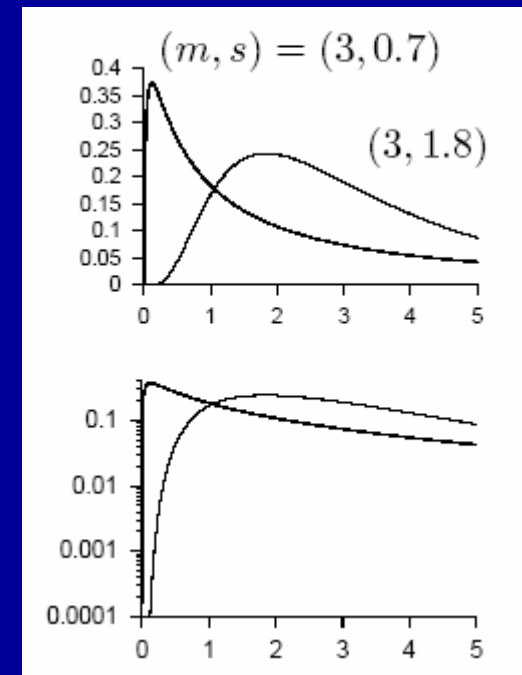
- If x has normal distribution, and $x = \gamma + \delta \ln(y)$, then y has a log-normal distribution

- p.d.f.:


$$dF = \frac{\delta}{\sqrt{2\pi}} e^{-(\gamma + \delta \ln y)^2 / 2} \frac{dy}{y}$$

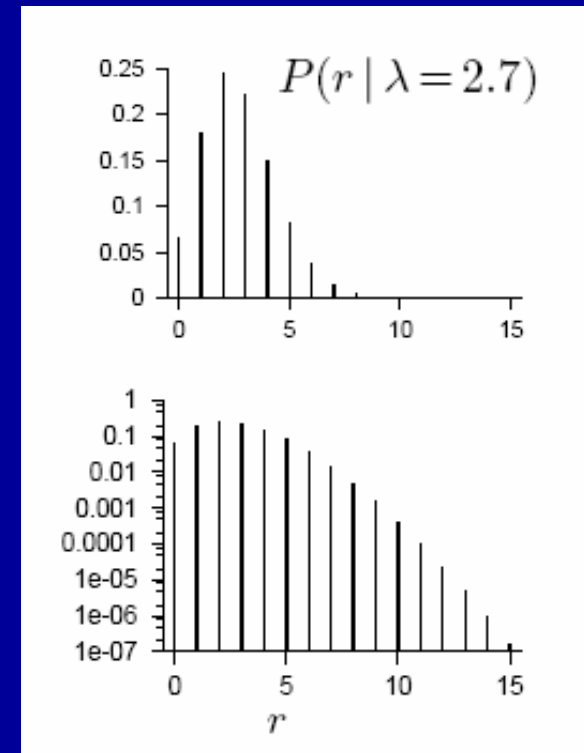
- applications:

- populations of gamma ray bursts
- coronal mass ejections
- size distributions of disk galaxies



Poisson Distribution

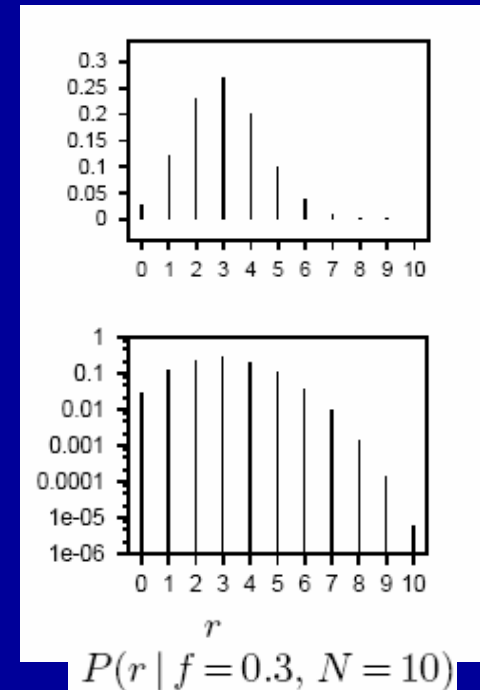
- Used for events which occur with a constant but small probability (classic example: radioactive decay )
- Definition:
$$p_n = \frac{\mu^n}{n!} e^{-\mu}$$
- Limiting case: when μ (average number of events in a time interval) becomes large, have $N(\mu, \mu)$



Binomial Distribution

- Process with two outcomes (e.g., unfair coin toss), with respective probabilities p and $q=1-p$, carried out n times – binomial tells probability of getting a certain outcome with p probability r times
- Definition:
- Limiting cases:
 - large n , small p (i.e., mean $\mu=np$ remains constant): Poisson distribution with mean μ
 - large n , p and q constant: $N(np, npq)$
- Comes up in random walk problems (see, e.g., Problem 12 in Lupton's book)

$$P(r) = \binom{n}{r} p^r q^{n-r}$$



Multinomial Distribution

- Generalization of binomial: N possible outcomes, probabilities p_i (e.g., weighted dice)
- Definition:

$$P(n_1, \dots, n_i, \dots, n_N) = \frac{n!}{n_1! \dots n_i! \dots n_N!} p_1^{n_1} \dots p_i^{n_i} \dots p_N^{n_N}$$

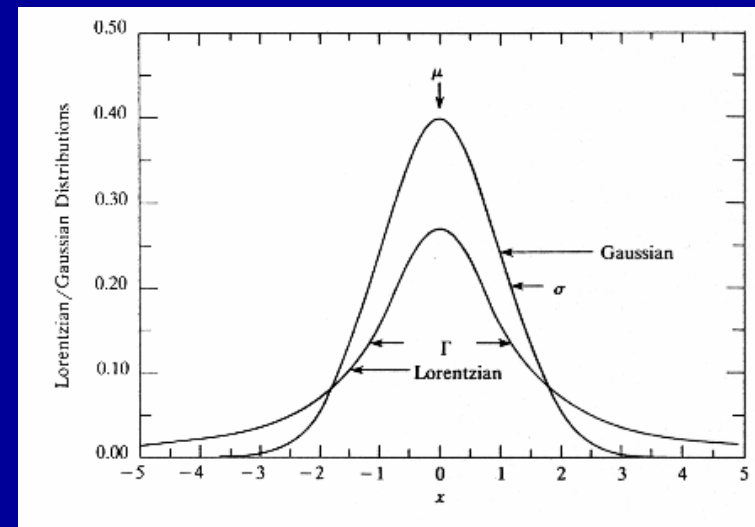
$$\sum n_i = n$$

Cauchy / Lorentzian Distribution

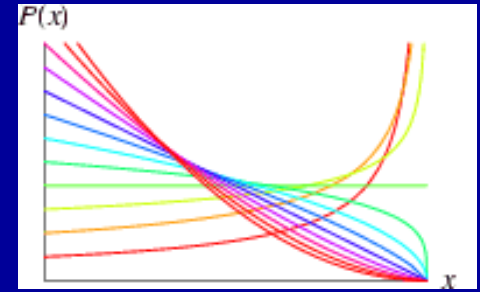
- Similar in shape to Gaussian but with heavier tails
- p.d.f.:

$$dF = \frac{1}{\pi} \frac{1}{1 + (x - \mu)^2} dx$$

- No moments!
- Applications: natural line profiles, collisional broadening, Fabry-Perot etalons, MCMC



Beta Distribution



- Parameters α , β can be adjusted to make a wide range of unimodal distributions
- p.d.f.:

$$dF(x; \alpha, \beta) = \begin{cases} \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} dx & \text{if } 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

- $\alpha = \beta = 1$ is a uniform distribution

Distributions Related to Gaussian

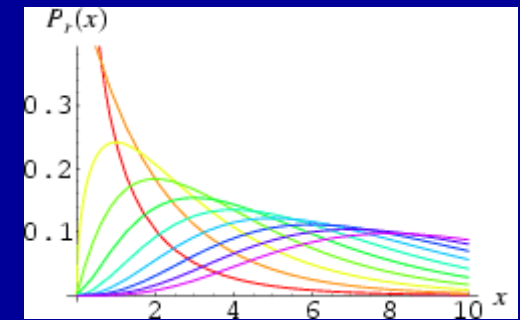
- χ^2
 - sum of squares of independent $N(0,1)$ variates
- Student's t
 - as $n \rightarrow \infty$, t distribution becomes $N(0,1)$
- F
 - distribution of ratio of variances estimated from two Gaussian populations

χ^2 Distribution

- *distribution of sum of squares of independent $N(0,1)$ variates*

$$dF = \frac{1}{2^{n/2}(n/2 - 1)!} e^{-X^2/2} X^{n-2} dX^2$$

$$dF = \frac{1}{2^{n/2-1}(n/2 - 1)!} e^{-X^2/2} X^{n-1} dX$$

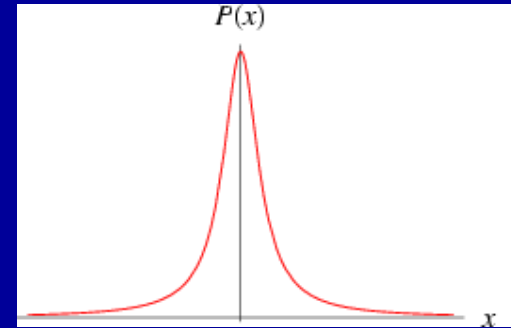


- n = number of degrees of freedom

χ^2 Distribution

- *distribution of sum of squares of independent $N(0,1)$ variates*
- *If $X_m^2 =$ sum of squares of m independent $N(0,1)$ variates; X_m^2 follows χ_m^2 distribution, and X_n^2 follows χ_n^2 distribution*
- *Then: $X_m^2 + X_n^2$ is a χ_{m+n}^2 variate*
- *If $Y_n^2 + Y_m^2$ follows a χ_{m+n}^2 distribution, and Y_n^2 follows χ_n^2 , and Y_m^2 and Y_n^2 independent, then Y_m^2 is χ_m^2 .*

Student's t Distribution



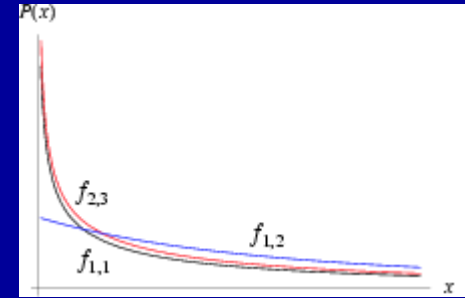
- *as $\nu \rightarrow \infty$, t distribution becomes $N(0, 1)$*

$$dF_t = \frac{(\nu/2 - 1/2)!}{(\nu\pi)^{1/2}(\nu/2 - 1)!} \frac{1}{(1 + t^2/\nu)^{(\nu+1)/2}} dt$$

- Arises in sampling: Take n independent measurements of x, get sample mean \bar{x} , sample variance s. Then the t-distribution describes the distribution of:

$$t \equiv \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

F Distribution



- *distribution of ratio of variances estimated from two Gaussian populations*

- Fisher's F / variance-ratio distribution:

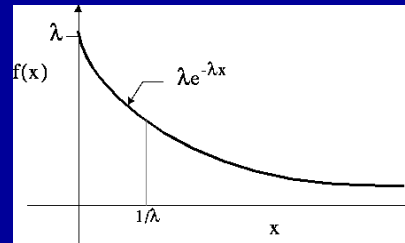
$$F_{\alpha\beta} = \frac{\chi_{\alpha}^2/\alpha}{\chi_{\beta}^2/\beta}$$

- p.d.f.:
$$dF_F = \frac{\alpha^{\alpha/2} \beta^{\beta/2} F^{\alpha/2-1}}{B(\alpha/2, \beta/2)(\alpha F + \beta)^{(\alpha+\beta)/2}} dF$$

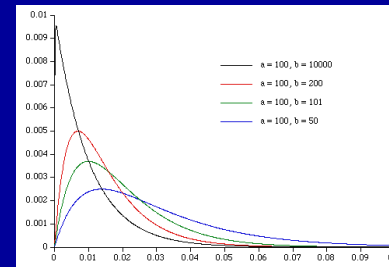
- Ratio, f , of estimation variances of two Gaussian populations follows F distribution; use departure of f from unity to test whether distributions have equal variance.

Additional Distributions

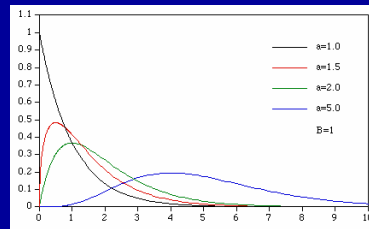
- exponential



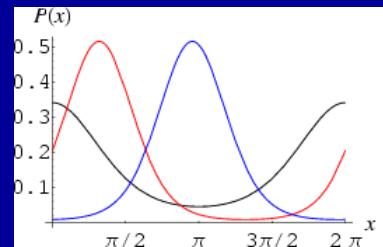
- bi-exponential



- gamma



- Von Mises



- Gaussian with wrap-around

- entropic

Weak Law of Large Numbers

- Also called the *Law of Averages*
- Basic idea: average sampled values of random variables approach the true mean as the number of samples goes to infinity

Let $S_n = X_1 + X_2 + \cdots + X_n$. Then for any $\epsilon > 0$,

$$P \left(\left| \frac{S_n}{n} - \mu \right| \geq \epsilon \right) \rightarrow 0$$

as $n \rightarrow \infty$.

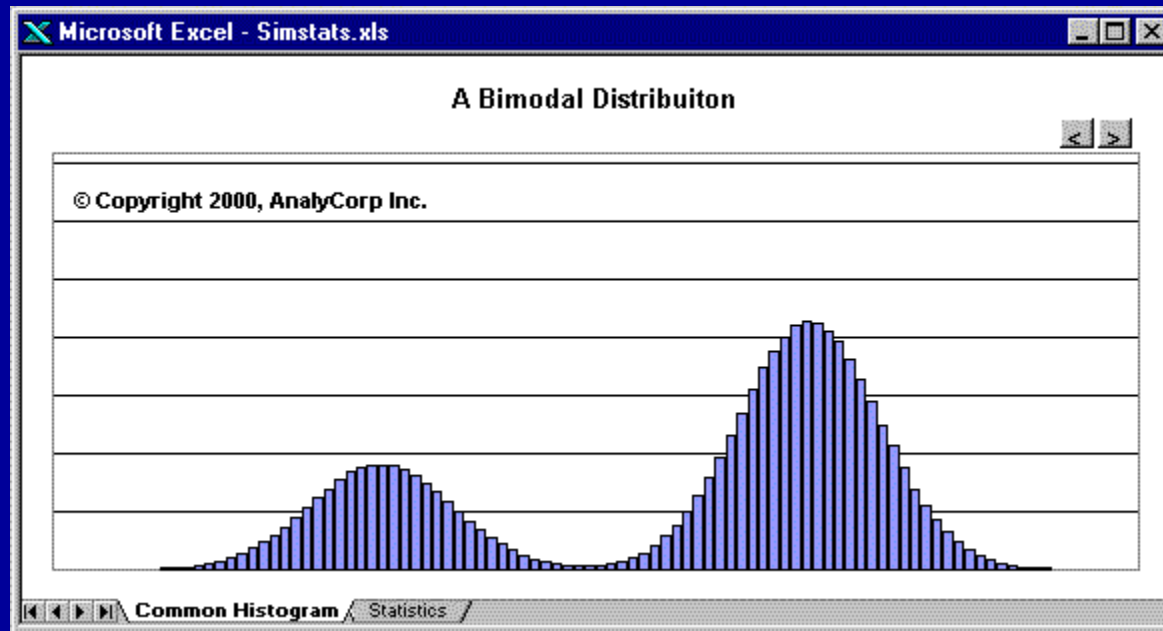
Central Limit Theorem

- Given any distribution with mean μ and variance σ^2 , the sampling distribution of the mean approaches a normal distribution with mean μ and variance σ^2/N as N increases
- More generally: For a set of N independent random variates X_1, X_2, \dots, X_N , each with arbitrary probability distribution $P(x_i, \dots, x_N)$ with mean μ_i and finite variance σ_i^2 , then the normal form variate

$$X_{\text{norm}} \equiv \frac{\sum_{i=1}^N x_i - \sum_{i=1}^N \mu_i}{\sqrt{\sum_{i=1}^N \sigma_i^2}}$$

is normally distributed as N goes to infinity.

Central Limit Theorem



Summary

- Tools for later use:
 - Probability Density Function (p.d.f.)
 - Mean, standard deviation, etc. (characterize p.d.f.)
 - Changing variables in p.d.f.s
 - Characteristic Functions (make life easier)
 - Weak Law of Large Numbers
 - Central Limit Theorem
 - The Gaussian and its relatives

