

## Chapter 4: Mathematical Expectation:

### 4.1 Mean of a Random Variable:

#### **Definition 4.1:**

Let  $X$  be a random variable with a probability distribution  $f(x)$ . The mean (or expected value) of  $X$  is denoted by  $\mu_X$  (or  $E(X)$ ) and is defined by:

$$E(X)=\mu_X = \begin{cases} \sum_{\text{all } x} x f(x); & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} x f(x) dx; & \text{if } X \text{ is continuous} \end{cases}$$

#### **Example 4.1:** (Reading Assignment)

#### **Example:** (Example 3.3)

A shipment of 8 similar microcomputers to a retail outlet contains 3 that are defective and 5 are non-defective. If a school makes a random purchase of 2 of these computers, find the expected number of defective computers purchased

#### **Solution:**

Let  $X$  = the number of defective computers purchased.

In Example 3.3, we found that the probability distribution of  $X$  is:

$x$	0	1	2
$f(x) = P(X=x)$	$\frac{10}{28}$	$\frac{15}{28}$	$\frac{3}{28}$

or:

$$f(x) = P(X = x) = \begin{cases} \frac{\binom{3}{x} \times \binom{5}{2-x}}{\binom{8}{2}}; & x = 0, 1, 2 \\ 0; & \text{otherwise} \end{cases}$$

The expected value of the number of defective computers purchased is the mean (or the expected value) of  $X$ , which is:

$$\begin{aligned} E(X) = \mu_X &= \sum_{x=0}^2 x f(x) \\ &= (0) f(0) + (1) f(1) + (2) f(2) \end{aligned}$$

$$\begin{aligned}
 &= (0) \frac{10}{28} + (1) \frac{15}{28} + (2) \frac{3}{28} \\
 &= \frac{15}{28} + \frac{6}{28} = \frac{21}{28} = 0.75 \text{ (computers)}
 \end{aligned}$$

**Example 4.3:**

Let  $X$  be a continuous random variable that represents the life (in hours) of a certain electronic device. The pdf of  $X$  is given by:

$$f(x) = \begin{cases} \frac{20,000}{x^3} & ; x > 100 \\ 0 & ; \textit{elsewhere} \end{cases}$$

Find the expected life of this type of devices.

**Solution:**

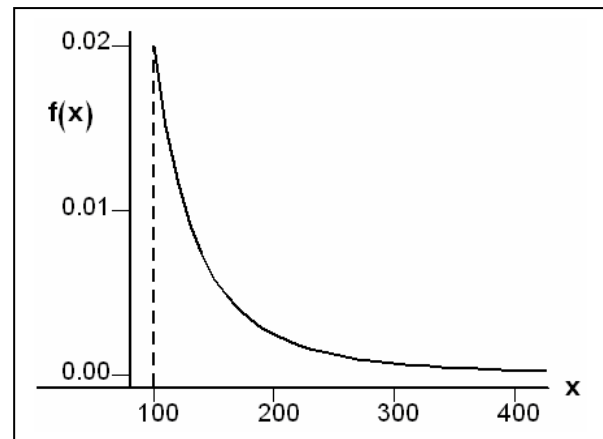
$$E(X) = \mu_X = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_{100}^{\infty} x \frac{20000}{x^3} dx$$

$$= 20000 \int_{100}^{\infty} \frac{1}{x^2} dx$$

$$= 20000 \left[ -\frac{1}{x} \Big|_{x=100}^{x=\infty} \right]$$

$$= -20000 \left[ 0 - \frac{1}{100} \right] = 200 \text{ (hours)}$$



Therefore, we expect that this type of electronic devices to last, on average, 200 hours.

**Theorem 4.1:**

Let  $X$  be a random variable with a probability distribution  $f(x)$ , and let  $g(X)$  be a function of the random variable  $X$ . The mean (or expected value) of the random variable  $g(X)$  is denoted by  $\mu_{g(X)}$  (or  $E[g(X)]$ ) and is defined by:

$$E[g(X)] = \mu_{g(X)} = \begin{cases} \sum_{\text{all } x} g(x) f(x) & ; \textit{if } X \textit{ is discrete} \\ \int_{-\infty}^{\infty} g(x) f(x) dx & ; \textit{if } X \textit{ is continuous} \end{cases}$$

**Example:**

Let  $X$  be a discrete random variable with the following probability distribution

$x$	0	1	2
$f(x)$	$\frac{10}{28}$	$\frac{15}{28}$	$\frac{3}{28}$

Find  $E[g(X)]$ , where  $g(X)=(X-1)^2$ .

**Solution:**

$$g(X)=(X-1)^2$$

$$\begin{aligned} E[g(X)] &= \mu_{g(X)} = \sum_{x=0}^2 g(x) f(x) = \sum_{x=0}^2 (x-1)^2 f(x) \\ &= (0-1)^2 f(0) + (1-1)^2 f(1) + (2-1)^2 f(2) \\ &= (-1)^2 \frac{10}{28} + (0)^2 \frac{15}{28} + (1)^2 \frac{3}{28} \\ &= \frac{10}{28} + 0 + \frac{3}{28} = \frac{13}{28} \end{aligned}$$

**Example:**

In Example 4.3, find  $E\left(\frac{1}{X}\right)$ . {note:  $g(X) = \frac{1}{X}$ }

**Solution:**

$$f(x) = \begin{cases} \frac{20,000}{x^3} & ; x > 100 \\ 0 & ; \text{elsewhere} \end{cases}$$

$$g(X) = \frac{1}{X}$$

$$\begin{aligned} E\left(\frac{1}{X}\right) &= E[g(X)] = \mu_{g(X)} = \int_{-\infty}^{\infty} g(x) f(x) dx = \int_{-\infty}^{\infty} \frac{1}{x} f(x) dx \\ &= \int_{100}^{\infty} \frac{1}{x} \frac{20000}{x^3} dx = 20000 \int_{100}^{\infty} \frac{1}{x^4} dx = \frac{20000}{-3} \left[ \frac{1}{x^3} \Big|_{x=100}^{x=\infty} \right] \\ &= \frac{-20000}{3} \left[ 0 - \frac{1}{1000000} \right] = 0.0067 \end{aligned}$$

## **4.2 Variance (of a Random Variable):**

The most important measure of variability of a random variable  $X$  is called the variance of  $X$  and is denoted by  $\text{Var}(X)$  or  $\sigma_X^2$ .

### **Definition 4.3:**

Let  $X$  be a random variable with a probability distribution  $f(x)$  and mean  $\mu$ . The variance of  $X$  is defined by:

$$\text{Var}(X) = \sigma_X^2 = E[(X - \mu)^2] = \begin{cases} \sum_{\text{all } x} (x - \mu)^2 f(x); & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx; & \text{if } X \text{ is continuous} \end{cases}$$

### **Definition:**

The positive square root of the variance of  $X$ ,  $\sigma_X = \sqrt{\sigma_X^2}$ , is called the standard deviation of  $X$ .

Note:

$$\text{Var}(X) = E[g(X)], \text{ where } g(X) = (X - \mu)^2$$

### **Theorem 4.2:**

The variance of the random variable  $X$  is given by:

$$\text{Var}(X) = \sigma_X^2 = E(X^2) - \mu^2$$

$$\text{where } E(X^2) = \begin{cases} \sum_{\text{all } x} x^2 f(x); & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{\infty} x^2 f(x) dx; & \text{if } X \text{ is continuous} \end{cases}$$

### **Example 4.9:**

Let  $X$  be a discrete random variable with the following probability distribution

$x$	0	1	2	3
$f(x)$	0.51	0.38	0.10	0.01

Find  $\text{Var}(X) = \sigma_X^2$ .

### **Solution:**

$$\begin{aligned} \mu &= \sum_{x=0}^3 x f(x) = (0) f(0) + (1) f(1) + (2) f(2) + (3) f(3) \\ &= (0) (0.51) + (1) (0.38) + (2) (0.10) + (3) (0.01) \\ &= 0.61 \end{aligned}$$

1. First method:

$$\begin{aligned}\text{Var}(X) &= \sigma_X^2 = \sum_{x=0}^3 (x - \mu)^2 f(x) \\ &= \sum_{x=0}^3 (x - 0.61)^2 f(x) \\ &= (0 - 0.61)^2 f(0) + (1 - 0.61)^2 f(1) + (2 - 0.61)^2 f(2) + (3 - 0.61)^2 f(3) \\ &= (-0.61)^2 (0.51) + (0.39)^2 (0.38) + (1.39)^2 (0.10) + (2.39)^2 (0.01) \\ &= 0.4979\end{aligned}$$

2. Second method:

$$\begin{aligned}\text{Var}(X) &= \sigma_X^2 = E(X^2) - \mu^2 \\ E(X^2) &= \sum_{x=0}^3 x^2 f(x) = (0^2) f(0) + (1^2) f(1) + (2^2) f(2) + (3^2) f(3) \\ &= (0) (0.51) + (1) (0.38) + (4) (0.10) + (9) (0.01) \\ &= 0.87 \\ \text{Var}(X) &= \sigma_X^2 = E(X^2) - \mu^2 = 0.87 - (0.61)^2 = 0.4979\end{aligned}$$

**Example 4.10:**

Let X be a continuous random variable with the following pdf:

$$f(x) = \begin{cases} 2(x-1) & ; 1 < x < 2 \\ 0 & ; \text{elsewhere} \end{cases}$$

Find the mean and the variance of X.

**Solution:**

$$\begin{aligned}\mu = E(X) &= \int_{-\infty}^{\infty} x f(x) dx = \int_1^2 x [2(x-1)] dx = 2 \int_1^2 x(x-1) dx = 5/3 \\ E(X^2) &= \int_{-\infty}^{\infty} x^2 f(x) dx = \int_1^2 x^2 [2(x-1)] dx = 2 \int_1^2 x^2 (x-1) dx = 17/6 \\ \text{Var}(X) &= \sigma_X^2 = E(X^2) - \mu^2 = 17/6 - (5/3)^2 = 1/18\end{aligned}$$

### 4.3 Means and Variances of Linear Combinations of Random Variables:

If  $X_1, X_2, \dots, X_n$  are  $n$  random variables and  $a_1, a_2, \dots, a_n$  are constants, then the random variable :

$$Y = \sum_{i=1}^n a_i X_i = a_1 X_1 + a_2 X_2 + \dots + a_n X_n$$

is called a linear combination of the random variables  $X_1, X_2, \dots, X_n$ .

#### **Theorem 4.5:**

If  $X$  is a random variable with mean  $\mu = E(X)$ , and if  $a$  and  $b$  are constants, then:

$$E(aX \pm b) = a E(X) \pm b$$

$$\Leftrightarrow$$

$$\mu_{aX \pm b} = a \mu_X \pm b$$

**Corollary 1:**  $E(b) = b$  (a=0 in Theorem 4.5)

**Corollary 2:**  $E(aX) = a E(X)$  (b=0 in Theorem 4.5)

#### **Example 4.16:**

Let  $X$  be a random variable with the following probability density function:

$$f(x) = \begin{cases} \frac{1}{3}x^2 & ; -1 < x < 2 \\ 0 & ; elsewhere \end{cases}$$

Find  $E(4X+3)$ .

**Solution:**

$$\mu = E(X) = \int_{-\infty}^{\infty} x f(x) dx = \int_{-1}^2 x \left[ \frac{1}{3}x^2 \right] dx = \frac{1}{3} \int_{-1}^2 x^3 dx = \frac{1}{3} \left[ \frac{1}{4}x^4 \Big|_{x=-1}^{x=2} \right] = 5/4$$

$$E(4X+3) = 4 E(X) + 3 = 4(5/4) + 3 = 8$$

Another solution:

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx \quad ; \quad g(X) = 4X+3$$

$$E(4X+3) = \int_{-\infty}^{\infty} (4x+3) f(x) dx = \int_{-1}^2 (4x+3) \left[ \frac{1}{3}x^2 \right] dx = \dots = 8$$

**Theorem:**

If  $X_1, X_2, \dots, X_n$  are  $n$  random variables and  $a_1, a_2, \dots, a_n$  are constants, then:

$$E(a_1X_1 + a_2X_2 + \dots + a_nX_n) = a_1E(X_1) + a_2E(X_2) + \dots + a_nE(X_n)$$

$$\Leftrightarrow$$

$$E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E(X_i)$$

**Corollary:**

If  $X$ , and  $Y$  are random variables, then:

$$E(X \pm Y) = E(X) \pm E(Y)$$

**Theorem 4.9:**

If  $X$  is a random variable with variance  $Var(X) = \sigma_X^2$  and if  $a$  and  $b$  are constants, then:

$$Var(aX \pm b) = a^2 Var(X)$$

$$\Leftrightarrow$$

$$\sigma_{aX \pm b}^2 = a^2 \sigma_X^2$$

**Theorem:**

If  $X_1, X_2, \dots, X_n$  are  $n$  independent random variables and  $a_1, a_2, \dots, a_n$  are constants, then:

$$\begin{aligned} Var(a_1X_1 + a_2X_2 + \dots + a_nX_n) \\ = a_1^2 Var(X_1) + a_2^2 Var(X_2) + \dots + a_n^2 Var(X_n) \end{aligned}$$

$$\Leftrightarrow$$

$$Var\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i^2 Var(X_i)$$

$$\Leftrightarrow$$

$$\sigma_{a_1X_1 + a_2X_2 + \dots + a_nX_n}^2 = a_1^2 \sigma_{X_1}^2 + a_2^2 \sigma_{X_2}^2 + \dots + a_n^2 \sigma_{X_n}^2$$

**Corollary:**

If  $X$ , and  $Y$  are independent random variables, then:

- $Var(aX + bY) = a^2 Var(X) + b^2 Var(Y)$
- $Var(aX - bY) = a^2 Var(X) + b^2 Var(Y)$
- $Var(X \pm Y) = Var(X) + Var(Y)$

**Example:**

Let  $X$ , and  $Y$  be two independent random variables such that  $E(X)=2$ ,  $\text{Var}(X)=4$ ,  $E(Y)=7$ , and  $\text{Var}(Y)=1$ . Find:

1.  $E(3X+7)$  and  $\text{Var}(3X+7)$
2.  $E(5X+2Y-2)$  and  $\text{Var}(5X+2Y-2)$ .

**Solution:**

1.  $E(3X+7) = 3E(X)+7 = 3(2)+7 = 13$   
 $\text{Var}(3X+7) = (3)^2 \text{Var}(X) = (3)^2 (4) = 36$
2.  $E(5X+2Y-2) = 5E(X) + 2E(Y) - 2 = (5)(2) + (2)(7) - 2 = 22$   
 $\text{Var}(5X+2Y-2) = \text{Var}(5X+2Y) = 5^2 \text{Var}(X) + 2^2 \text{Var}(Y)$   
 $= (25)(4) + (4)(1) = 104$

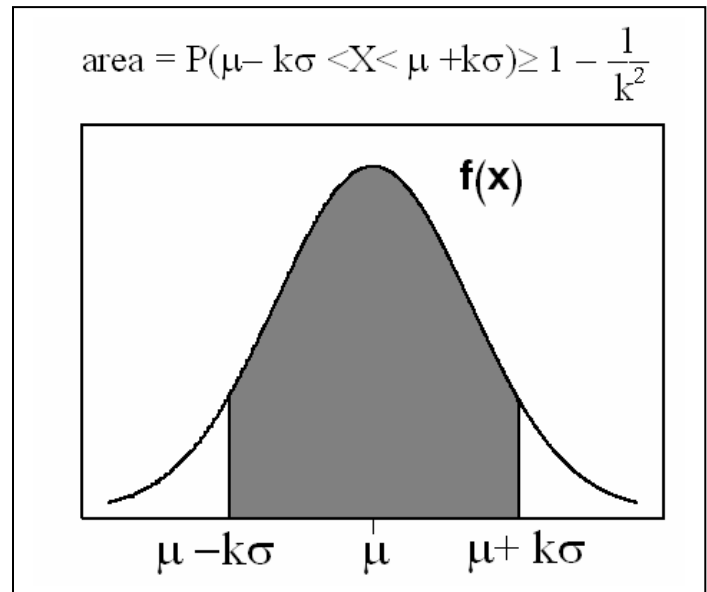


#### 4.4 Chebyshev's Theorem:

\* Suppose that  $X$  is any random variable with mean  $E(X)=\mu$  and variance  $\text{Var}(X)=\sigma^2$  and standard deviation  $\sigma$ .

\* Chebyshev's Theorem gives a conservative estimate of the probability that the random variable  $X$  assumes a value within  $k$  standard deviations ( $k\sigma$ ) of its mean  $\mu$ , which is  $P(\mu - k\sigma < X < \mu + k\sigma)$ .

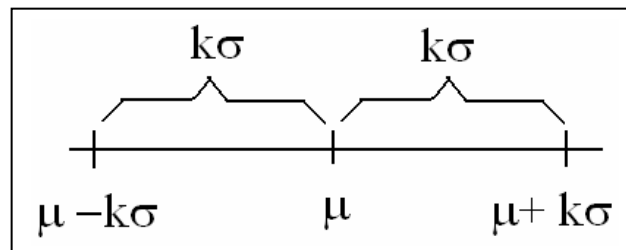
$$* P(\mu - k\sigma < X < \mu + k\sigma) \approx 1 - \frac{1}{k^2}$$



#### Theorem 4.11:(Chebyshev's Theorem)

Let  $X$  be a random variable with mean  $E(X)=\mu$  and variance  $\text{Var}(X)=\sigma^2$ , then for  $k>1$ , we have:

$$P(\mu - k\sigma < X < \mu + k\sigma) \geq 1 - \frac{1}{k^2} \Leftrightarrow P(|X - \mu| < k\sigma) \geq 1 - \frac{1}{k^2}$$



#### Example 4.22:

Let  $X$  be a random variable having an unknown distribution with mean  $\mu=8$  and variance  $\sigma^2=9$  (standard deviation  $\sigma=3$ ). Find the following probability:

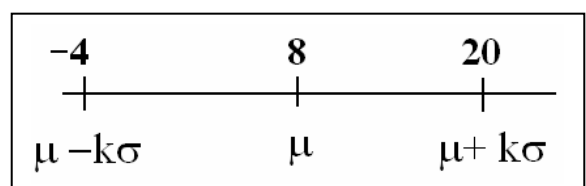
(a)  $P(-4 < X < 20)$

(b)  $P(|X-8| \geq 6)$

#### Solution:

(a)  $P(-4 < X < 20) = ??$

$$P(\mu - k\sigma < X < \mu + k\sigma) \geq 1 - \frac{1}{k^2}$$



$$(-4 < X < 20) = (\mu - k\sigma < X < \mu + k\sigma)$$

$$\begin{aligned} -4 = \mu - k\sigma &\Leftrightarrow -4 = 8 - k(3) \quad \text{or} \quad 20 = \mu + k\sigma \Leftrightarrow 20 = 8 + k(3) \\ &\Leftrightarrow -4 = 8 - 3k && \Leftrightarrow 20 = 8 + 3k \\ &\Leftrightarrow 3k = 12 && \Leftrightarrow 3k = 12 \\ &\Leftrightarrow \mathbf{k=4} && \Leftrightarrow \mathbf{k=4} \end{aligned}$$

$$1 - \frac{1}{k^2} = 1 - \frac{1}{16} = \frac{15}{16}$$

Therefore,  $P(-4 < X < 20) \geq \frac{15}{16}$ , and hence,  $P(-4 < X < 20) \approx \frac{15}{16}$   
(approximately)

$$(b) P(|X - 8| \geq 6) = ??$$

$$P(|X - 8| \geq 6) = 1 - P(|X - 8| < 6)$$

$$P(|X - 8| < 6) = ??$$

$$P(|X - \mu| < k\sigma) \geq 1 - \frac{1}{k^2}$$

$$(|X - 8| < 6) = (|X - \mu| < k\sigma)$$

$$6 = k\sigma \Leftrightarrow 6 = 3k \Leftrightarrow \mathbf{k=2}$$

$$1 - \frac{1}{k^2} = 1 - \frac{1}{4} = \frac{3}{4}$$

$$P(|X - 8| < 6) \geq \frac{3}{4} \Leftrightarrow 1 - P(|X - 8| < 6) \leq 1 - \frac{3}{4}$$

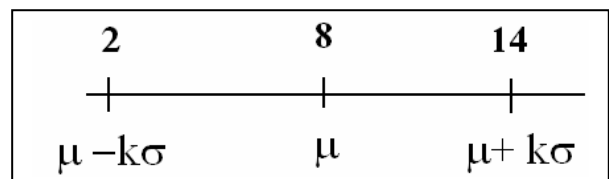
$$\Leftrightarrow 1 - P(|X - 8| < 6) \leq \frac{1}{4}$$

$$\Leftrightarrow P(|X - 8| \geq 6) \leq \frac{1}{4}$$

Therefore,  $P(|X - 8| \geq 6) \approx \frac{1}{4}$  (approximately)

Another solution for part (b):

$$\begin{aligned} P(|X - 8| < 6) &= P(-6 < X - 8 < 6) \\ &= P(-6 + 8 < X < 6 + 8) \\ &= P(2 < X < 14) \end{aligned}$$



$$(2 < X < 14) = (\mu - k\sigma < X < \mu + k\sigma)$$

$$2 = \mu - k\sigma \Leftrightarrow 2 = 8 - k(3) \Leftrightarrow 2 = 8 - 3k \Leftrightarrow 3k = 6 \Leftrightarrow \mathbf{k=2}$$

$$1 - \frac{1}{k^2} = 1 - \frac{1}{4} = \frac{3}{4}$$

$$P(2 < X < 14) \geq \frac{3}{4} \Leftrightarrow P(|X-8| < 6) \geq \frac{3}{4}$$

$$\Leftrightarrow 1 - P(|X-8| < 6) \leq 1 - \frac{3}{4}$$

$$\Leftrightarrow 1 - P(|X-8| < 6) \leq \frac{1}{4}$$

$$\Leftrightarrow P(|X-8| \geq 6) \leq \frac{1}{4}$$

Therefore,  $P(|X-8| \geq 6) \approx \frac{1}{4}$  (approximately)